NAVAL POSTGRADUATE SCHOOL MONTEREY, CALIFORNIA



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THESIS

US ARMY'S DELAY ENTRY PROGRAM:
A SURVIVAL STUDY

by

Jeffrey S. Vales

June 1994

Thesis Advisor:

Lyn Whitaker

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US Army's Delay Entry Program:
A Survival Study

by

Jeffrey S. Vales
Lieutenant, United States Navy Reserve
B.S., University of South Florida, 1985

Submitted in partial fulfillment of the requirements for the degree of

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June 1994

Approved by:

Lyn Whitaker, Thesis Advisor

I'riplious Assuptions Second Reader

Siriphong Lawphongsanich, Second Reader

Peter Purdue, Chairman

Department of Operations Research

ABSTRACT

The Delayed Entry Program (DEP) has served a variety of roles in the recruiting process. In general, it acts as an inventory system of recruits which can be used to smooth out seasonal fluctuations in demand for basic and advanced individual training. It also serves to address the routine seasonal fluctuations in the recruiting process itself. losses, or those individuals that renege on the agreement made with the Army to attend basic training, are a costly aspect of Several studies have examined a number of the program. aspects associated with DEP losses. When setting recruiting goals, Army analyst must consider the pool of individuals who are already in DEP. This thesis attempts to provide a method to estimate how many of those in DEP will survive to the end of the contract and enter basic training. Specifically, this thesis estimates survival probabilities as a function of time spent in DEP, contract length and other pertinent variables.

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EXECUTIVE SUMMARY

Background:

The US Army along with other military services, government organizations and private enterprises need qualified "entry level" individuals to meet present and future manning levels. The Army must actively recruit these individuals to meet target levels. The Delayed Entry Program (DEP) has served a variety of roles in the recruiting process. In general, it acts as an inventory system of recruits which can be used to smooth out seasonal fluctuations in demand for basic and advanced individual training. It also serves to deal with the routine seasonal fluctuations in the recruiting process itself. DEP losses, or those individuals that renege on the agreement made with the Army to attend basic training, are a costly aspect of recruiting. Because of DEP losses, the Army must be careful when setting goals for its recruiters.

Purpose:

When setting recruiting goals, Army analysts must consider the pool of individuals who are already in DEP. Therefore, the question is posed: How many of those in the DEP pool will survice to the end of the DEP and enter basic training? We look for a methodology for estimating how many individuals in the DEP inventory pool will survive the program and access

into the Army. An accurate estimate of DEP survivors (accessions) would assist analyst in setting recruiter goals.

Framework:

In pervious research, demographic characteristics believed to be related to DEP survival (or DEP loss) have be studied. This thesis takes into account some of these demographic characteristics in addressing DEP survival as a function of time already spent in DEP. The logit model applied in this thesis was derived from the statistical literature on analyses of ordinal categorical data.

Approach:

Demographic characteristics, DEP contract length and time in DEP when lost were obtained from the United States Recruiting Command (USAREC) Mini-master file. The modeling effort was restricted to male high-school graduates who score on the top 50 percentile of the Armed Forces Qualification Test (AFQT). USAREC considers this individual a GMA candidate, the most populous recruit category.

Because DEP losses do not occur uniformly over the length of a contract, obtaining the general structure of losses was imperative. A loglinear regression model assisted in determining a method for combining months with few losses. This greatly reduced the number of parameters needed to estimate survival probabilities and provided the basis for

fitting the continuation-ratio model (which employs a cumulative logit) used to estimate the survival probabilities directly.

Conclusion:

The fitted values from the continuation-ratio model can be used to estimate the probabilities necessary to find the expected number of accessions for a given inventory of candidates. These values provide USAREC analysts with a method to determine accession probabilities as a function of the time survived in DEP and other demographic factors. The probabilities obtained can be used to provide USAREC planners a quantitative basis for recruit accession projections needed for recruiter goal setting. While the model was applied to a specific recruit population, the modeling effort can be extended to other groups as well.

I. INTRODUCTION

The United States Army Recruiting Command (USAREC) is responsible for recruiting candidates for the Army's regular and reserve enlistment programs. This thesis pertains to the recruiting process of the former. In particular, it focuses on a widely used tool in military recruiting called the Delayed Entry Program (DEP). Below we give a brief description of the program.

A. DELAYED ENTRY PROGRAM

After an individual decides to join the Army and signs a contract at a Military Entrance Processing (MEP) Station, he/she does not proceed directly to basic training. The time period between the signing of the contract and the beginning of basic training is referred to as the Delayed Entry Program or DEP. DEP generally lasts from one to twelve months. The minimum of one month allows USAREC to perform the necessary administrative processing, background check and Drug and Alcohol Testing (DAT). In general, DEP acts like an inventory of recruits which can be used to smooth out the seasonal fluctuations in recruiting as well as seasonal demand in basic and advanced individual training. Moreover, DEP also allows USAREC to expand its "recruiting market" to include seniors who are still waiting to graduate from high school. After

signing enlistment contracts, these seniors are automatically placed in DEP until after graduation. While in DEP, recruits play a role in attracting candidates for enlistment programs. Recruits in DEP provide additional leads for new candidates and influence others among their peer groups to join the Army.

DEP allows recruits more career choices. While at MEP stations, recruits must meet an Army career councilor who provides them with their assignment options. By varying the length of DEP, desirable assignments may become available at some future dates. Once the desirable assignment is obtained, the counselor then reserves and, thereby, guarantees the assignment for the recruits.

However, DEP comes at a cost to USAREC. Historically, 15 percent of the recruits in DEP renege on their contract and do not access into the Army. Based on the resources spent on recruiting an individual to sign an enlistment contract, the processing cost at the MEP and the administrative cost of DEP, USAREC estimates it spends approximately \$5,000 per recruit who does not access into the Army, or a DEP loss [Ref. 1]. This amounts to an annual cost of \$67 million that USAREC has to spend on DEP losses.

B. PROBLEM STATEMENT

In an effort to reduce the cost of DEP losses, USAREC has commissioned many studies [Ref. 2-5] to analyze the factors

that affect DEP losses. More recently Burris [Ref. 1] studies DEP losses as a function of recruits contracted to be in DEP. To utilize this information in planning, he develops an optimization model to assist USAREC analyst in setting the monthly recruiting goals. In this research, we continue to work towards providing USAREC with a tool to assist in setting monthly recruiting goals.

Prior to the beginning of each quarter, USAREC analyst must set the recruiting goals for each of the next three These goals simply tell the recruiting brigades, months. battalions and stations how many individuals they need to recruit to sign enlistment contracts each month (Individuals who sign enlistment contract are simply referred to as "contracts"). When setting these recruiting goals, the analyst must consider the pool of individuals who are already in DEP. These individuals have been in DEP for various lengths of time and they are also scheduled to be sent to basic training (thereby, accessing into the Army) at different future dates. The problem that faces the analyst at this point is to determine how many individuals in the current DEP pool will eventually enter basic training and access into the Army. We can refer to these individuals as those who survive the DEP. If the expected number of survivors is small, the recruiting goals must be set higher to offset the expected losses. On the other hand, if they are higher, the goals can be adjusted downward. In the current budget environment,

there is an acute need for an accurate estimate of DEP survivors. Thus, it is the goal of this thesis to provide a methodology to estimate the number of DEP survivors. The estimate will be based on the length of time the individual has already spent in DEP, the length of time contracted to be in DEP and other pertinent demographic factors. The main focus of the thesis will be to estimate the probability of DEP survival for the largest recruiting category. This category composes of male high-school graduates with Armed Forces Qualification Test (AFQT) scores in the upper 50 percentile. However, the same methodology can be applied to other recruiting categories as well.

C. THESIS ORGANIZATION

Chapter II gives an overview of the data set used for the research, a description of the variables used in modelling and an exploratory analysis conducted to develop the final model. In addition, contained within chapter II is a discussion of generalized linear models used for our model development. Chapter III concerns the development of the continuation-ratio model used to estimate the desired probabilities. Chapter IV gives the results of the model fit and uses of these results. Finally, Chapter V provides recommendations.

II. PRELIMINARY ANALYSIS

Before attempting to model DEP survival estimates a preliminary study of the data is in order. Understanding which variables influence DEP survival and the general structure of how DEP losses occur helps develop a strategy for model development. This Chapter includes a description of the data base, justification for considering male high school graduates with AFQT scores in the upper 50th percentile apart from other recruits, and motivation for including various explanatory variables in the model. Also contained is the justification for pooling the number of losses in the middle months of contracts with a DEP of 5 or more months.

A. DATA BASE

USAREC keeps extensive records on all candidates that have signed contracts. These records are contained in the Minimaster file and include a variety of information ranging from age and social security number to number of dependents and years of education. A labeling system encoding a candidates gender, AFQT score and education level is also included in the records. USAREC classifies recruits into 22 possible categories or mission boxes. Figure 1 provides the percentages of the predominantly recruited groups or mission boxes for FY 1988 through 1992. Of the mission boxes shown in

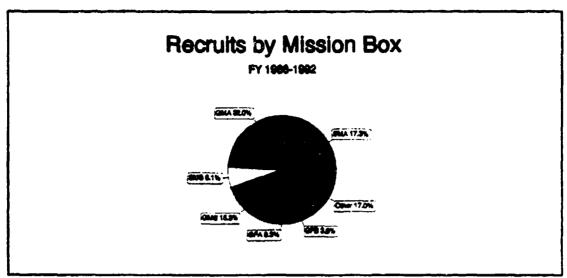


Figure 1 RECRUITS BY MISSION BOX

Figure 1, four are male and two female groups. For the male recruits, these are GMA, graduates with mental category A (upper 50th percentile on the AFQT), SMA, high school seniors with mental category A, GMB, graduates and with mental category B (between the 32 and 50 percentile on the AFQT) and SMB, seniors with mental category B. Female mission boxes or categories are high school graduates and are classified as GFA and GFB which corresponds to graduate female mental category A and B. Burris provides a complete listing of all 22 USAREC mission boxes [Ref 1].

USAREC supplied a subset of the Mini-master database records from FY 1988 though FY 1993. Although records for FY 1993 were available, they were not used due to a significant number of open records. Recruits with open records are still in DEP and therefore the outcome of the contract not known

(loss or accessed). This is particularly true of records for recruits contracted for longer DEP periods.

To confine the scope of our model to typical DEP cases, only those records where the actual time in DEP (based on the accession or loss date) less than or equal to contracted DEP length were considered. For example, an individual with a DEP contract of 5 month with an accession or loss date in month 12 represented an unusual observation and is not considered. Records that indicate contracts of more than twelve months are not used because contracting for this length of time is not army policy. Without this criteria for data selection the number of possible combinations becomes unmanageable. Records not used in the study comprise those that reflect unusual DEP observations or have errors in one or more of the fields. Because our study concentrates on the GMA mission box the number of records removed from the data set are summarized in Table I.

Table I DATA SUMMARY

YR	TOTAL RECORDS	TOTAL GMA RECORDS	GMA REC NOT USED	USED GMA RECORDS
88	131,724	34,251	775	33,476
89	136,184	34,538	1,099	33,439
90	96,941	35,331	1,055	34,276
91	86,048	37,045	821	36,244
92	59,825	22,649	570	22,079

B. EXPLANATORY VARIABLES

DEP length, the outcome of the contract (loss ox accession) and if a loss, when it occurred are needed for the model. In addition, a set of demographic variables is included to add depth to the modelling effort. These variables are listed below in Table II and were selected with the assistance of USAREC analysts.

Table II SELECTION OF VARIABLES

- 1. RACE
- 2. MISSION BOX
- 3. CHILDREN
- 4. MARITAL STATUS
- 5. RECRUITING BRIGADE

Mission box contains 22 different categories and therefore as many levels for our model. Rather than including mission

¹Based on phone conversations with USAREC analysts in November 1993.

boxes an exploratory variable, our approach was to fit a single model to the GMA mission box. If all 22 levels are included in the model, the resulting model would require many estimated parameters and could prove difficult to fit. Our study therefore concentrates on GMA's.

When analyzing the data we note that GMA recruits can not be combined with other mission boxes because of the significant differences in the proportion of DEP losses between them. For instance, from Figure 2, it is clear that the proportion of DEP losses for the GFA mission box is consistently higher than for the GMA.

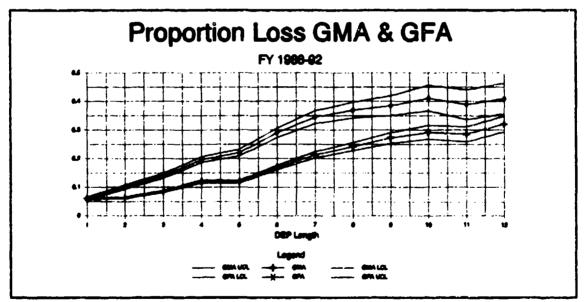


Figure 2 PROP LOSSES GMA & GFA WITH 95% CONF INTERVAL

Figures 3 and 4 graph the proportion of losses as a function of DEP length for the six recruited mission boxes.

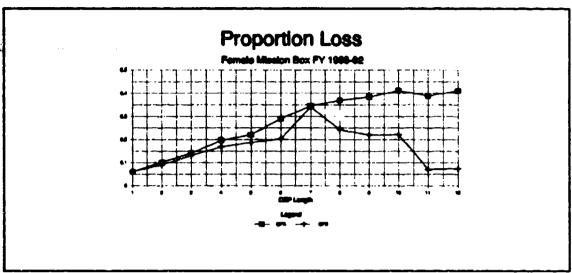


Figure 3 PROP LOSSES FOR SOME FEMALE MISSION BOXES

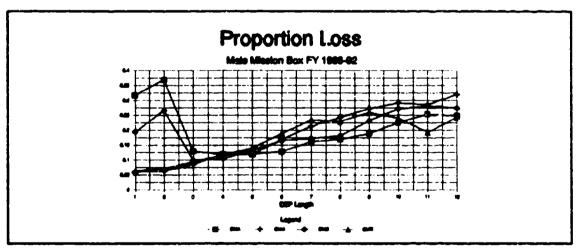


Figure 4 PROP LOSSES FOR SOME MALE MISSION BOXES

At present USAREC targets those individuals that are high school seniors or graduates and score above the 50th percentile on the AFQT. During FY 1992, nearly 37% of all candidates were GMA, the most of any recruit mission box. This is above historical percentages for all fiscal years except FY 1991 where 43 percent of recruits fell into this category. Figure 5 gives a historical summary of the percentages of

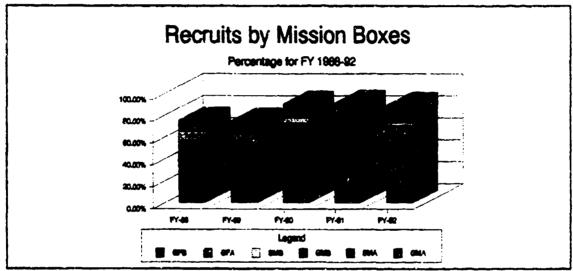


Figure 5 RECRUITS BY MISSION BOXES

recruits for 6 mission boxes for FY 1988-92 (These account for over 80% of all recruits).

