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RECENT METHODOLOGICAL DEVELOPMENTS IN MAGNITUDE DETERMINATION AND YIELD ESTIMATION WITH APPLICATIONS TO SEMIPALATINSK EXPLOSIONS

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This technical report has been reviewed and is approved for publication.

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## **Table of Contents**

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Table of Contents
Summary v
Section I 1
I.1 Abstract 1
1.2 Introduction
I.3 Definition of a New Magnitude, m <sub>2.9</sub>
I.4 Underlying Model and Advantages of The New Magnitude 29
I.5 Magnitude: Yield Relationship at Semipalatinsk with $m_{2.9}$
I.6 Geophysical Interpretation of Our Near-source Corrections 40
I.7 Discussion and Conclusions 43
I.8 Acknowledgements 44
I.9 References 44
Appendix A. Geotech's Maximum-Likelihood Network mb: GLM91A
Appendix B. Estimates of Test Site Bias with m <sub>b</sub> (GLM91A)62
Section II
II.1 Contract No. and II.2 Project Objectives
II.3 Research Accomplished During Contract Period
II.3.1 Upgrading of Unbiased Network mb Estimator
II.3.2 m <sub>b</sub> -yield Regression Routine with Censored Yields: MLE-CY
II.3.3 $m_b$ -yield Regression Routine with Uncertain Data: DWLSQ
II.3.4 Expansion of Geotech's WWSSN mb Database
II.3.5 Magnitude-yield Relationship at Semipalatinsk
II.3.6 Reports, Presentations, and Publications
II.4 Conclusions and Recommendations
II.5 Acknowledgements
II.6 References
Distribution List

# RECENT METHODOLOGICAL DEVELOPMENTS IN MAGNITUDE DETERMINATION AND YIELD ESTIMATION WITH APPLICATIONS TO SEMIPALATINSK EXPLOSIONS

R.-S. Jih and R. A. Wagner Teledyne Geotech Alexandria Laboratory 314 Montgomery Street Alexandria, VA 22314-1581

#### SUMMARY

This report includes two parts. The first section discusses in detail the research that was done since the submittal of our first annual report *(GL-TR-90-0107)*, and the second part gives a perspective overview of the whole project. Note that the same contract number (F19628-89-C-0063) is also used by two other totally independent tasks. This volume covers only the work performed under Task 1.

An improved magnitude determination procedure is described and tested in Section I. This procedure accounts for the near-source focusing/defocusing effects as well as the receiver effects with empirically determined correction terms. Although no *a priori* geophysical information is required in deriving these receiver and near-source terms, the inferred corrections turn out to show fair correlation with the tectonics underneath the receivers as well as the visible geological structures near the source region. For 79 out of 82 Semipalatinsk events in our WWSSN database, the new scheme provides more stable  $m_b$  measurements across the whole recording network with a reduction in the fluctuational variation by a factor of up to 3. The 3 events which do not show significant improvement could have been detonated in environments with different focusing patterns. The scatter in the network-averaged  $m_b$  based on the new scheme versus log(yield) is smaller than that for conventional GLM or LSMF  $m_b$ , if the standard and/or the rounding errors in the  $m_b$  and Soviet-published yields (Bocharov *et al.*, 1989) are included in the regressions.

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**Section II** summarizes the research accomplished during the contract period, including those major results already reported in our first annual report (*GL-TR-90-0107*) as well as those discussed in Section I. The most important developments in the seismic yield estimation methodology under this project are:

- a refined *m<sub>b</sub>* determination scheme (*cf.* Section I),
- an algorithm "MLE-CY" which utilizes the bounded yields in the m<sub>b</sub>-yield regression based on the maximum-likelihood approach,
- an algorithm "DWLSQ" which permits both variables to be imprecise due to either rounding or standard measurement errors.

We have fully tested these techniques with Soviet-published yields and the WWSSN  $m_b$  database established at Geotech, and the results all appear to be very encouraging in improving our remote monitoring capability.

Also included in this report is a complete listing of 192 event  $m_b$  values measured off 21547 WWSSN recordings (Appendix A). We have updated the "yield-dependent" test site bias estimate using these GLM91A  $m_b$  values (Appendix B).

#### **SECTION I**

# A REFINED NETWORK m<sub>b</sub> DETERMINATION SCHEME INCORPORATING NEAR-SOURCE EFFECTS

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## **I.1 ABSTRACT**

An improved magnitude determination procedure is presented to account for the empirical near-source focusing/defocusing effects. For 79 out of 82 Semipalatinsk events in our WWSSN database, the new scheme provides more stable  $m_b$  measurements across the whole recording network with a reduction in the fluctuational variation by a factor of up to 3. The standard error in the resulting network-averaged  $m_b$  is typically around 0.02 m.u., similar to that of  $RMS L_g$  based on in-country regional recordings. The 3 events which do not show significant improvement could have been detonated in environments with different focusing patterns. The scatter in the network-averaged  $m_b$  based on the new scheme versus log(yield) is smaller than that for conventional GLM or LSMF  $m_b$ , if the standard and/or the rounding errors in the  $m_b$  and Soviet-published yields are included in the regressions.

### **I.2 INTRODUCTION**

The main problem with the conventional  $m_b$  is that it is a rather nebulous parameter; simply, it is a function of the largest peak-to-peak amplitude in the first few seconds of *P* wave motion with adjustment for the period of the arriving phase. The

- 1 -

parameter  $m_b$  was adapted from the need to order systematically the size of earthquakes. The measure itself has inherent impreciseness as the measure is not related to the physics of the source *per* se but is the largest constructive interference of waves originating at the source, source region, propagation path, receiver region, and receivers (Butler, 1981; Johnson, 1981). To relate the  $m_b$  to the seismic yield, all effects not due to the source must naturally be corrected. It is often difficult, however, to separate these effects. In fact, the effects of source and propagation are often indistinguishable, and unless one is known the other can not be uniquely determined (Johnson, 1981). Consequently, it was reported to be difficult to make  $m_b$  measurements that are internally consistent within 0.1  $m_b$  with the conventional  $m_b$  (Bache, 1982), simply because some of the aforementioned effects are not accounted for accurately.

Von Seggern (1973) showed that including station corrections typically halves the standard deviation of  $m_b$  from North American LRSM stations recording NTS events. Even better results, with the standard deviation reduced by a factor of 3 or so, can be obtained for a network with all stations beyond 30°. Applying station corrections to the  $m_b$  determination or the network spectra averaging has become a standard procedure in this community (*e.g.*, Lilwall *et al.*, 1988; Murphy *et al.*, 1989; Sykes and Ekstrom, 1989; Jih and Shumway, 1989; and many others). The station effects are strongly dependent on azimuth (Chang and von Seggern, 1980), which led Bache (1982) and many others to believe that statistical station corrections will not be nearly so effective in reducing  $m_b$  variance in the multiple source region problem.

Marshall *et al.* (1979) attempted to correct several important factors that can bias  $m_b$ . They used bulletin log(A/T) data. These data are corrected for receiver-station attenuation differences, and the resulting magnitude is called  $m_2$ . Correcting  $m_2$  for source-region attenuation gives  $m_3$ , and correcting  $m_3$  for source depth gives  $m_0$ . (The network averages of these  $m_b$ s are denoted by  $\overline{m}_1$ ,  $\overline{m}_2$ ,  $\overline{m}_3$ , and  $\overline{m}_0$ , respectively.) The major change in this scheme is associated with the source-region

- 2 -

correction, which can be as different as 0.4 m.u. between sites or a factor of about 2.5 in yield estimates. Essentially this approach is based on the the discovery by Marshall and Springer (1976) and Douglas *et al.* (1981) that LRSM amplitude residuals correlate with  $P_n$  velocity near the stations, and under the assumption that such correlation between the attenuation and  $P_n$  is valid elsewhere as well. However, it turns out that the standard deviation of the  $m_0$  is not less than that for the  $m_1$  and  $m_2$  from the same data set, indicating that the attenuation correction in Marshall *et al.* (1979) is not correlated with the residuals from these stations (Bache, 1982).

It is obvious that the only way to reduce the statistical fluctuation is to obtain fundamental causal knowledge of the focusing and defocusing beneath the source and receiver. We expect teleseismic *P*-wave amplitudes to vary as source location changes within a test site. The  $m_b$  residuals (with respect to the best-fitting  $m_b$ -yield curve) of NTS events show systematic trends that are consistent with local tectonics (Minster *et al.*, 1981). At Yucca Flat, the residuals are positive to the west and negative to the east of the north-south trending normal fault system that bisects the valley. At Pahute Mesa, the spatial pattern is less clear, but the residuals tend to be negative toward the center of the buried Silent Canyon Caldera and positive toward the edges. An attractive explanation is that these variations are due to focusing/defocusing effects that are not averaged out over the network, although the possibility of systematic source-coupling difference has not been eliminated.

Figure 1 illustrates the  $m_b$  residual patterns of E. Kazakh explosions as seen from various directions. For each event, the residual is the (maximum-likelihood) average of all station residuals (*viz*, with network  $m_b$  and the station term removed) in that quadrant. All three subsites exhibit very different azimuthal variation. For instance, Murzhik events are enhanced in the NE and SW directions and reduced in the NW and SE directions, whereas Degelen events are reduced towards the SW direction. There seems to be some weak distinction between NE and SW subregions of Balapan test site along SE and NW directions. Figure 2 gives the azimuthal pattern with the

- 3 -

four quadrants rotated by 45°. The initial *P* waves from the three adjacent test sites have virtually the same incident angle at any particular teleseismic station, and anything in common across all events (such as the crustal amplification as well as the upper mantle attenuation underneath the receiver) would have been lumped into the constant station term. Thus the station residuals averaged over all events from the same test site would correlate very little with the receiver. Instead, they should reveal more site-dependent information about the focusing/defocusing pattern underneath E. Kazakhstan.

In this study, we present an improved scheme to determine the network  $m_b$  with both the station terms and near-source focusing/defocusing effects corrected. Examples are given to illustrate that such procedure can reduce the random fluctuation in the station  $m_b$  values to about 0.15 m.u. or lower. It is shown that, by applying this scheme to worldwide explosions, it is possible to have a consistent base line in estimating the absolute magnitudes, which is crucial in estimating the test site bias, while the precision in the resulting network  $m_b$  values can be maintained as well as could be achieved by the single-test-site approach.



#### mb(Pmax) RESIDUALS AS SEEN FROM VARIOUS DIRECTIONS

Figure 1. The  $m_b$  residuals of E. Kazakh explosions as seen from various directions. For each event, the residual is the (maximum-likelihood) average of all station residuals (viz with network  $m_b$  and the station term removed) in that quadrant. All three subsites exhibit very different azimuthal variation. For instance, Murzhik events are enhanced in the NE and SW directions, and reduced in the NW and SE directions, whereas Degelen events are reduced towards the SW direction. There seems to be some weak distinction between NE and SW subregions of Balapan test site along SE and NW directions.



mb(Pmax) RESIDUALS AS SEEN FROM VARIOUS DIRECTIONS

Figure 2. The azimuthal patterns with the four quadrants rotated by 45°. Most Semipalatinsk events are enhanced in the north, and reduced in the east.

### 1.3 DEFINITION OF A NEW MAGNITUDE, m29

The conventional definition of station magnitude is computed as

$$m_b = \log_{10}(A/T) + B(\Delta)$$
<sup>[1]</sup>

where A is the displacement amplitude (in nm) and T is the predominant period (in sec) of the *P* wave. The B( $\Delta$ ) is the distance-correction term that compensates for the change of *P*-wave amplitudes with distance (*e.g.*, Gutenberg and Richter, 1956; Veith and Clawson, 1972).  $m_b$  in [1] is also denoted as  $m_1$  in Marshall *et al.* (1979). The ISC bulletin  $m_b$  is just the network average of these raw station  $m_b$  values without any further adjustment.

Consider  $N_e$  explosions detonated at  $N_F$  source regions that are recorded at some or all of  $N_s$  stations. The GLM91A network  $m_b$  (*cf.* Appendix A) is the (maximum-likelihood) average of the "station-corrected" magnitudes:

$$m_{2,2}(i,j) \equiv m_1(i,j) - S(j)$$
 [2]

where S(j) is the "statistical" receiver correction at the j-th station. In Marshall *et al.* (1979), *a priori* information about the  $P_n$  velocity underneath each station is used to determine its associated "deterministic" receiver correction, S(j), and the resulting magnitude is called  $m_2$ . Our receiver corrections (Figure 3), however, are inferred jointly from a suite of event-station pairs, and no *a priori* geophysical or geological condition is assumed (and hence the different notation  $m_{2,2}$ ). It turns out when the azimuthal coverage is broad enough, receiver corrections derived by such statistical approach do reveal the average tectonic structure underneath the recording stations, as many earlier studies have reported (*e.g.*, North, 1977): the station terms are positive in shields regions such as Australia, India, Canada, and Scandinavia; and they are negative in the east Africa rift valleys, island arcs (*e.g.*, Japan and Taiwan), and Indonesia. The high correlation between the tectonic type and the station terms suggests that the station corrections do reflect the upper mantle conditions underneath the receivers. The result also supports the claim of a marginal superiority of WWSSN

over ISC data. For instance, Pacific island arcs are believed to have high attenuation, low  $P_n$  velocities as well as negative station terms. However, North's (1977) station corrections based on 38316 ISC recordings of earthquakes around the world do not show this phenomenon because ISC does not have so wide an azimuthal coverage and uniform spatial sampling as does WWSSN.

#### WWSSN & ISC STATION CORRECTIONS



Figure 3. The station terms derived with WWSSN and ISC recordings. Our station terms (top; used in this study) are based on the GLM/MLE joint inversion of 192 worldwide explosions recorded at 122 "good" WWSSN stations. Only paths within 20 and 95 degrees are used. For each station, our GLM/LSMF joint inversion scheme puts any component constant across all events as the "station term", similar to the Douglas' (1966) LSMF approach. Both GLM (top) and LSMF (middle) station corrections exhibit a good correlation with the tectonics. The high correlation between the tectonic type and the station terms suggests that the station corrections do reflect the upper mantle conditions underneath the receivers. The result also supports the claim of a marginal superiority of WWSSN over ISC data, such as the 38316 ISC recordings used by North (1977) (bottom).

We now define a new magnitude,  $m_{2.9}$ , to account for the near-source focusing and defocusing effects:

$$m_{2,9}(i,j) \equiv m_1(i,j) - S(j) - F(k(i),j) = m_{2,2}(i,j) - F(k(i),j)$$
[3]

At the j-th station, F(k(\*),j) is a constant for all events detonated in the same k-th "geologically and geophysically uniform region". Partitioning a single nuclear test site into several "regions" may be necessary in order to account accurately for the focusing/defocusing effects. This  $m_{2.9}$  is very similar to the  $m_3$  in Marshall *et al.* (1979) except that, again, a priori attenuation information of the source region is used in Marshall *et al.* (1979) to determine the correction term, whereas we invert for the near-source effects from the data empirically. (The correlation between our statistical focusing/defocusing corrections and the geological structures will be verified in a later section.) As a result, the source-region corrections used by Marshall *et al.* (1979) are constants (for all explosions in the same source region) regardless of the location of the seismic stations, whereas our near-source corrections are dependent on the source-station paths.

Table 1 lists the station corrections (which are invariant for any explosion from any test site at any direction) as well as the near-source corrections associated with each subsite of Soviet's Semipalatinsk nuclear test ground.<sup>1</sup> Figure 4 plots out the lower-hemisphere equal-area projection of these "secondary corrections". The spatial map of these corrections are shown in Figure 5.

<sup>&</sup>lt;sup>1</sup>The near-source corrections for other test sites can be similarly derived. Our explosion data set needs to be expanded, however, to warrant a reasonable partitioning at other source regions.



# EQUAL-AREA PLOTS OF STATION MEAN mb(Pmax) ANOMALIES

**Figure 4.** Lower-hemisphere equal-area projection of station-corrected  $m_b(P_{m,n})$  anomalies for Soviet's 3 test sites in Eastern Kazakhstan. For each station, the mean  $m_b$  anomaly is defined as the mean residual averaged over all explosions from the same test site. The different patterns more or less reflect the focusing/defocusing underneath the source region.



