



SOFTWARE RELIABILITY ESTIMATION UNDER CONDITIONS OF INCOMPLETE INFORMATION

University of Utah

RADC-TR-79-230 Final Technical Report

October 1979

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EVALUATION

The increased importance of software for embedded avionics systems has led to an increasing desire to insure that avionics software meets very strict reliability and quality goals. However, a significant problem in assuring such goals are met is the inability of Government personnel to accurately predict the reliability of an avionics software development project. This problem has been expressed at several Government and industry sponsored conferences, as well as in documents such as the Joint Logistics Commanders Software Reliability Working Group Report (November 1975) and the Joint Logistics Commanders Software Quality Management Workshop Report (July 1979). As a result, efforts have been initiated to develop and validate mathematical models for predicting the reliability and error content of a software system. However, models developed to date have not adequatley addressed the unique features of avionics software developments.

This effort was initiated in response to the need for developing software reliability prediction models applicable to avionics software developments, and fits into the goals of RADC TPG No. 5, Software Cost Reduction, in the subthrust of Software Quality (Software Modeling). This report summarizes the development of a mathematical model for predicting the reliability and mean-time-to-failure of a software development under the assumptions of incomplete information available on error correction, and discrete versions of the software being developed. The report also describes the modified nonlinear search algorithm developed for finding model parameters and an accompanying computer program for operating the model. The importance of this model development is that the assumptions underlying this model more closely reflect the actual avionics software development process than prior model developments.

The theory and model algorithm developed under this effort will lead to much needed predictive measures for use by software managers of avionics software developments in adequately tracking those developments in terms of reliability and mean-time-to-failure objectives. More importantly, the measures developed under this effort will be applicable to current avionics software developments and thus help to produce the high quality, low cost avionics software needed for today's aircraft.

alan N. ALAN N. SUKERT Project Engineer

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1.0 Introduction

1.1 Problem Statement

As the cost and complexity of computer software continue to increase, there is a growing need for accurate determination of software reliability. Before a software package is put into operation, there is a testing period during which errors are detected and corrected. The problem with which we are concerned is the estimation of certain reliability parameters from the error data generated during the test phase. Specifically, we wish to estimate the number of errors remaining in the software package at any time, and the mean time to failure (MTTF). Accurate determination of these parameters could reduce the cost associated with excessive testing, and could increase the confidence with which the package is used.

In order to estimate software reliability, it is necessary to develop an appropriate model describing the error detection and correction processes, and to develop procedures for estimating the parameters of this model from observed error data. Our intention is to generalize certain models which have previously been used for this purpose in order to depict more accurately an actual testing environment. In addition, we will consider a somewhat different approach to the estimation of the parameters of this generalized model.

1.2 Previous Work

A substantial body of work now exists on the application of statistical modeling and estimation techniques to the determination of

- 1.1 -

software reliability. We make no attempt to describe all this work, but rather restrict ourselves to those efforts which are directly related to our own. For a comprehensive review and bibliography, see [1] or [2].

One of the most widely-used error models was developed by Jelinski and Moranda [3]. A similar model has been considered by Shooman [4] a others. The assumptions about the error-detection and error-correction processes which underlie this model are the following:

- (a) The error-detection process is a Poisson process whose detection rate is constant between error detections.
- (b) The error-detection rate at the time prior to the detection of the ith error is a function of i; it is denoted by z_i. It is commonly assumed that z_i is proportional to the number of errors in the program at detection time. This can be written as:

$$z_{i} = \phi \left(N_{o} - i + 1 \right)$$
(1.1)

where N is the initial number of errors and ϕ is a positive constant. An alternative assumption is that the detection rate forms a geometric progression

$$z_{i} = \lambda a^{i}$$
(1.2)

with both λ and a being positive constants. It should be noted that the main justification for (1.2) is the improved

- 1.2 -

convergence of the resulting estimator equations [5].

- (c) Error detection is followed by an immediate correction. Consequently, upon detection of the ith error, the number of remaining errors drops to $(N_0 - i)$.
- (d) The debugging process is perfect and no new errors are generated by the correction process.

These assumptions, although restrictive, were initially adopted by most of the researchers in the field. The estimation of the reliability parameters was based on the above assumptions, and employed the maximum likelihood (ML) criterion to derive the best estimates.

It is realized now that the assumptions given above are quite restrictive and unrealistic in most cases, and steps have been taken to make the model more realistic. The model assumptions have been changed to comply more closely with the real process.