C. INFLUENCE OF VARIABLES

Many of the studies centered on DEP loss examine individual and system characteristics believed to be associated with DEP loss [Ref. 2-5]. Cross-tabular analyses has been used to examine the relationship between many of

these characteristics. The studies have reported statistically significant differences between loss rates associated with these characteristics. Our attempt is to view the influences of the variables to be used in the model from a practical and statistical point of view.

Preliminary analysis of the variables considered in our study was conducted by obtaining the proportion of accessions for GMA's by recruiting brigade', race and whether or not the candidate had one or more children. Formal hypothesis tests performed for differences in the proportion of accesses between groups result in significant differences. For example, from Appendix C we can obtain the necessary data to test for differences in the proportion of DEP accessions between blacks recruited from brigade I and III. This results in a test statistic of -3.03 and an associated significance level of 0.0012. If we let P₁ and P₂ correspond respectively to the proportion of black recruits reported as accessions from brigade I and III, then the hypotheses tested are:

 $H_0: P_1 = P_2$

 $H_a: P_1 \neq P_2$

with test statistic:

Presently USAREC has four recruiting brigades down from six. Brigades II and V have been consolidated into the other four. Candidates recruited from these brigades II and V were considered part of the other four for this study. USAREC analyst appropriately accounted for this when providing the data set.

$$Z = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\frac{\bar{p}_1(1 - \bar{p}_1)}{n_1} + \frac{\bar{p}_2(1 - \hat{p}_2)}{n_2}}} , \qquad (1)$$

where \hat{p}_1 and \hat{p}_2 are the estimated proportion of accessions and n_1 and n_2 are the sample sizes for the two brigades. In Table III we summarize the hypothesis tests for black recruits between recruiting brigades. All tests are significant at the 5% level of significance except for proportions between brigade I and IV.

The fact that the significance level is small is due in part to the large number of observations. Tests performed on data with very large sample sizes often result in statistical significance even when there is no practical difference between the magnitudes of the proportions. Figures 6-9 are

Table III HYPOTHESIS TEST FOR BLACK RECRUITS

BRIG	POP.	PROPOR	RTION	SAMPLE	E SIZE	TEST	
1	2	p ₁ p ₂		n,	n,	Z	p-value
1	III	.876	.892	6648	7725	-3.03	.00120
I	IV	.876	.874	6648	3842	0.283	.61141
I	VI	.876	.835	6648	1624	4.009	.00005
III	IV	.892	.874	7725	3842	2.820	.00240
III	VI	.892	.835	7725	1624	5.732	0
IV	VI	.874	.835	3842	1624	3.606	.00015

provided summarizing the results of the accession proportions. Each figure separates by brigade the accession proportions for all candidates and those with and without children. From these figures we can see there is a noticeable difference between the magnitude of the proportions. Thus, this motivates the inclusion of race, brigade, children in the model to estimate DEP accessions.

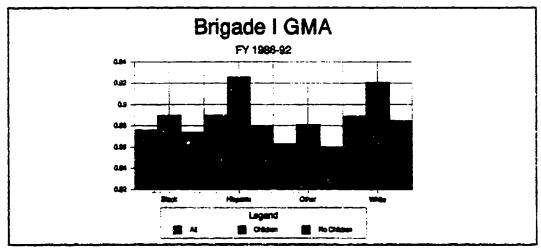


Figure 6 PROPORTION OF GMA ACCESSIONS FOR BRIGADE I

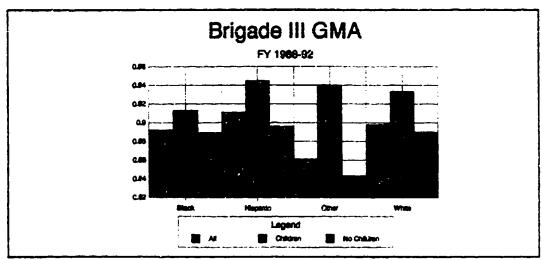


Figure 7 PROPORTION OF GMA ACCESSIONS FOR BRIGADE III

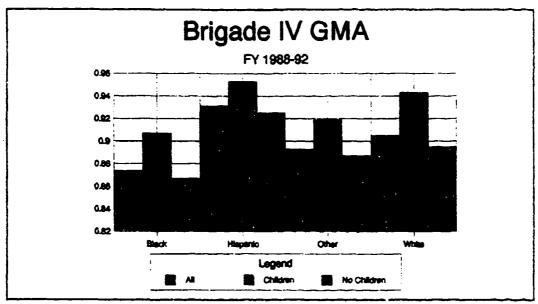


Figure 8 PROPORTION OF GMA ACCESSIONS FOR BRIGADE IV

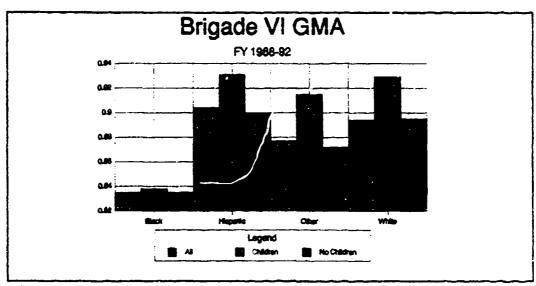


Figure 9 PROPORTION OF GMA ACCESSIONS FOR BRIGADE VI

D. EXPLANATORY ANALYSIS

When examining how losses occur a function of contract length and time in DEP, we note that most of the losses occur late in DEP. The last months of DEP exhibit most of the losses while all other months account for relatively few. This pattern of losses is true for FY 1988 through 1992 and may be indicative of either recruits in DEP waiting to renege their contract or recruiters not reporting an individual a loss until the last minute. Table IV shows how DEP losses for all mission boxes during FY 1992 are distributed over the contract lengths for a typical year. Similar loss patterns can be seen in Appendix A for other years. The numbers contained within the table correspond to the percentage of losses incurred per month of DEP. For a one month DEP contract all losses occur in month one. From the table we see for a two month DEP, 38% of losses occur in the first month while the remainder are realized in the last month.

Table IV DEP LOSS PERCENTAGES FY 1992 - ALL MISSION BOXES

	LOSS MONTH												
		1	2	3	4	5	6	7	8	9	10	11	12
	1	100											
С	2	38	62										
0	3	21	18	60									
N	4	17	4	21	58								
T	5	17	2	6	21	54							
R	6	16	1	2	6	21	54						
λ	7	14	1	3	3	6	26	47					
С	8	8	1	2	2	2	7	25	53				
T	9	5	0	1	1	2	6	9	26	50			
	10	3	0	0	1	1	4	7	10	30	45		
	11	3	0	0	1	1	1	4	5	8	28	50	
	12	2	1	0	0	0	0	3	5	2	10	28	49

The fact that losses occur either at the very beginning or toward the end of DEP suggest that several months could be combined together. This should result in a more parsimonious model requiring fewer estimated parameters. This notion is explored in the next section. The loglinear model discussed in the next section is used to investigate the viability of a combining months together. A family of models called generalized linear models, of which the loglinear and logit models are members, is offered next along with a discussion of the loglinear model we use in our exploratory analysis.

R. GENERALIZED LINEAR MODELS

The classical linear model is of the form

$$E(Y) = \alpha + \beta^2 x \quad , \tag{2}$$

where Y is the response variable, α is the intercept parameter, B is a column vector of unknown coefficients and $\mathbf{x}=(\mathbf{x}_1,\ldots,\mathbf{x}_n)^T$ is a vector of m explanatory variables. The response variable is assumed to have a Normal distribution with constant variance. Although this approach is fairly robust with respect to the normality assumption, there are many instances where it is not appropriate. For example, if Y is a binary variable, equation (1) can lead to an expected value $\mathbf{E}(\mathbf{Y})$ that is not between 0 and 1. Also, the variance of binary responses is not constant, but changes with the expected value of the response.

Generalized linear models (GLM) effectively address these issues. They do this by reparametertizing the model so that the expected response is a linear function of the coefficients. These models contain three parts:

1. A random component, associated with the distribution of the response variable (e.g. Binomial, Poisson, Normal). This gives the basis for the variance function used to describe how the variance change with the mean.

- 2. The systematic component, B'x which gives the linear function of the explanatory variables used for prediction.
- 3. The link component g(E(Y)) which describes how the mean depends on the linear predictors (for Normally distributed responses the link function g(.) is the identity function).

The power of the GLM approach is that when the distribution of Y belongs to an exponential family and the link function g(.) is chosen appropriately (i.e. g(.) is the canonical link), the Maximum Likelihood Estimators (MLE) of the parameters can be approximated using an iterative least squares algorithm. For large samples, inference for parameters of GLM's is much the same as those in the usual linear model setting.

In logistic regression, the response variable Y is binary. Let $\pi(x)$ represent the conditional expectation of Y given a set of explanatory variables $x = (x_1, ..., x_j)$; the canonical link function in this setting is the logit function

$$\ln\left[\frac{\pi(x)}{1-\pi(x)}\right] = \alpha + \beta^2 x \quad , \tag{3}$$

where $\beta^{\tau} = (\beta_1, \dots, \beta_j)$ is the coefficient vector and α is the constant intercept. Equivalently,

$$\pi(x) = \frac{e^{\alpha \cdot \beta^T x}}{1 + e^{\alpha \cdot \beta^T x}} \quad . \tag{4}$$

Equation (4) guarantees that the probability $\pi(x)$ will be between 0 and 1.

When the response variable is Poisson, such as the number of counts falling into a particular cell of contingency table, then the canonical link is

$$\ln E(Y) = \alpha + \beta^T x \quad , \tag{5}$$

giving us loglinear models.

1. Loglinear model

Preliminary analysis included fitting a loglinear model. We consider a candidates race (Black, Hispanic, Other, White), recruiting brigade (I,III,IV,VI), Marital status (Single, Married), DEP length (1,...,12), DEP loss month (1,...,12) or accession in the model. The response variable for a loglinear model is the number of observations that fall into a particular cell or a contingency table defined by the variables considered.

By fitting such a model we are able to gain insight into advantageous ways to combine DEP length with DEP loss month. We defined one variable, DEP Outcome, with 90 levels and thereby consolidate two variables into one. For example, a

three month DEP has a possible outcome of a loss in first, second or third month or accession (levels 6 - 9, Table V).

One can see that this arrangement results in 90 possible outcomes - 78 for losses and 12 accessions.

Table V DEP OUTCOME MATRIX (90 LEVELS)

						1	OFFE O	r der i	LOSS					
	_	1	2	3	4	5	6	7	8	9	10	11	12	A
	1	1												2
D	2	3	4											5
t	3	6	7	8										9
?	4	10	11	12	13									14
	5	15	15	16	17	18								19
L	6	21	22	23	24	25	26							27
1	7	28	29	30	31	32	33	34						35
ı	8	3′.	37	38	39	40	41	42	43					44
Ç	9	45	46	47	48	49	50	51	52	53				54
Y	10	55	56	57	58	59	60	61	62	63	64			65
1	11	66	67	68	69	70	71	72	73	74	75	76		77
	12	78	79	80	81	82	83_	84	85	86	87	88	89_	90

NOTE: A-Accession.

The model initially fit using all 90 levels proved to be numerically unstable. This can be attributed to the sparse nature of the contingency table with relatively few counts in the "middle" months of DEP.

Several combining strategies were attempted to reduce the number of levels in DEP Outcome by joining together adjacent

Table VI COMBINED DEP OUTCOME MATRIX (37 LEVELS)

			MONTE OF DEP LOSS											
L		1	2	3	4	5	6	1	8	9	10	11	12	1
	1	1		·•							_			2
D	2	3	4											5
1	3	6	7	8		-								9
,	4	10	11	12	13		-							14
	5	15	15	16	17	18								19
L	6	21	22	23	24	25	26		_					27
I	7	28	29	30	31	32	33	34						35
,	8	36	37	38	39	40	41	42	43		_			44
ç	,	45	46	47	48	49	50	5 1	52	53		_		54
7	10	55	56	57	58	59	60	61	62	63	64		_	65
1	11	66	67	68	69	70	71	72	73	74	75	76	1	11
	12	78	79	80	81	82	83	84	85	86	87	88	89	90

NOTE: L-loss; A-Accession.

cells (months) to form a composite level. Table VI indicates composite successful combining strategy. The table outlines each composite level by dashed lines. By combining the levels representing losses in the middle DEP months (that had few counts), we reduce the number of estimated parameters, bring more stability to the model and do not sacrifice model fit. In doing so we reduce the number of levels from 90 to 37. The

hypothesis test between the null model with all 90 levels and the alternative model combining levels as indicated in table VI produced a χ^2 statistic of 4420 with 4770 degrees of freedom. The associated p-value of .9998 indicates no practical difference between the fit of the model using no combining and the model combined as in Table VI. Attempts to further combine levels give models which do not fit the data as well.

A first look at the data shows which explanatory variables can not be ignored when estimating the probabilities of DEP loss. An important step to finding estimates of the DEP survival (accession) is to find the most parsimonious model for estimating these probabilities. Without it the number of parameters in any reasonably well fitting model becomes unmanageable. We do this in two ways. We restrict attention to the most heavily recruited group (GMA) and by using a loglinear model, we justify combined levels of a variable that takes into account DEP length and number of months served in DEP. This provides the basis for developing the final model used for estimating the conditional probabilities in the next Chapter.

III. ESTIMATING CONDITIONAL PROBABILITIES

In this chapter, we estimate the conditional probability of accession given an individual has survived DEP for a certain period of time. The loglinear regression approach of the previous section gives estimates of the expected number of individuals which fall into each cell of a contingency table defined by the explanatory variables. From the loglinear model, all conditional probabilities for a set of explanatory variables could be estimated. The difficulty with using this model is that it gives the joint probability mass function of all the variables. Because we seek the conditional probabilities, the loglinear approach forces us into estimating too many parameters.

An alternative approach is to use a GLM to estimate the conditional probabilities directly. If the response variable is strictly binary, where an individual is an accession or loss, then a standard logistic regression model would apply. However, in this case the response variable Y has several levels that indicate when the individual is lost or survives DEP and becomes an accession. In addition, this variable has a sequential order to it and is therefore ordinal in nature. When the levels of an ordinal response variable are stages or time frames through which an individual, item or process may progress, such as a fixed series of attempts seeking a

success, a technique called continuation-ratio modeling is used [Ref. 8].

