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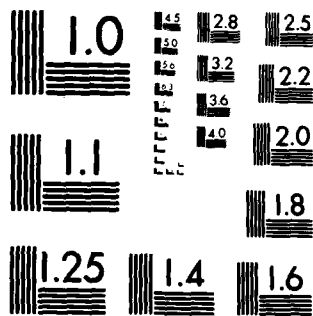
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(6) A SEQUENTIAL CONFIDENCE INTERVAL
FOR THE ODDS RATIO.

BY

(10) D. SIEGMUND

(9) TECHNICAL REPORT NO. 7
(11) 2 FEBRUARY 1980

(14) TR-7
TR-141

(12) 19

PREPARED UNDER CONTRACT

(15) N00014-77-C-0306, [REDACTED] NSF-MCS77-16
FOR THE OFFICE OF NAVAL RESEARCH

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by

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Technical Report No. 7 ✓

February 2, 1980

Prepared under Contract

✓N00014-77-C-0306 (NR-042-373)

for the Office of Naval Research

D. Siegmund, Project Director

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STANFORD UNIVERSITY

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Also issued as Technical Report No. 141 under National Science
Foundation Grant MCS77-16974 - Department of Statistics,
Stanford University.

ABSTRACT

A sequential fixed width confidence interval is proposed for the log odds ratio of a 2×2 table. It is shown that the proposed interval has asymptotically the correct coverage probability and is asymptotically efficient uniformly in the unknown parameters.

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AMS 1970 Subject Classification. Primary 62L10

Key Words and Phrases. Fixed width confidence interval, log odds ratio, stopping rule, uniformly asymptotically normal.

A SEQUENTIAL CONFIDENCE INTERVAL FOR THE ODDS RATIO

1. Introduction

For $i = 1, 2$ let s_{in_i} and $f_{in_i} = n_i - s_{in_i}$ be respectively the numbers of successes and failures in n_i independent Bernoulli trials with constant success probability p_i on each trial. A simple large sample approximate confidence interval for the log odds ratio, $\log(p_1 q_2 / p_2 q_1)$, is

$$(1) \quad \log(s_{1n_1} f_{2n_2} / s_{2n_2} f_{1n_1}) \pm z_{\alpha} [n_1 / s_{1n_1} f_{1n_1} + n_2 / s_{2n_2} f_{2n_2}]^{1/2},$$

where $\int_{z_{\alpha}}^{\infty} (2\pi)^{-1/2} \exp(-x^2/2) dx = \alpha/2$ (Cox, 1970, p. 35). The confidence coefficient, $1 - \alpha$, is asymptotically correct for fixed p_1, p_2 as $\min(n_1, n_2) \rightarrow \infty$.

These intervals have two defects when p_1 and p_2 may be near 0 or 1. On the one hand the rate of approach to normality can be very slow, so that use of asymptotic theory is questionable. More importantly, however, even with exact calculations, no fixed sample size design will permit one to estimate the log odds ratio by an interval of preassigned width in these boundary cases.

For one binomial population with success probability p , Robbins and Siegmund (1974) proposed a sequential scheme for obtaining approximately a confidence interval of preassigned width for $\log(p/q)$. However, they do not consider the question of the uniformity of their procedure for p near 0 or 1, when a sequential procedure would presumably be of greatest value.

The purpose of this paper is to consider the two population analogue of the procedure of Robbins and Siegmund. The procedure will be seen to attain asymptotically the required coverage probability and to be asymptotically efficient uniformly in $0 < p_1, p_2 < 1$.

In Section 2 the one-population case is reviewed, and the results of Robbins and Siegmund are appropriately strengthened to provide the tools for the two-population problem. It is also shown that Robbins and Siegmund's uncritical acceptance of Haldane's (1955) modification of the empirical log odds ratio is inappropriate in the sequential case.

Section 3 is concerned with the case of two populations. Remarks about further extensions are collected in Section 4.