## **MEAN STATION-CORRECTED mb(Pmax) ANOMALIES**

Figure 5. Same as Figure 4 except the mean  $m_b$  anomalies are superimposed by the map. We propose to regard these (site-dependent) mean station  $m_b$  anomalies as the "secondary corrections" to be applied in the procedure of magnitude determination.

	Table 1. Receiver and Near-source Corrections of WWSSN Stations									
Station Term Near-source Term, F				m, F	Station					
Code	S	Balapan	Degelen	Murzhik	Longitude	Latitude	Description			
AAE	-0.352	-0.387	-0.076	-0.127	38.766	9.029	Addis Ababa, Ethiopia			
AAM	0.205	0.202	-0.073	-0.247	-83.656	42.300	Ann Arbor, Michigan			
ADE	-0.044				138.709	-34.967	Adelaide, south Australia			
AFI	-0.135				-171.777	-13.909	Afiamalu, Samoa Islands			
AKU	-0.048	0.300	0.311	0.212	-18.107	65.687	Akureyri, Iceland			
ALQ	-0.028				-106.457	34.943	Albuquerque, New Mexico			
ANP	-0.349	-0.167	0.402	0.183	121.517	25.183	Anpu, Taiwan			
ANT	0.040				-70.415	-23.705	Antofagasta, northern Chile			
AQU	-0.133	-0.195	0.039	-0.024	13.403	42.354	Aquila, central Italy			
ARE	0.223				-71.491	-16.462	Arequipa, southern Peru			
ATL	0.141				-84.338	33.433	Atlanta, Georgia			
ATU	0.128	0.191	-0.156	0.034	23.717	37.972	Athens Univ., Greece			
BAG	-0.027	0.076	-0.009	-0.103	120.580	16.411	Baguio City, Luzon Island			
BDF	0.067				-47.903	-15.664	Brasilia Array, Brazil			
BEC	-0.114	0.090	0.058	-0.092	-64.681	32.379	Bermuda-Columbia, Atlantic			
внр	-0.219				-79.558	8.961	Balboa Heights, Panama			
BKS	0.089	-0.014	0.101	0.229	-122.235	37.877	Byerly, central California			
BLA	0.057	-0.233	-0.177	-0.299	-80.421	37.211	Blacksburg, West Virginia			
BOG	0.032				-74.065	4.623	Bogota, Colombia			
BOZ	0.188	-0.325	-0.025	-0.180	-111.633	45.600	Bozeman, Montana			

	Table 1. Receiver and Near-source Corrections of WWSSN Stations									
Station	n Term	Near	-source Ter	m, F	Station					
Code	S	Balapan	Degelen	Murzhik	Longitude	Latitude	Description			
BUL	0.003	0.014	-0.266	-0.037	28.613	-20.143	Bulawayo, Rhodesia			
CAR	0.190				-66.928	10.507	Caracas, Venezuela			
CHG	-0.140	0.240	0.106	0.045	98.977	18.790	Chiengmai, southeast Asia			
СМС	-0.178	0.114	0.375	0.602	-115.083	67.833	Copper Mine, Canada			
COL	0.065	0.181	0.188	0.051	-147.793	64.900	College Outpost, Alaska			
COP	0.127	-0.003	0.159	-0.276	12.433	55.683	Copenhagen, Denmark			
COR	0.155	0.132	0.183	0.172	-123.303	44.586	Corvallis, Oregon			
СТА	0.153	-0.072	0.003	-0.073	146.254	-20.088	Charters Towers, Australia			
DAG	0.036	-0.052	0.086		-18.770	76.770	Danmarkshavn, Greenland			
DAL	0.202				-96.784	32.846	Dallas, central Texas			
DAV	-0.320	-0.264	-0.053		125.575	7.088	Davao, Mindanao Island			
DUG	0.149	0.038	0.371	0.352	-112.813	40.195	Dugway, Utah			
EIL	0.004	-0.117	-0.228	-0.103	34.950	29.550	Eilat, Arabic Peninsula			
EPT	-0.023				-106.506	31.772	El Paso, Texas-Mexico border			
ESK	0.048	-0.042	0.162	-0.327	-3.205	55.317	Eskdalemuir, Scotland			
FLO	0.000	-0.294	-0.093	-0.446	-90.370	38.802	Florissant, eastern Missouri			
FVM	0.034	-0.038	0.045		-90.426	37.984	French Village, eastern Missouri			
GDH	-0.147	0.094	0.015	0.336	-53.533	69.250	Godhavn, western Greenland			
GEO	-0.003	0.038	-0.016	-0.051	-77.067	38.900	Georgetown, Washington D.C.			
GIE	-0.175				-90.300	-0.733	Galapagos Islands			

	Table 1. Receiver and Near-source Corrections of WWSSN Stations									
Station	Station Term Near-source Term, F					Station				
Code	s	Balapan	Degelen	Murzhik	Longitude	Latitude	Description			
GOL	-0.204	0.114	0.199	-0.028	-105.371	39.700	Golden, Colorado			
GRM	-0.285	-0.075			26.573	-33.313	Grahamstown, southern Africa			
GSC	0.057	-0.072	0.131	-0.064	-116.805	35.302	Goldstone, central California			
GUA	-0.088	-0.05 <del>9</del>			144.912	13.538	Guam, Mariana Islands			
нкс	-0.188	-0.130	-0.200	-0.026	114.172	22.304	Hong Kong			
HLW	-0.150	-0.096	-0.354		31.342	29.858	Helwan, Arabic Peninsula			
HN-ME	0.175				-67.986	46.162	Houlton, New Brunswick			
HNR	0.238	-0.278			159.947	-9.432	Honiara, Solomon Islands			
IST	0.148	0.129	-0.195	-0.090	28.996	41.046	Istanbul, Turkey			
JCT	0.133				-99.802	30.479	Junction City, central Texas			
JER	0.011	-0.035	-0.159	-0.054	35.197	31.772	Jerusalem, Dead Sea region			
KBL	0.023				69.043	34.541	Kabul, Afghanistan			
KBS	-0.213	-0.429	0.043	-0.284	11.924	78.918	Kingsbay, Svalbard region			
KEV	-0.119	0.139	0.122	-0.029	27.007	69.755	Kevoa, Finland			
KIP	0.110				-158.015	21.423	Kipapa, Hawaii			
кор	0.196	0.075	0.015	0.402	77.467	10.233	Kodaikanal, India			
KON	0.046	0.271	0.141	-0.243	9.598	59.649	Kongsberg, southern Norway			
KRK	-0.004		0.233	0.150	30.062	69.724	Kirkenes, Sandinavia			
КТG	-0.208	0.041	0.101	0.132	-21.983	70.417	Kap Tobin, eastern Greenland			
LAH	0.469				74.333	31.550	Lahore, India-Pakistan border			

Table 1. Receiver and Near-source Corrections of WWSSN Stations									
Station	Station Term Near-source Term, F				Station				
Code	S	Balapan	Degelen	Murzhik	Longitude	Latitude	Description		
LEM	-0.499	0.123	-0.535		107.617	-6.833	Lembang, Java		
LON	-0.044	-0.119	0.159	0.077	-121.810	46.750	Longmire, Washington		
LOR	0.095	-0.362	-0.165	0.056	3.851	47.267	Lormes, France		
LPA	0.357				-57.932	-34.909	La Plata, Uruguay		
LPB	0.064				-68.098	-16.533	La Paz, Peru-Bolivia border		
LPS	-0.068				-89.162	14.292	La Palma, Guatemala		
LUB	0.191				-101.867	33.583	Lubbock, west Texas		
MAL	-0.010	0.005	0.013	-0.145	-4.411	36.728	Malaga, Straits of Gibraltar		
MAN	0.276	0.212	-0.181		121.077	14.662	Manila, Luzon island		
МАТ	-0.188	-0.508	-0.167	0.015	138.207	36.542	Matsushiro, Honshu, Japan		
MDS	-0.043	-0.031	0.302		-89.760	43.372	Madison, Wisconsin		
MNN	0.153	-0.037			-93.190	44.914	Minneapolis, Minnesota		
MSH	0.202				59.588	36.311	Meshed, Iran-USSR border		
MSO	-0.045	-0.019	-0.047		-113.941	46.829	Missoula, Montana		
MUN	0.172	0.184	0.076	0.026	116.208	-31.978	Mundaring, western Australia		
NAI	-0.112	0.075	-0.108	-0.075	36.804	-1.274	Nairobi, Kenya		
NAT	0.118	·			-35.033	-5.117	Natal, Brazil		
NDI	0.1 <b>58</b>	0.216	0.286	0.618	77.217	28.683	New Delhi, northern India		
NHA	-0.134		-0.010	0.019	109.212	12.210	Nhatrang, southeast Asia		
NIL	-0.030				73.252	33.650	Nilore, Pakistan		

	Table 1. Receiver and Near-source Corrections of WWSSN Stations									
Station	Term	Near	-source Ter	rm, F			Station			
Code	S	Balapan	Degelen	Murzhik	Longitude	Longitude Latitude Description				
NNA	-0.171				-76.842	-11.988	Nana, Peru			
NOR	-0.240	0.086	0.154	0.335	-16.683	81.600	Nord, north coast of Greenland			
NP-NT	0.107				-119.372	76.252	North Pole, Queen Elizabeth Islands			
NUR	0.090	0.473	-0.076	0.041	24.651	60.509	Nurmijarvi, Finland			
OGD	-0.167	-0.061	-0.214	-0.263	-74.596	41.088	Ogdensburg, New York			
OXF	0.293				-89.409	34.512	Oxford, Mississippi			
PDA	0.017	0.018	-0.167		-25.663	37.747	Ponta Delgada, Azores Islands			
PEL	0.029				-70.685	-33.144	Peldehue, Chile-Argentina border			
PMG	0.151	0.101	0.060	0.253	147.154	-9.409	Port Moresby, New Guinea			
POO	0.076	-0.041	0.082	0.251	73.850	18.533	Poona, India			
PRE	-0.089	0.118	-0.210	-0.017	28.190	-25.753	Pretoria, south Africa			
PTO	-0.172	-0.163	-0.166	-0.029	-8.602	41.139	Porto, Serro Do Portugal			
QUE	-0.446	0.484	0.107	0.190	66.950	30.188	Quetta, Pakistan			
QUI	0.023				-78.501	-0.200	Quito, Ecuador			
RAB	0.022	0.090	-0.347	-0.002	152.170	-4.191	Rabaul, New Britain region			
RAR	-0.093				-159.773	-21.212	Rarotonga, Cook Islands region			
RCD	0.370	-0.217	-0.139		-103.208	44.075	Rapid city, South Dakota			
RIV	0.300				151.158	-33.829	Riverview, SE Australia			
RK-ON	-0.013				-93.672	50.839	Red Lake, Ontario			
SCP	-0 005	0.037	-0.129	-0.229	-77.865	40.795	State College Pennsylvania			
SDB	0.053	0.114	0.130	0.172	13.572	-14.926	Sa Da Bandeira, Angola			

Table 1. Receiver and Near-source Corrections of WWSSN Stations								
Statio	ion Term Near-source Term, F				Station			
Code	S	Balapan	Degelen	Murzhik	Longitude	gitude Latitude Description		
SEO	-0.132	-0.159	-0.377	-0.304	126.967	37.567	Seoul, South Korea	
SHA	0.297				-88.143	30.694	Spring Hill, Mississippi	
SHI	0.237	-0.273	-0.027	-0.105	52.520	29.638	Shiraz, southern Iran	
SHK	-0.286	-0.063	-0.387	-0.160	132.678	34.532	Shiraki, southern Honshu, Japan	
SHL	-0.001	0.106	0.024	-0.088	91.883	25.567	Shillong, India-Bangladesh border	
SJG	-0.129				-66.150	18.112	San Juan, Puerto Rico	
SNG	-0.003	0.099	-0.055	0.115	100.620	7.173	Songkhla, Malay Peninsula	
SPA	-0.630				0.000	-90.000	South Pole, Antarctica	
STU	0.067	-0.127	0.230	0.170	9.195	48.772	Stuttgart, Germany	
TAB	0.234	0.189	0.363	0.325	46.327	38.068	Tabriz, Iran-USSR border	
TAU	-0.137				147.320	-42.910	Tasmania Univ., Tasmania	
TOL	0.169	-0.135	-0.143	-0.027	-4.049	39.881	Toledo, Spain	
TRI	-0.144	0.027	0.215	0.016	13.764	45.709	Trieste, northern Italy	
TRN	0.064				-61.403	10.649	Trinidad, Trinidad	
TUC	0.008				-110.782	32.310	Tucson, eastern Arizona	
UMF	0.144	0.385	0.058	0.040	20.237	63 815	Umea, Sweden	
UNM	-0.266				-99.178	19.329	Nat Univ. of Central Mexico	
VAL	-0.027	-0.070	0.237	0.050	-10.244	51.939	Valentia, Eire	
WEL	0.107				174.768	-41.286	Wellington, New Zealand	
WES	-0.216	-0.014	-0.309	-0.141	-71.322	42.385	Weston, New England	
WIN	-0.154	0.151	0.009	-0.055	17.100	-22.567	Windhoek, South-West Africa	

Table 2 lists the results of applying the receiver and near-source corrections (as shown in Table 1) to the 82 Semipalatinsk explosions in our database. For 79 out of 82 Semipalatinsk events used in this study, the final  $\sigma$  is typically around or below the same level that a "single-event MLE with primary correction only" could achieve. The only three events which do not show reduction in the variance are 770730D, 730723B, and 880914B (Table 2). A plausible explanation is that perhaps these three events are located in very different geological or geophysical environments from other events in the same testing area. Or, perhaps the 26 WWSSN stations for which the filmchips of JVE were available only cover a small portion of the focal sphere, and hence the focusing/defocusing effect is not fully accounted for. Most of the  $m_{2.9}$  in Table 2 have a standard error around 0.02 m.u., which is about the same as that for *RMS L<sub>g</sub>* values inferred from an in-country regional network (*e.g.*, Israelson, 1991; Hansen *et al.*, 1990).

Another observation is that the resulting  $m_b$  values,  $\overline{m}_{2.9}$  are essentially the same as those inferred from the GLM or the single-event MLE,  $\overline{m}_{2.2}$ . Figures 6 through 10 plot out the three different  $m_b$ s for five arbitrarily selected events. The solid line and the dashed lines represent the mean network-averaged  $m_b$  and the associated standard deviation of station  $m_b$ . The upward arrows represent the lower bounds of the station  $m_b$ , which came from those clipped measurements. The "Y" symbols are the upper bounds of the station  $m_b$  which are resulted from those noisy measurements. Both the "uncensored" (shown in filled circles) and "censored" station  $m_b$ s are used in computing the mean station residuals with the maximum-likelihood scheme described in Jih and Shumway (1989).