Goel [6] has considered a nonideal debugging process in which the probability of correcting an error is p. Based on this assumption, an analysis of the resulting model is performed. Further generalization is suggested by Shooman [7], who modified both assumptions c and d above concerning the error-correction process. According to the modified model of Shooman, the correction process does not necessarily proceed identically to the detection process, and new errors may be introduced. Denote by $r_d(t)$, $r_c(t)$, and $r_g(t)$ the rates of error detection, correction, and new error generation, respectively. The models suggested by Shooman assume different relationships between these rates. The main models are:

- 1.3 -

Model 1

$$r_{c}(t) = \beta r_{d}(t) \qquad (1.3)$$

$$r_{g}(t) = \alpha r_{c}(t) \qquad (1.4)$$

and

Model 2

$$r_{d}(t) = b r_{d}(t)$$
 (1.5)

$$r_{g}(t) = a n(t) r_{d}(t)$$
 (1.6)

where n(t) represents the number of errors in the program.

These models and others have been studied by Shooman, and the results are described [7].

Another generalization of the original model concerns the assumption that the corrections are implemented continuously. This is not consistent with actual practice in which a program is replaced by a newer version at discrete times. Between the replacement times, the program undergoing the test is the same and the number of errors in it is constant. A possible solution for this discrepancy is that rediscovery of errors should not be counted. However, this requires the analysis of the source of errors in order to determine whether the error sources are the same, and this is not always practical. A modified model in which this generalization was implemented was discussed by

- 1.4 -

Tal [5] and by Sukert [2], and estimator equations for use with this model were developed.

These generalizations, along with some additional ones, will be incorporated into a new model. The new model, we believe, more accurately describes an actual testing environment. We will first discuss the behavior of this model as a function of its parameters under the simplifying assumption that the error processes are deterministic rather than random.

After presenting certain results for deterministic processes, we will then show results using simulated random error data. We have developed a least-squares search procedure for estimating the model parameters, and will discuss its convergence behavior. Recommendations are made toward increased utility, and toward closer coupling of the algorithm to information in real test data.

The true test of the usefulness of the model will lie in its ability to describe real software tests. Thus, there remains for subsequent work the application of the model to enough real cases to draw conclusions concerning validity.

One of the difficulties encountered by researchers in the past has been the inadequacy, incompleteness, and ambiguity of available test data. We found some of these same problems with the data available to us during this work. Hence, we include comments regarding data requirements.

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2.0 Development and Analysis of the Model

2.1 Assumptions

In order to develop a generalized model to describe the error detection and correction processes, we make the following assumptions:

(a) The error detection process is a Poisson process whose average rate of occurrence is proportional at any time to the number of errors present in the software package. Denoting the number of errors present at time t by N(t) and the average error occurrence rate by $r_d(t)$, we have

$$\mathbf{r}_{d}(t) = \phi \mathbf{N}(t) \tag{2.1}$$

where ϕ is a fixed constant of proportionality.

- (b) No attempt is made to correct detected errors at the time of detection. Instead, a new and corrected version of the program is provided to the testing group at discrete ("tape replacement") times t₁, t₂, ..., t_j, Thus, the number of errors present in the program at time t, t_j < t < t_{j+1}, is constant and equal to N(t_j). This is illustrated in Fig. 1.
- (c) Of the detected errors reported to the correcting group, some are corrected and some are not. In addition, new errors are generated. Denote the cumulative number of errors corrected to time t in the program being tested by $N_c(t)$, and the cumulative number of newly-generated errors by $N_g(t)$. Both N_c and N_g are piecewise constant because of the assumption

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Fig. 1. Number of errors in program as a function of time.

that a new version of the program is provided only at the discrete times t_1, t_2, \ldots . A key feature of our model is that many errors may be detected, corrected, and generated between update times. At any time t during testing, we have

$$N(t) = N_{o} - N_{c}(t) + N_{g}(t)$$
(2.2)

where N_0 is the initial number of errors in the program. (d) The error correction rate $r_c(t)$ depends on both the error detection rate $r_d(t)$ and the error backlog $N_b(t)$, where

$$N_{b}(t) = N_{d}(t) - N_{c}(t).$$
 (2.3)

For simplicity, we assume a linear relationship

$$r_{c}(t) = \alpha r_{d}(t) + \beta N_{b}(t).$$
 (2.4)

The addition of the second term in (2.4) represents a generalization of the model of Shooman [7].

(e) The rate of generation of new errors is proportional to the error-correction rate:

$$r_{g}(t) = \gamma r_{c}(t). \qquad (2.5)$$

(f) The error-detection process $N_d(t)$ is precisely known, but the error-correction process $N_c(t)$ is unknown. This appears to be a realistic assumption in view of the way

- 2.3 -

error correction is actually performed. The error generation process is also unknown.