2. One Population

Let $x_1, x_2, \dots$ be independent with $P\{x_j = 1\} = p$, $P\{x_j = 0\} = q = 1 - p$ ($j = 1, 2, \dots$). Let $s_n = x_1 + \dots + x_n$ and $f_n = n - s_n$. For large n $\log(s_n/f_n)$ is approximately normally distributed with mean $\log(p/q)$ and variance $1/(npq)$. Hence to find a confidence interval for $\log(p/q)$ of preassigned width, or equivalently in large samples to estimate $\log(p/q)$ by an estimator with preassigned variance $1/c$, Robbins and Siegmund (1974) define

$$(2) \quad T = \inf\{n : s_n f_n > nc\}.$$

They propose estimating $\log(p/q)$ by

$$(3) \quad \log[(s_T + \frac{1}{2})/(f_T + \frac{1}{2})],$$

which they show is asymptotically normally distributed with mean $\log(p/q)$ and variance $1/c$ as $c \rightarrow \infty$. The modification of the empirical log odds by adding $1/2$ to numerator and denominator was originally suggested by Haldane (1955) as a bias reducing device in the fixed sample case. Robbins and Siegmund also show that $ET \sim c/(pq)$ as $c \rightarrow \infty$. This may be interpreted as showing that their procedure is asymptotically efficient in the sense of requiring asymptotically about the same number of observations as a fixed sample procedure chosen to be appropriate for a value p_0 which happens to be the actual value of p .

In this section it is shown that the asymptotic normality of (3) holds uniformly over $0 < p < 1$. This is in marked contrast to the fixed sample case, as was noted in the Introduction. It will

also be shown that the analogue of Haldane's bias reducing device in this sequential context is to subtract $\frac{1}{2}$ from numerator and denominator of the empirical odds ratio. However, for simplicity and because the appropriate modification for the two sample case is unknown, in most of what follows only the unmodified empirical odds ratio is considered.

The main result of this section is Theorem 1. Lemma 1, which was obtained by Robbins and Siegmund (1974), is of interest in its own right. It says that the asymptotic efficiency of (2) is uniform in $0 < p < 1$. Repeated use will be made of the algebraic identity

$$(4) \quad s_n f_n / n = (q - p)(s_n - np) + npq - (s_n - np)^2 / n .$$

Theorem 1. For the stopping rule T defined in (2), uniformly in $0 < p < 1$

$$\lim_{c \rightarrow \infty} P\{c^{-1/2} [\log(s_T/f_T) - \log(p/q)] \leq x\} = \Phi(x) ,$$

where

$$\Phi(x) = (2\pi)^{-1/2} \int_{-\infty}^x \exp(-u^2/2) du .$$

The proof utilizes the following lemmas. For the simple proof of Lemma 1 based on (4), see Robbins and Siegmund (1974).

Lemma 1. $c < pq$ $ET < (c+1)/[1 - (4c)^{-1}]$.

Lemma 2. There exists a c_0 such that for all $c \geq c_0$ and all $0 < p < 1$

$$(pq)^2 E(T - c/pq)^2 \leq 4c .$$

Lemma 3. For each $0 < \epsilon < 1$ and $c \geq c_0$, where c_0 is defined as in Lemma 2,

$$P\{|s_T - pT| \geq \epsilon pq T\} < \kappa/\epsilon^2 c ,$$

where κ does not depend on ϵ or c .

Proof of Lemma 2. Squaring (4) gives

$$\begin{aligned} (s_n f_n / n - c)^2 &= (q - p)^2 (s_n - np)^2 + (pq)^2 (n - c/pq)^2 + (s_n - np)^4 / n^2 \\ &+ 2\{(q - p)(s_n - np)(pqn - c) - (q - p)(s_n - np)^3 / n - (npq - c)(s_n - np)^2 / n\} . \end{aligned}$$

By the Schwarz inequality and Wald's second moment identity

$$|E\{(s_T - pT)(T - c/pq)\}| \leq \{pqET E(T - c/pq)^2\}^{1/2} .$$

Hence, since $(s_T f_T / T - c)^2 < 1$, Wald's second moment identity yields

$$\begin{aligned} 1 &> (q - p)^2 pq E(T) + (pq)^2 E(T - c/pq)^2 - 2pq|q - p|\{pq E(T)E(T - c/pq)^2\}^{1/2} \\ &- 2|q - p|pq ET - 2(pq)^2 ET , \end{aligned}$$

or

$$\begin{aligned} (pq)^2 E(T - c/pq)^2 - 2pq|q - p|\{pqET E(T - c/pq)^2\}^{1/2} + pq(q - p)^2 ET \\ < 1 + 2pq ET . \end{aligned}$$