A. CONTINUATION-RATIO MODEL

In analyzing DEP survival, we use the case of a four month contract to illustrate the continuation-ratio model. We begin with a fixed number of individuals. Some of these individuals will drop out during the first, second, third or fourth month while others will access. For each individual, the response variable Y falls into one of five levels where Y = 1,...,4 represents a loss in month 1, month 2, month 3 and month 4 respectively. We use Y = 5 to represent an accession. The continuation-ratio model uses "cumulative" or "accumulated" logits [Ref. 9] that take into account category order. Because we are ultimately interested in the probability of a recruit accessing given he survives in DEP for a number of months, we can use

$$L_{j} = Logit[Px(Y)j \mid Y \ge j)] \qquad for j = 1, ..., 4 . \qquad (6)$$

For instance, in the context of our example, L_2 this the logit of probability of surviving past the second month given survival of two or more months. An alternative expression for L_1 is

$$L_{j} = -\ln \left[\frac{Px(Y > j)}{1 - Px(Y \le j)} \right] . \tag{7}$$

The continuation-ratio model is similar to Cox's proportional hazard model and approaches it as the number of stages increases [Ref. 8].

In the continuation-ratio model explanatory variables are incorporated by letting

$$L_1 = \alpha + \beta^2 X \quad , \tag{8}$$

where α is the intercept, β^{r} is the vector of unknown coefficients and X is a vector of explanatory variables. Using the continuation-ratio model we estimate the conditional probability $P(Y>j|Y\geq j)$ for $j=1,\ldots,4$. We can now estimate the probability of accession given the recruit survives for a number of months in DEP (i.e. $P(Y>4|Y\geq j)$ for $j=1,\ldots,4$).

B. THE MODEL CONSTRUCTION PROCESS

S-PLUS version 3.1, by Statistical Science, Incorporated was used on a Hewlett Packard 730 work station to fit the model. The exploratory variables are the same as those previously used for the loglinear model discussed in Chapter II. In addition, we establish an ordinal variable, Zone, that takes into account the strategy for combining months successfully exploited for the loglinear model. Table VII itemizes these and their associated levels.

Table VII LIST OF CONTINUATION-RATIO MODEL VARIABLES

EXPLANATORY VARIABLE	SYMBOL	LEVELS	TYPE
DEP CONTRACT	D	1,2,,12	ORDINAL
ZONE	z	1,,5	ORDINAL
Marital Status	М	MARRIED, SINGLE (1,2)	CATEGORICAL
RACE	R	BLACK, HISPANIC, OTHER, WHITE (1,,4)	CATEGORICAL
CHILDREN	С	YES, NO (1,2)	CATEGORICAL
BRIGADE	В	I, III, IV, VI (1,,4)	CATEGORICAL

Zone represents the stages through which a recruit must survive to become an accession. The number of zones vary with the length of contract. For all DEP lengths, reaching zone 5 represents an accession while zone 1 represents the first month of contract. A one month DEP has two levels; zone 1 for a loss and zone 5 for an accession (Figure 10). Three levels exist for a two month contract where losses can occur in either the first or second month represented by zone 1 and 2. Similarly, an accession is labeled as zone 5. A three month contract accordingly calls for 4 levels with losses possible in zones 1 through 3. Finally, for DEP lengths of 4 or more months 5 levels are needed. Zone 2 for DEP lengths greater than 4 months is a composite of all "middle" months. The

extent to which these "middle" months are combined into zone 2 will vary according to the length of DEP. Zone 2 will consist of all months but the first, last and next to last for

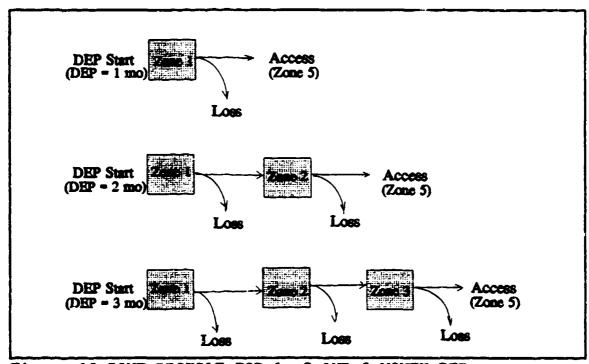


Figure 10 ZONE PROFILE FOR 1, 2 AND 3 MONTH DEP

contracts of 4 or more months. Figure 11 summarizes the zones for contracts greater than 3 months.

Combining months into zones as motivated by our loglinear model reduces the number of cells in the contingency table defined by the variables to 2688. Accordingly, the estimated parameters for the model are reduced.

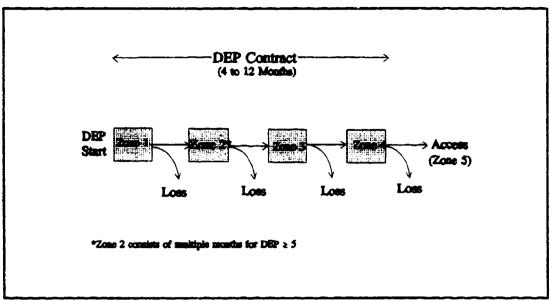


Figure 11 ZONE PROFILE FOR DEP ≥ 4 MONTHS

C. FITTING

S-PLUS glm function with a binomial link component was used to run the model. Recall, for GLM models, the link component (p.19) describes how the mean of the response variable depends on the linear predictors. Since we are using a logit model the binomial link is appropriate.

Our assessment of model adequacy was centered on the deviance explained by the model. Based on the outcome of the goodness-of-fit test terms were added or removed while keeping with the discipline of hierarchical models. Our model was kept hierarchical, where all lower-order terms are included if they involve those in higher-order. Because our model consists of no more than two-way interactions between

exploratory variables, keeping a model hierarchical requires the main effects to be kept for any two way interactions in which they appear.

In linear regression, summary measures of the distance between the observed and fitted values are a function of the residual definer as the difference between the fitted and observed value. In logistic regression, there are several possible measures of the fitted and observed values. The residual deviance D is one of these measures and is a most reliable and routinely used measure. The asviance is -2 times the difference of the log-likelihood of the fitted and saturated models. Under the null hypothesis that a model fits and when the sample size is large, D has an approximately χ^2 distribution with D degrees of freedom equal to the number of observations N less the number of parameters p. Therefore, if the model fit is good, asymptotically the expected deviance is N-p.

In logistic models with many $\pi(\mathbf{x}_1)$'s very close to zero or, as in our case, close to one, the distribution of the deviance is not close to being χ^2 with N-p degrees of freedom. Instead D tends to be much smaller than N-p. Thus, when using D our overall test of model fit results in very large p-values whether the models fit or not. The asymptotic distributions do hold for testing between two nested models. Specifically, the test statistic of H_0 : model 1 versus H_1 : model 2 is D_1-D_2 which is asymptotically χ^2 with $\nu_1-\nu_2$ degrees of freedom.

Here D_1 and D_2 are the total deviance of model 1 and 2 while v_1 and v_2 correspond to each of the model's residual degree of freedom.

In linear regression, a way to measure the contribution of an independent variable toward predicting a dependant variable is by calculating the coefficient of determination (R²). This coefficient takes values between 0 and 1 with a high value suggesting a good fit. Because logit models have binary dependent variables, they have low R² values regardless of how well they fit [Ref. 10]. A method analogous to the coefficient of determination that avoids the problem with binary dependent variables is the likelihood ratio index [Ref 10]. This index provides a measure of the ratio of explained deviation in the model and is defined as:

$$\rho = 1 - \frac{Explained Deviance}{Total Deviance} . (9)$$

Where the total deviance is the deviance from fitting the logit model with only a constant term (no explanatory variables).

D. FINAL MODEL

In the final model, the logit of the probability of surviving zone j given survival up to zone j can be parameterized as

$$L_{j} = \alpha + \beta_{j}^{g} + \beta_{j}^{D} + \beta_{k}^{H} + \beta_{j}^{H} + \beta_{m}^{B} + \beta_{jj}^{DB}. \tag{10}$$

Here β_j^2 j = 1,...,4 are the parameters corresponding to zones 1 through 4, β_i^D for i = 1,...,12 are the parameters for DEP of 1 to 12 months, β_k^M k = 1,2 are the parameters corresponding to the marital status married and single respectively, β_i^B l = 1,...,4 are the parameters corresponding to the respective races, black, hispanic, other and white and β_k^B m = 1,...,4 are the parameters corresponding to the brigades I, III, IV, VI respectively. Finally, the two - way interaction parameters between zone and DEP length are β_{ij}^{DZ} for i = 1,...,12; j = 1,...,4. This model is simple and fits the data well by explaining 87% of the total deviance (Table VIII).

Table VIII REVIEW OF MODELS

MODEL	RESIDUAL DEVIANCES	RESIDUAL DF	Q
Null	16,616	2,687	0
Main Effects	3,148	2,114	.82
Final	2,053	2,088	.87
Main Effects with Two Way Interactions	1,654	1,953	.99

A model using all main effects and two-way interactions provides a nearly perfect fit. However, our aim was to obtain the best fit with the fewest possible parameters. To this end, the model was simplified to include only DEP length and zone as a two-way interaction term. The analysis of deviance given Table IX provides a summary of the deviances for the final fitted model. In it is summarized the contribution each or the terms makes toward the final model fit. The 87% of the total deviance explained and a χ^2 goodness-of-fit significance level of .6814 associated with 2088 total deviance and 2057 degrees of freedom provides an adequate fit to calculate the probabilities we need.

Table IX ANALYSIS OF DEVIANCE FOR FITTED MODEL

TERMS	DF	DEV	RESID DF	RESID DEV	ρ	SIGNIF
NULL			2687	16616	0	.000
DEP LENGTH	562	1844	2125	14771	.11	.000
ZONE	3	11114	2122	3658	.77	.000
MARRIED	1	360	2121	3298	.80	.000
RACE	3	124	2118	3174	.81	.000
BRIGADE	3	16	2115	3158	.82	.001
øp Length : Zone	27	1101	2088	2057	.87	.000

Plots of residual deviance are often used in regression modelling to check a model's adequacy. We note that with the large number of cases where $\pi(\mathbf{x}_i)$ is nearly one, plots of standardized residuals to investigate model adequacy are of small help. In fact, all these plots reveal, is for the given data the estimated proportions are near one. For this reason we do not use plots of standardized residuals here.

In this chapter we use the continuation-ratio model to estimate conditional probabilities directly. The fitted values allow us to directly calculate probabilities we ultimately desire. The final model is simple and explains most of the total deviance.

IV. MODEL RESULTS

The fitted continuation-ratio model for GMA's provide the values needed to estimate the expected probability of accessions given a certain survival has occurred. To illustrate, take the case of fitted values obtained by the model for MARRIED, BLACK and BRIGADE I (Table X). The fitted values indicate a single, black male recruited from brigade I has a 95% chance of surviving the first month and thereby

Table X P(x>zone i|x≥zone i) FITTED VALUES FOR BLACK, MARRIED, BRIGADE I

Condi	Conditional Probability of Surviving Past Zone (1)													
	ZONE													
1		1	2	3	4									
•	1	.9502												
	2	.9782	.9607											
	3	.9826	.9903	.9456										
D	4	.9832	.9964	.9822	.9246									
E	5	.9861	.9948	.9817	.9183									
P	6	.9897	.9883	.9727	.8782									
	7	.988	.9853	.9659	.8434									
•	8	.9923	.9719	.9593	.8289									
	9	.9932	.9694	.9471	.8029									
i	10	.995	.9612	.9471	.7881									
	11	.9946	.9689	.9528	.7756									
	12	.9942	.9516	.9531	.8102									

accessing into the Army under a one month DEP contract. A similar recruit with a DEP of 10 months has a 99.5% chance of surviving past the initial month of DEP, a 96.1% chance of surviving past months 2 through 8 (zone 2) and a 94.7% chance of making it past the next to last month of the contract. Meanwhile, there exist a 78.8% chance of making it past the last month and thereby accessing. A complete set of fitted probabilities is included in Appendix E.

A. CONDITIONAL PROBABILITIES

The probability an individual X survives past zone 4,

can be expressed as a function of conditional probabilities of surviving previous zones. In other words, the probability of surviving past zone 4 and becoming an accession is the probability of surviving past zone 4 given the individual survived past zone 3 multiplied by the probability of surviving past zone 3. Equation 11 is equivalent to

$$Pr(X > Some 4|X > Some 3) Pr(X > Some 3)$$
 . (12)

In general terms for i=2,...,4

$$P(X > sone i) = Pr(X > sone i|X > sone i-1) Pr(X > sone i-1)$$
(13)

From Table XI we can obtain the probability of accession (surviving zone 4) given an individual has survived past some time. For example, the probability of an individual accessing given he has survived past zone 2 is:

Pr(X > sone 4 | X > sone 3)Pr(X > sone 3 | X > sone 2).

XI the estimated probability Thus from Table 0.74641=(0.7881)(0.9471). Therefore, a married, black male, recruited from brigade I under a 10 month contract who has survived past zone 2 is expected to access 74.6% of the time. Table XI contains the estimated probability of surviving DEP conditioned on time survived. The value of .7468 calculated above is accordingly listed with a 10 month contract and under zone 3. Notice in Table XI all numbers increase from left to right. This should be the case since the more a recruit survives the more apt he is to access. A complete listing of all fitted values and their appropriate conditional probability is provided in Appendix E.

Table XI CONDITIONAL PROBABILITIES FOR TABLE XI

A STATE OF THE STA

Prob. c	Access	ion Given	Survived	Zone (i-	1)										
		ZONE													
** 1s		1	2	3	4										
·	1	.9502													
	2	.93976	.9607												
	3	.92013	.93643	.9456											
ם	4	.88967	.90487	.90814	.9246										
E	5	.88434	.89681	.9015	.9183										
P	6	.83554	.84423	.85423	.8782										
	7	.79303	.80266	.81464	.8434										
	8	.76687	.77282	.79516	. 8289										
	9	.73214	.73716	.76043	.8029										
	10	.71386	.71745	.74641	.7881										
	11	.71214	.71601	.73899	.7756										
	12	.73057	.73483	.77220	.8102										

B. USING PROBABILITIES TO ESTIMATE ACCESSIONS

USAREC planners may use the conditional probabilities to estimate accessions at some future date. By multiplying the appropriate conditional probability by the number of recruits in a given category a point estimate of the number of anticipated accessions is obtained. To illustrate, we borrow

from table XI discussed in the previous section and use a simple example. If we have 100 single, black males recruited from brigade I. Of the 100 recruits, 50 have signed 6 month contracts and the rest are under 9 month DEP. Assume that of the 50 recruits in each contract group, 13 are in zone 1 and 13 are in zone 4. Of the remainder, 12 are in zone 2 and 12 in zone 3. To estimate the number accessing at some future date we multiply the number of individuals in DEP by their associated conditional probability. Table XII provides the results of our example. In similar fashion, if we desired to

Table XII ESTIMATING ACCESSIONS

Cont Length	Prob	Number in DEP	Product	Expected accessions
6	.8355	13	10.68	11 in 6 months
6	.8442	12	10.13	10 in 2 to 4 months
6	.8542	12	10.25	10 next month
6	.8782	13	11.41	11 this month
9	.7321	13	9.53	10 in 9 months
9	.7371	12	8.846	9 in 2 to 7 months
9	.7604	12	9.125	9 next month
9	.8029	13	10.4838	10 this month

estimate the number of accessions in one month, we would multiply the total number of recruits with a projected accession date for next month by the appropriate last zone conditional probability. Thus, next months estimates are obtained.