In the above we tacitly equate software "failure" to coding "faults". In effect, we include in α and ϕ the proportionality between the two, and call them both "errors".

2.2 The Model

The model which we develop is actually a deterministic model which relates the expected values of the various random processes involved. The required connection between the observed sample functions of the random processes involved and the deterministic model is established by means of an estimation algorithm which operates on the observed data to estimate model parameters. The deterministic model will be described first, followed by a discussion of the estimation procedure.

Taking expected values of (2.1)-(2.4) yields the equations

$$r_{d}(t) = \phi n(t),$$
 (2.6)

$$r_{c}(t) = \alpha r_{d}(t) + \beta n_{b}(t),$$
 (2.7)

$$n(t) = N_{o} - n_{c}(t) + n_{g}(t),$$
 (2.8)

$$n_{b}(t) = n_{d}(t) - n_{c}(t),$$
 (2.9)

$$n_{g}(t) = \gamma n_{c}(t),$$
 (2.10)

- 2.4 -

where a lower-case n denotes the expected value of the process repre-. sented by the corresponding upper-case N.

It follows from the relationship between $r_d(t)$ and $n_d(t)$ that

$$n_{d}(t) = \int_{0}^{t} r_{d}(t) du.$$
 (2.11)

Similarly,

$$n_{c}(t) = \int_{0}^{t} r_{c}(u) du.$$
 (2.12)

The model represented by the above equations can be viewed as a linear system with sampling and feedback as shown in Fig. 2. Our problem is now one of system identification: Given $N_d(t)$, estimate the parameters of the system shown in Fig. 2. Revisions to the software are applied at time instants t_k , between which times n(t) remains constant. The system therefore is treated as a discrete-time system. We employ the usual notation k in place of the argument t_k .

The four system equations (2.6-2.9) can be reduced to two:

$$r_{d}(k) = \phi \left[N_{o} - (1 - \gamma) n_{c}(k) \right]$$
 (2.13)

$$r_{c}(k) = \alpha \phi \left[N_{o} - (1 - \gamma) n_{c}(k) \right] + \beta \left[n_{d}(k) - n_{c}(k) \right]$$
 (2.14)

- 2.5 -



Fig. 2. Block diagram representation of the proposed model for the error-detection and the error-correction processes.

and the number of parameters reduced to four:

$$\mathbf{r}_{d}(\mathbf{k}) = \phi_{a} \left[N_{a} - n_{c}(\mathbf{k}) \right]$$
(2.15)

$$\mathbf{r}_{c}(\mathbf{k}) = \alpha \phi_{a} \left[N_{a} - n_{c}(\mathbf{k}) \right] + \beta \left[n_{d}(\mathbf{k}) - n_{c}(\mathbf{k}) \right]$$
(2.16)

where

$$\phi_{a} = (1 - \gamma)\phi, \qquad N_{a} = N_{o}/(1 - \gamma).$$
 (2.17)

The application of Laplace transform techniques and some algebraic manipulation similarly lead to the equivalent block diagram shown in

- 2.6 -



Fig. 3. Simplified model.

Fig. 3. The identification problem reduces to the estimation of the four parameters N_a , ϕ_a , α , and β . Note that N_a is the sum of the initial errors N_o and all errors which are subsequently generated during the correction process. Note further that the ultimately sought reliability factor, mean-time-to-failure, is:

MTTF =
$$\frac{1}{r_d(k)} = \frac{1}{\phi_a n_a(k)} = \frac{1}{\phi_a N_a - n_c(k)}$$
. (2.18)

Defining the discrete state

It is noted that the dynamics of the system can be studied using the even simpler nondimensionalized three-parameter system, using (ϕ T), (β T), and (n/N_a), with unit step input.

$$\vec{n}(k) = \begin{pmatrix} n_{d}(k) \\ \\ \\ n_{c}(k) \end{pmatrix}$$
(2.19)

and using the usual approximation, which in our case is exact,

$$\vec{r}(k) = \frac{\vec{n}(k+1) - \vec{n}(k)}{T(k)}$$
, $T(k) = t(k+1) - t(k)$ (2.20)

the model becomes

$$\vec{n}(k+1) = \vec{Ln}(k) + N_a \phi_a B$$
 (2.21)

where

$$L = \begin{pmatrix} 1 & -\phi_{a}T(k) \\ & & \\ \beta T(k) & 1 - T(k)\left[\beta + \alpha\phi_{a}\right] \end{pmatrix}, \quad B = \begin{pmatrix} T(k) \\ \\ \alpha T(k) \end{pmatrix} \quad (2.22)$$