Taking square roots in this expression, then rearranging terms and squaring yields

$$\begin{aligned} (pq)^2 E(T - c/pq)^2 &\leq \{(pqET)^{\frac{1}{2}} + (1 + 2pqET)^{\frac{1}{2}}\}^2 \\ &\leq 1 + 3pqET \leq 1 + 3(c+1)/(1 - 1/4c) , \end{aligned}$$

where the last inequality follows from Lemma 1. This completes the proof.

Proof of Lemma 3. Let $0 < \delta < 1$ and $n_0 = c/pq$. By Lemma 2

$$P\{|T - n_0| > \delta c/pq\} \leq (\delta c)^{-2} (pq)^2 E(T - n_0)^2 \leq 4/\delta^2 c .$$

Hence, by Wald's lemma for the second moment and Lemma 1

$$\begin{aligned} P\{|s_T - pT| > \epsilon pqT\} &\leq 4/\delta^2 c + P\{|s_T - pT| > \epsilon pqT, |T - n_0| \leq \delta c/pq\} \\ &\leq 4/\delta^2 c + P\{|s_T - pT| > \epsilon(1 - \delta)c\} \leq 4/\delta^2 c + E(s_T - pT)^2 / \epsilon^2 (1 - \delta)^2 c^2 \\ &\leq 4/\delta^2 c + 2/[\epsilon^2 (1 - \delta)^2 c] . \end{aligned}$$

Proof of Theorem 1. From the mean value theorem follows

$$\begin{aligned} c^{\frac{1}{2}} [\log(s_T/f_T) - \log(p/q)] &= c^{\frac{1}{2}} (s_T - pT)/pqT + c^{\frac{1}{2}} [(s_T - pT)/pqT] \left[\frac{pq}{\eta_T(1 - \eta_T)} - 1 \right] \\ &= n_0^{\frac{1}{2}} (s_T - pT)/(pq)^{\frac{1}{2}} T + n_0^{\frac{1}{2}} [(s_T - pT)/(pq)^{\frac{1}{2}} T] \left[\frac{pq}{\eta_T(1 - \eta_T)} - 1 \right] , \end{aligned}$$

where $|\eta_T - p| \leq |T^{-1} s_T - p|$, and as before $n_0 = c/pq$. Hence it suffices to show that uniformly in $0 < p < 1$

$$\lim_{c \rightarrow \infty} P\{n_0^{1/2}(s_T - pT)/(pq)^{1/2} T \leq x\} = \Phi(x)$$

and

$$pq/\eta_T(1 - \eta_T) \xrightarrow{P} 1.$$

The second statement follows easily from Lemma 3, and the first may be obtained by minor modifications in the standard proof of Anscombe's theorem (e.g., Rényi, 1966, p. 390).

An asymptotically more precise approximation to $E(T)$ than that provided by Lemma 1, although one which is decidedly not uniform in p , is

$$(5) \quad pq ET = c + \frac{1}{2}(p - q)^2 + \frac{1}{2} pq + o(1) \quad (c \rightarrow \infty),$$

which is valid for all p for which $(p/q)^2$ is irrational. This result follows easily from (4) and Theorem 2 of Lai and Siegmund (1979).

As an estimator of $\log(p/q)$, Haldane (1955) considered $\log\{(s_n + a)/(f_n + a)\}$ and showed by a Taylor series expansion that the choice of a minimizing the asymptotic bias of this estimator is $a = \frac{1}{2}$. The following heuristic calculation shows that $a = -\frac{1}{2}$ is appropriate in the present context. The machinery for justifying this calculation may be found in Pollak and Siegmund (1975). It should be noted that this result is appropriate for the stopping rule T defined by (2). It does not carry over to the two-population case discussed in Section 3.