- 19 -

Table 2. Comparison of Network-Averaged m <sub>b</sub> with Various Corrections							
Event	# of signals	Without correction	Station corrected	Near-source corrected			
Date		<i>Π</i> 1, σ	<i>Μ</i> 2.2, σ	<i>Π</i> <sub>2.9</sub> , σ			
661218M	51 2 1	5.747±0.038 0.281	5.753±0.033 0.239	5.738±0.022 0.161			
670916M	36 18 2	5.097±0.041 0.305	5.110±0.036 0.268	5.095±0.018 0.137			
670922M	35 20 1	5.033±0.036 0.271	5.048±0.029 0.214	5.029±0.017 0.125			
671122M	7 52 0	4.169±0.054 0.413	4.318±0.039 0.299	4.231±0.013 0.099			
690531M	30 21 0	4.965±0.050 0.357	5.006±0.039 0.276	5.026±0.015 0.110			
691228M	45 2 3	5.666±0.043 0.306	5.665±0.035 0.250	5.660±0.018 0.125			
700721M	38 12 1	5.182±0.043 0.307	5.199±0.033 0.236	5.184±0.018 0.125			
701104M	38 12 1	5.267±0.042 0.303	5.279±0.032 0.232	5.249±0.020 0.145			
710606M	38 6 2	5.323±0.040 0.272	5.341±0.031 0.210	5.319±0.015 0.099			
710619M	41 6 0	5.294±0.040 0.278	5.311±0.030 0.208	5.287±0.013 0.086			
711009M	27 9 3	5.165±0.037 0.233	5.187±0.028 0.174	5.136±0.016 0.100			
711021M	32 6 0	5.348±0.049 0.301	5.383±0.036 0.224	5.341±0.021 0.127			
720826M	29 10 2	5.168±0.045 0.290	5.186±0.030 0.193	5.163±0.017 0.111			
720902M	15 25 0	4.615±0.049 0.312	4.635±0.038 0.243	4.602±0.017 0.107			
651121D	48 12 1	5.384±0.035 0.271	5.394±0.024 0.188	5.381±0.019 0.152			
660213D	51 2 10	6.073±0.038 0.305	6.092±0.027 0.218	6.088±0.014 0.114			
660320D	49 6 8	5.848±0.041 0.322	5.865±0.030 0.239	5.853±0.009 0.074			
660507D	9 23 1	4.517±0.043 0.250	4.559±0.031 0.181	4.456±0.014 0.078			
661019D	51 8 5	5.534±0.035 0.276	5.542±0.025 0.201	5.525±0.014 0.111			
670226D	4876	5.826±0.044 0.341	5.843±0.035 0.274	5.854±0.011 0.086			

Table 2. Comparison of Network-Averaged m <sub>b</sub> with Various Corrections							
Event	# of signals	Without correction	Station corrected	Near-source corrected			
Date		<i>Μ</i> <sub>1</sub> , σ	m <sub>2.2</sub> , σ	<i>m</i> <sub>2.9</sub> , σ			
680929D	50 4 6	5.610±0.035 0.275	5.642±0.026 0.202	5.641±0.018 0.138			
690723D	38 17 1	5.172±0.041 0.306	5.201±0.029 0.220	5.186±0.013 0.100			
690911D	19 35 0	4.533±0.053 0.392	4.605±0.040 0.295	4.634±0.025 0.186			
710322D	43 11 3	5.498±0.040 0.305	5.528±0.032 0.245	5.530±0.017 0.126			
710425D	37 3 0	5.764±0.052 0.327	5.793±0.042 0.267	5.826±0.020 0.127			
711230D	16 2 0	5.522±0.060 0.255	5.556±0.062 0.262	5.556±0.035 0.148			
720328D	28 15 0	4.955±0.048 0.313	4.997±0.034 0.226	4.984±0.021 0.141			
720816D	23 20 1	4.908±0.044 0.293	4.931±0.033 0.221	4.921±0.025 0.165			
721210D	30 6 5	5.512±0.045 0.287	5.543±0.033 0.209	5.555±0.029 0.183			
770329D	25 12 0	4.974±0.055 0.332	5.015±0.042 0.256	4.992±0.042 0.254			
770730D	21 14 0	4.858±0.051 0.301	4.881±0.049 0.287	4.855±0.052 0.306			
780326D	25 4 0	5.461±0.052 0.279	5.519±0.037 0.198	5.476±0.018 0.096			
780422D	21 7 0	4.985±0.057 0.301	5.032±0.044 0.231	5.010±0.023 0.120			
780728D	36 7 6	5.506±0.042 0.291	5.525±0.028 0.198	5.494±0.012 0.086			
800522D	36 20 1	5.138±0.033 0.250	5.156±0.025 0.193	5.131±0.012 0.089			
650115B	46 1 2	5.896±0.040 0.280	5.893±0.032 0.223	5.861±0.022 0.155			
680619B	28 3 2	5.276±0.047 0.270	5.285±0.035 0.203	5.229±0.026 0.151			
691130B	51 0 0	5.950±0.044 0.313	5.973±0.031 0.222	5.945±0.026 0.184			
710630B	31 18 1	5.045±0.043 0.301	5.075±0.035 0.247	5.043±0.027 0.192			
720210B	34 8 2	5.297±0.042 0.278	5.329±0.029 0.195	5.289±0.018 0.121			

Table 2. Comparison of Network-Averaged m <sub>b</sub> with Various Corrections							
Event	# of signals	Without correction	Station corrected	Near-source corrected			
Date		<i>m</i> <sub>1</sub> , σ	<i>m</i> <sub>2.2</sub> , σ	<i>m</i> <sub>2.9</sub> , σ			
721102B	42 0 15	6.173±0.045 0.339	6.183±0.034 0.256	6.160±0.023 0.177			
721210B	45 1 11	6.006±0.037 0.282	6.013±0.029 0.223	5.983±0.022 0.169			
730723B	54 0 1	6.174±0.042 0.314	6.202±0.030 0.220	6.179±0.031 0.231			
731214B	50 7 6	5.750±0.044 0.348	5.769±0.036 0.287	5.749±0.030 0.241			
750427B	18 1 1	5.457±0.099 0.444	5.480±0.089 0.397	5.436±0.072 0.322			
760704B	38 0 5	5.819±0.062 0.404	5.849±0.048 0.313	5.829±0.027 0.180			
761207B	17 2 1	5.571±0.107 0.480	5.611±0.099 0.442	5.550±0.078 0.347			
780611B	17 0 1	5.882±0.048 0.205	5.873±0.042 0.179	5.804±0.038 0.160			
780915B	37 1 6	5.826±0.057 0.379	5.850±0.043 0.284	5.828±0.030 0.199			
790623B	40 2 3	6.015±0.052 0.349	6.071±0.040 0.267	6.069±0.036 0.240			
790804B	40 4 20	6.064±0.037 0.295	6.086±0.027 0.212	6.103±0.016 0.126			
791028B	44 6 13	5.909±0.036 0.284	5.941±0.024 0.190	5.933±0.019 0.151			
791223B	41 3 17	6.111±0.033 0.260	6.128±0.020 0.154	6.131±0.018 0.141			
800914B	35 4 6	6.006±0.062 0.417	6.041±0.051 0.345	6.053±0.043 0.287			
811018B	41 3 7	5.954±0.039 0.279	5.976±0.029 0.210	5.996±0.023 0.164			
840526B	31 0 3	5.966±0.058 0.338	5.993±0.043 0.252	6.009±0.022 0.130			
880914B	25 0 1	6.004±0.037 0.191	6.032±0.023 0.117	6.026±0.033 0.168			
761123B	22 0 0	5.577±0.075 0.354	5.626±0.065 0.305	5.687±0.053 0.249			
780829B	16 0 0	5.869±0.079 0.315	5.905±0.075 0.302	5.936±0.045 0.182			
781129B	2800	5.840±0.068 0.358	5.880±0.054 0.287	5.895±0.026 0.138			
790707B	30 0 0	5.734±0.048 0.261	5.800±0.043 0.236	5.827±0.031 0.169			

	Table 2. Comparison of Network-Averaged m <sub>b</sub> with Various Corrections							
Event	# of signals	Without correction	Station corrected	Near-source corrected				
Date		<i>π</i> 1, σ	m <sub>2.2</sub> , σ	<i>m</i> <sub>2.9</sub> , σ				
790818B	28 0 0	6.023±0.065 0.344	6.087±0.052 0.273	6.088±0.031 0.166				
791202B	15 0 0	5.807±0.112 0.435	5.874±0.089 0.344	5.892±0.067 0.259				
801012B	23 0 0	5.804±0.087 0.417	5.828±0.070 0.334	5.872±0.047 0.228				
801214B	29 0 0	5.873±0.053 0.288	5.911±0.047 0.252	5.935±0.030 0.162				
801227B	24 0 0	5.855±0.056 0.273	5.896±0.042 0.208	5.892±0.034 0.165				
810422B	25 0 0	5.826±0.070 0.350	5.865±0.057 0.287	5.922±0.035 0.174				
810913B	1700	6.014±0.093 0.383	6.024±0.064 0.265	6.060±0.023 0.096				
811227B	23 0 0	6.187±0.079 0.380	6.207±0.060 0.287	6.189±0.036 0.174				
820425B	1400	5.924±0.099 0.372	5.944±0.072 0.269	5.982±0.040 0.150				
820704B	21 0 0	6.073±0.053 0.245	6.089±0.048 0.222	6.095±0.032 0.145				
821205B	26 0 0	6.043±0.064 0.325	6.093±0.053 0.271	6.109±0.037 0.191				
830612B	16 0 0	5.921±0.091 0.365	5.943±0.069 0.276	5.934±0.046 0.186				
831006B	25 0 0	5.885±0.064 0.318	5.942±0.048 0.238	5.935±0.032 0.162				
831026B	1800	5.933±0.070 0.298	5.941±0.053 0.225	5.998±0.034 0.146				
840425B	21 0 0	5.850±0.099 0.453	5.892±0.082 0.374	5.905±0.043 0.196				
840714B	23 0 0	5.920±0.088 0.424	5.999±0.074 0.357	6.054±0.043 0.207				
841027B	1900	6.141±0.078 0.340	6.150±0.066 0.289	6.191±0.034 0.146				
841202B	22 0 0	5.630±0.065 0.305	5.693±0.057 0.266	5.720±0.044 0.206				
841216B	1500	5.911±0.107 0.415	5.993±0.070 0.272	6.043±0.035 0.135				
841228B	1900	5.853±0.077 0.335	5.916±0.053 0.230	5.945±0.037 0.162				
850615B	1500	6.016±0.078 0.301	6.069±0.049 0.191	6.060±0.035 0.134				





Figure 6. Scatter plot of 3 different types of station  $m_b$ s for Murzhik explosion 710606. The 38 good recordings, 6 noise, and 2 clips are shown with filled circles, Y-shaped downward arrows and upward arrows, respectively. The raw station  $m_b$ s (top) has a standard deviation of 6.27 mu. Applying the "primary" station corrections reduces the scatter to 0.21 m.u. Applying the proposed "secondary" station corrections to count for the near-source focusing/defocusing effects would further reduce the scatter down to 0.1 m.u. The dashed lines around the network-averaged  $m_b$  clearly illustrate the remarkable reduction of fluctuation across the recording stations. The mean event  $m_b$  itself is not significantly changed, however.



VARIOUS WWSSN MAGNITUDES OF EVENT 710619K

Figure 7. Same as Figure 6 except for Murzhik event 710619.



VARIOUS WWSSN MACNITUDES OF EVENT 711009K

Figure 8. Same as Figure 6 except for Murzhik event 711009.



VARIOUS WWSSN MACNITUDES OF EVENT 660320D

Figure 9. Same as Figure 6 except for Degelen event 660320. The near-source correction proposed in this study not only reduced the  $m_b$  scatter at stations that reported the good signals, but also improved the data consistence of the censored recordings.



VARIOUS WWSSN MAGNITUDES OF EVENT 710425D

Figure 10. Same as Figure 6 except for Degelen event 710425.

### **I.4 UNDERLYING MODEL AND ADVANTAGES OF THE NEW MAGNITUDE**

We now examine the fundamental difference between the present scheme and the previous ones. In LSMF and the standard GLM scheme (Douglas, 1966; Blandford and Shumway, 1982; Jih and Shumway, 1989; Murphy *et al.*, 1989), it is assumed that the observed station  $m_b(i,j)$  is the sum of the true source size of the i-th event, E(i), the receiver term of the j-th station, S(j), and the random noise, v(i,j):

$$m_b(i,j) = E(i) + S(j) + v(i,j)$$
 [4]

The receiver term, S(j), is constant with respect to all explosions from many azimuths, and hence it would inherently reflect the "averaged" receiver effect --- provided the paths reaching the station have broad azimuthal coverage. These receiver corrections correlate with the upper mantle property underneath the receivers (North, 1977). When world-wide explosions are used, the standard deviation of the noise v in [4] is typically about 0.3 m.u. or larger.

If LSMF or GLM is applied to events within a smaller area of source region, then the  $\sigma$  could reduce to 0.15 or 0.2 m.u. Unfortunately, there are severe drawbacks associated with such "single-test-site GLM" approach. First of all, the station corrections will not necessarily represent the attenuation underneath the receiver side. They could be contaminated or even overwhelmed by the near-source effects shared by the explosions confined in a narrow azimuthal range. This explains the phenomenon Butler (1981) and Burdick (1981) reported that using Soviet explosions exclusively may fail to discern the attenuation differential between the eastern and western U.S. Secondly, when the "single-test-site GLM" inversion is applied to several test sites separately, there may not be a consistent baseline for magnitude comparison or absolute yield estimation, since the station terms are inherently inconsistent.

In the present scheme ([3]), however, we reformulate the whole model as

$$m_b(i,j) = E(i) + S(j) + F(k(i),j) + v(i,j)$$
 [5]

where F(k(i),j) is the correction term at the j-th station for the near-source

- 29 -

focusing/defocusing effect, which is constant for all events in the k-th "geologically and geophysically uniform region". For each seismic station, this F can be regarded as its azimuthal variation around the mean station term S. However, as we already explained, it would be more appropriate to consider F the near-source term because the back azimuths at the station could be nearly identical for adjacent test sites (such as Degelen and Murzhik), and yet the "F" terms could be very different. By incorporating the F term into the model, the  $\sigma$  for world-wide explosions is reduced to about 0.2, roughly the same level that which a "single-test-site GLM" could achieve. Intuitively, the present scheme (Equation [5]) provides a more detailed (and better) model than that of Equation [4] in describing the whole propagation path from the source towards the receiver. Simply put, Equation [4] yields a stronger fluctuation in the source terms, E, as well as a larger standard deviation of v because each term in the right-hand side of Equation [4] would have to "absorb" part of the missing F term in [5]. This is exactly the same reason why  $m_{2.2}$  has smaller variation than  $m_1$  since the latter would be interfered by the missing station term S in Equation [1].

For actual implementation, the present scheme can be replaced with an equivalent multi-stage procedure as follows. First, a set of station corrections (the so-called "primary corrections") is determined with one GLM. Then the "secondary correction" at each station is defined as the mean of all residuals of all events from the same test site (or the same geologic/geophysical regime) recorded at this particular station.

#### 1.5 MAGNITUDE: YIELD RELATIONSHIP AT SEMIPALATINSK WITH m29

To further demonstrate that  $m_{2,9}$  would provide more precise yield estimates, the 19 Semipalatinsk explosions for which the yields are published by Bocharov *et al.* (1989) are used as a test case. Table 3 gives the date, various  $m_b$  values, and the associated standard errors. Table 4 lists the Soviet-published yields and the

- 30 -
postulated uncertainties. We assume that these yields are subject only to 10% standard errors (S.E.) and/or the rounding. For each  $(m_b, yield)$  pair, we use a random number generator to produce a perturbed ( $m_b$ , yield) pair according to their uncertainty distribution. A standard least-squared regression is performed for each data set of perturbed samples. The procedure is repeated for several hundred iterations, and the resulting calibration curves (*i.e.*, the straight best-fitting lines) are shown as the darkened bundle in Figures 11 through 13. The detail of this generalized "doublyweighted least-squares scheme" is discussed in Jih (1991). Here we only summarize the results to illustrate the advantages of  $m_{2,9}$  relative to the more conventional source measures. For comparison, regression result using RMS  $L_{a}$  reported at NORSAR (Ringdal, 1990) is also included in Tables 5 and 6 (Figure 14). Note that for the Soviet JVE shot (880914B), the yield is assumed to be 119 kt after Gordan (1988) (see also Sykes and Ekstrom, 1989; Priestley et al., 1990). The regression result based on  $\overline{m}_{2,9}$  has a smaller  $m_b$  scatter around the mean calibration curve (and hence a smaller uncertainty factor in the yield estimates) than those based on  $\overline{m}_1$  and  $\overline{m}_{2,2}$ , as expected. It has a precision very close to that based on NORSAR RMS  $L_{\alpha}$  over a wide range of yields. The precision is also very similar to what Patton (1988) found for 69 NTS explosions below the water table with  $L_{a}$  recorded at LLNL digital network. Note that the uncertainty factor in the yield estimates is yield dependent. It is smaller near the centroid of the data set (*i.e.*, around 50 kt in our case) and larger at both ends. This is contrary to a general perception that we might know yields much better at higher values around 150 kt. For yields below 10 kt there is no data point in NORSAR'S RMS  $L_q$  data set, and hence there is a much larger uncertainty than that based on  $m_b$ .