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V. FINAL SUMMARY

The unique structure of DEP losses necessitates exploring avenues to address the modelling challenge associated with sparse contingency table. The use of the loglinear model to explore viable strategies for combining levels and thereby reducing the size of the contingency table (increasing the cell counts) is valuable toward this end. We note that the sparsity of the contingency table may be largely due to the inherent nature of the DEP loss reporting process or to a recruit's hesitation to renege on his contract early on. Recruiters are reluctant to report a candidate as a loss until no doubt exists of the individuals intention to renege on the contract. Recruits, on the other hand, may feel compelled to rescind only at the very end of the DEP. Despite this, a loss can only be realized when it is reported. Without major policy changes the actual loss date may never be available.

A. CONCLUSIONS AND RECOMMENDATIONS

The model building effort provides a methodology to estimate probabilities of accession based on time survived in DEP. The probabilities can provide USAREC planners a quantitative basis for accession projections for GMA recruits. By entering the model's fitted values into a commercial spreadsheet or database USAREC analysts can quickly estimate

recruit accession numbers at some future date based on the present DEP inventory. By applying this modelling methodology to the other recruit categories (mission boxes) a complete analysis tool can be obtained. In doing so, a means to estimate recruit accessions may be realized which can ultimately assist in setting recruiter goals.

							obabili	·						└
		mo1	mo2	mo3	mo4	mo5	mo6	mo7	mo8	mo9	mo10	mo11	mo12	acc
	mo1	0.059	L			<u> </u>	Ĺ							0.94
	mo2	0.042	0.058										L	0
<u>c_</u>	ma3	0.027	0.028	0.061									<u>.</u>	0.88
0	mo4		0.006						l					0.86
	mo5		0.003											0.8
t	mo6	0.009	9E-04	0.002	0.008	0.035	0.123							0.8
r	mo7	0.008	0.001	0.002	0.003	0.008	0.039	0.123						0.8
A	moß	0.01	5E-04	0.001	0.001	0.004	0.007	0.032	0.125					0.
c	mo9	0.009	0.001	0.001	0.001	0.001	0.003	0.009	0.04	0.128		İ		0.8
t	mo10	0.009	3E-04	5E-04	0.001	0.001	0.002	0.007	0.012	0.042	0.153			0.7
	mo11	0.009	0.002	2E-04	0	1E-03	7E-04	0.002	0.005	0.011	0.041	0.179		0.
	mo12	0.01	9E-04	5E-04	3E-04	5E-04	5E-04	0.002	0.002	0.004	0.01	0.047	0.195	0.7
	 	 								ļ				
			L <u></u>	L		Perce	at of La	es YR	1988				-	-
		mo1	mo2	mo3	mo4	mo5	mo6	mo7	mo8	mo9	mo10	mo11	mo12	
	mo1	100%												
	mo2	42%	58%											
$\widehat{\mathbf{C}}$	mo3	23%	24%	53%										
0	mo4	17%	4%	20%	58%									
p	mo5	11%	2%	3%	19%	65%								
t	mo6	5%	1%	1%	5%	19%	69%							
r	mo7	4%	1%	1%	1%	5%	21%	67%						
	mo8	6%	0%	1%	1%	2%	4%	18%	69%					
c	mo9	4%	1%	1%	1%	1%	1%	4%	21%	66%				
t	mo10	4%	0%	0%	0%	1%	1%	3%	5%	19%	67%			
	mo11	4%	1%	0%	0%	0%	0%	1%	2%	5%	16%	71%		
	mo12	4%	0%	0%	0%	0%	0%	1%	1%	1%	4%	17%	71%	
	-	L	Prot	DEP	Loss in	month	i Sur	vived t	o mont	h i - 1)	for Y	R 1988		
	1	mo1	mo2	mo3	mo4	mo5	mo6	mo7	mo8			mo11	mo12	
	mo1		1						ļ				1	
	mo2	0.061				 								
$\overline{\mathbf{c}}$	mo3		0.064						 					
•	mo4		0.098						<u> </u>			 	1	
 B	mo5	0.132		0.125		<u> </u>			 					Τ-
<u>. </u>	mo6	0.17		0.168					 		 		1	1
<u>. </u>	mu7						0.131	 	1			 	 	
	mo8		0.171			0.166		0.132	 	 	 	 	1	
-	mo9								0.137		 	 		+
-	mo10								0.202		 			
`	mo11								0.235			1	 -	+
		U.270			0.264	L V.E-1	7.27	1 4.20	1 4.500		U. 102	1		1

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		T				Ī.	oss Pro	babilir	v YR 1	989					
	 	 	mo1	mo2	fiio3	mo4					mo9	mo10	mo11	mo12	acc
		mo1	0.073									-			0.927
. godržana		mo2	0.029	0.036								ı			∩ 935
	C	mo3	0.021	0.007	U.059										0.913
	0	mo4	0.018	0.003	0.014	0.088									0.878
	0	เกอร์		0.001											0.853
	t	mo6		9 04											0.821
· ·	r	mo7		(.01								Ĺ			0.822
	8	mo8		7E-04											0.8
* •*	c	mo9								0.022					0.726
	<u>t</u>	mo10								0.026					0.726
	<u></u>	mo11		2E-04						0.073					0.70
	<u></u>	mo12	0.005	0.001	0.001	0.001	0.002	0.002	0.007	0.02	<i>ს</i> .063	0.094	0.031	0.087	0.685
	<u> </u>	-	<u> </u>			L	<u></u>						ļ		
	<u> </u>	 	1				ercent			,		- 10		- 40	
			mo1		mo3	mo4	mo5	mo6	mo7	mo8	mo9	molO	m:011	mo12	
	<u> </u>	mo1	100%						<u> </u>	·					
	_	mo2	45%		609/				 			⊢—	├ -		
	<u>C</u> _	mo3	24%		68%	72%	 		ļ			<u> </u>	ļ		
	0	mo4	11%		11% 3%	15%			 -			 			
	-	mo5 mo6	6%		2%	7%		e:/%							
	<u> </u>	mo7	8%		0%	7%		14%				 	 		
	-	mo8	4%		1%	12%		9%	13%			1	1	 	
	-	mo9	2%		1%	6%		21%	10%		34%	 	 		
	1	mo10	3%	0%	0%			14%	22%		_		}		
	<u> </u>	mo11	2%		0%	0%		4%	10%		19%				
	-	mo12	2%		0%			1%						28%	
-		1						.,,		0.0		00.0	1070	- 22.75	
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			mo1	mo2	mo3	mo4	mo5	mo6	mo7	то8	mo9	mo10	mo11	mo12	i
		mo1			ļ										
		mo2	0.037												
	C	mo3		0.061											
	0	mo4		0.092											
		mo5		0.131											
	1	mo6		0.169											
	r	mo7		0.165						↓ — —					
	8	mo8		0.193							l		L		
	<u>c</u>	mo9								0.115					
	t	mo10								0.133					
		mo11								0.194					
		mo12	0.31	0.31	0.300	0.208	0.307	0.205	0.301	0 286	0 238	0 147	0 113		i

L									ity YR	1990		<u> </u>	<u> </u>	
		mo1		mo3	mo4	mus	mo6	mo7	mo8	mo9	mo10	mo11	mo12	acc
	mo1	0.049	I											0.9
	mo2	0.022	0.05											0.9
C	mo3	0.017	0.019	0.073							}			0.8
0	mo4	0.015	0.008	0.038	0.082						1	 -		8.0
	mo5	0.016	0.007	0.011	0.025	0.106						[0.8
t	mo6	0.016	0.017	0.013	0.008	0.026	0.122							0.7
r	mo7	0.01	0.014	0.021	0.006	0.004	0.034	0.127						0.7
2	mo8	0.006	0.002	0.025	0.01	0.002	0.007	0.032	0.123					0.7
c	mo9	0.005	7E-04	0.004	0.004	0.002	0.009	0.016	0.057	0.134				0.7
t	mo10	C.003	8E-04	4E-04	4E-04	0.002	0.005	0.014	0.022	0.053	0.148			0.7
	mo11	0.004	4E-04	0	4E-04	0	0.001	0.009	0.011	0.021	0.064	0.156		0.7
	mo12		6E-04						0.003					0
_									1				-	
	1		·	<u> </u>	P	ercent	of Loss	YR 1	990		<u> </u>			
_		mo1	mo2	mo3	mo4			mo7	mo8	mo9	mo10	mo11	mo12	\vdash
	mo1	100%		-	1	_					1000		1	
	mo2	31%	69%	 		 		 	 		ļ		 	\vdash
C	mo3	16%		67%	 					-	 -	 	 	-
6	mo4	10%		27%		 	-		 				 	\vdash
	mos	10%	5%	6%	15%		 		ļ <u> </u>		 			-
-	mo6	8%		6%	4%	13%		-	 	 			 	-
-	mo7	4%		10%	3%	2%		59%	 			 	 	-
: -	mo8	3%		12%	5%			15%		 	 		 	
-	mo9	2%		2%		1%		7%			├		 	
-	mo10	1%				1%		5%		21%				-
H	mo11	1%		0%		0%		3%		8%		58%	 	
	mo12	1%						0%						⊢
-	moiz	1.70	078	0%	076	070	0%	070	170	376	1170	2376	30%	├
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-		<u> </u>	Prob	/ DFP	I oce in	month	i i Sme	vived t	o mont	h i - 1)	for VI	2 1000		\vdash
\vdash	- 	mo1	mo2		mo4		mo6		mo8	mo9		mo11		├-
-	mo1	1	11102	11.00	1110-1				11100	11100	111010		111012	├
-	mo2	0.051	 		 	 	├	 	 -	 		 -		├~
C	mo3		0.076	 	 	 	 	┼─-~	 	 		 	 	-
-	mo3		0.078				┼──	 -	 					-
<u> -</u>			0.107			 		 	 	 	 	 	 	
-	mo5						 -		 		-		{	
<u>-</u>	mo6		0.175					 	 	 	-	 	 	<u> </u>
-	mo7		0.196						 			 		-
2	mo8		0.201				0.163				<u> </u>		 	<u> </u>
c	mo9								0.149		-	ļ	 	
11-	mo10								0.211					<u> </u>
I	mo11	1 0.264	0.264	0.264	0. 264	10.264	I 0.263	1 0.256	0.248	1 0.231	0.175	}	}	1
-		0.259												┿

A second		APPI	MDIX	At 1	LOSS 1	PARCE	IT A	ND PRO	OPORT.	TOWN	(ALL	MIRRI	ON BO	(ES)	
and the second s															
		1					Loss Pr	obabili	ty YR	1991					
			mo1	mo2	mo3	mo4	mo5	mo6	mo7	mo8	mo9	mo10	mo11	mo12	acc
•		mo1	0.058												0.942
		mo2	0.049	0.041											0.91
en e regisk film i de film.	C	mo3	0.023	0.04	0.07										0.867
:	0	mo4	0.014	0.006	0.04	0.09									0.85
	n	mo5	0.012	0.001	0.006	0.041	0.071								0.869
	t	mo6	0.009	0.001	0.003	0.01	0.065	0.071							0.841
	F	mo7	0.015	0.001	0.001	0.009	0.018	0.084	0.121		Ĺ				0.752
		moß	0.011	3E-04	0	0.005	0.012	0.017	0.077	0.113					0.76
	C	mo9	0.01	9E-04	0	0.004	0.007	0.006	0.018	0.08	0.098				0.777
	t	mo10	0.02	0	0	0.003	0.008	0.008	0.015	0.018	0.103	0.145			0.68
		mo11	0.006	5E-04	8E-04	8E-04	5E-04	0.008	0.008	0.017	0.028	0.068	0.115		0.747
		mo12	0.005	6E-04	0.001	2E-04	0.001	0.001	0.003	0.006	0.018	0.033	0.059	0.095	0.776
-															
						P	ercent	of Loss	YR I	991					
			mo1	mo2	mo3	mo4	mo5	mo6	mo7	mo8	mo9	mo10	mo11	mo12	
		mo1	100%						!						
		mo2	54%	46%											
	C	mo3	17%	30%	52%										
	0	mo4	9%	4%	27%	60%									
	12	mo5	9%	1%	5%	31%	54%								
	1	moß	6%	1%	2%	6%	41%	45%							
•	F	mo7	6%	1%	0%	4%	7%	34%	49%						
		mo8	5%	0%	0%	2%	5%	7%	33%	48%					
•	c	mo9	4%	0%	0%	2%	3%	3%	8%	36%	44%				
	1	mo10	6%	0%	0%	1%	2%	3%	5%	6%	32%	45%			
		mo11	2%	0%	0%	0%	0%	3%	3%	7%	11%	27%	46%		
		mo12	2%	0%	1%	0%	0%	1%	1%	3%	8%	15%	26%	43%	
										{					
		T		Prot	X DEP	Lees in	month	i Sur	vived t	o mont	h i - 1)	for Y	R 1991		
			mo1	mo2	mo3	mo4	mo5	moß	mo7	mo8	mo9	mo10	mo11	mo12	
		mo1	}		I										
		mo2	0.044												
	C	mo3	0.112	0.074											
	0_	mo4	0.138	0.115	0.083										
	•	mo5	0.121	0.119	0.114	0.075									
`•	1	mo6				0.139									
	r	mo7	0.237	0.236	0.235	0.229	0.214	0.138							
		mo8	0.227	0.226	0.226	0.222	0.213	0.199	0.129						
	c	mo9	0.215	0.214	0.214	0.212	0.206	0.201	0.186	0.112					
	t	mo10	0.306	0.308	0.306	0.304	0.299	0.292	0.282	0.287	0.176				
		mo11	0.249	0.248	0.248	0.247	0.247	0.241	0.234	0.22	0.197	0.134			
		mo12	0.22	0.22	0.210	0 210	0 210	0 217	0.215	0.209	0.106	0 100	0.100	ĭ	T