A two term Taylor series expansion gives

$$\begin{aligned} \log\{(s_T + a)/(f_T + a)\} - \log(p/q) &= (s_T - pT + a)/pT - (f_T - qT + a)/qT \\ &\quad - (s_T - pT)^2/2(pT)^2 + (f_T - qT)^2/2(qT)^2 + o_p(c^{-1}) . \end{aligned}$$

Since $T \stackrel{P}{\sim} c/pq$ and hence $E\{(s_T - pT)^2/T^2\}$

$$\sim (pq c^{-1})^2 E\{(s_T - pT)^2\} = (pq c^{-1})^2 pq ET \sim (pq)^2/c ,$$

one obtains

$$\begin{aligned} &E[\log\{(s_T + a)/(f_T + a)\}] - \log(p/q) \\ (6) \quad &\sim E\{(s_T - pT)/pqT\} + c^{-1}(q - p)(a - \frac{1}{2}) . \end{aligned}$$

It is shown below that

$$E\{(s_T - pT)/T\} \sim c^{-1} pq(q - p) \quad (c \rightarrow \infty) ,$$

which shows that the right hand side of (6) is $\sim c^{-1}(q - p)(a + \frac{1}{2})$,
leading to the optimal choice $a = -\frac{1}{2}$.

Let $\xi_T = s_T f_T / T - c$. By (4) and Taylor expansions, one obtains

$$\begin{aligned} (s_T - pT)/T &= (q - p)^{-1}(c + \xi_T)\{pq c^{-1} - (pq c^{-1})^2(T - c/pq) \\ &\quad + (pq c^{-1})^3(T - c/pq)^2 + \dots\} - (q - p)^{-1} pq + (q - p)^{-1} (s_T - pT)^2/T^2 . \end{aligned}$$

It is easy to see from (4) that $c + E\xi_T = pq ET - pq + o(1)$; and

Robbins and Siegmund (1974) have obtained $E(T - c/pq)^2 = (q-p)^2 c/(pq)^2 + O(1)$. Hence by the asymptotic independence of ξ_T and $c^{-1/2}(T - c/pq)$ (Lai and Siegmund, 1977),

$$\begin{aligned} E\{(s_T - pT)/T\} &= (q-p)^{-1}(c + E\xi_T)\{pq c^{-1} - (pq c^{-1})^2[E\xi_T/pq + 1] \\ &\quad + (pq c^{-1})^3(q-p)^2 c/(pq)^2\} - (q-p)^{-1} pq + (pq)^2/(q-p)c + o(c^{-1}) \\ &\sim c^{-1} pq(q-p) \quad , \end{aligned}$$

as claimed.

3. Two Populations

Consider again the two population case described in the introduction and suppose that observations are taken in pairs, one from each population, so $n_1 = n_2 = n$, say. This restriction is stronger than necessary, but it simplifies the subsequent analysis. It is easy to modify the results to accommodate the case in which observations are taken from the two populations in an arbitrary fixed ratio. It seems possible to achieve a slight reduction in the total expected sample size by choosing the sampling rates adaptively, but the fairly small improvement seems not to be worth the considerable complication in analysis.

The obvious analogue of the stopping rule (2) is (cf. (1))

$$(7) \quad T = \inf\{n: n\left(\frac{1}{s_{1n}f_{1n}} + \frac{1}{s_{2n}f_{2n}}\right) \leq \frac{1}{c}\}.$$

The main results of this section are Theorems 2 and 3, which correspond respectively to Lemma 1 and Theorem 1 in the single population case. Theorem 2 shows that T defined by (7) is uniformly asymptotically efficient and Theorem 3 shows that it asymptotically provides the correct coverage probability uniformly in p_1, p_2 .

Theorem 2. Uniformly in $0 < p_1, p_2 < 1$,

$$ET \sim c\{(p_1q_1)^{-1} + (p_2q_2)^{-1}\} \quad (c \rightarrow \infty).$$

The inequality in one direction is a consequence of the following trivial lemma.