Table 3. Various Network-Averaged $m_b$ of 19 Special Events							
Event	# of signals	Without correction	Station corrected	Near-source corrected			
Date		<i>π</i> <sub>1</sub> , σ	<i>̄</i> m <sub>2.2</sub> , σ	<i>̄</i> m <sub>2.9</sub> , σ			
651121D	48 12 1	5.384±0.035 0.271	5.394±0.024 0.188	5.381±0.019 0.152			
660213D	51 2 10	6.073±0.038 0.305	6.092±0.027 0.218	6.088±0.014 0.114			
660320D	49 6 8	5.848±0.041 0.322	5.865±0.030 0.239	5.853±0.009 0.074			
670922M	35 20 1	5.033±0.036 0.271	5.048±0.029 0.214	5.029±0.017 0.125			
680929D	50 4 6	5.610±0.035 0.275	5.642±0.026 0.202	5.641±0.018 0.138			
690723D	38 17 1	5.172±0.041 0.306	5.201±0.029 0.220	5.186±0.013 0.100			
691130B	51 0 0	5.950±0.044 0.313	5.973±0.031 0.222	5.945±0.026 0.184			
691228M	45 2 3	5.666±0.043 0.306	5.665±0.035 0.250	5.660±0.018 0.125			
710425D	37 3 0	5.764±0.052 0.327	5.793±0.042 0.267	5.826±0.020 0.127			
710606M	38 6 2	5.323±0.040 0.272	5.341±0.031 0.210	5.319±0.015 0.099			
711009M	2793	5.165±0.037 0.233	5.187±0.028 0.174	5.136±0.016 0.100			
711021M	32 6 0	5.348±0.049 0.301	5.383±0.036 0.224	5.341±0.021 0.127			
720210B	34 8 2	5.297±0.042 0.278	5.329±0.029 0.195	5.289±0.018 0.121			
720816D	23 20 1	4.908±0.044 0.293	4.931±0 033 0.221	4.921±0.025 0.165			
720902M	15 25 0	4.615±0.049 0.312	4.635±0.038 0.243	4.602±0.017 0.107			
721102B	42 0 15	6.173±0.045 0.339	6.183±0.034 0.256	6.160±0.023 0.177			
721210B	45 1 11	6.006±0.037 0.282	6.013±0.029 0.223	5.983±0.022 0.169			
880914B	25 0 1	6.004±0.037 0.191	6.032±0.023 0.117	6.026±0.033 0.168			

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Table 6. 95% Confidence Scatter in m <sub>b</sub>								
(assuming y	rields are si	ubject to rou	inding and 1	0% S.E.)	• ,			
m <sub>b</sub> used 1 kt 10 kt 50 kt 100 kt 150 k								
m <sub>1</sub>	0.27	0.16	0.14	0.15	0.17			
m <sub>2.2</sub>	0.24	0.14	0.11	0.13	0.14			
m <sub>2.9</sub>	0.21	0.11	0.09	0.11	0.13			
NORSAR RMS Lg"	0.27	0.13	0.08	0.09	0.11			
(assum	ing yields a	re subject to	o 10% S.E.	only)				
m <sub>1</sub>	0.24	0.16	0.13	0.15	0.16			
<i>₩</i> 2.2	0.21	0.13	0.11	0.13	0.14			
<i>m</i> <sub>2.9</sub>	0.18	0.10	0.08	0.10	0.11			
NORSAR RMS Lg	0.27	0.13	0.08	0.09	0.11			

\*) 19 Semipalatinsk shots in Table 3 used as calibration events.

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\*\*) 9 Semipalatinsk events with RMS  $L_g$  reported in Ringdal (1990) and Ringdal and Marshall (1989).





Figure 11. Regressing the simple network-averaged  $m_b$  (*i.e.*,  $m_1$ ) on the 19 Soviet-published yields. The yields are assumed to be subject to 10% standard errors. The uncertainties in the  $m_b$ s and the yields are taken into account through 800 bootstrap resamplings. The darkened bundle is actually the collection of all 800 regressions, each produced by a possible realization of 19 perturbed ( $m_b$ , yield) pairs. The 95% confidence band (shown as 2 curves around the darkened bundle) is narrower near the centroid and wider towards both ends, as expected. The individual 95% confidence intervals of the two inferred parameters (*i.e.*, the slope and the intercept of the calibration curve) are shown with the dashed line in the scatter plot (bottom). Note that the dashed rectangle is not the joint 90% confidence interval, however, due to the highly correlated nature of the two parameters.



Figure 12. Same as Figure 11 except the  $m_b$ s are those with station corrections applied, namely the  $m_{2,2}$ . Note that the yields are assumed to have 10% standard error as in Figure 11, and the reduction in the scatter is due to the better precision of  $m_{2,2}$ .



Scatter Plot of Inferred Parameters



[97.5% quantile of t(17. D.o.F.), 2.110, used]

Figure 13. Same as Figure 11 except the  $m_b$ s are those with both the station corrections and near-source focusing terms applied, namely the  $m_{2,9}$ . The improved  $m_b$  precision has direct impact on the regression, as compared to Figures 11 and 12. The yield of future Semipalatinsk explosions can be reliably predicted using this  $m_{2,9}$  yield calibration curve for a precision similar to what RMS  $L_q$  could provide (cf. Figure 14).





Figure 14. Regressing 9 RMS  $L_g$  values (Ringdal, 1990) on Soviet-published yields. RMS  $L_g$  recorded at NORSAR has very high S/N ratio and hence very stable source measure for Semipalatinsk explosions above 10KT. There is no calibration data below 10KT and hence the extrapolation for future events with RMS  $L_g$  would have inherently larger uncertainty, as illustrated by the much wider 95% confidence band. Thus either teleseismic records based on P phases or  $L_g$  measured at in-state stations must be used for low-yield events.

### **I.6 GEOPHYSICAL INTERPRETATION OF OUR NEAR-SOURCE CORRECTIONS**

If we remove the globally-averaged source sizes from the station-corrected magnitudes, all three test sites would exhibit different azimuthal and radial amplitude variations (Figure 5): Degelen and Murzhik events are systematically enhanced in the western U.S. and reduced in eastern U.S., whereas Balapan events are all reduced in the whole U.S. Degelen events are reduced in Indonesia and southeast Africa, whereas Balapan events are enhanced in these regions. Murzhik events are reduced in Scandinavia, but Balapan and Degelen events get enhanced there. Such highly direction-dependent, distance-dependent, and site-dependent patterns of the amplitude fluctuation could be a diagnostic for the path effects in the proximity of the test sites. Back projections (e.g., Lynnes and Lay, 1990) of the m<sub>b</sub> residuals onto the upper mantle and the lower crust reveal that similar m<sub>b</sub> residuals come into alignment in several regions partitioned by known geological features (Figure 15). Murzhik events recorded in the western U.S. and in northeast Asia, Degelen events in the western U.S., and SW Balapan events at western European stations must pass through the area between Chinrau fault and Chingiz-Kalba shear zone. All these paths show positive  $m_b$  residuals. The north of Chinrau fault might have smaller  $P_n$ velocity and higher heat flow (Bonham et al., 1980; Leith, 1987a, 1987b) and has negative mean  $m_b$  residuals on the back projections. Paths from NE Balapan to North America and many continental European stations must cross this area or even travel along the Chinrau fault before entering deeper mantle, and hence the complexity in the waveforms is inevitable. It seems that the mean  $m_b - L_g$  separation of 0.14±0.02 m.u. (e.g., Ringdal and Hokland, 1987; Ringdal and Marshall, 1989; Richards et al., 1990; Jih and Wagner, 1990) between the NE and SW subregions of Balapan could be due in part to the path effects --- in addition to the difference of source medium postulated previously by Marshall et al. (1984). Path effects can also explain why the SW Balapan waveforms tend to be more complex at YKA than those recorded at WRA, EKA, and GBA arrays (Jih and Wagner, 1991).

The initial *P* waves from the three adjacent test sites have virtually the same incident angle at each teleseismic station, and anything in common across all events (such as the crustal amplification as well as the upper mantle attenuation underneath the receiver) would have been lumped into the constant station term. Thus the station residuals averaged over all events from the same test site would correlate very little with the receiver. Instead, they should reveal more site-dependent information about the focusing/defocusing pattern underneath E. Kazakhstan (Figure 15).

The largest and prominent fault in the region is the southeast-trending Chingiz right-lateral strike-slip fault that passes about 10 km southwest of Degelen Mountain and right across the Murzhik test area (Rodean, 1979; Bonham *et al.*, 1980; Leith, 1987b). Soviets reported that this fault has a very steep dip, which is consistent with its linear expression over large distance as seen on Landsat imagery (Bonham *et al.*, 1980). A distinct fault-line scarp is developed along much of the oldest metamorphic rocks. Chingiz Fault extends for a total length of about 700 km. Soviet reports postulate that this fault extends down to the boundary of the granite layer of the crust and possibly into the upper mantle. For Murzhik explosions, the propagation of  $P_n$  and  $L_g$  waves could be affected by this fault significantly, which results in a radiation pattern such as we are observing. More specifically, the rays towards NW direction could be reflected or diffracted to other quadrants, due to its post-critical incidence angles. Such relatively distant crustal structure should have little impact on the first P waves of Balapan explosions at teleseismic distances, however.



GRIDDED LS-AVERAGED mb[Pmax] RESIDUALS OF E. KAZAKH SHOTS

**Figure 15.** Back-projected  $m_b$  residuals averaged over a grid of every 0.2 degree by 0.2 degree underneath Soviet's Semipalatinsk nuclear test sites. The averaged residual pattern show some association with local geological features if the residuals are projected to depths down to the lower crust and the upper mantle. Thus even the teleseismic recordings might reveal some useful information about the path effects on the propagation of regional phases such as  $P_a$  and  $L_g$ .

### **1.7 DISCUSSION AND CONCLUSIONS**

The new magnitude determination scheme (Equation [3]) presented in this study significantly reduces the fluctuational variation across the recording stations. It is shown that, by applying this scheme to worldwide explosions, it is possible to have a consistent base line in estimating the absolute magnitudes (which is crucial in estimating the test site bias) while the precision in the resulting network  $m_b$  values can be maintained as well as could be achieved by the single-test-site approach. The standard error in most  $\overline{m}_{2.9}$  values is 0.02 m.u., about the same as that for *RMS Lg* inferred from in-country regional network recordings reported by Israelson (1991) and Hansen *et al.* (1990). Most Murzhik events show nearly identical residual patterns, suggesting a common focusing/defocusing structure. Further partitioning of Balapan test site seems necessary, however.

The most detailed description of the wave propagation model would naturally suggest that yet another term could be added into Equation [3] to count for the sourceregion attenuation. So far such source region bias in  $m_b$  has always been inferred with other information such as the yields (as in this study),  $M_S$  (e.g., Evernden and Marsh, 1987) or  $P_n$  velocity (e.g., Marshall *et al.*, 1979) *etc.*. The hypothesis that such attenuation differential could be directly discerned from  $m_b$  alone by further improving Equation [3] is worth testing.

Digital signals recorded on seismic instruments at regional distance will be critical for monitoring low yield explosions below 10 kt. Obviously the regression routines developed in this study can well be applied to other source measures which are based on regional phases. On the other hand, teleseismic data such as the WWSSN data used in this study still carry invaluable information that is worthy of further exploitation, beyond simply calculating the "unified yield".

Throughout this study, our emphasis has been to reveal the site-dependent characteristics from the observations exclusively so that in the future the inferred results can be critically examined and compared with those derived by other means.

- 43 -

We suggest that the follow-up research be accompanied by well-constrained forward modeling studies using realistic structures. The upgraded linear finite-difference code which incorporates boundary conditions for topographical free-surface of arbitrary shape (Jih *et al.*, 1988) in addition to the "strain filter" and "marching grid" features as outlined in Jih *et al.* (1989) can be utilized in the future to improve our understanding of the fundamental issues of seismic energy partitioning on the focal sphere as well as their implications for yield determination.

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#### **APPENDIX A**

# GEOTECH'S MAXIMUM-LIKELIHOOD NETWORK m<sub>h</sub>, GLM91A

Short-period WWSSN vertical recordings (SPZ) of body waves from Soviet nuclear explosions detonated at the Semipalatinsk Test Site, Eastern Kazakhstan, USSR, are being measured and added to our database to determine the optimal network magnitudes using the maximum-likelihood estimator (MLE), which accounts for the effects of data censoring due to clipping and to noise (Blandford and Shumway, 1982; Jih and Shumway, 1989). As of now, our WWSSN database has been expanded to 192 events (totaling 515 usable "a", "b", and "max" event phases) from a variety of test sites. Only the stations at teleseismic distance (20 to 95 degrees) which recorded 7 or more good signals were used in the network  $m_b$  determination. Although we have also included some measurements made off LRSM tapes and CDSN recordings, most of those data failed to meet the criteria aforementioned. The 12170 good signals, 8047 noise measurements, and 1330 clipped recordings yield a  $\hat{\sigma}_{MLE}$  of 0.300.

The 192 events in Table A.2 are grouped by test sites. 111 events were measured before 1/1/90 under various contracts during the past decade. 25 Balapan events (with prefix "SAF") were based on the raw WWSSN station magnitudes distributed by DARPA in 1988. The three numbers under the column "# of signals" represent the number of signals, noise, and clips associated with the  $P_{max}$  phase of each event. S.E.M. is the "standard error in the mean" of the event magnitudes. Except for the U.S. and French Sahara explosions which have specific code names, all the remaining events are identified with the dates and abbreviated test site codes shown below:

Table A.1. Geotech's m <sub>b</sub> Database							
Code	Number of	Events	Nuclear Test Site				
	1/1/90	7/15/91					
	19	37	Nevada Test Site, U.S.A.				
	6	6	Outside Nevada Test Site, U.S.A.				
	3	3	Amchitka Island, Aleutians, U.S.A.				
AZG	11	11	Azgir, U.S.S.R.				
PNE	1 2		"PNE", U.S.S.R.				
MEK	0	14	Murzhik (Konystan), E. Kazakh, U.S.S.R.				
DEK	9	21	Degelen Mountain, E. Kazakh, U.S.S.R.				
SEK	12	22	Balapan (Shagan River), E. Kazakh, U.S.S.R.				
SAF	0	25	Balapan (Shagan River), E. Kazakh, U.S.S.R.				
NNZ	18	18	Northern Novaya Zemlya, U.S.S.R.				
SNZ	6	6	Southern Novaya Zemlya, U.S.S.R.				
	9	9	Ahaggar, French Sahara				
TU	11	11	Tuamoto Islands, France				
RAJ	1	1	Rajasthan, India				
СН	6	6	Lop Nor, Sinkiang, China				

Table A.2. Geotech's Maximum-Likelihood Network mb						
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	$m_b(P_a)$	
ALMENDRO	26 0 2	0.057	6.229	6.021	5.732	
BANEBERRY	14 30 0	0.045	4.869	4.547	4.373	
BENHAM	42 1 7	0.042	6.392	6.140	5.829	
BILBY	36 3 0	0.048	5.706	5.453	5.201	
BOURBON	18 31 0	0.043	4.931	4.751	4.621	
BOXCAR	32 0 4	0.050	6.443	6.220	5.887	
CALABASH	36 17 0	0.041	5.551	5.357	5.180	
CAMBRIC	14 35 0	0.043	4.576	4.310	4.012	
CARPETBAG	37 7 1	0.045	5.806	5.585	5.352	
CHANCELLOR	15 11 1	0.058	5.360	5.201	4.924	
CHARTREUSE	31 16 1	0.043	5.260	5.029	4.909	
CHATEAUGAY	17 28 2	0.044	5.080	4.898	4.509	
COMMODORE	31 5 1	0.049	5.794	5.585	5.361	
CORDUROY	18 14 0	0.053	5.324	5.131	5.013	
DISCUSTHROWER	12 39 1	0.042	4.677	4.451		
DURYEA	23 29 0	0.042	5.049	4.874	4.723	
FLASK	36 8 0	0.045	5.509	5.221	5.038	
GREELEY	49 2 2	0.041	6.340	6.143	5.909	
HALFBEAK	43 2 2	0.044	6.113	5.811	5.583	
HANDCAR	16 33 0	0.043	4.650	4.516	4.345	