						Loss Pr	Hdado	ty YR	1992					
		mo1	mo2	mo3	mo4	mo5	mo6	mo7	mo8	mo9	mo10	mo11	mo12	acc
	mo1	0.066												0.934
	mo2	0.028	0.046											0.926
C	mo3	0.022	0.019	0.063										0.896
0	mo4	0.025	0.006	0.031	0.085									0.853
	mo5	0.029	0.003	0.009	0.036	0.091								0.832
t	moß	0.032	0.002	0.005	0.013	0.043	0.109							0.797
2	mo7	0.032	0.003	0.006	0.006	0.014	0.06	0.106						0.773
	mo8	0.017	0.001	0.004	0.005	0.005	0.016	0.056	0.119					0.777
c_	mo9	0.013	8E-04	0.002	0.002	0.005	0.015	0.022	0.062	0.121				0.758
t	mo10	0.006	3E-04	9E-04	0.001	0.002	0.009	0.016	0.023	0.069	0.104			0.768
	mo11	0.007	6E-04	0	0.002	0.001	0.003	0.01	0.012	0.019	0.07	0.124		0.751
	mo12	0.003	0.002	0	0	0	7E-04	0.005	0.01	0.004	0.018	0.052	0.093	0.811
														·
						Percen	t of Lo	s YR	1992					
		mo1	mo2	mo3	mo4	mo5	тов	mo7	mo8	mo9	mo10	mo11	mo12	
	rno1	100%												
	mo2	38%	62%											
C	mo3	21%	18%	60%										
•	mo4	17%	4%	21%	58%									
	mc5	17%	2%	6%	21%	54%					\ <u> </u>			
t	mo6	16%	1%	2%	6%	21%	54%	-			10 15 1			
r	mo7	14%	1%	3%	3%	6%	26%	47%						
A	mo8	8%	1%	276	2%	2%	7%	25%	53%					!
c	mo9	5%	0%	1%	1%	2%	6%	9%	26%	50%				
t	mo10	3%	0%	0%	1%	1%	4%	7%	10%	30%	45%			
	mo11	3%	0%	0%	1%	1%	1%	4%	5%	8%	28%	50%	0%	
	mo12	2%	1%	0%	0%	0%	0%	3%	5%	2%	10%	28%	49%	
	 													
									1					
		<u> </u>	Prob	(DEP	Loes in	month	i Sur	vived to	o mont	k (-1)	for YI	R 1992	· · · · · ·	
		mo1	mo2	mo3	mo4	mo5	mo6	mo7	mo8	mo9	mo10	mo11	mo12	
	mo1													
	mo2	0.047												
C	mo3	0.084	0.066		i			i	1			;		
0	mo4	0.125	0.104	0.078										
0	mo5				0.098					<u> </u>				
t	moß			0.172		0.12								
r	mo7						0.121							
	mo8				0.201		0.184							
c	mo9		0.231		0.228				0.137	1				
1	mo10	<u> </u>	0.227		0.225				0.184		 			
	mo11				0.241		0.238		0.221			— —		
	mo12				0.184				0.171					

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÷			GMA BEX	i-LSUM	4AEY_			
		ACCES	SION			LOSS		
	BLACK	HISP	OTHER	WHITE	B! ACK	HISP	OTHER	WHITE
Child	865	188	126	2985	107	15	17	258
No Child	5783		736	22134	837	85	120	2886
Total	6648	807	862	25090	944	100	137	
Overall	0.875859	0.889746	0.862863	0.888743				
w/Child	0.889918	0.926108	0.881119	0.920522				
who Child	0.873585	0.879261	0.859813	0.884652				
	No Child Total Overall w/Child	Child 885 No Child 5783 Total 6848 Overall 0.875859 w/Child 0.889918	ACCES: BLACK HISP Child 865 188 No Child 5783 619 Total 5848 807 Overall 0.875859 0.889748 w/Child 0.889918 0.928108	ACCESSION BLACK HISP OTHER Child 885 188 126 No Child 5783 619 738 Total 5848 807 862 Overall 0.875859 0.889746 0.862863 w/Child 0.889918 0.928108 0.881119	ACCESSION BLACK HISP OTHER WHITE Child 865 188 126 2965 No Child 5783 619 736 22134 Total 5848 807 862 25098 Overall 0.875659 0.889746 0.862863 0.888745 wChild 0.889918 0.926108 0.881119 0.920522	ACCESSION BLACK HISP OTHER WHITE BLACK Child 885 188 126 2985 107 No Child 5783 619 738 22134 837 Total 5848 807 882 25099 944 Overall 0.875859 0.889746 0.862863 0.888745 w/Child 0.889918 0.928108 0.881119 0.920522	ACCESSION LOSS BLACK HISP OTHER WHITE BLACK HISP Child 865 188 126 2965 107 15 No Child 5783 619 738 22134 837 85 Total 5848 807 862 25098 944 100 Overall 0.875659 0.889746 0.862663 0.888745 w/Child 0.889918 0.926108 0.881119 0.920522	ACCESSION LOSS BLACK HISP OTHER WHITE BLACK HISP OTHER Child 885 188 126 2985 107 15 17 No Child 5783 619 738 22134 837 85 120 Total 5848 807 882 25099 944 100 137 Overall 0.875869 0.889746 0.862863 0.888743 w/Child 0.889918 0.928108 0.881119 0.920522

GMA BRIG - III SUMMARY

		ACCES	SION			LOSS	3	
	BLACK	HISP	OTHER	WHITE	BLACK	HISP	OTHER	WHITE
Child	1072	412	94	6517	102	24	. 6	467
No Child	6653	840	376	28003	834	98	70	3450
Total	7725	1252	470	34520	936	122	76	3917
Overall	0.891929	0.911208	0.860806	0.898093				
w/Child	0.913118	0.944954	0.94	0.933133				
w/o Child	0.888607	0.895522	0.843049	0.890313				

GMA BRIG - IV SUMMARY

		ACCES	SION			LOSS		
	BLACK	HISP	OTHER	WHITE	BLACK	HISP	OTHER	WHITE
Child	645	602	147	5981	66	30	13	362
No Child	3197	2008	650	23449	489	163	83	2741
Total	3842	2610	797	29430	555	193	96	3103
Overall	0.873778	0.931145	0.892497	0.90462				
w/Child	0.907173	0.952532	0.91875	0.942929				
w/o Child	0.867336	0.924919	0.886767	0.895342				

GMA BRIG - VI SUMMARY

		ACCES	SION			LOSS	<u> </u>	
	BLACK	HISP	OTHER	WHITE	BLACK	HISP	OTHER	WHITE
Child	212	363	257	3396	41	27	24	260
No Child	1412	2150	1813	18986	279	241	265	2391
Total	1624	2513	2070	22364	320	268	289	2651
Overall	0.835391	0.903632	0.87749	0.894024				
w/Child	0.837945	0.930769	0.914591	0.928923				
wo Child	0.835009	0.899205	0.872474	0.888046				

APPENDIX D: COEFFICIENTS FOR FITTED MODEL

सर्थ कर्षाः ह	Torm	Coefficient	T-value
	Intercept	2.9489	70.8386
	DEP 2	0.8564	20.7224
The state of the s	DEP 3	1.0873	21.6483
a an an is an	DEP 4	1.1244	21.0712
	DEP 5	1.3172	21.9937
	DEP 6	1.6258	16.7283
	DEP 7	1.4663	12.9993
	DEP 8	1.9191	11.3314
	DEP 9	2.0457	8.8271
	DEP 10	2.3308	6.9678
	DEP 11	2.2630	6.3743
	DEP 12	2.4008	6.3331
	ZONE 2	-2.3392	-5.8631
	ZONE 3	-2.2915	-5.7102
	ZONE 4	-4.1370	-10.7285
	MARRIED	-0.4653	-17.3229
	HISPANIC	0.3939	8.7309
	OTHER	0.0611	1.2649
	WHITE	0.2063	9.2504
:	BRIG III	0.0474	2.1839
•	BRIG IV	0.0737	3.1831
	BRIG VI	-0.0093	-0.3860

NOTE: 1. First level of each term set to 0 by S-Plus.

APPENDIX D: COEFFICIENTS FOR FITTED MODEL

Term	Coefficient	T-Value
DEP 2:ZONE	2 1.7314	4.3146
DEP 3:ZONE	2 2.9350	7.2405
DEP 4:20NB	2 3.9049	9.4196
DEP 5: ZONE	2 3.3384	8.0940
DEP 6:ZONE	2 2.2073	5.2636
DEP 7:ZONE	2 2.1321	5.0068
DEP 8:ZONE	2 1.0163	2.3004
DEP 9:ZONE	2 0.8031	1.6946
DEP 10:ZONE	2 0.2682	0.5015
DEP 11:ZONE	2 0.5655	1.0118
DEP 12:ZONE	2 0	NA ²
DEP 3:ZONE	3 1.1121	2.7489
DEP 4:ZONE	3 2.2299	5.4832
DEP 5:20NE	3 2.0110	4.9312
DEP 6:ZONE	3 1.2935	3.1066
DEP 7:ZONE	3 1.2213	2.8970
DEP 8:ZONE	3 0.5839	1.3224
DEP 9:ZONE	3 0.1823	0.3870
DEP 10: ZONE	3 -0.1064	-0.1994
DEP 11: ZONE	3 0.0817	0.1483
DEP 12: ZONE	3 0	NA ²

NOTE: 2. The model is overdetermined and this coefficient had multiple least square solutions. The coefficient does not contribute to the model.

APPENDIX D: COEFFICIENTS FOR FITTED MODEL

	Term	Coefficient	T-value
	DEP 4:ZONE 4	2.5705	6.6027
्रा १८५५ - वृत्रियम् राष्ट्राच्या	DEP 5:ZONE 4	2.2905	5.8678
	DEP 6:ZONE 4	1.5381	3.8617
	DEP 7:ZONE 4	1.4057	3.4893
	DEP 8: ZONE 4	0.8468	2.0022
	DEP 9:ZONE 4	0.5469	1.2074
	DEP 10: ZONE 4	0.1949	0.3786
	DBP 11: ZONE 4	0.1891	0.3574
	DEP 12: ZONE 4	0	NA ²

NOTE: 2. The model is overdetermined and this coefficient had multiple least square solutions. The coefficient does not contribute to the model.

						CASE: M	ARRIED, BL	ACK, BRI	Gi		
		Condit	tional Pro	b (fitted) d	of Survivin	g Past Zone (i)	Prob of A	ccession	Given Su	rvived Zoi	ne (i-1)
3	ed Heren	Aga +		ZOI	NE			•	ZOI	NE	
			1	2	3	4		1	2	3	4
	*	1	0.9502				1	0.9502			
•		2	0.9782	0.9607			2	0.9398	0.9607		
		3	0.9826	0.9903	0.9456		3	0.9201	0.9364	0.9456	
	KM: D	4	0.9832	0.9964	0.9822	0.9246	4	0.8897	0.9049	0.9081	0.9246
	E	5	0.9861	0.9948	0.9817	0.9183	5	0.8843	0.8968	0.9015	0.9183
	P	6	0. 989 7	0.9883	0.9727	0.8782	6	0.8355	0.8442	U. 8542	0.8782
		7	0.988	0.9853	0.9659	0.8434	7	0.793	0.8027	0.8146	0.8434
		. 8	0.9923	0.9719	0.9593	0.8289	8	0.7669	0.7728	0.7952	0.8289
		9	0.9932	0.9694	0.9471	0.3029	9	0.7321	0.7372	0.7604	0.8029
		10	0.995	0.9612	0.9471	0.7881	10	0.7139	0.7174	0.7464	0.7881
		11	0.9946	0.9689	0.9528	0.7756	11	0.7121	0.716	0.739	0.7756
		12	0.9942	0.9516	0.9531	0.8102	12	0.7306	0.7348	0.7722	0.8102
						CASE: S	INGLE, BLA	-			
		Condi	tional Pro			ng Past Zone (i)	Prob of A	ccession			ne (i-1)
				ZO 1					ZO		
:			1	2	3	4		1	2	3	4
		1	0.9229				1	0.9229			
l		2	C.9657	0. 938 9			2	0.9067	0.9389		
	D	3	0.9726	0.9847	0.9161		3	0.8774	0.9021	0.9161	
	E	4	0.9736	0.9943	0.9719	0.885	4	0.8327	0.8552	0.8601	0.885
	P	5	0.9781	0.9918	0.9712	0.8759	5	0.8252	0.8437	0.8507	0.8759
		6	0.9638	0.9816	0.9573	0.8191	6	0.7572	0.7697	0.7841	0.8191
		7	0.9611	0.9768	0.9468	0.7718	7	0.7003	0.7138	0.7307	0.7718
		8	0.9679	0.958	0.9367	0.7526	8	0.8658	0.6739	0.705	0.7526
ļ		9	0.9893	0.9522	0.9183	0.7189	9	0.6219	0.6286	0.8602	0.7189
		10	0.9919	0.9395	0.918	0.7052	10	0.6033	0.6082	0.6474	0.7052
l		11 12	0. 9913 0. 9924	0. 95 13 0. 927 2	0. 9266 0. 9304	0. 6897 0.6787	11 12	0. 6027 0. 58 1	0. 608 0. 5855	0. 639 1 0. 63 15	0. 6897 0. 678 7
		12	0.5924	0.8212	0.550-	0.0707	12	0.501	0.3000	0.0013	0.0707
							IARRIED, HI				
		Condi	itional Pro			ng Past Zone (i)	Prob of	Accessio	n Given S		one (i-1)
				ZO				_	ZO		
ŀ		_	1	2	3	4		1	2	3	4
		1	0.9658				1	0.9658	A A704		
ŀ		2	0.9852	0.9731	0.0007		2	0.9587	0.9731	0.9627	
ł		3	0.9882	0.9934	0.9827	0.0470	3	0.9451	0. 9563 0. 9341	0.9363	0.9478
ľ	D	4	0.9886	0.9978	0.9879	0.9478	4	0.9234 0.9197	0.9284	0.9317	0.9434
	E	5	0. 9906 0. 993	0. 996 5 0. 992 1	0. 9876 0. 9814	0. 9434 0. 9145	5 6	0.8842	0.8904	0.8975	0.9145
	-	6	0.993	0.9921	0.9768	0.8863	7	0.8502	0.8571	0.8657	0.8863
I		7	0.9948	0.9809	0.9723	0.8752	8	0.8304	0.8347	0.851	0.8752
I		8	0.9955	0.9792	0.9638	0.8 55	9	0.8033	0.8069	0.824	0.855
		10	0.9966	0.9735	0.9637	0.8465	10	0.7915	0.7942	0.8158	0.8465
		11	0.9533	0.9788	0.9877	0.8367	11	0.7896	0.7925	0.8097	0.8367
l		12	0.996	0.9868	0.9679	0.8632	12	0.8045	0.8078	0.8355	0.8632
		12	V.##O	U.,JUU	V. 3 010	3.000		J.JU-10	5.5010		-,