Lemma 4. For all $0 < p_1, p_2 < 1$ and all c

$$ET \geq c\{(p_1 q_1)^{-1} + (p_2 q_2)^{-1}\} .$$

Proof. From (4), Wald's identity and Jensen's inequality one obtains

$$\begin{aligned} c^{-1} &\geq E\{T(1/s_{1T} f_{1T} + 1/s_{2T} f_{2T})\} \geq \{E(s_{1T} f_{1T}/T)\}^{-1} \\ &\quad + \{E(s_{2T} f_{2T}/T)\}^{-1} = \{p_1 q_1 ET - E[(s_{1T} - p_1 T)^2/T]\}^{-1} \\ &\quad + \{p_2 q_2 ET - E[(s_{2T} - p_2 T)^2/T]\}^{-1} \geq (ET)^{-1} \{(p_1 q_1)^{-1} + (p_2 q_2)^{-1}\} . \end{aligned}$$

To obtain asymptotic upper bounds on $E(T)$ it is useful to define (cf. (2))

$$T_1(c) = \inf\{n : n/s_{in} f_{in} \leq 1/c\} .$$

Since $s_{in} f_{in}/n$ increases with n , for all $\alpha > 1$ and $\beta > 1$ with $1/\alpha + 1/\beta = 1$,

$$(8) \quad T \leq \max(T_1(\alpha c), T_2(\beta c)) .$$

In what follows $\alpha = (p_1 q_1 + p_2 q_2)/p_2 q_2$ and $\beta = (p_1 q_1 + p_2 q_2)/p_1 q_1$, so

$$(9) \quad \alpha/p_1 q_1 = \beta/p_2 q_2 = (p_1 q_1 + p_2 q_2)/(p_1 q_1 p_2 q_2) = (p_1 q_1)^{-1} + (p_2 q_2)^{-1} .$$

With these fixed values of α and β there is no ambiguity in writing

T_1 for $T_1(\alpha c)$ and T_2 for $T_2(\beta c)$.

It is now possible to complete the proof of Theorem 2.

Obviously from (8)

$$(10) \quad E(T) \leq E\{\max(T_1, T_2)\} = \int_{\{T_1 \leq T_2\}} T_2 dP + \int_{\{T_2 < T_1\}} T_1 dP.$$

Let $\epsilon > 0$ be arbitrary. Then

$$(11) \quad \int_{\{T_1 \leq T_2\}} T_2 dP \leq \int_{\{T_1 \leq T_2, T_2 \leq \beta c(1+\epsilon)/p_2 q_2\}} T_2 dP + \int_{\{T_2 > \beta c(1+\epsilon)/p_2 q_2\}} T_2 dP$$

$$\leq (p_2 q_2)^{-1} \beta c(1+\epsilon) P\{T_1 \leq T_2\} + (p_2 q_2)^{-1} \beta c P\{T_2 > (p_2 q_2)^{-1} \beta c(1+\epsilon)\}$$

$$+ \int_{\{T_2 > (p_2 q_2)^{-1} \beta c(1+\epsilon)\}} |T_2 - (p_2 q_2)^{-1} \beta c| dP.$$

By Lemma 2

$$(p_2 q_2)^{-1} \beta c P\{T_2 > (p_2 q_2)^{-1} \beta c(1+\epsilon)\} \leq 4(p_2 q_2)^{-1} \epsilon^{-2};$$

and by the Schwarz inequality and Lemma 2 again

$$\int_{\{T_2 > (p_2 q_2)^{-1} \beta c(1+\epsilon)\}} |T_2 - (p_2 q_2)^{-1} \beta c| dP \leq [(p_2 q_2)^{-2} E|p_2 q_2 T_2 - \beta c|^2 P\{T_2 > (p_2 q_2)^{-1} \beta c(1+\epsilon)\}]^{1/2}$$

$$\leq 4(p_2 q_2)^{-1} \epsilon^{-1}.$$

Putting these inequalities together with (9), (10), and (11) yields

$$ET \leq c\{(p_1 q_1)^{-1} + (p_2 q_2)^{-1}\}(1 + \epsilon + 8/\epsilon^2 c),$$

which completes the proof, as ϵ is arbitrarily small.