Table A.2. Geotech's Maximum-Likelihood Network mb (Continued)							
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	m <sub>b</sub> (P <sub>a</sub> )		
HANDLEY	41 1 1	0.046	6.519	6.345	6.100		
HARZER	31 5 1	0.049	5.549	5.327	5.031		
KANKAKEE	24 27 0	0.042	4.875	4.624	4.390		
KNICKERBOCKER	28 21 0	0.043	5.253	4.976	4.778		
MAST	29 1 0	0.055	6.040	5.800	5.465		
MINIATA	37 7 0	0.045	5.491	5.176	4.908		
NASH	31 21 0	0.042	5.166	4.939	4.789		
PALANQUIN	200	0.212	3.942				
PILEDRIVER	40 12 2	0.041	5.480	5.243	4.996		
PURSE	900	0.100	5.880	5.571	5.296		
REX	16 35 1	0.042	4.778	4.442	3.952		
SCAUP	210	0.173	4.625	4.305	4.247		
SCHOONER	790	0.075	4.389	4.371	3.869		
SCOTCH	38 8 1	0.044	5.643	5.386	5.133		
SCROLL	200	0.212	4.077	3.642			
STARWORT	21 6 0	0.058	5.474	5.162	4.937		
STILTON	700	0.114	5.839	5.663	5.455		
CANNIKIN	49 0 20	0.036	6.957	6.710	6.463		
LONGSHOT	71 4 3	0.034	5.873	5.494	5.137		
MILROW	52 0 4	0.040	6.544	6.245	6.000		

Table A.2. Geotech's Maximum-Likelihood Network mb (Continued)						
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	$m_b(P_a)$	
FAULTLESS	47 1 3	0.042	6.497	6.193	5.869	
GASBUGGY	11 37 0	0.043	4.690	4.438	4.197	
RIOBLANCO	15 20 0	0.051	4.831	4.568	4.127	
RULISON	9 37 0	0.044	4.595	4.287	4.161	
SALMON	6 33 0	0.048	4.200	3.989	3.484	
SHOAL	16 27 0	0.046	4.776	4.497	4.346	
AZG22APR66	3 10 0	0.083	4.225	4.144	3.919	
AZG01JUL68	44 10 3	0.040	5.542	5.245	4.932	
AZG22DEC71	12 0 2	0.080	6.181	5.845	5.490	
AZG25APR75	1 16 0	0.073	3.986	3.948		
AZG29JUL76	41 5 7	0.041	5.877	5.594	5.133	
AZG30SEP77	21 30 1	0.042	4.855	4.619	4.092	
AZG17OCT78	705	0.087	6.108	5.733	5.294	
AZG18DEC78	903	0.087	6.155	5.780	5.406	
AZG17JAN79	10 0 4	0.080	6.170	5.881	5.524	
AZG14JUL79	10 0 1	0.091	5.725	5.396	4.866	
AZG24OCT79	306	0.100	5.942	5.678	4.865	

Table A.2. Geotech's Maximum-Likelihood Network mb (Continued)							
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	$m_b(P_a)$		
PNE21MAY68	41 9 1	0.042	5.301	5.089	4.858		
PNE29AUG74	27 18 0	0.045	4.753	4.433	4.041		
KON18DEC66	55 8 1	0.038	5.726	5.511	5.280		
KON16SEP67	36 29 2	0.037	5.086	4.852	4.558		
KON22SEP67	35 31 1	0.037	5.013	4.757	4.447		
KON22NOV67	7 64 0	0.036	4.317	4.001			
KON31MAY69	30 31 0	0.038	4.990	4.775	4.398		
KON28DEC69	45 9 3	0.040	5.652	5.468	5.192		
KON21JUL70	38 21 1	0.039	5.178	4.933	4.592		
KON04NOV70	38 22 1	0.038	5.242	5.053	4.844		
KON06JUN71	38 12 2	0.042	5.321	5.119	4.793		
KON19JUN71	41 13 0	0.041	5.297	5.076	4.783		
KON09OCT71	27 12 3	0.046	5.165	4.977	4.742		
KON210CT71	32 9 0	0.047	5.359	5.139	4.795		
KON26AUG72	29 15 2	0.044	5.155	4.934	4.606		
KON02SEP72	15 29 0	0.045	4.602	4.330	4.079		
DEK21NOV65	48 15 1	0.038	5.378	5.169	4.894		
DEK13FEB66	51 4 10	0.037	6.089	5.898	5.652		
DEK20MAR66	49 9 8	0.037	5.854	5.638	5.353		
DEK07MAY66	9 26 1	0.050	4.495	4.243	4.016		

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Table A.2. Geotech's Maximum-Likelihood Network mb (Continued)						
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	m <sub>b</sub> (P <sub>a</sub> )	
DEK19OCT66	51 10 5	0.037	5.539	5.370	5.112	
DEK26FEB67	48 9 6	0.038	5.833	5.610	5.368	
DEK29SEP68	50 8 6	0.038	5.631	5.439	5.138	
DEK23JUL69	38 21 1	0.039	5.183	4.943	4.628	
DEK11SEP69	19 39 0	0.039	4.603	4.265	4.013	
DEK22MAR71	43 14 3	0.039	5.519	5.347	5.052	
DEK25APR71	37 5 0	0.046	5.783	5.594	5.331	
DEK30DEC71	16 3 0	0.069	5.553	5.377	5.020	
DEK28MAR72	28 17 0	0.045	4.979	4.747	4.380	
DEK16AUG72	23 23 1	0.044	4.905	4.650	4.361	
DEK10DEC72	30 7 5	0.046	5.542	5.340	4.990	
DEK29MAR77	25 14 0	0.048	5.004	4.723	4.329	
DEK30JUL77	21 16 0	0.049	4.877	4.630	4.230	
DEK26MAR78	25 6 0	0.054	5.507	5.284	4.963	
DEK22APR78	2190	0.055	5.020	4.771	4.480	
DEK28JUL78	36 9 6	0.042	5.524	5.313	5.002	
DEK22MAY80	36 23 1	0.039	5.129	4.926	4.671	
SEK15JAN65	46 1 2	0.043	5.894	5.746	5.511	
SEK19JUN68	28 3 2	0.052	5.282	5.022	4.651	
SEK30NOV69	50 0 0	0.042	5.965	5.787	5.401	

- 56 -

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Tabl	Table A.2. Geotech's Maximum-Likelihood Network $m_b$ (Continued)						
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	$m_b(P_a)$		
SEK30JUN71	31 19 1	0.042	5.062	4.794	4.499		
SEK10FEB72	34 8 2	0.045	5.319	5.089	4.825		
SEK02NOV72	42 1 15	0.039	6.185	5.944	5.608		
SEK10DEC72	44 2 11	0.040	6.020	5.794			
SEK23JUL73	53 1 1	0.041	6.191	6.006	5.763		
SEK14DEC73	49 8 6	0.038	5.760	5.564	5.261		
SEK27APR75	18 1 1	0.067	5.494	5.254	4.917		
SEK04JUL76	38 0 5	0.046	5.848	5.601	5.236		
SEK07DEC76	17 2 1	0.067	5.615	5.420	4.976		
SEK11JUN78	17 0 1	0.071	5.879	5.572	5.294		
SEK15SEP78	37 1 6	0.045	5.851	5.698	5.447		
SEK23JUN79	40 3 3	0.044	6.060	5.860	5.631		
SEK04AUG79	40 5 20	0.037	6.093	5.861	5.594		
SEK28OCT79	44 5 13	0.038	5.946	5.706	5.467		
SEK23DEC79	41 3 17	0.038	6.145	5.894	5.599		
SEK14SEP80	34 5 6	0.045	6.033	5.771	5.459		
SEK18OCT81	41 4 7	0.042	5.979	5.754	5.478		
SEK26MAY84	30 0 3	0.052	6.002	5.915	5.590		
SEK14SEP88	25 0 1	0.059	6.034	5.762	5.509		

Table	Table A.2. Geotech's Maximum-Likelihood Network mb (Continued)						
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	$m_b(P_a)$		
SAF23NOV76	22 0 0	0.064	5.626				
SAF29AUG78	16 0 0	0.075	5.905				
SAF29NOV78	28 0 0	0.057	5.880				
SAF07JUL79	30 0 0	0.055	5.799				
SAF18AUG79	28 0 0	0.057	6.087				
SAF02DEC79	15 0 0	0.078	5.874				
SAF12OCT80	23 0 0	0.063	5.828				
SAF14DEC80	29 0 0	0.056	5.911				
SAF27DEC80	24 0 0	0.061	5.896				
SAF22APR81	25 0 0	0.060	5.865				
SAF13SEP81	1700	0.073	6.024				
SAF27DEC81	23 0 0	0.063	6.207				
SAF25APR82	14 0 0	0.080	5.944				
SAF04JUL82	21 0 0	0.066	6.089				
SAF05DEC82	26 0 0	0.059	6.093				
SAF12JUN83	1600	0.075	5.943				
SAF06OCT83	25 0 0	0.060	5.942				
SAF26OCT83	18 0 0	0.071	5.941				
SAF25APR84	21 0 0	0.066	5.892				
SAF14JUL84	23 0 0	0.063	5.999				

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Table A.2. Geotech's Maximum-Likelihood Network m <sub>b</sub> (Continued)							
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	$m_b(P_a)$		
SAF27OCT84	19 0 0	0.069	6.150				
SAF02DEC84	22 0 0	0.064	5.693				
SAF16DEC84	15 0 0	0.078	5.993				
SAF28DEC84	19 0 0	0.069	5.916				
SAF15JUN85	15 0 0	0.078	6.069				
NNZ27OCT66	56 0 14	0.036	6.447	6.305	6.075		
NNZ21OCT67	53 5 3	0.038	5.783	5.611	5.424		
NNZ07NOV68	59 1 5	0.037	6.042	5.847	5.602		
NNZ14OCT69	59 2 7	0.036	6.144	5.972	5.778		
NNZ14OCT70	35 0 22	0.040	6.820	6.640	6.436		
NNZ27SEP71	23 0 21	0.045	6.629	6.487	6.276		
NNZ28AUG72	32 0 11	0.046	6.383	6.261	6.008		
NNZ12SEP73	23 0 21	0.045	6.770	6.677	6.356		
NNZ29AUG74	25 0 18	0.046	6.583	6.402	6.141		
NNZ21OCT75	23 0 17	0.048	6.548	6.344	6.110		
NNZ23AUG75	27 0 12	0.048	6.495	6.376	6.128		
NNZ20OCT76	25 34 0	0.039	4.680	4.369	4.056		
NNZ01SEP77	25 2 2	0.056	5.572	5.433	5.126		
NNZ10AUG78	39 3 18	0.039	5.867	5.637	5.414		
NNZ11OCT80	42 4 6	0.042	5.674	5.460	5.202		

Table A.2. Geotech's Maximum-Likelihood Network mb (Continued)						
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	$m_b(P_a)$	
NNZ01OCT81	43 4 5	0.042	5.666	5.505	5.251	
NNZ18AUG83	30 4 5	0.048	5.721	5.542	5.339	
NNZ25OCT84	22 3 4	0.056	5.610	5.439	5.174	
SNZ27SEP73	32 3 1	0.050	5.754	5.518	5.227	
SNZ27OC73A	14 0 24	0.049	7.082	6.864	6.645	
SNZ27OC73B	9 28 0	0.049	4.189	4.037		
SNZ27OC73C	4 34 0	0.049	3.951	3.928	3.587	
SNZ02NOV74	12 0 29	0.047	7.001	6.784	6.502	
SNZ18OCT75	21 0 21	0.046	6.838	6.527	6.245	
BERYL	1160	0.073	5.017	4.815	4.455	
CORUNDON	11 41 0	0.042	4.247	3.951	3.852	
EMERAUDE	14 25 0	0.048	4.596	4.269		
GRENAT	32 31 1	0.038	4.787	4.524	4.332	
OPALE	3 50 0	0.041	3.950	3.909	3.827	
RUBIS	45 5 0	0.042	5.434	5.185	4.863	
SAPHIR	55 5 5	0.037	5.725	5.479	5.196	
TOURMALINE	27 39 0	0.037	4.671	4.463	4.158	
TURQUOISE	11 53 0	0.038	4.258	3.986		

Table A.2. Geotech's Maximum-Likelihood Network mb (Continued)					
Event	# of Signals	S.E.M.	$m_b(P_{\max})$	$m_b(P_b)$	m <sub>b</sub> (P <sub>a</sub> )
TU19FEB77	16 27 0	0.046	4.665	4.413	
TU19MAR77	20 5 1	0.059	5.682	5.475	5.175
TU24NOV77	32 0 0	0.053	5.702	5.437	5.102
TU30NOV78	38 7 2	0.044	5.635	5.251	4.862
TU25JUL79	18 0 0	0.071	5.917	5.624	5.158
TU23MAR80	27 14 3	0.045	5.394	5.141	4.713
TU19JUL80	38 2 2	0.046	5.559	5.202	4.939
TU03DEC80	31 10 0	0.047	5.371	5.020	4.739
TU25JUL82	22 13 0	0.051	5.252	5.078	4.717
TU19APR83	21 1 0	0.064	5.533	5.228	5.011
TU25MAY83	18 0 0	0.071	5.764	5.446	5.163
RAJ18MAY74	7 23 0	0.055	4.595	4.341	4.081
CH22SEP69	30 12 0	0.046	5.190	4.801	4.409
СН27ОСТ75	12 24 0	0.050	4.655	4.467	4.223
CH17OCT76	12 33 0	0.045	4.610	4.298	4.134
CH06OCT83	16 12 1	0.056	5.207	4.997	4.740
CH03OCT84	10 12 0	0.064	5.020	4.769	4.489
CH19DEC84	3 10 0	0.083	4.424	4.077	4.101

### **APPENDIX B**

# ESTIMATES OF TEST SITE BIAS WITH m<sub>b</sub>(GLM91A)

Since the near-source correction procedure presented in this study has not been applied to regions other than Semipalatinsk, some experiments using the more complete  $m_b$  (GLM91A) values may be interesting. The event  $m_b$  values in Table A.2 are corrected for the station terms, and hence are very similar to the  $\overline{m}_{2.2}$  discussed in Section I.3. The only difference is that  $\overline{m}_{2.2}$  is computed as the (maximum-likelihood) averaged  $m_{2.2}$  across those and only those stations which reported the amplitude, whereas  $m_b$  (GLM91A) is inverted jointly with all 515 events and 122 stations simultaneously, and hence the missing stations are also included in the (maximumlikelihood) averaging. The discrepancy in the two  $m_b$  values is insignificant, however (*cf.* Tables 2 and A.2).

Figure 16 shows the results of regressing  $m_b$  (GLM91A) on the published yields of Semipalatinsk and NTS high-coupling explosions with 10% S.E. in yields assumed. U.S. and U.S.S.R. have released the yields of roughly equally many events (Springer and Kinaman, 1971, 1975; Bocharov *et al.*, 1989; Vergino, 1989). Note that the NTS calibration curve based on  $P_{max}$  phase of  $m_b$  (GLM91A) (top of Figure 16) has the same slope and intercept as those based on  $\overline{m}_{2.9}$  (Figure 13). Also note that the NTS curve based on the first arrivals (*i.e.*,  $P_a$  phase) has a smaller slope, and there seems to be a lot of scatter for the low yields. Blandford (written communication, 1991) pointed out that this might occur if the low-yield  $m_b(P_a)$  were biased high and had a lot of scatter due to the noise wavelets interfering with signal wavelets.



Figure 16.  $m_b$  (GLM) versus published yields of Semipalatinsk and NTS high-coupling explosions (with 10% S.E. in yields assumed). U.S. and U.S.S.R. have released the yields of roughly equally many events. The mean KTS-NTS test site bias is larger than 0.35 m.u. if  $P_{max}$  or  $P_b$  is used, and the larger scatter in NTS calibration curve is ignored. The bias would be about 0.15-0.2 m.u. if only granitic shots at NTS are used.

Table B.1 lists the mean KTS-NTS test site bias at three different yield levels. Results based on "b" and "max" phases suggest that the bias is yield-dependent and appears to be larger at the lower yield end. The "mean" KTS-NTS  $m_b$  bias is slightly larger than 0.35 m.u., although Figure 16 also indicates that the NTS granite events would seem to be about 0.15-0.2 m.u. below the Semipalatinsk curve.