When tape replacement occurs at uniform time intervals, T is constant over k and the system is seen to be stationary, and the equations can be solved immediately by successive evaluation:

$$\vec{n}(1) = N_{a}\phi_{a}B, \quad \vec{n}(0) = 0$$

$$\vec{n}(2) = N_{a}\phi_{a}(L+1)B$$

$$\vec{n}(3) = N_{a}\phi_{a}(L^{2}+L+1)B$$

$$\vdots$$

$$\vec{n}(k) = N_{a}\phi_{a}\sum_{j=0}^{k-1}L^{j}B$$

- 2.8 -

and applying the familiar procedure for the geometric sum,

$$Ln(k) - n(k) = N_a \phi_a \left(L^k B - B \right)$$

which gives for the state at the kth tape replacement time,

$$\vec{n}(k) = N_a \phi_a (L - I)^{-1} (L^k - I) B$$
 (2.23)

The increment $\vec{\delta}(k) \equiv \vec{n}(k) - \vec{n}(k-1)$ at the kth tape replacement time is given by:

$$\vec{\delta}(k) = N_{a}\phi_{a} (L - 1)^{-1} (L^{k} - L^{k-1}) B$$

= $N_{a}\phi_{a} L^{k-1} B$ (2.24)

2.3 Model Behavior

Note from the discrete state equations above that the parameter N_a is simply a scale factor on the state \vec{n} . Recall also that the initial slope of $n_d(k)$ is $N_a\phi_a$, and that of $n_c(k)$ is $\alpha N_a\phi_a$, regardless of the value of β . Furthermore, for $\beta = 0$, $n_d(k)$ and $n_c(k)$ maintain the constant ratio $n_c(k)/n_d(k) = \alpha \le 1$, and, of course, coincide as $\alpha \ne 1$.

The effect of $\beta > 0$ is to increase the error correction rate, and therefore increase $n_c(k)$, especially for the larger differences $n_d(k) - n_c(k)$ (backlog) which tend to occur later in the test program. The resulting decrease in remaining errors $N_a - n_c(k)$ causes the detected

- 2.9 -

error curve $n_d(k)$ to be bent downward. Thus the effect of β is to draw the two curves together. Figures 4 and 5 display this effect for $0 \leq \beta$ ≤ 0.5 . The "bending" of the curves due to β , together with the effects of the discrete nature of the model, are expected to occur in real data.

2.4 The Data Simulator

RELY I contains a data simulator for the purposes of study and experimentation. The simulator is an optional source of input data to the estimator (see Appendix C). The simulator reads from input cards the nominal parameter values, α , β , ϕ_a , N_a , the time interval T, the number of test intervals K, and an input initial random number (RRR), and computes the associated software test history $\Delta_d(k)$. The random number RRR is changed by the investigator when he wishes a different sample of the random data set $\Delta_d(k)$ (see Appendix A, RNDTA, RANDEX).



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3.0 The Estimation Algorithm

3.1 The Estimation Problem and Method

Having described the model, we turn to the parameter estimation algorithm which estimates the values of the model parameters corresponding to a given set of real test data. The resulting parameter estimates provide the reliability information sought regarding the tested software package.

Though the model is linear in the sense that the equations are linear in the state \vec{n} , the model equations are nevertheless nonlinear in the parameters \vec{e} (i.e., in α , β , ϕ_a , N_a). Determining the parameter values corresponding to a given set of real test data $\vec{\Delta}(k)$ is then a nonlinear estimation problem.

Nonlinear parameter estimation methods, in general, are iterative procedures in which the estimate is approached from some initial guess for the parameter values, in steps which successively decrease a cost functional J. Since our purpose is to determine the parameter values $\vec{\theta}$ for which the solution $\delta_d(k)$ of the model equations approximates the measured function (or sequence) $\Delta_d(k)$, we choose the cost functional J to be the sum of the squares of the residuals, $\delta_d(k) - \Delta_d(k)$, viz.,

$$J = \sum_{k=1}^{K} \left[\delta_{d}(k) - \Delta_{d}(k) \right]^{2}$$
(3.1)

Minimizing this cost functional, then, minimizes the difference between the observed function $\Delta_d(k)$ and its expected value $\delta_d(k)$ in the least squares sense.