Theorem 3. For T defined by (7), uniformly in $0 < p_1, p_2 < 1$

$$\lim_{c \rightarrow \infty} P\{\sqrt{c} [\log(s_{1T} f_{2T} / s_{2T} f_{1T}) - \log(p_1 q_2 / p_2 q_1)] \leq x\} = \Phi(x).$$

With the help of Lemma 5 below, the proof of Theorem 3 may be carried out along the same lines as the proof of Theorem 1.

Lemma 5. Let $\mu = \{(p_1 q_1)^{-1} + (p_2 q_2)^{-1}\}^{-1}$. For all $\epsilon > 0$ and all large c (not depending on ϵ)

$$P\{|\mu T - c| > c\epsilon\} \leq 8/c\epsilon^2.$$

Proof. The proof of Theorem 2 shows that

$$\begin{aligned} P\{T > c(1+\epsilon)\mu^{-1}\} &= P\{T_1 < T_2, T_2 > (p_2 q_2)^{-1} \beta c(1+\epsilon)\} \\ &+ P\{T_2 \leq T_1, T_1 > (p_1 q_1)^{-1} \alpha c(1+\epsilon)\} \leq 4/c\epsilon^2. \end{aligned}$$

The same upper bound for $P\{T < c(1-\epsilon)\mu^{-1}\}$ follows by a similar calculation and the observation that $T \geq \min(T_1(\alpha c), T_2(\beta c))$.

4. Remarks

(a) Unpublished numerical computations of H. Levene in the one-sample case show that the asymptotic theory of Section 2 provides good approximations for $c \geq 10$ and reasonable ones for c as small as

3. It seems likely that similar results hold for two populations.

(b) The heuristic principle which suggests the stopping rules (2) and (7) is quite common in the literature of fixed precision estimation (e.g., Anscombe, (1953)), and it leads to reasonable stopping rules for more complicated log linear models. However, the uniform asymptotic theory developed here seems to require new ideas for very simple extensions.

One important generalization is a set of 2×2 tables with equal odds ratios. Appropriate asymptotic theory might involve a large number of observations from each of a small number of tables or a large number of tables.

Another interesting variation is log linear regression. In this case one might also wish to consider sequential design in selecting values of the independent variable.

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4. TITLE (and Subtitle) A SEQUENTIAL CONFIDENCE INTERVAL FOR THE ODDS RATIO		5. TYPE OF REPORT & PERIOD COVERED Technical Report
		6. PERFORMING ORG. REPORT NUMBER
7. AUTHOR(s) D. SIEGMUND		8. CONTRACT OR GRANT NUMBER(s) N00014-77-C-0306
9. PERFORMING ORGANIZATION NAME AND ADDRESS DEPARTMENT OF STATISTICS STANFORD UNIVERSITY STANFORD, CALIFORNIA		10. PROGRAM ELEMENT, PROJECT, TASK AREA & WORK UNIT NUMBERS (NR-042-373)
11. CONTROLLING OFFICE NAME AND ADDRESS OFFICE OF NAVAL RESEARCH STATISTICS & PROBABILITY PROGRAM CODE 436 ARLINGTON, VA 22217		12. REPORT DATE February 2, 1980
		13. NUMBER OF PAGES 15
14. MONITORING AGENCY NAME & ADDRESS (if different from Controlling Office)		15. SECURITY CLASS. (of this report) UNCLASSIFIED
		16a. DECLASSIFICATION/DOWNGRADING SCHEDULE
16. DISTRIBUTION STATEMENT (of this Report) APPROVED FOR PUBLIC RELEASE AND SALE: DISTRIBUTION UNLIMITED.		
17. DISTRIBUTION STATEMENT (of the abstract entered in Block 20, if different from Report)		
18. SUPPLEMENTARY NOTES Also issued as Technical Report No. 141 under National Science Foundation Grant MCS77-16974 - Department of Statistics, Stanford University.		
19. KEY WORDS (Continue on reverse side if necessary and identify by block number) Fixed width confidence interval, log odds, ratio, stopping rule, uniformly asymptotically normal.		
20. ABSTRACT (Continue on reverse side if necessary and identify by block number) A sequential fixed width confidence interval is proposed for the log odds ratio of a 2 x 2 table. It is shown that the proposed interval has asymptotically the correct coverage probability and is asymptotically efficient uniformly in the unknown parameters.		

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