Table B.1. Estimated Test Site Bias (from Nuttli's earlier studies)						
Test Sites	Magnitude	Description	Description 10kt		150kt	
Balapan - NTS	m <sub>b</sub> (ISC), L <sub>g</sub>	Nuttli (1987)	0.35	0.35	0.35	
Degelen - NTS	m <sub>b</sub> (ISC), L <sub>g</sub>	Nuttli (1987)	0.58	0.58	0.58	
Estimated Test Site Bias (from this study)						
Test Sites	Magnitudes	Description	10kt	100kt	150kt	
KTS' - NTS"	$m_b(P_{\max})$	Marshall	0.55	0.50	0.48	
KTS - NTS	$m_b(P_a)$	TG	0.38	0.39	0.40	
KTS - NTS	$m_b(P_b)$	TG	0.51	0.47	0.47	
KTS - NTS	$m_b(P_{\rm max})$	TG	0.47	0.42	0 4 1	
KTS - NTS	$m_b(P_{max})$	TG, S-Cubed	0.36	0.36	0.36	

\*) Combining all UK/AWE's Balapan, Degelen, and Murzhik m, values as listed in Vergino (1989)

\*\*) UK/AWE's NTS m<sub>b</sub> as distributed in 1987

\*\*\*) Murphy (1981): m<sub>b</sub>(S-Cubed) = 3.92 + 0.81 log(W) for NTS high-coupling events

That NTS granite events might lie above typical wet-tuff or rhyolite events on the  $m_b$ -yield calibration curve can be further illustrated by some simple calculations with the three events Dougals (1987) analyzed. The announced yields of events 680619B, 710630B, and PILEDRIVER are <20 kt, <20 kt, and 56 kt, respectively. Our yield estimates (*cf.* the  $P_{max}$  calibration curves shown in Figure 16), however, are 17 kt, 9 kt,

and 100 kt, respectively, based on their corresponding  $m_b$  (GLM91A,  $P_{max}$ ) of 5.282, 5.062, and 5.480. (Note that the  $m_b$  discrepancy between  $\bar{m}_{2.9}$  and  $m_b$  (GLM91A) is insignificant.) At NTS, a high-coupling 17-kt shot has an expected  $m_b$  of 4.82, which is 0.46 m.u. below that of 680619B at Balapan. Likewise, a 9-kt high-coupling shot at NTS would be expected to have a  $m_b$  around 4.58, about 0.48 m.u. smaller than that of 710630B. At KTS, the expected  $m_b$  for 100 kt and 56 kt would be 5.9 and 5.69, respectively; which are 0.42 and 0.43 m.u. larger than typical NTS shots at the corresponding yields. The KTS-NTS bias of 0.48 (9 kt), 0.46 (17 kt), 0.43 (56 kt), and 0.42 (100 kt) resemble that yield dependency as shown in Table B.1, as expected. PILEDRIVER's  $m_b$  (GLM91A), 5.480, is about 0.22 m.u. larger than that of a 56-kt shot at NTS, *i.e.*, 5.26. This result seems to match very well with Ryall's (1985) inference of the attenuation differential between Semipalatinsk and NTS using earthquake data recorded at seismic stations in these two region.

Murphy (1981) points out that the "statistically significant"  $m_b$ -yield relationship for the wet tuff/rhyolite explosions at Pahute Mesa and Yucca Flat is

$$m_b = 3.92 + 0.81 \log(W)$$
 [1]

which happens to be about 0.36 m.u. below our inferred calibration curve for historical Semipalatinsk explosions. This could be simply accidental. Nevertheless, an interesting speculation can be offered to explain the coincidence. It is not impossible that the Soviets are fully aware of Equation [1] and the commonly quoted KTS-NTS bias of 0.35 m.u. (*e.g.*, OTA, 1988). Perhaps the Soviets have purposefully released a subset of their historical explosions which would fit a prescribed curve roughly 0.35 m.u. above Equation [1]. If this was indeed the case, probably the released 19 Eastern Kazakhstan events were not "fudged" otherwise ---- although whether they are truly representative of the whole explosion population would still remain open (*cf.* the discussion in Gray *et al.*, 1990). We could also argue that, if the aforementioned speculation were valid, then Geotech's  $m_b$  measurements of Soviet events must correlate very well with the magnitudes which the Soviet seismologists have used in determining

their own yields.

Using 20 NTS tuff/rhyolite events, Marshall et al. (1979)'s  $\overline{m}_2$  give

$$\overline{m}_2 = 3.71 + 0.89 \log(W).$$
 [2]

This is not significantly different from our result for 21 NTS tuff/rhyolite events (Figure 16):

$$m_b(\text{GLM91A}) = 3.76 + 0.86 \log(W)$$
 [3]

Combining our GLM/MLE-derived WWSSN  $m_b$  with RMS  $L_g$  values measured at NORSAR (Ringdal and Marshall, 1989), the  $m_b$ - $L_g$  residuals for E. Kazakh explosions show a strong difference among these three test sites (Figure 17). The SW subregion of Balapan test site excites slightly larger  $m_b$  (relative to  $L_g$ ), whereas all the remaining regions of E. Kazakh test site have negative residuals. All studies of the intrasite  $m_b$ - $L_g$  bias of Balapan explosions lead to a very consistent estimate, namely 0.14±0.02 m.u. (Table B.2 and Jih and Wagner, 1991).

As noted in Appendix A, DARPA distributed the WWSSN station  $m_b$  values of 39 large Balapan explosions furnished by AFTAC in the spring of 1988. Lilwall *et al.* (1988) supplemented this data set with some ISC recordings in their analysis, and they found that the event  $m_b$  values based on Blacknest's Joint Maximum-Likelihood (JML) method are not significantly different from those based on LSMF. We have incorporated these AFTAC's  $m_b$  values into our database (*cf.* pages 58-59), with a compensating correction for the different B( $\Delta$ ) factors. The LSMF results of the 39 AFTAC-measured Balapan events show a mean  $m_b$ - $L_g$  bias of 0.14 between SW and NE subregions of Balapan test site (Table B.2).

Table B.2. Mean $m_b - l_g$ of Balapan Explosions				
Reference	SW	TZ	NE	SW-NE
Ringdal and Hokland (1987)	0.112±0.009(?)		-0.059±0.014(?)	0.17
Marshall (1987) + NORSAR	0.116±0.009(26)	0.041±0.012(10)	-0.042±0.014(14)	0.16
Ringdal and Marshall (1989)	0.05±0.007(46)	-0.02±0.009(20)	-0.10±0.012(30)	0.15
TGAL + NORSAR	0.02010.015(20)	-0.071±0.017(8)	-0.112+0.017(8)	0.13
AFTAC + NORSAR	-0.012±0.010(18)	-0.113±0.020(6)	-0.153±0.018(7)	0.14
Marshall (1987) + Nuttli	-0.012±0.015(17)	-0.060±0.026(4)	-0.118±0.014(14)	0.11

Table B.3 lists the mean  $m_b-L_g$  values at three test sites of Eastern Kazakhstan. Murzhik events have smaller relative  $m_b$  excitation, as compared to Balapan and Degelen explosions. Since Balapan and Murzhik events essentially followed the same depth-yield scaling (Jih and Shumway, 1991; Jih, 1990), the relatively strong  $L_g$  excitation at Murzhik could probably be due to the smaller size (and hence shallower depth of burial) at Murzhik, or due to the source medium (Figure 17).

Table B.3. Mean $m_b$ - $L_g$ of Eastern Kazakh Explosions				
Reference	SR	DM	МК	
TGAL + NORSAR	-0.030±0.014(36)	-0.047±0.034(5)	-0.128±0.034(3)	
AFTAC + NORSAR	-0.063±0.014(31)	(?)	(?)	



### SPATIAL PATTERN OF mb-Lg RESIDUALS OF E. KAZAKH SHOTS

Figure 17. The spatial pattern of  $m_b$ - $L_g$  residuals of Semipalatinsk explosions with TG's  $m_b$ (GLM) and RMS  $L_g$  values reported at NORSAR. The residual pattern of Balapan events strongly indicates significant difference in the source medium across the Chinrau fault separating the northeastern and southwestern portion of the test site, as reported by Ringal and Marshall (1989) and Marshall *et al.* (1984). The mean  $m_b$ - $L_g$  bias between SW and NE Balapan is about 0.13 m.u.

### mb: Geotech's WWSSN GLM/MLE Pmax (GLM91A, 515 events)

(122 stations, each recorded 7 signals or more) NORSAR RMS Lg (Ringdal, 1990; Ringdal and Marshall, 1989) Surface geology: Bonham et al. (1980), Leith (1987). Balapan, SW region, 20 events: mb(TG)= mb(Lg) + 0.020(0.015) Balapan, TZ region, 8 events: mb(TG)= mb(Lg) - 0.071(0.017) Balapan, NE region, 8 events: mb(TG)= mb(Lg) - 0.112(0.017) Balapan, 36 events: mb(TG)= mb(Lg) - 0.030(0.014) Degelen, 5 events: mb(TG)= mb(Lg) - 0.047(0.034) Murzhik, 3 events: mb(TG)= mb(Lg) - 0.128(0.034) Sedimentary & volcanic rocks Devonian & Carboniferous rock Granitic rocks



Limestone
### SECTION II PROJECT OVERVIEW

# RECENT METHODOLOGICAL DEVELOPMENTS IN MAGNITUDE DETERMINATION AND YIELD ESTIMATION WITH APPLICATIONS TO SEMIPALATINSK EXPLOSIONS

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II.1 CONTRACT NO.: F19628-89-C-0063, Task 1 (expired July 1991)

## **II.2 OBJECTIVES**

The primary objective is to study the technical issues relating to the seismic estimation of the yield of remote underground explosions. Our approach is to improve both the numerical and statistical modeling tools as much as we can, and then apply the upgraded numerical tools to the excitation/propagation study of the teleseismic and regional phases, and apply the upgraded statistical tools to the magnitude determination as well as the yield estimation problems with emphasis on Soviet explosions.

## **II.3 RESEARCH ACCOMPLISHED DURING CONTRACT PERIOD**

## II.3.1 Upgrading of Unbiased Network m<sub>b</sub> Estimator

The body wave magnitudes used in developing and applying the magnitude-yield relationship are network  $m_b$  values which are some "average" of the station  $m_b$ .

Previously a major objective of magnitude-calculation research was to determine the network  $m_h$  that is not biased by the sample truncation due to the limited range of the seismometers. Ringdal (1976) introduced the maximum-likelihood estimator [MLE] to correct for the statistical bias introduced by data censoring from non-detection. Von Seggern and Rivers (1978) pointed out the importance of accounting for the data censoring due to signal clipping. Blandford and Shumway (1982) derived the general linear model [GLM] in the presence of data censoring using the Expectation Maximization [EM] algorithm. They simultaneously estimated event magnitudes and station corrections in a maximum-likelihood sense. Jih and Shumway (1989) re-examined and documented the GLM algorithms, and they discussed the uncertainty assessment in the censoring situation. It is concluded that in the multi-parameter linear regression problem with censored data, the scaling of  $\sigma/\sqrt{\text{degrees of freedom still provides an}}$ extremely good approximation of the uncertainty associated with each parameter. In the case of non-censoring, such approximation can be proved to be "exact". The methodological similarities and differences between the iterative least squares [ILS] and the maximum-likelihood estimator [MLE] were also identified in Jih and Shumway (1989).

A recent breakthrough in magnitude determination is the development of a procedure to account for the near-source focusing/defocusing effects. Jih and Wagner (1991b) propose to compute the new station magnitude  $m_{2.9}$  for the i-th event recorded at the j-th station as

$$m_{2,9}(i,j) = \log_{10}[A(i,j)/T(i,j)] + B(\Delta(i,j)) - S(j) - F(k(i),j)$$
<sup>[1]</sup>

where A(i,j) is the displacement amplitude (in millimicrons) and T(i,j) is the period (in seconds) of the *P* wave. The B( $\Delta$ ) is the distance-correction term. S(j) is the station correction, and F(k(i),j) is the near-source focusing correction for explosions from the k(i)-th source region. Some of the S terms for explosions from any test site and the F terms for the Semipalatinsk area are listed in Table 1. A complete list can be found in Jih and Wagner (1991b). This new magnitude is called  $m_{2,9}$  to avoid confusion with

Table 1. Receiver and Near-source Corrections for WWSSN Stations (partial listing)							
Station Term Near-source Term, F			Station				
Code	S	Balapan	Degelen	Murzhik	Longitude	Latitude	Description
AAE	-0.352	-0.387	-0.076	-0.127	38.766	9.029	Addis Ababa, Ethiopia
AAM	0.205	0.202	-0.073	-0.247	-83.656	42.300	Ann Arbor, Michigan
AKU	-0.048	0.300	0.311	0.212	-18.107	65.687	Akureyri, Iceland
ANP	-0.349	-0.167	0.402	0.183	121.517	25.183	Anpu, Taiwan
AQU	-0.133	-0.195	0.039	-0.024	13.403	42.354	Aquila, central Italy
ATU	0.128	0.191	-0.156	0.034	23.717	37.972	Athens Univ., Greece
BAG	-0.027	0.076	-0.009	-0.103	120.580	16.411	Baguio City, Luzon Island
BEC	-0.114	0.090	0.058	-0.092	-64.681	32.379	Bermuda-Columbia, Atlantic
BKS	0.089	-0.014	0.101	0.229	-122.235	37.877	Byerly, central California
BLA	0.057	-0.233	-0.177	-0.299	-80.421	37.211	Blacksburg, West Virginia
BOZ	0.188	-0.325	-0.025	-0.180	-111.633	45.600	Bozeman, Montana
BUL	0.003	0.014	-0.266	-0.037	28.613	-20.143	Bulawayo, Rhodesia
CHG	-0.140	0.240	0.106	0.045	98.977	18.790	Chiengmai, southeast Asia
СМС	-0.178	0.114	0.375	0.602	-115.083	67.833	Copper Mine, Canada
COL	0.065	0.181	0.188	0.051	-147.793	64.900	College Outpost, Alaska
COP	0.127	-0.003	0.159	-0.276	12.433	55.683	Copenhagen, Denmark
COR	0.155	0.132	0.183	0.172	-123.303	44.586	Corvallis, Oregon
СТА	0.153	-0.072	0.003	-0.073	146.254	-20.088	Charters Towers, Australia
DAG	0.036	-0.052	0.086		-18.770	76.770	Danmarkshavn, Greenland
DAV	-0.320	-0.264	-0.053		125.575	7.088	Davao, Mindanao Island
DUG	0.149	0.038	0.371	0.352	-112.813	40.195	Dugway, Utah
EIL	0.004	-0.117	-0.228	-0.103	34.950	29.550	Eilat, Arabic Peninsula
ESK	0.048	-0.042	0.162	-0.327	-3.205	55.317	Eskdalemuir, Scotland
FLO	0.000	-0.294	-0.093	-0.446	-90.370	38.802	Florissant, eastern Missouri

the  $m_3$  defined in Marshall *et al.* (1979) that corrects for the source-region attenuation and station terms solely based on published  $P_n$  velocity. Testing results indicate that the procedure described in [1] has the following advantages:

- [A] For 79 out of 82 Semipalatinsk events we have tested, Equation [1] provides more stable  $m_b$  measurements across the whole recording network, as compared to the conventional GLM or LSMF procedure which only corrects for the station terms (*cf.* Table 2). The reduction in the standard deviation of network  $m_b$  from  $\overline{m}_1$  to  $\overline{m}_{2.9}$  could reach a factor of 3 (*cf.* Table 2). Events which do not show improvements in precision could have been detonated in environments with different focusing patterns.
- [B] The resulting network  $m_b$  values are not significantly different from the GLM results. Thus if the mean network  $m_b$  values derived by GLM or LSMF are unbiased, so are the refined results.
- [C] The scatter in  $\overline{m}_{2.9}$  versus log(yield) is smaller than that for other  $m_b$  (cf. Table 3).