	Condi	tional Pm	b (fitted) d	of Survivin	ig Past Zone (i)	NGLE, HISP Prob of			urvived Za	one (i-1)
	-	40110	ZOI		19 1 as (20 110 (1)	1 100 017	100000	ZOI		Sile (1-1)
i ve tig	. *.	1	2	3	4		1	2	3	4
	1	0.9467	-	•	-	1	0.9467	•		- 4
	2	0.9766	0.9579			2	0.9355	0.9579		
	3	0.9813	0.9896	0.9418		3	0.9146	0.932	0.9418	
- D	4	0.982	0.9961	0.9809	0.9194	4	0.8822	0.8983	0.9018	0.9194
THE STATE	5	0.9851	0.9944	0.9804	0.9127	5	0.8765	0.8898	0.8948	0.9127
P	.: 6	0.989	0.9875	0.9708	0.8704	6	0.8252	0.8344	0.845	0.8704
•	7	0.9873	0.9843	0.9636	0.8304	7	0.7776	0.7876	0.8002	0.8304
	8	0.9919	0.97	0.9566	0.8149	8	0.75	0.7561	0.7795	0.8149
	9	0.9928	0.9673	0.9436	0.7874	9	0.7135	0.7187	0.743	0.7874
	10	0.9945	0.9584	0.9432	0.7801	10	0.7013	0.7052	0.7358	0.7801
	11	0.9942	0.9667	0.9495	0.7629	11	0.6962	0.7003	0.7244	0.7629
	12	0.9937	0.9482	0.9498	0.7991	12	0.7151	0.7197	0.759	0.7991
	12	0.9301	0.5402	0.5457	0.7301	12	0.7131	0.7137	0.739	0.7331
					CASE:MA	RRIED, OT	HER ,BR	IG I		
	Condi	tional Pro			ng Past Zone (i)	Prob of	Accessio	n Given S	urvived Z	one (i-1)
			ZO	NE				ZO	NE	
		1	2	3	4		1	2	3	4
	1	0.953				1	0.953			
	. 2	0.9799	0.9632			2	0.9438	0.9632		
,	3	0.9837	0.9909	0.9487		3	0.9248	0.9401	0.9487	
D	4	0.9842	0.9967	0.9833	0.9288	4	0.8958	0.9102	0.9132	0.9288
E	5	0.987	0.9952	0.9828	0.9228	5	0.8908	0.9025	0.9069	0.9228
P	6	0.9904	6. 989 1	0.9744	0.8846	6	0.8443	0.8525	0.862	0.8846
	7	0.9889	0.9862	0.968	0.8483	7	0.8009	0.8099	0.8212	0.8483
	8	0.9929	0.9736	0.9619	0.8309	8	0.7727	0.7782	0.7992	0.8309
	9	0.9939	0.9714	0.9505	0.8051	9	0.7388	0.7433	0.7652	0.8051
	10	0.9953	0.9835	0.9501	0.7982	10	0.7273	0.7307	0.7584	0.7982
	11	0.9951	0.9708	0.9557	0.7821	11	0.7221	0.7257	0.7475	0.7821
	12	0.9946	0.9544	0.9558	0.8194	12	0.7434	0.7475	0.7832	0.8194
					CASE: SI	NGLE, OTH	HER . BRI	G I		
	Condi	itional Pro	ob (fitted)	of Survivi	ng Past Zone (i)				urvived Z	one (i-1)
			ZO	NE				ZO	NE	
		1	2	3	4		1	2	3	4
	1	0.929				1	0.929			
	2	0.9677	0.9423			2	0.9119	0.9423		
	3	0.9742	0.9856	0.9208		3	0.8841	0.9075	0.9208	
D	4	0.9751	0.9947	0.9736	0.8912	4	0.8416	0.863	0.8676	0.8912
E	5	0.9794	0.9923	0.9729	0.8824	5	0.8344	0.8519	0.8585	0.8824
P	6	0.9848	0.9827	0.9598	0.828	6	0.7691	0.781	0.7947	0.820
	7	0.9822	0.9782	0.9498	0.7824	7	0.714	0.727	0.7432	0.7824
	8	0.9886	0.9586	0.9403	0. 7638	8	0.6806	0.6884	0.7182	0.763
	8	0.9901	0.9551	0.9231	0. 7264	9	0.6341	0.6405	0.6706	0.726
					0.7129	10	0.6159	0.6205	0.6579	0.712
	10	0.9926	0. 943 1	0.9228	U. 7 128	10	V.U (Jø	0.0203	U.QJ/ B	0.7 123
	10 11	0.9926	0.9431	0.9226	0.6977	11	0.6148	0.6198	0.6495	0.6977

CASE: M	arried,	MHL	TE .E	3RIG	ı
7000 (i)	D	-4 4 -			٠

	Condi	tional Pro	b (fitted)	of Survivir	ng Past Zone (i)	Prob of Accession Given Survived Zone (i-1)					
			ZO	NE		ZONE					
		1	2	3	4		1	2	3	4	
	1	0.9591				1	0.9591				
	2	0.9822	0.9678			2	0.9506	0.9678			
	3	0.9858	0.9921	0.9554		3	0.9344	0.9478	0.9554		
D	4	0.9863	0.9971	0.9855	0.9378	4	0.9089	0.9215	0.9242	0.9378	
E	5	0.9887	0.9958	0.9851	0.9325	5	0.9045	10.9148	0.9186	0.9325	
P	8	0.9917	0.9905	0.9778	0.8986	6	0.8631	0.8703	0.8787	0.8986	
	. 7	0.9903	0.988	0.9721	0.8888	7	0.8263	0.8345	0.8446	0.8688	
	8	0.9938	0.9771	0.9667	0.8562	8	0.8037	0.8087	0.8277	0.8582	
	9	0.9945	0.975	0.9566	0.8335	9	0.7732	0.7774	0.7973	0.8335	
	10	0.9959	0.9683	0.9566	0.8206	10	0.7569	C.76	0.7849	0.8206	
	11	0.9956	0.9745	0.9612	0.8131	11	0.7583	0.7617	0.7816	0.8131	
	12	0.9962	0.9615	0.9632	0.8052	12	0.7428	0.7457	0.7756	0.8052	
					CASE: CII	NOIE WH	TE BOK	2.1			

CASE: SINGLE, WHITE, BRIG I

	Condi	tional Pro	rb (fitted) d	of Survivir	ng Past Zone (i)	Prob of	Accessio	n Given S	urvived Zi	on e (i-1)
			ZO	NE						
		1	2	3	4		1	2	3	4
	1	0.9364				1	0.9364			
	2	0.972	0.9497			2	0.9231	0.9497		
	3	0.9776	0.9875	0.9307		3	0.8986	0.9191	0.9307	
D	4	0.9784	0.9954	0.9771	0.9045	4	0.8607	0.8797	0.8837	0.9045
E	5	0.9821	0.9933	0.9765	0.8967	5	0.8542	0.8698	0.8756	0.8967
P	6	0.9868	0.985	0.965	0.8477	6	0.7952	0.8058	0.8181	0.8477
	7	0.9846	0.9811	0.9563	0.8061	7	0.7447	0.7563	0.7709	0.8061
	8	0.9901	0.964	0.9479	0.789	8	0.7139	0.721	0.7479	0.789
	9	0.9913	0.9608	0.9326	0.7587	9	0.6739	0.6798	0.7075	0.7587
	10	0.9934	0.9503	0.9323	0.7463	10	0.6568	0.6611	0.6958	0.7463
	11	0.993	0.9601	0.9395	0.7321	11	0.6557	0.6603	0.6878	0.7321
	12	0.9939	0.94	0.9426	0.7219	12	0.6357	0.6397	0.6805	0.7219

CASE: MARRIED, BLACK, BRIG II

	Condi	tional Pro	b (fitted)	of Survivir	ng Past Zone (i)	Prob of	Accessio	n Given S	urvived Z	one (i-1)
			201	NE				ZOI	NE	
		1	2	3	4		1	2	3	4
	1	0.9524				1	0.9524			
•	2	0.9792	0.9625			2	0.9425	0.9625		
	3	0.9834	0.9908	0.9481		3	0.9238	0.9393	0.9481	
D	4	0.984	0.9966	0.983	0.9279	4	0.8945	0.909	0.9121	0.9279
E	5	0.9868	0.9951	0.9826	0.9201	5	0.8877	0.8996	0.904	0.9201
P	6	0.9903	0.9889	0.974	0.8832	6	0.8425	0.8507	0.8603	0.8832
	7	0.9888	0.9861	0.9676	0.8465	7	0.7986	0.8077	0.8191	0.8465
	8	0.9928	0.9733	0.9613	0.8322	8	0.7731	0.7786	0.8	0.8322
	9	0.9937	0.9709	0.9496	0.8066	9	0.739	0.7437	0. 76 6	0.8066
	10	0.9952	0.963	0.9495	0.796	10	0.7243	0.7278	0.7558	0.796
	11	0.9948	0.9703	0.9548	0.7878	11	0.726	0.7298	0.7521	0.7878
	12	0.9945	0.9538	0.9552	0.8174	12	0.7406	0.7447	0.7808	0.8174

			APP	andir e	: FITT	LED ANTOES VND	COMD	PROBABI	LITIES	
	1				_					
			L 10m - 11 -			ASE: SINGLE, BLAC			d 7 .	// 4\
-	Condi	tional Pro	b (fitted) d Z O I		g Past 40	ne (I) Propor A	VCC655101	Given Si ZON	urvived Zo	ou e (1-1)
		1	201	3	4		1	2	3	4
	1	0.9263	2	3	7	1	0.9263	_	•	=
in the second	2	0.923	0.9416			2	0.9108	0.9416		
	3	0.9739	0.9854	0.9198		3	0.8827	0.9064	0.9198	
Ð	4	0.9748	0.9946	0.9732	0.8898	4	0.8397	0.8614	0.866	9.8898
E	5	0.9791	0.9922	0.9726	0.881	5	0.8324	0.8501	0.8568	0.881
	6	0.98 **	0.9824	0.9593	0.8261	6	0.7665	0.7785	0.7924	U.8261
rrager is a lista	·	0.982	0.9779	0.9492	0.7801	7	0.711	0.7241	0.7404	0.7801
•	. 8	0.9885	0.958	0.9395	0.7813	8	0.6773	0.6852	0.7153	0.7613
	9	0.9892	0.9544	0.9218	0.7237	9	0.6302	0.8367	0.6671	0.7237
	10	0.9923	0.9422	0.9216	0.715	10	0.6161	0.6208	0. 6589	0.715
	11	0.9918	0.9535	0.9298	0.6948	11	0.6109	0.616	0.646	0. 6948
	12	0.9928	0.9304	0.9334	0.6889	12	0.594	0.5982	0.643	0.6889
•										
						ASE: MARRIED, HI				
	Cutidi	tional Pro	b (fitted) d		ng Past Zo	one (i) Prob of A	Accession		urvived Zo	one (i-1)
			ZOI				_	ZOI		
	_	1	2	3	4	•	1	2	3	4
	1	0.9682				1	0.9682	0 074E		
	2	0.9862	0.9745	0.0041		2	0.9611 0.9473	0.9745 0.958	0.9641	
	3	0. 9888 0. 9892	0. 993 7 0. 997 7	0.9641 0.9665	0.949	3	0.9258	0.936	0.9381	0.945
D E	4 5	0.9911	0.9967	0.9882	0.9459	5	0.9233	0.9316	0.9347	0.9459
P	6	0.9935	0.9925	0.9824	0.9163	6	0.8877	0.8935	0.9002	0.9163
	7	0.9924	0.9906	0.9779	0.891	7	0.8566	0.8631	0.8714	0.891
	8	0.9952	0.9818	0.9736	0.8803	8	0.8374	0.8415	0.857	0.8803
	9	0.9957	0.0802	0.9655	0.8608	9	0.8112	0.8147	0.8311	0.8608
	10	0.9968	0.9747	0.9654	0.8526	10	0.7997	0.8023	0.8231	0.8526
	11	0.9966	0. 3798	0.9692	0.8431	11	0.7979	0.8006	0.8171	0.8431
	12	0.9983	0.9624	0.9694	0.8687	12	0.8124	0.8154	0.8421	0.8687
						ASE: SINGLE, HIS			_	
	Condi	itional Pr	ob (fitted)		ng Past Zo	one (i) Prob of	Accessio		urvived Z	on e (i-1)
			ZO					ZO		
		1	2	3	4	4	1	2	3	4
	1	0.9491	0.0000			1	0. 949 1 0. 938 4	0.0500		
	2	0.9777	0. 959 3 0. 990 1	0.9444		2	0.9384	0. 9598 0. 935 1	0.9444	
	3	0.9822 0.9829	0.9964	0.9818	0.9229	4	0.8874	0.9028	0.9061	0.9229
D E	5	0.9858	0.9947	0.9813	0.9165	5	0.882	0.8946	0.8994	U.9165
P	6	0.9897	0.9881	0.9723	0.873	6	0.8301	0.8387	0.8488	0.873
_	7	0.9878	0.985	0.9651	0.837	7	0.786	0.7957	0.8078	0.837
	8	0.9922	0.9713	0.9584	0.8255	8	C.7624	0.7684	0.7911	0.8255
	9	0.9932	0.9689	0.9461	0.7952	9	0.724	0.729	0.7524	0.7952
	10	0.9948	0.9603	0.9459	0.7841	10	0.7086	0.7123	0.7417	0.7841
	11	0.9945	0.9682	0.9518	07714	11	0.707	0.7109	0.7342	0.7714
	12	0.9941	0.9506	0.9521	0.8066	12	0.7257	0.73	0.768	0.8066
İ										