Table 2. Various Network-Averaged m <sub>b</sub> of 19 Special Events								
Event # of signals		Without correction	Station corrected	Near-source corrected				
Date		<i>̄m</i> <sub>1</sub> , σ	<i>̄</i> m <sub>2.2</sub> , σ	<i>̄</i> m <sub>2.9</sub> , σ				
651121D	48 12 1	5.384±0.035 0.271	5.394±0.024 0.188	5.381±0.019 0.152				
660213D	51 2 10	6.073±0.038 0.305	6.092±0.027 0.218	6.088±0.014 0.114				
660320D	49 6 8	5.848±0.041 0.322	5.865±0.030 0.239	5.853±0.009 0.074				
670922M	35 20 1	5.033±0.036 0.271	5.048±0.029 0.214	5.029±0.017 0.125				
680929D	50 4 6	5.610±0.035 0.275	5.642±0.026 0.202	5.641±0.018 0.138				
690723D	38 17 1	5.172±0.041 0.306	5.201±0.029 0.220	5.186±0.013 0.100				
691130B	51 0 0	5.950±0.044 0.313	5.973±0.031 0.222	5.945±0.026 0.184				
691228M	45 2 3	5.666±0.043 0.306	5.665±0.035 0.250	5.660±0.018 0.125				
710425D	37 3 0	5.764±0.052 0.327	5.793±0.042 0.267	5.826±0.020 0.127				
710606M	38 6 2	5.323±0.040 0.272	5.341±0.031 0.210	5.319±0.015 0.099				
711009M	27 9 3	5.165±0.037 0.233	5.187±0.028 0.174	5.136±0.016 0.100				
711021M	32 6 0	5.348±0.049 0.301	5.383±0.036 0.224	5.341±0.021 0.127				
720210B	34 8 2	5.297±0.042 0.278	5.329±0.029 0.195	5.289±0.018 0.121				
720816D	23 20 1	4.908±0.044 0.293	4.931±0.033 0.221	4.921±0.025 0.165				
720902M	15 25 0	4.615±0.049 0.312	4.635±0.038 0.243	4.602±0.017 0.107				
721102B	42 0 15	6.173±0.045 0.339	6.183±0.034 0.256	6.160±0.023 0.177				
721210B	45 1 11	6.006±0.037 0.282	6.013±0.029 0.223	5.983±0.022 0.169				
880914B	25 0 1	6.004±0.037 0.191	6.032±0.023 0.117	6.026±0.033 0.168				

\*) 19 Semipalatinsk explosions for which the yields were published.

\*\*)  $m_1$  = network average of raw  $m_b$  without any correction --- equivalent to ISC bulletin  $m_b$ .  $m_{2,2}$  = network average of  $m_b$  with GLM station corrections applied.  $\overline{m}_{2,9}$  = network average of  $m_b$  with both station terms and near-source focusing terms removed.

Table 3. 95% Confidence Factor of Semipalatinsk Yield Estimate								
(assuming yields are subject to rounding and 10% S.E.)								
m <sub>b</sub> used	1 kt	10 kt	50 kt	100 kt	150 kt			
m <sub>1</sub>	4.74	2.58	2.19	2.42	2.61			
т <sub>2.2</sub>	3.98	2.20	1.89	2.12	2.29			
<b>m</b> <sub>2.9</sub>	3.23	1.82	1.67	1.88	2.04			
Ringdal's <i>RMS</i> L <sub>g</sub> <sup>1</sup>	5.74	2.39	1.65	1.80	1.98			
Israelson's $RMS L_g^2$	2.82	1.81	1.72	1.87	1.97			
(assuming yields are subject to 10% S.E. only)								
$\overline{m}_1$ 3.96 2.44 2.13 2.38 2								
m <sub>2.2</sub>	3.45	2.09	1.86	2.07	2.23			
m <sub>2.9</sub>	2.83	1.74	1.60	1.77	1.91			
Ringdal's RMS L <sub>g</sub>	5.74	2.39	1.65	1.80	1.98			
Israelson's <i>RMS L<sub>g</sub></i>	2.70	1.84	1.78	1.86	1.95			

1) RMS L<sub>o</sub> of 9 Semipalatinsk events furnished by Ringdal (1990) with NORSAR and GRF data.

2) RMS  $L_g$  of 16 Semipalatinsk events furnished by Israelson (1991b) with hand-digitized Soviet analog seismograms.

## II.3.2 m<sub>b</sub>-Yield Regression Routine with Censored Yields: MLE-CY

In general there are four types of yield data available: [0] the yield, W, is known as  $y_0$  kt, [1] W is left censored, *i.e.*, the exact value of W is only known to be less than certain level, [2] W is right censored, *i.e.*, the exact value of W is only known to be larger than certain level, and [3] W is only known to lie between two bounds. The majority of Soviet yields recently published by Bocharov *et al.* (1989) and Vergino (1989) are of type 3. The problem of estimating the yield of an explosion from the estimated seismic magnitude has been handled traditionally using the linear or piecewise linear model

$$X = \alpha + \beta \log(W) + v = \alpha + \beta Y + v$$
[2]

where X is the measured magnitude (*e.g.*,  $m_b$  or  $M_S$ ),  $\alpha$  and  $\beta$  are intercept and slope estimators, W is the yield in kiloton [kt], and v is an error term. v is assumed to be a Gaussian random variable with mean zero and standard deviation  $\sigma$ . One may then collect a number of "calibration events", estimating  $\alpha$  and  $\beta$  by least squares using a number of known yields and measured magnitudes. This classical calibration approach leads to predicting a future log-yield Y at magnitude =  $\hat{X}$  by inverting Equation [2], *i.e.*,  $\hat{Y} = (\hat{X} - \hat{\alpha})/\hat{\beta}$ .

The geometrical interpretation of "regressing X on Y" is that the  $(\hat{\alpha}, \hat{\beta})$  thus estimated will be the optimal solution that minimizes the sum of the squared X residuals,  $\sum (X - \hat{\alpha} - \hat{\beta}Y)^2$ . Implicitly, here we have made an assumption that the independent variable Y has nearly perfect accuracy and precision as compared to X. Alternately, one can estimate  $\lambda$  and  $\eta$  in the inverse regression model

$$Y = \lambda + \eta X + v'$$
[3]

and then predict a future log-yield directly as  $\hat{Y} = \hat{\lambda} + \hat{\eta} \hat{X}$ . Likewise, here one is implicitly assuming that X has perfect accuracy and precision, and hence the optimal estimate  $(\hat{\lambda}, \hat{\eta})$  is the one that minimizes the sum of squared Y residuals,  $\sum (Y - \hat{\lambda} - \hat{\eta}X)^2$ . Thus either the yield or the magnitude must be regarded as an error-free independent variable in these two models. Symbolically, these two conventional regression models are based on the following two extreme assumptions:  $\sigma(X)/\sigma(Y) = \infty$  and  $\sigma(X)/\sigma(Y) = 0$ , respectively.

In reality, both the  $m_b$  and the yield measurements are subject to error. At NTS,  $\sigma(m_b) >> \sigma(\log \text{ yield})$  could be a reasonable assumption to justify the regression of  $m_b$ on the yields. However, this may not be the case in general. Note that [3] can be rewritten in a form similar to [2]:  $X = \alpha' + \beta'Y + v''$  with the transformations  $\alpha' = -\lambda/\eta$ ,

- 75 -

 $\beta'=1/\eta.$ 

Elegant maximum-likelihood theory can be derived for "regressing censored Y on X" (Jih *et al.*, 1990a). Suppose there are  $n_0$ ,  $n_1$ ,  $n_2$ , and  $n_3$  events for each type, respectively. The conditional likelihood function of the censored observations ( $y_0$ ,  $t_1$ ,  $t_2$ ,  $t_3$ ) given the intercept  $\alpha$ , slope  $\beta$ , and  $\sigma$  is

$$L (\mathbf{y}_{0}, \mathbf{t}_{1}, \mathbf{t}_{2}, \mathbf{t}_{3} | \alpha, \beta, \sigma) = \prod_{j=1}^{n_{0}} P(\mathbf{Y}_{j} = \mathbf{y}_{0j} | \alpha, \beta, \sigma) * \prod_{j=1}^{n_{1}} P(\mathbf{Y}_{j} < t_{1j} | \alpha, \beta, \sigma) * [4]$$
$$\prod_{j=1}^{n_{2}} P(\mathbf{Y}_{j} > t_{2j} | \alpha, \beta, \sigma) * \prod_{j=1}^{n_{3}} P(t_{aj} < \mathbf{Y}_{j} < t_{bj} | \alpha, \beta, \sigma)$$

and the log-likelihood function is

In L ( 
$$\mathbf{y}_0$$
,  $\mathbf{t}_1$ ,  $\mathbf{t}_2$ ,  $\mathbf{t}_3 \mid \alpha, \beta, \sigma$ ) =  $-\frac{n_0}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{j=1}^{n_0} (\mathbf{y}_{0j} - \frac{\mathbf{x}_{0j} - \alpha}{\beta})^2 + [5]$   
$$\sum_{j=1}^{n_1} \ln \Phi(\mathbf{z}_{1j}) + \sum_{j=1}^{n_2} \ln \Phi(-\mathbf{z}_{2j}) + \sum_{j=1}^{n_3} \ln [\Phi(\mathbf{z}_{bj}) - \Phi(\mathbf{z}_{aj})]$$

where  $z_i \equiv (\alpha + \beta t_i - x_i)/\beta\sigma$ ;  $y_0$ ,  $t_1$ ,  $t_2$ , and  $t_3$  are the collection of announced yields.

Solving  $\frac{\partial \ln L}{\partial \sigma} \equiv 0$  implies immediately that the  $\hat{\sigma}$  must satisfy the following necessary condition:

$$\sigma^{2} = \frac{\sum_{j=1}^{n_{0}} (y_{0j} - \frac{x_{0j} - \alpha}{\beta})^{2}}{n_{0} + \sum_{j=1}^{n_{1}} \frac{\phi(z_{1j})}{\Phi(z_{1j})} z_{1j} - \sum_{j=1}^{n_{2}} \frac{\phi(z_{2j})}{\Phi(-z_{2j})} z_{2j} + \sum_{j=1}^{n_{3}} \frac{\phi(z_{bj}) z_{bj} - \phi(z_{aj}) z_{aj}}{\Phi(z_{bj}) - \Phi(z_{aj})}}$$
[6]

Solving  $\frac{\partial \ln L}{\partial A} \equiv 0$  implies that the sum of the "refined residuals" should be zero. Solving  $\frac{\partial \ln L}{\partial B} \equiv 0$  implies that the vector of refined residuals should be orthogonal to the vectors of means. It follows that the optimal estimate of A and B can be obtained by the "standard least squares" inversion with the censored data all replaced by their

conditional expectations, *i.e.*, the "refined observations". Thus  $\sigma$  can be solved iteratively with [6] along with  $\alpha$  and  $\beta$  using the EM algorithm. In the non-censored case, this "MLE-CY" code gives results identical to those derived by the standard least squares.

## II.3.3 A General $m_b$ -Yield Regression Routine with Uncertain Data: DWLSQ

Even the 19 Semipalatinsk explosions for which the "exact" yields were published would inevitably be subject to many sources of error. The Soviets might have rounded 8 of the announced 19 yields to the nearest 5 kt or 10 kt. An announced yield of 100 kt (*e.g.*, 660320D) could mean something actually measured between 95 kt and 104 kt. It could also indicate that 100 kt was the designed energy release, and the actual yield was somewhere nearby. Likewise, the "real yield" of 2 kt (*e.g.*, 720902M) could be something between 1.5 kt and 2.4 kt. Below 100 kt, the rounding errors could overwhelm the presumed standard measurement error --- assuming the announced yields are not otherwise "fudged".

A more general regression routine has been developed to take the rounding and standard errors in the yields into account (Jih, 1991). For each ( $m_b$ , yield) pair, we use a random number generator to produce a perturbed ( $m_b$ , yield) pair according to their uncertainty distribution. A standard least-squared regression is then performed for each data set of 19 perturbed pseudo-observations. The procedure is repeated for several hundred iterations, and all the resulting calibration curves are then used to infer the ensemble behavior. This "doubly-weighted least-squares scheme" is an extension to the "ordinary weighted least-squares" in which only errors in the  $m_b$  would be used to adjust the inferred parameters.

The "upper 95% confidence limit" of the predicted  $m_b$  at a given log(yield) level (say, Y<sub>0</sub>) can be computed as follows:

- 77 -

$$\hat{m}_{b}(max) + t(D.O.F., 0.975)[\sigma^{2}(m_{b}) + \sigma^{2}(regression)(\frac{1}{N} + \frac{(Y_{0} - \overline{Y})^{2}}{\sum(Y_{i} - \overline{Y})^{2}})]^{0.5}$$
 [7]

where N = number of data points used in the regression, D.O.F. = N-2,  $\sigma(m_b)$  = the mean S.E. in the network  $m_b$  used in the regression,  $\sigma(\text{regression})$  = the  $\sigma$  of residuals,  $m_b(\text{max})$  = estimate of the largest possible mean  $m_b$  at the given log(yield) level,  $\overline{Y}$  is the mean log(yield) used in the regression, and t(D.O.F., 0.975) is the 97.5 percentile of Student's *t* distribution at "D.O.F." degrees of freedom. The "lower 95% confidence limit" can be computed in a similar way.

## II.3.4 Expansion of Geotech's WWSSN m<sub>h</sub> Database

Our database of station  $m_b$  values based on the short-period vertical-component (SPZ) WWSSN [World Wide Standard Seismograph Network] recordings of body waves has been expanded to 192<sup>1</sup> events from a variety of regions including the N.T.S. (U.S.), French Sahara, Azgir (U.S.S.R.), Urals (U.S.S.R.), Murzhik (E. Kazakh, U.S.S.R.), Degelen Mountain (E. Kazakh, U.S.S.R.), Balapan (E. Kazakh, U.S.S.R.), Novaya Zemlya (U.S.S.R.), Tuamoto Islands (France), Rajasthan (India), and Lop Nor (Sinjiang, China). This database consists of 515 usable "a" (*i.e.*, zero-crossing to first peak), "b" (*i.e.*, first peak to first trough), and "max" (*i.e.*, max peak-to-trough or trough-to-peak in the first 5 seconds) event phases. Jih *et al.* (1990b) reported evidences which indicate that the GLM network  $m_b$  values inferred from these WWSSN recordings are better than many other magnitude measurements.

<sup>&</sup>lt;sup>1</sup>111 events were measured by R. A. Wagner, M. E. Marshall, R. O. Ahner, and J. A. Burnetti under previous contracts during the past decade. R. A. Wagner added 56 more events to Geotech's m<sub>b</sub> database during FY90-91. The remaining 25 events were adapted from the data that DARPA distributed in 1988.

### II.3.5 Magnitude-Yield Relationship at Semipalatinsk Area

A systematic comparative analysis of the magnitude-yield relationship at three regions (Balapan [= Shagan River], Degelen, and Murzhik [= Konystan]) of Semipalatinsk, Eastern Kazakhstan, U.S.S.R. has been conducted using miscellaneous unclassified magnitudes as well as the recently published yields of 96 Soviet explosions. Only the most noteworthy observations are summarized here:

- 1 MLE-CY and DWLSQ vs. Ericsson's code. Including the censored yields in the regression does generally improve the accuracy of the yield estimates slightly, if the magnitudes are "consistent". The "MLE-CY" code is robust in detecting the inherent inconsistency of the magnitudes by utilizing the censored information. In reality, both the magnitude and the yield measurements are subject to error. Pending the determination as to which of the two extreme hypotheses, namely  $\sigma(m_b)/\sigma(Y) = 0$  and  $\sigma(m_b)/\sigma(Y) = \infty$ , is closer to the real situation, we also applied Ericsson's (1971) curve-fitting method<sup>2</sup> to the regressions of non-censored yields using various  $\sigma(m_b)/\sigma(Y)$  ratios. As expected, we can see the smooth transition of estimated parameters (*i.e.*, the slope and the intercept) as  $\sigma(m_b)/\sigma(Y)$  varies. Thus the censored cases with nontrivial  $\sigma(m_b)/\sigma(Y)$  values could also be "interpolated" accordingly (Jih et al., 1990b). Note that Ericsson's (1971) method allows different variances in both the independent and the dependent variables in the regression. However, it can be applied to the non-censored case only. "MLE-CY" and Ericsson's methods represent two different directions in extending the standard least squares. "DWLSQ" is even more flexible than Ericsson's code in that it permits the errors in each  $(m_b, yield)$  pair to be arbitrary.
- 2 Rounding Errors vs. Gaussian Errors in the Yields. There has been some concern about the accuracy and precision of the Soviet published yields. It turns out that, so long as the best  $m_b$  (such as  $\overline{m}_{2,9}$ ) is used, the uncertainty factor in the predicted yield of future Semipalatinsk events is not very sensitive to the

<sup>&</sup>lt;sup>2</sup>Code provided by R. H. Shumway and T. M. McEllresh

postulated uncertainty (precision) in the published yields of the 19 events. Also, the yields of future underground explosions in the Semipalatinsk area can be estimated seismically with a capability much better than the factor-of-2 uncertainty that is commonly reported. For instance, a factor of 1.5 could be a reasonable uncertainty estimate at around the 50-kt level if  $\overline{m}_{2.9}$  is used as the source measure. At yields below 10 kt small variations of the physical environment may produce greater uncertainty. Therefore, the uncertainty may be inherently greater at such low yield level.

Table 4. Inferred Calibration Parameters and Associated Uncertainty Factors								
Uncertainty in yield	Slope	Intercept	1kt	10kt	50kt	100kt	150kt	
R.E. + 20% S.E.	0.793±0.031	4.308±0.050	4.14	2.04	2.09	2.55	2.89	
R.E. + 10% S.E.	0.804±0.022	4.288±0.038	3.23	1.82	1.67	1.88	2.04	
R.E. + 5% S.E.	0.805±0.020	4.286±0.035	3.11	1.78	1.53	1.64	1.77	
R.E. + 2% S.E.	0.806±0.020	4.285±0.035	3.33	1.81	1.53	1.69	1.82	
R.E. + 1% S.E.	0.805±0.020	4.287±0.035	3.04	1.80	1.49	1.60	1.75	
R.E. Only	0.807±0.019	4.284±0.034	3.04	1.83	1.54	1.60	1.72	
20% S.E.	0.794±0.031	4.306±0.049	4.08	2.16	1.88	2.23	2.50	
10% S.E.	0.807±0.018	4.282±0.027	2.83	1.74	1.60	1.77	1.91	
5% S.E.	0.811±0.012	4.277±0.018	2.41	1.62	1.51	1.61	1.71	
2% S.E.	0.811±0.009	4.276±0.014	2.33	1.57	1.48	1.57	1.64	
1% S.E.	0.812±0.009	4.275±0.014	2.32	1.57	1.49	1.57	1.64	
0.1% S.E.	0.812±0.009	4.275±0.014	2.32	1.57	1.49	1.57	1.64	

R.E.: Rounding Error ; S.E.: Standard Error

3 Yield estimates of recent Balapan explosions. Ringdal and Hokland (1987) noted that, for Balapan explosions detonated after 1976, the largest peak of clustered NORSAR RMS L<sub>g</sub> values was 6.06. It corresponds to 138 kt on the calibration curve derived with "DWLSQ":

$$RMS L_{q} = 4.54(\pm 0.050) + 0.71(\pm 0.028) \log(W)$$
[8]

which is almost identical to the mean yield of  $139\pm7$  kt computed by Sykes and Ruggi (1989) using the 9 largest Balapan shots during 1976-1985.

 4 Depth-yield scaling at Eastern Kazakhstan. The Soviet-announced burial depths [DOB] indicate a strong tendency to detonate Semipalatinsk explosions at a "scale depth", Ds, (*i.e.*, the DOB scaled to a yield of 1 kt) of 117 meters, based on the relation

DOB (meters) = 
$$117 \cdot [W(kt)]^{0.25}$$
 [9]

[9] is determined using 18 Semipalatinsk events (4 Balapan, 6 Murzhik, and 8 Degelen) of known yields and DOB. The yields are assumed to be perturbed by 10% standard error, and the DOB are subject to 0.1% error. Deleting the 8 Degelen events gives very similar result with smaller scatter and slightly larger Ds:

For explosions from Murzhik and Balapan test sites, the focal depths correlate with the announced yields even better than  $m_b$ (NEIS),  $m_b$ (EKA),  $m_b$ (4 UK arrays), and log( $M_o$ , 4 UK arrays) (*cf.* Table 5D of Jih *et al.*, 1990b). Both [9] and [10] strongly suggest that the DOBs of historical Semipalatinsk explosions appeared to be proportional to the **quartic root** of the yields instead of the **cubic root** as observed at NTS. Degelen explosions tend to be underburied except the event 660507. The Soviets seem to have been pushing the DOBs to the shallow limit (and thereby accepting the possible containment risks) at Degelen Mountain during the period 1961-1972.

• 5 Test site bias. There have been several studies which indicate that Nuttli's (1987) "Degelen puzzle" (Table 5) could be invalid because of the relatively poorer quality  $m_b$ (ISC) used (jih *et al.*, 1990). The updated  $m_b$  bias estimate between Eastern Kazakhstan and NTS is systematically larger than 0.35 if  $m_b$  values reported by TG, S-cubed, or UK/AWE are used (Table 6; see also Evernden and Marsh, 1987). The WWSSN data reveal a  $m_b$  bias of 0.13 m.u. (relative to NORSAR's *RMS*  $L_g$ ) between the SW and NE subregions of Balapan Test Site, which confirms what Ringdal and Marshall (1989) found with ISC and NORSAR data. Relative to  $m_b$ ,  $L_g$ -yield relationship does appear to be more transportable, as indicated by the insignificant difference between KTS and NTS calibration curves using *RMS*  $L_g$  (Table 6).

Table 5. Nuttli's (1987) Estimates of $m_b$ Bias (Relative to $m_b(L_g)$							
Test Sites Magnitude Used 10 kt 100 kt 150 kt							
Balapan - NTS	$m_b$ (ISC), $m_b(L_g)$	0.35	0.35	0.35			
Degelen - NTS	$m_b(ISC), m_b(L_g)$	0.58	0.58	0.58			

Table 6. Updated Estimates of Test Site Bias							
Magnitud	KTS-NTS Bias						
KTS NTS		10 kt	100 kt	150 kt			
Marshall <i>et al.</i> <sup>1</sup>	Marshall et al. <sup>2</sup>	0.55	0.48	0.46			
Jih and Wagner <sup>3</sup>	Jih and Wagner <sup>4</sup>	0.47	0.42	0.41			
Jih and Wagner <sup>3</sup>	Murphy <sup>5</sup>	0.36	0.36	0.36			
Murphy <sup>6</sup>	Murphy <sup>5</sup>	0.47	0.41	0.40			
Ringdal <sup>7</sup>	Patton <sup>8</sup>	0.09	0.05	0.04			
Israelson <sup>9</sup>	Patton <sup>8</sup>	0.04	0.06	0.07			

1) Combining all UK/AWE's Balapan, Degelen, and Murzhik m<sub>b</sub> values as listed in Vergino (1989).

2) UK/AWE's NTS m<sub>b</sub> values as distributed in 1987.

3) m<sub>b</sub> = 0.81 log(W) + 4.28 using 19 KTS explosions with published yields (Jih and Wagner, 1991).

4) m<sub>b</sub> = 0.86 log(W) + 3.76 for NTS high-coupling events (Jih and Wagner, 1991).

5) m<sub>b</sub>(S-cubed) = 0.75 log(W) + 4.45 based on the network-averaged spectra (Murphy, 1990).

6) m<sub>h</sub>(S-Cubed) = 3.92 + 0.81 log(W) for NTS high-coupling events (Murphy, 1981).

7) RMS L<sub>a</sub> = 0.71 log(W) + 4.54 with NORSAR RMS L<sub>a</sub> furnished by Ringdal (1990).

8)  $m_{b}(L_{a}) = 0.76 \log(W) + 4.40$  with LLN data (Patton, 1988).

9) RMS  $L_a = 0.78 \log(W) + 4.42$  based on RMS  $L_a$  values furnished by Israelson (1991b) (Jih, 1991).

- 6 Cratering to non-cratering correction. The new calibration curve for Balapan explosions also provides an alternative and straightforward approach to derive the *m<sub>b</sub>* adjustment converting cratering shots to contained explosions of the same yield (*cf.* Table 8A of Jih *et al.*, 1990b). The correction derived by this approach matches that by other studies rather well.
- 7 Distinct features of Degelen and Balapan sites. Degelen Mountain is the only test site that has a decreasing  $\log(P_{max}/P_a)$  and  $\log(P_b/P_a)$  with increasing yields (*cf.* Table 8A of Jih *et al.*, 1990b). It is also the only test site for which the phase "a" shows the smallest scatter around the calibration curve, as compared to the phases "b" and "max" (*cf.* Table 6C of Jih *et al.*, 1990b). Both the mountainous topography (which causes complex pP interference) as well as the testing practice (*e.g.*, the abnormally shallow shot depths and the usage of tunnels) could be responsible. At Balapan, the phase "b" has the smallest scatter around the

calibration curve (*cf.* Table 5C of Jih *et al.*, 1990b). These observations confirm the conjecture that the first cycle could give better results than does the "max" phase in a proper environment.

- 8 Benchmark of various magnitudes. For Balapan events,  $RMS L_g$  reported at NORSAR (Ringdal and Marshall, 1989; Ringdal and Hansen, 1989) and  $m_{2.9}$  based on WWSSN provide the smallest scatter around the calibration curve. In fact, even  $\overline{m}_{2.2}$  would seem to be better than almost all other unclassified magnitudes based on the teleseismic P waves or  $\log(\Psi_{\infty})$  in terms of yield estimation as well as the  $m_b$  scaling against Ringdal's  $RMS L_g$ .
- 9 Abnormal events detected by MLE-CY. The Balapan cratering event 650115 (100-150 kt, 178 meters) and four events at Degelen Mountain (641116, 194 meters; 660629, 187 meters; 661019, 185 meters; and 671017, 181 meters) were rejected by the "MLE-CY" code as outliers. All 4 of these Degelen events were said to have yields between 20 and 150 kt. However, a fully contained explosion at Balapan or Murzhik regions with a shallow DOB of 180 meters or so would be expected to have a yield near 2 kt (*e.g.*, Murzhik event 720902) rather than any value between 20 and 150 kt. Another interesting event is Degelen event 660507 which has an announced yield of 4 kt and a remarkably deep DOB (as compared to Balapan and Murzhik explosions). Perhaps an experiment with a much larger yield was planned for that explosion.

## **II.3.6 REPORTS, PRESENTATIONS, AND PUBLICATIONS**

The publications and presentations generated during the contract period are listed as follows:

(1989) Iterative network magnitude estimation and uncertainty assessment with noisy and clipped data, *Seism. Res. Let.*, **60-1**, 28, presented at 1989 Annual SSA Meeting, Victoria, British Columbia, Canada.

- (1989) Finite-difference simulations of near-regional propagation --- preliminary results, presented at *MIT/ERL 3rd Annual Workshop on Seismic Wave Propagation and Inversion in Heterogeneous Media* (July 31 August 3, 1989, Boston, MA.)
- (1989) Iterative network magnitude estimation and uncertainty assessment with noisy and clipped data, *Bull. Seismo. Soc. Am.*, **79**, 1122-1141.
- (1989) Iterative network magnitude estimation and uncertainty assessment with noisy and clipped data, presented at *Mid-Atlantic Regional Probability and Statistics Meeting* (October 21, 1989, N.I.S.T., Gaithersburg, MD.)
- (1989) Simultaneous modeling of teleseismic and near regional phases with linear finite-difference method, *EOS, Trans. Am. Geophys. Union*, **70-43**, 1189 (1989 Fall AGU Meeting, San Francisco, CA.)
- (1990) Magnitude-yield relationship at various nuclear test sites --- a maximumlikelihood approach using heavily censored explosive yields, *Report GL-TR-90-0107 (=TGAL-90-03)*, Geophysics Laboratory, Hanscom AFB, MA. (ADA223490)
- (1990) Maximum-likelihood magnitude-yield regression with heavily censored data, *EOS, Trans. Am. Geophys. Union*, **71-17**, 566 (1990 Spring AGU Meeting, Baltimore, MD.)
- (1990) Geotech's magnitude-yield study during 1989-1990, Proceedings of 12th DARPA/AFGL Seismic Research Symposium, 281-287 (18-20 Sept 1990, Key West, FL.) (Eds J. Lewkowicz and J. McPhetres), Report GL-TR-90-0212, Geophysics Laboratory, Hanscom Air Force Base, MA. (ADA226635)
- (1990) m<sub>b</sub> bias between Balapan and Degelen Test Sites, U.S.S.R, as revealed by direct regression of WWSSN data on Soviet-published censored and uncensored yields, EOS, Trans. Am. Geophys. Union, 71-43, 1477 (1990 Winter AGU Meeting, San Francisco, CA.)
- (1991) Azimuthal variation of m<sub>b</sub> residuals of E. Kazakh explosions and assessment of the path effects, EOS, Trans. Am. Geophys. Union, 72-17, 193 (1991 Spring AGU Meeting, Baltimore, MD.)
- (1991) A refined network  $m_b$  determination scheme incorporating near-source effects, in *Report PL-TR-91-2212 (=TGAL-91-05)*, Phillips Laboratory, Hanscom Air Force base, MA.
- (1991) Recent methodological developments in magnitude determination and yield estimation with applications to Semipalatinsk explosions (Section A of "Explosion source size determination, discrimination, and spectral characteristics"), Proceedings of 13th DARPA/PL Seismic Research Symposium, (8-10 Oct 1991, Keystone,

CO.) (Eds J. Lewkowicz and J. McPhetres), *PL-TR-91-2208*, Phillips Laboratory, Hanscom Air Force base, MA. (ADA241325)

#### **II.4. CONCLUSIONS AND RECOMMENDATIONS**

The new magnitude determination scheme (Equation [1]) significantly reduces the fluctuational variation across the recording stations, as illustrated in this study. It is shown that, by applying this scheme to worldwide explosions, it is possible to have a consistent base line in estimating the absolute magnitudes (which is crucial in estimating the test site bias) while the precision in the resulting network  $m_b$  values can be maintained as well as could be achieved by the single-test-site approach. The standard error in most  $\overline{m}_{2.9}$  values is 0.02 m.u., about the same as that for *RMS*  $L_g$  inferred from in-country regional recordings reported by Israelson (1991a) and Hansen *et al.* (1990).

The most detailed description of the wave propagation model would naturally suggest that yet another term could be added into the Equation [1] to count for the source-region attenuation. So far such source region bias in  $m_b$  has always been inferred with other information such as the yields (as in this study),  $M_S$  (e.g., Evernden and Marsh, 1987) or  $P_n$  velocity (e.g., Marshall *et al.*, 1979) *etc.*. The hypothesis that such attenuation differential could be directly discerned from  $m_b$  alone by further improving Equation [1] is worth testing.

The two regression routines developed in this study (*i.e.*, "MLE-CY" and "DWLSQ") represent two very different directions in extending the standard least-squares regression. They should be merged together for a more general and robust tool.

Digital signals recorded on in-country seismic instruments at regional distance will be critical for monitoring low yield explosions below 10 kt. Obviously the regression routines developed in this study can well be applied to other source measures which are based on regional phases. On the other hand, teleseismic data such as the WWSSN data used in this study still carry invaluable information that are worthy of further exploitation, beyond simply calculating the "unified yield".

Throughout this study, our emphasis has been to reveal the site-dependent characteristics from the observations exclusively so that in the future the inferred results can be critically examined and compared with those derived by other means. We suggest that the follow-up research be accompanied by well-constrained forward modeling studies using realistic structures. The upgraded LFD code which incorporates boundary conditions for topographical free-surface of arbitrary shape (Jih *et al.*, 1988) in addition to the "strain filter" and "marching grid" features as outlined in Jih *et al.* (1989) can be utilized in the future to improve our understanding of the fundamental issues of seismic energy partitioning on the focal sphere as well as their implications for yield determination.

## **II.5 ACKNOWLEDGEMENTS**

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Bonham *et al.* (1980). Wilmer Rivers ported the code "MLE-CY" to classified computer facility at CSS, and he also reviewed all manuscripts generated under this project. Richard Baumstark, Tom McElfresh, and Mary Ann Brennan always provide prompt answers and assistance on our questions about more efficient use of UNIX software. This research was supported under DARPA contract F19628-89-C-0063, monitored by Phillips Laboratory. The views and conclusions contained in this paper are those of the author and should not be interpreted as representing the official policies, either expressed or implied, of the Defense Advanced Research Projects Agency or the U.S. Government.

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4

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