•						CASE: M	ARRIED, O	THER RI	31G 11		
٠.		Condi	ional Da	h (Swad) .	ما جيمهم	ng Past Zone (i)		Accession		unimed 7	nna (i.1)
1.7 2		CONOR	DOILES IS TO	ZO (ig P as (2014 (I)	P100 01	MCC635101	ZOI)
	W		1	2	3	4		1	2	3	4
·	• • •	1	0.9551		•		1	0.9551			
		2	0.9804	0.9648			2	0.946	0.9648		
		3	0.9845	0.9913	0.9506		3	0.9276	0.9423	0.9506	
	D	4	0.9852	0.9968	0.9841	0.9303	4	0.8991	0.9126	0.9155	0.9303
-	E	5	0.9876	0.9954	0.9836	0.9261	5	0.8954	0.9067	0.9109	0.9261
.15 \	P	6	0.991	0.9896	0.9756	0.887	6	0.8487	0.8584	0.8654	0.887
	.•-	7	0.9894	0.9869	0.9695	0.8543	7	0.8087	0.8173	0.8282	0.8543
		8	0.9933	0.9748	0.9635	0.8406	8	0.7842	0.7895	0.8099	0.8406
		9	0.9942	0.9727	0.9527	0.8124	9	0.7484	0.7528	0.774	0.8124
		10	0.9956	0.9652	0.9525	0.802	10	0.7341	0.7374	0.7639	0.802
		11	0.9952	0.9721	0.9575	0.7939	11	0.7354	0.739	0.7602	0.7939
		12	0.9939	0.9552	0.956	0.8631	12	0.7833	0.7881	0.8251	0.8631
							-				
						CASE: SI	NGLE, OTH	IERBR#	GII		
		Condi	tional Pro	b (fitted) d	of Survivir	ng Past Zone (i)	Prob of	Accession	n Given S	urvived Z	one (i-1)
	ZONE								ZO	NE	
			1	2	3	4		1	2	3	4
		1	0.9304				1	0.9304			
		2	0.9692	0.9448			2	0.9157	0.9448		
		3	0.9754	0.9863	0.9241		3	0.889	0.9115	0.9241	
	D	4	0.9763	0.9949	0.9748	0.8957	4	0.8481	0.8687	0.8731	0.8957
	Ε	5	0.9803	0.9927	0.9742	0.8848	5	0.8389	0.8557	0.862	0.8848
	P	6	0.9855	0.9835	0.9616	0.8347	6	0.7779	0.7893	0.8026	0.8347
		7	0.9833	0.9792	0.9522	0.7864	7	0.721	0.7333	0.7488	0.7864
		8	0.9891	0.9604	0.9429	0.7723	8	0.6917	0. 6993	0.7282	0.7723
		9	0.9906	0.9571	0.9264	0.7358	9	0.6462	0.6524	0.6816	0.7358
		10	0.9929	0.9456	0.9262	0.7226	10	0.6283	0.6328	0.6692	0.7226
		11	0.9924	0.9562	0.934	0.7076	11	0.6271	0.6319	0.6609	0.7076
		12	ი.9918	0.9323	0.9344	0.7501	12	0.6481	0.6535	0.7009	0.7501
						CASE. M	ADDIED W	ucc oo			
		Coodi	محالا احماد	h (fished)	ad Canadia	CASE: Manage Past Zone (i)	ARRIED, W	Accessio		undered 7	(i.1)
		Condi	uonai Pio	ZO		ing rest zone (i)	PIOD OI	ACCUSSIO	ZO		(ו-ו) שווכ
	•		1	2	3	4		1	20	3	4
		1	0.9609	_	3	-	1	0.9609	_	•	7
		2	0.983	0.9693			2	0.9528	0.9693		
		3	0.9865	0.9925	0.9573		3	0.9373	0.9502	0.9573	
	0	4	0.987	C.9972	0.9862	0.9405	4	0.9129	0.9249	0.9275	0.9405
	D	5	0.9892	0.996	0.9358	0.9354	5	0.9086	0.9185	0.9222	0.9354
	P	6	0.9921	0.991	0.9788	0.9029	6	0.8688	0.8757	0.8837	0.9029
	•	7	0.9907	0.9886	0.9734	0.8741	7	0.8333	0.8411	0.8508	0.8741
		8	0.9941	0.9781	0.9681	0.862	8	0.8114	0.8162	0.8345	0.862
		9	0.9948	0.9762	0.9585	0.84	9	0.7818	0.5162	0.8051	0.84
		10	0.9951	0.9696	0.9583	0.8308	10	0.769	0.772	0.7962	0.8308
		11	0.9959	0.9758	0.963	0.8167	11	0.7643	0.7674	0.7865	0.8367
		12	0.9955	0.9736	0.9633	0.846	12	0.7805	0.7874	0.7865	0.846
		12	U. 88333	U.502 I	U. 5U.33	U. 9-FJ	12	U. / Q UO	0.704	U.0 148	U. 040

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•													
	•				CASE: S	INGLE, WHITE, BRIG II							
	Condi	tional Pro	b (fitted) (of Survivir	ng Past Zone (i)				n Survive	d Zone (i-1	١		
			ZOI					ZOI			,		
		1	2	3	4		1	2	3	4			
	1	0.9392		•	·	1	0.9392	_	•	•			
	2	0.9732	0.9519			2	0.9265	0.9519					
	3	0.9786	0.9881	0.9337		3	0.9029	0.9226	0.9337				
D	4	0.9794	0.9956	0.8781	0.9085	4	0.8665	0.8847	0.8886	0.9085			
E	5	0.983	0.9937	0.9776	0.901	5	0.8603	0.8752	0.8808	0.901			
P	6	0.9874	0.9857	0.9686	0.8538	6	0.8032	0.8134	0.8252	0.8538			
٠.,	7	0.9853	0.982	0.9583	0.8134	7	0.7541	0.7654	0.7795	0.8134			
	8	0.9906	0.9656	0.9502	0.7988	8	0.7242	0.7311	0.7571	0.7968			
	9	0.9917	0.9626	0.9355	0.7673	9	0.6852	0.6909	0.7178	0.7673			
	10	0.9937	0.9525	0.9352	0.7551	10	0.6685	0.6727	0.7062	0.7551			
÷.,	11	0.9934	0.9619	0.9424	0.7367	11	0.8834	0.6678	0.6942	0.7367			
	12	0.9942	0.9426	0.9452	0.7313	12	0.6477	0.8515	0.6912	0.7313			
					CASE: N	MARRIED. BL	ACK, BR	IG III					
	Condi	tional Pro	b (fitted)	of Survivir	ng Past Zone (i)	Prob of	Accession	n Given S	urvived Z	one (i-1)			
			ZO	NE				ZO	NE				
		1	2	3	4		1	2	3	4			
	1	0.9536				1	0.9536						
	2	0.9798	0.9634			2	0.9439	0.9634					
	3	0.9839	0.991	0.9493		3	0.9256	0.9408	0.9493				
D	4	0.9844	0.9967	0.9835	0.9296	4	0.897	0.9112	0.9142	0.9296			
E	5	0.9871	0.9952	0.983	0.9237	5	0.892	0.9037	0.908	0.9237			
P	6	0.9905	0.9892	0.9747	0.8859	6	0.8461	0.8542	0.8635	0.8859			
	7	0.9891	0.9864	0.9684	0.8499	7	0.803	0.8119	0.8231	0.8499			
	8	0.9929	0.9739	0.9621	0.8359	8	0.7777	0.7832	0.8042	0.8359			
	9	0.9939	0.9717	0.9509	0.8107	9	0.7444	0.749	0.7709	0.8107			
	10	0.9954	0.9639	0.9507	0.8002	10	0.7299	0.7333	0.7608	0.8002			
	11	0.995	0.9711	0.9561	0.7882	11	0.7282	0.7318	0.7536	0.7882			
	12	0.9956	0.9562	0.9582	0.8213	12	0.7492	0.7525	0.7869	0.8213			
					CASE. S	INGLE, BLA	OV DDV	S 114					
	Condi	ional Dw	h (finad)	مراجع المراجعة	ng Past Zone (i)		Accession			nn			
	Silu:	OCHE FIC	Z O !		ig =25t 20il# (i)	F100 01	MOCES IO	ZOI		OII W (I-1)			
		1	2	3	4		1	2	3	4			
	1	0.9281	-	•	7	1	0.9281	2	3	-			
	2	0.9681	0.943			2	0.913	0.943					
	3	0.9745	0.9658	0.9217		3	0.8855	0.9086	0.9217				
D	4	0.9754	0.9948	0.9739	0.8924	4	0.8433	0.8646	0.8691	0.8924			
E	5	0.9797	0.9924	0.9732	0.8837	5	0.8362	0.8536	0.8801	0.8837			
P		0.985	0.9829	0.9603	0.8298	6	0.7715	0.7832	0.7969	0.8298			
-	7	0.9624	0.9785	0.9504	0.7846	7	0.7168	0.7296	0.7457	0.7846			
	8	0.9888	0.9591	0.941	0.7861	8	0.6836	0.6914	0.7209	0.7661			
	9	0.9902	0.9556	0.924	0.7289	9	0.6374	0.6437	0.6735	0.7289			
	10	0.9926	0.9437	0.9237	0.7155	10	0.6192	0.6238	0.661	0.7155			
	11	0.992	0.9546	0.9315	0.7053	11	0.6222	0.8272	0.657	0.7053			
	12	0.0015	0.9301	0 9333	0.7436	12	0.6393	0.8448	0.6032	0.7436			

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12 0.6393 0.6448 0.6932 0.7436

12 0.9915 0.9301 0.9322 0.7436

					CASE: MA	VRRIED, HI	SPANIC,E	BRIG III		
	Condi	tional Pro	•		g Past Zone (i)		Accession	Given Sı		on e (i-1)
			ZOI		_		_	ZOI		•
Ç .	_	1	2	3	4	_	1	2	3	4
= -	1	0.9682			•	1	0.9682			
	2	0.9863	0.975			2	0.9616	0.975		
_	3	0.9891	0.9939	0.9653		3	0.9489	0.9594	0.9653	0.054.4
D	4	0.9895	0.9978	0.9888	0.9514	4	0.9288	0.9387	0.9407	0.9514
£	5	0.9913	0.9968	0.9885	0.9472	5	0.9252	0.9333	0.9363	0.9472
. P	. 6	0.9937	0.9927	0.9829	0.9183	6	0.8903	0.896	0.9026	0.9183
	. 7	0.9926	0.9908	0.9785	0.8936	7	0.8599	0.8663	0.8743	0.8936
	8	0.9953	0.9823	0.9742	0.8831	8	0.8411	0.8451	0.8803	0.8831
	9	0.9958	0.9807	0.9663	0.8639	9	0.8154	0.8188	0.8349	0.8639
	10	0.9969	0.9754	0.9662	0.8559	10	0.8041	0.8066	0.827	0.8559
	11	0.9967	0.9803	0.9699	0.8466	11	0.8023	0.805	0.8211	0.8466
	12	0.9957	0.9683	0.9689	0.9003	12	0.841	0.8446	0.8723	0.9003
					CASE: SI	NGLE, HISI	PANIC,BF	AG III		
	Condi	tional Pro	b (fitted)	of Survivir	ng Past Zone (i)	Prob of	Accessio			one (i-1)
			ZO	NE				ZOI	NE	
		1	2	3	4		1	2	3	4
	1	0.9503				1	0.9503			
	2	0.9783	0.9608			2	0.94	0.9606		
	3	0.9827	0.9904	0.9458		3	0.9205	0.9367	0.9458	
D	4	0.9633	0.9965	0.9823	0.9248	4	0.8901	0.9052	0.9084	0.9248
Ε	5	0.9862	0.9949	0.9818	0.9185	5	0.8848	0.8972	0. 90 18	0.9185
P	6	0.9898	0.9884	0.9729	0. 8759	6	0.8337	0.8423	0.8522	0. 8759
	7	0.9681	0.9854	0.966	0.8437	7	0.7936	0.8032	0.8151	0.8437
	8	0.9924	0.972	0.9594	0. 8293	8	0.7674	0.7733	0.7956	0.8293
	8	0.9934	0.9696	0.9474	0. 799 5	9	0.7296	0.7345	0.7575	0.7995
	10	0.9949	0.9613	0.9471	0.7925	10	0.7178	0.7215	0.7506	0.7925
	11	0.9946	0. 969	0.9528	0.7802	11	0.7183	0.7202	0.7433	0. 7802
-	12	0.994%	0.9518	0.9533	0.8107	12	0.7313	0.7355	0.7728	0.8107
					CASE: M	ARRIED, O	THER .B	RIG III		
	Cond	itional Pro	b (fitted)	of Survivi	ng Past Zone (i)		Accessio		urvived Z	one (i-1)
			žó		•			ZO	NE	
		1	2	3	4		1	2	3	4
	1	0.9573				1	0.9573			
	2	0.9809	0.9655			2	0.9471	0.9655		
	3	0.9849	0.9915	0.9518		3	0.9294	0.9437	0.9518	
Đ	4	0.9856	0.9969	9.9845	0.932	4	0.9015	0.9147	0.9176	0.932
Ε	5	0.9879	0.9955	0.984	0.9279	5	0.8979	0.909	0.9131	0.9279
P	6	0.9912	0.9899	0.9762	0.8896	6	0.8521	0.8597	0.8685	0.8896
	7	0.9897	0.9872	0.9702	0.8575	7	0.8129	0.8214	0.832	0.8575
	8	0.9933	0.9754	0.9843	0.8472	8	0.7915	0.7968	0.8169	0.8472
	9	0.9942	0.9733	0.9537	0.8199	9	0.7566	0.761	0.7819	0.8199
	10	0.9957	0.9661		0.8062	10	0.7396	0.7428	0. 7688	0.8062
	11	0.9953	0.9728		0.7982	11	0.7408	0.7443	0. 765 1	0.7982
	12		0.9563		0.8662	12	0.788	0.7928	0.829	0.8662

Marie and the second se

						SINGLE, OTH				
•	Condi	tional Pro			ng Past Zone (i)	Prob of	Accession		urvived Z	one (i-1)
•			ZOI					ZOI		
		1	2	3	4		1	2	3	4
<i>∞</i> •• · ·	1	0.932				1	0.932			
	2	0.97	0.9462			2	0.9178	0.9462		
	3	0.976	0. 986 6	0.926		3	0.8917	0.9136	0. 92 6	
. D	4	0.9769	0.9951	0.9754	0.8981	4	0.8516	0.8717	0.876	0.8981
Ε	5	0.9808	0.9929	0.9748	0.8899	5	0.8447	0.8612	0.8674	0.8899
P	- 6	0.9859	0.9839	0.9625	0.8383	6	0.7827	0.7939	0.8089	0.8383
	7	0.9835	0.9797	0.9532	0.7947	7	0.7299	0.7422	0.7575	0.7947
** * *	8	0.9894	0.9614	0.9443	0.7769	8	0.6978	0.7053	0.7336	0.7769
	9	0.9908	0.9582	0.9282	0.7408	9	0.6528	0.6589	0.6876	0.7408
	10	0.9831	0.9469	0.9279	0.7278	10	0.6351	0.6395	0.6754	0.7278
	11	0.9926	0.9573	0.9356	0.713	11	0.6339	0.6386	0.6671	0.713
	12	0.992	0.934	0.936	0.7549	12	0.6547	0. 66	0.7086	0.7549
					CASE:	MARRIED, W	HITE, BR	IG III		
	Condi	itional Pro	b (fitted)	of Survivi	ng Past Zone (i)	Prob of	Accessio	n Given S	urvived Z	on e (i-1)
			ZO	NE				ZOI	NE	
		1	2	3	4		1	2	3	4
	1	0.9619				1	0.9619			
	2	0.9835	0.9701			2	0.954	0.9701		
	3	0.9868	0.9027	0.9584		3	0.9389	0.9514	0.9584	
Ð	4	0.9873	0.9973	0.9865	0.942	4	0.915	0.9268	0.9293	0.942
E	5	0.9895	0.9961	0.9862	0.937	5	0.9108	0.9204	0.924	0.937
P	6	0.9923	0.9912	0.9793	0.9052	6	0.8718	0.8786	0.8864	0.9052
	7	0.9909	0.9889	0.974	0.877	7	0.8371	0.8447	0.8542	0.877
	8	0.9942	0.9787	0.969	0.8651	8	0.8156	0.8203	0.8382	0.8651
	9	0.9949	0.9788	0.9595	0.8435	9	0.7866	0.7906	0.8094	0.8435
	10	0.9962	0.9704	0.9595	0.8312	10	0.771	0.7739	0.7978	0.8312
	11	0.996	0.9784	0.964	0.8206	11	0.7692	0.7723	0.791	0.8206
	12	0.9956	0.963	0.9642	0.8493	12	0.7852	0.7887	0.8189	0.8493
					CASE:	SINGLE, WH	ITE, BRIG	1 I I I		
	Cond	itional Pro	b (fitted)	of Survivi	ng Past Zone (i)	Prob of	Accessio	n Given S	iurvived Z	on e (i-1)
			ZO	NE				ZO	NE	
		1	2	3	4		1	2	3	4
	1	0.9407				1	0.9407			
	2	0.9739	0.9531			2	0.9263	0.9531		
	3	0.9792	0.9684	0.9353		3	0.9053	0.9245	0.9353	
D	4	0.9799	0.9957	0.9787	0.9107	4	0.8696	0.8874	0.8912	0.9107
E	5	0.9834	0.9638	0.9781	0.9033	5	0.8635	0.8781	0.8835	0.9033
P	6	0.9877	0.986	0.9674	0.857	6	0.8075	0.8175	0.8291	0.857
	7	0.9857	0.9824	0.9593	0.8174	7	0.7593	0.7703	0.7841	0.8174
	8	0.9908	0.9664	0.9514	0.801	8	0.7298	0.7366	0.7621	0.801
	9	0.9919	0.9635	0.937	0.772	9	0.6913	0.697	0.7234	0.772
	10	0.9939	0.9536	0.9368	0.76	10	0.6748	0.679	0.712	0.76
	11	0.9935	0.9628	0.9436	0.7483	11	0. 673 6	0. 678	0.7042	0.7463
	12	0.9943	0.944	0.9442	0.7805	12	0.6917	0.6957	0. 7389	0. 7805

		IG VI	ACK, BR	RRIED, BL	CASE: MA							
one (i-1)	urvived Za	n Given S	Accession	Prob of	ng Past Zone (i)	of Survivir	b (fitted) d	tional Pro	Dondi			
	NE	ZOI			ZONE							
4	3	2	1		4	3	2	1				
			0.9498	1				0.9498	1			
		0.9604	0.9393	2			0.9604	0.978	2			
	0.9447	0.9355	0.9191	3		0.9447	0.9902	0.9825	3			
0.924	0.9074	0.9041	0.8889	4	0.924	0.9821	0.9964	0.9831	4			
0.9158	0.899	0.8943	0.8821	5	0.9158	0.9817	0.9948	0.9863	5			
0.8772	0.8532	0.8432	0.8345	6	0.8772	0.9725	0.9883	0.9697	6			
0.839	0.8103	0.7983	0.7889	7	0.839	0.9658	0.9853	0.9681	37			
0.8276	0.7936	0.7712	0.7652	8	0.8276	0.959	0.9717	0.9923	8			
0.7976	0.7552	0.732	0.7272	9	0.7976	0.9469	0.9693	0.9933	9			
0.7886	0.7447	0.7156	0.712	10	0.7866	0.9467	0.9609	0.995	10			
0.774	0.7372	0.7141	0.7103	11	0.774	0.9524	0.9687	0.9946	11			
0.8486	0.8068	0.7663	0.761	12	0.8486	0.9508	0.9498	0.9931	12			

CASE: SINGLE, BLACK , BRIG VI Conditional Prob (fitted) of Surviving Past Zone (i) Prob of Accession Given Survived Zone (i-1) Z O N E 1 2 3 4 1 2 3 4

D E

		1	2	3	4		1	2	3	4
	1	0.9223				1	0.9223			
	2	0.9655	0.9384			2	0.906	0.9384		
	3	0.9724	0.9846	0.9155		3	0.8765	0.9014	0.9155	
D	4	0.9734	0.9943	0.9717	0.8841	4	0.8315	0.8543	0.8591	0.8841
Ε	5	0.9779	0.9918	0.971	0.8749	5	0.8239	0.8425	0.8495	0.8749
P	6	0.9837	0.9814	0.957	0.8178	6	0.7555	0.7681	0.7826	0.8178
	7	0.9812	0.9767	0.9466	0.7 659	7	0.6949	0.7081	0.725	0.7659
	8	0.968	0.9558	0.9364	0.7464	8	0.66	0.668	0.699	0.7464
	9	0.9892	0.9518	0.9177	0.7171	9	0.6196	0.6263	0.658	0.7171
	10	0.992	0.9392	0.9177	0.8983	10	0. 59 7	0.6019	0.6408	0. 698 3
	11	0.9915	0.951	0.9263	0.6826	11	0.5962	0.6014	0.6323	0.6826
	12	0.9024	0.9266	0.9298	0. 676 6	12	0. 578 5	0.5829	0.6291	0.6766

CASE: MARRIED. HISPANIC, BRIG VI

	Condi	tional Pro	b (fitted)	of Survivi	ng Past Zone (i)	Prob of	Accessio	n Given S	urvived Zi	one (i-1)
			ZO	NE				ZOI	NE	
		1	2	3	4		1	2	3	4
	1	0.9656				1	0.9656			
	2	0.9651	0.9729			2	0.9584	0.9729		
	3	0.9881	0.9934	0.9624		3	0.9447	0.958	0.9624	
D	4	0.9885	0.9976	0.9678	0.8474	4	0.9229	0.9336	0.9359	0.9474
E	5	0.9905	0.9965	0.9875	0.9429	5	0.9191	0.9279	0.9311	0.9429
P	6	0.993	0.9921	0.9813	0.9138	6	0.8834	0.8896	0.8967	0.9138
	7	0.992	0.99	0.9767	0.8854	7	0.8492	0.8561	0.8647	0.8854
	8	0.9949	0.9808	0.9721	0.8742	8	0.8292	0.8335	0.8498	0.8742
	9	0.9954	0.979	0.9634	0. 8568	9	0.8044	0.8082	0.8255	0.8568
	10	0.9966	0.9733	0.9634	0.8453	10	0.79	0.7927	0.8144	0.8453
	11	0.9964	0.9787	0.9674	0.8355	11	0.7881	0.791	0.8083	0.8355
	12	0.9953	0.9856	0.9662	0.8926	12	0.6289	0.8328	0.8625	0.8926

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	CASE							INGLE, HISPANIC,BRIG VI					
		Condit	tional Pro		of Survivin	i) Prob of	Prob of Accession Given Survived Zone (i-1)						
	-		_	ZOI		•		•	ZOI		•		
3 k. 3		±	1	2	3	4	4	1	2	3	4		
		1	0.9463	0.0536			1	0.9463	0.0576				
		2	0.9784	0.9576	0.0414		2	0. 93 5 0. 91 4	0. 9576 0. 93 15	0.9414			
	8	3	0.9812 0.9819	0. 9896 0. 9962	0.9414 0.9608	0.9188	4	0.8814	0.8977	0.9011	0.9188		
	D	4	0.985	0.9944	0.9803	0.9121	5	0.8757	0.8891	0.894	0.9121		
	P	5 8	0.9889	0.9874	0.9706	0.8894	6	0.8239	0.8331	0.8438	0.8694		
		7	0.9873	0.9842	0.9633	0.8291	7	0.7761	0.7861	0.7987	0.8291		
		8	0.9919	0.9697	0.9562	0.8136	8	0.7483	0.7544	0.778	0.8136		
		9	0.9927	0.987	0.9429	0.7898	9	0.7149	0.7202	0.7448	0.7898		
		10	0.9946	0.9581	0.943	0.7744	10	0.6959	0.6997	0.7302	0.7744		
		11	0.9942	0.9664	0.9491	0.7613	11	0.6942	0.6983	0.7225	0.7613		
		12	0.9937	0.9478	0.9494	0.7977	12	0.7133	0.7178	0.7574	0.7977		
						CASE	: MARRIED, O	THER R	31G VI				
		Condi	tional Pro	b (fitted)	of Survivir	g Past Zone (urvived Zo	one (i-1)		
				ZO					ZOI		(,		
			1	2	3	4		1	2	3	4		
		1	0.9526	_	_		1	0.9526					
		2	0.9793	0.9627			` 2	0.9427	0.9627				
•		3	0.9835	0.9908	0.9483		3	0.9241	0.9396	0.9483			
	D	4	0.9841	0.9966	0.9831	0.9282	4	0.8949	0.9094	0.9125	0.9282		
	E	5	0.9668	0.9951	0.9827	0.9221	5	0.8898	0.9017	0.9061	0.9221		
	P	6	0.9903	0.989	0.9741	0.8837	6	0.8431	0.8513	0.8606	0.8837		
		7	0.9688	0.9861	0.9677	0.8471	7	0.7993	0.8084	0.8196	0.8471		
		8	0.9928	0.9733	0.9613	0.8361	8	0.7767	0.7823	0.8038	0.8361		
		9	0.9936	0.971	0.9497	0.811	9	0.743	0.7478	0.7702	0.811		
		10	0.9953	0.9632	0.9497	0. 796 7	10	0.7253	0.7287	0.7586	0. 796 7		
		11	0.9949	0.9705	0.9551	0.7845	11	0.7235	0.7272	0.7493	0.7845		
		12	0.9945	0.954	0.9554	0.818	12	0.7415	0.7456	0.7816	0.818		
						CASE		SINGLE, OTHER, BRIG VI					
		Condi	tional Pro	-		ng Past Zone	(i) Prob of	Accessio		urvived Z	one (i-1)		
			1	ZO:	NE 3	4		1	Z O 1	NE 3	4		
		4	0.9266	2	•	7	1	0.9266	-	•	-		
		1 2	0.9675	0.9418			2	0.9112	0.9418				
		3	0.974	0.9855	0.9201		3	0.8831	0.9067	0.9201			
	D	4	0.9749	0.9947	0.9734	0.8903	4	0.8403	0.8619	0.8665	0.8903		
	E	5	0.9796	0.9923	0.9728	0.8789	5	0.831	0.8484	0.855	0.8789		
	P	6	0.9847	0.9825	0.9594	0.8267	6	0.7673	0.7793	0.7932	0.8287		
	-	7	0.982	0.978	0.9494	0.7808	7	0.712	0.725	0.7413	0.7808		
		8	0.9885	0.9582	0.9397	0.7578	8	0.6745	0.6824	0.7121	0.7578		
		9	0.99	0.9547	0.9224	0.7246	9	0.6317	0.8381	0.6684	0.7246		
		10	0.9925	0.9426	0.9222	0.711	10	0.6134	0.618	0.6557	0.711		
		11	0.9918	0.9537	0.9301	0.7007	11	0.6165	0.6216	0.6517	0.7007		
		12	0.9913	0.9287	0.9308	0.7394	12	0.6336	0.6392	0.6883	0.7394		

CASE: MARRIED, WHITE, BRIG VI

	Condi	ditional Prob (fitted) of Surviving Past Zone (i) Prob of Accession Given Survived Zone (i-1)								
			ZOI	NE			NE	Ē		
_	•	1	2	3	4		1	2	3	4
. •·	1	0.9588				1	0.9588			
	2	0.9821	0.9675			2	0.9502	0.9675		
	3	0.9857	0.9921	0.955		3	0.9338	0.9474	0.955	
D	4	0.9862	0.9971	0.9854	0.9373	4	0.9082	0.9208	0.9235	0.9373
E	5	0.9886	0.9958	ി.985	0.9319	5	0.9036	0.914	0.9179	0.9319
P	8	0.9916	0.9904	0.9776	0.8978	6	0.882	0.8692	0.8776	0.8978
	7	0.9902	0.9879	0.9719	0.8677	7	0.825	0.8331	0.8433	0.8677
	8	0.9937	0.9768	0.9664	0.8551	8	0.8021	0.8072	0.8263	0.8551
	9	0. 994 6	0.9749	0.9563	0.8289	9	0. 768 6	0.7728	0.7927	0.8289
	10	0.9959	0.968	0.9562	0.8192	10	0.7551	0.7582	0.7833	0.8192
	11	0.9956	0.9744	0.961	0.8081	11	0.7533	0.7566	0.7785	0.8081
	12	0.9961	0.9611	0. 9629	0.8037	12	0.7409	0.7438	0.7739	0.8037

CASE: SINGLE, WHITE, BRIG VI

	Condi	tional Pro	Prob (fitted) of Surviving Past Zone (i) Prob of Accession Given Survived Zone (i-1)							
			ZO	NE				ZOI	NE	
		1	2	3	4		1	2	3	4
	1	0.9359				1	0.9359			
-	2	0.9717	0.9493			2	0.9224	0.9493		
	3	0.9774	0.9874	0.9301		3	0.8977	0.9184	0.8301	
D	4	0.9782	0.9954	0.9769	0.9037	4	0.8595	0.8787	0.8828	0.9037
E	5	0.982	0.9933	0.9763	0.8958	5	0.853	0.8687	0.8745	0.8958
P	6	0.9867	0.9849	0.9647	0.8465	6	0.7936	0.8043	0.8167	0.8465
	7	0.9844	0.9809	0.9559	0.8047	7	0.7428	0.7545	0.7692	0.8047
	8	0.99	0.9636	0.9475	0.7875	8	0.7118	0.719	0.7461	0.7875
	9	0.9912	0.9605	0.932	0.757	9	0.6717	0.6776	0.7055	0.757
	10	0.9934	0.9498	0.9317	0.7445	10	0.6545	0.8589	0.8937	0.7445
	11	0.993	0.9598	0.9392	0.7255	11	0.8495	0.6541	0.6814	0.7255
	12	0.9938	0.9395	0.9421	0.72	12	0.6334	0.6373	0.6783	0.72

LIST OF REFERENCES

- 1. Bryan D. Burris, Captain US Army, MISSIONING WITH MINIMAL DELAY ENTRY PROGRAM (DEP) LOSS, Master's Thesis, Naval Postgraduate School, September 1993.
- Jeanna F. Celete and Michael J. Wilson, <u>An Even History</u>
 <u>Analysis of DEP Contract Losses</u>, Research Report, Westat Inc., Rockville, MD, 1981.
- 3. Ray A. Zimmerman, Dona C. Zimmerman, and Mary Ellen Lathrop, Study of Factors Related to Army Delayed-Entry Program Attrition, USAREC Study Report 85-3, Research and Study Division, Program Analysis and Evaluation Directorate, Fort Knox, Kentucky, 1985.
- 4. Jeanna F. Celete, <u>An Even History Analysis of DEP Contract Losses</u>, Research Report, Westat Inc., Rockville, MD,1985.
- 5. Cyril E. Kearl and Abraham Nelson, <u>DEP Attrition: A Microdata Model</u>, Research Report, US Army Research Institute, Alexandria, VA, 1990
- 6. Daniel C. Bunning, Captain US Army. <u>US ARMY'S DELAY ENTRY PROGRAM: ATTRITION MODELING</u>, Master's Thesis, Naval Postgraduate School, September 1991.
- 7. Alan Agresti, <u>Categorical Data Analysis</u>. John Wiley and Sons, 1990.
- 8. Alan Agresti, Analysis of Ordinal Categorical Data. John Wiley and Sons, 1984.
- 9. Stephen E. Fienberg, <u>The Analysis of Cross-Classified Categorical Data</u>. MIT Press, 1980.
- 10. Robert S. Pindyck, Daniel L. Rubinfeld. <u>Economic Models</u> and <u>Economic Forecasts</u>, Third Edition. McGraw-Hill Inc, 1991.

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