- 3.1 -

Of the numerous methods described in the literature, both direct search methods (Fletcher [9]) and gradient methods (Bard [8]), the gradient methods are generally preferred when the computation of the derivatives of J is not prohibitive. Gradient methods, in principle, step from one point $\vec{\theta}_i$ in parameter space to the next $\vec{\theta}_{i+1}$ according to

$$\vec{\theta}_{i+1} = \vec{\theta}_i - \tau_i R_i \vec{g}_i$$
(3.2)

where \vec{g}_i is the gradient of J evaluated at $\vec{\theta}_i$, R_i is some matrix which operates on the gradient to define the ith step direction $R_i \vec{g}_i$, and τ_i is a scalar which determines the step size. The methods differ in what each employs for R_i , i.e., in the step direction each takes relative to the gradient. The method of steepest descent, for example, uses the identity matrix for R_i , so that the step direction is opposite to that of the gradient. This is "the steepest way down" locally but tends to be less efficient and therefore less desirable than methods which use second order information about the surface $J(\vec{\theta})$.

The Newton-Raphson method uses for R_i the inverse of the Hessian, the matrix of the second partial derivatives,

$$H_{mn} = \frac{\partial^2 J}{\partial \theta_m \partial \theta_n}, \qquad (3.3)$$

of the cost functional.

Notice that the Taylor series expansion of J to second order terms,

- 3.2 -

$$J = J_{i} + \vec{g}_{i}^{T} \left(\vec{\theta} - \vec{\theta}_{i} \right) + \frac{1}{2} \left(\vec{\theta} - \vec{\theta}_{i} \right)^{T} H_{i} \left(\vec{\theta} - \vec{\theta}_{i} \right)$$

has an extremum,

$$\frac{\partial J}{\partial \dot{\theta}_{i}} = g_{i} + H_{i} \left(\stackrel{\rightarrow}{\theta} - \stackrel{\rightarrow}{\theta}_{i} \right) = 0,$$

at

$$\vec{\theta} = \vec{\theta}_i - H_i^{-1}g_i$$
 (H nonsingular)

so if $R_i = H_i^{-1}$, $\rho_i = 1$, and J is quadratic, then $\vec{\theta}_{i+1}$ coincides with the extremum. The Newton-Raphson method in this case converges in a single iteration. This method is quite efficient even for nonquadratic J, but only if H_i is positive definite. This latter condition is the principal weakness of the method. The Marquardt method meets this weakness by guarantee-ing positive definiteness in R_i by adding to H_i (or to some convenient approximation of H_i) a variable amount of a positive definite matrix C_i^2 :

$$R_{i} = \left(H_{i} + \lambda_{i}c_{i}^{2}\right)^{-1}$$
(3.4)

and suggests C_{i}^{2} be a matrix of the diagonal elements of H_{i} , viz.,

$$C_{i,ss}^2 = |H_{i,ss}|.$$
 (3.5)

For sufficiently large λ_i , R_i then is positive definite, even when H_i is not. The Marquardt method behaves as the Newton-Raphson for small

- 3.3 -

 λ_i , but where larger λ_i is necessary it steps nevertheless in some acceptable (downward) direction. A step is said to be acceptable if it decreases J. If λ_i is large and H_i has low condition number (eigenvalues of near-equal magnitude), the method approximates that of steepest descent. The Marquardt method varies from step to step according to λ_i , between the behavior of the Newton-Raphson method and that of steepest descent.

3.2 Description of the Search Algorithm

The program, RELY I, uses the above Marquardt R_i , i.e.,

$$\theta_{i+1} = \theta_i - \tau_i \left(H_i + \lambda_i c_i^2 \right)^{-1} g_i$$
 (3.6)

and selects τ_i or λ_i from step to step according to the procedure described below. Essentially the program progresses in one or the other of two modes. In mode A, λ_i is fixed while the largest τ_i (0.0001 < $\tau_i \leq 1$) is sought which results in an acceptable step size. If the sought τ_i is found, the program continues in mode A preferring smaller and smaller values of λ (more nearly Newton-Raphson). If at any point insufficient progress is being made in mode A, the routine moves to mode B, in which τ_i is initially fixed, and λ_i is successively increased until an acceptable step direction is reached. In mode B, when a sufficiently large λ_i is reached, then the program steps in that direction until J begins to increase, or until, for large J, J has decreased more than 10 percent, at which point the routine returns to mode A. In short, when progress is slow in mode A, the program resorts to mode B to move to a different

- 3.4 -

"locality". "Progress" in mode B is deliberately restricted for large J due to experience which indicates that mode B for large J tends to settle into local minima. The program terminates when J becomes less than a predetermined value (ERR), or upon a time limit for machine computation.

More specifically, the estimator proceeds as follows: Given initial guess $\vec{\theta}_i$ and $\lambda_i = 1$, i = 0:

Mode A

- 1. Compute cost J_i and step direction $R_i g_i$.
- If J < ERR terminate, otherwise determine an acceptable step size in the following way:
 - a. Compute a τ_i such that twice the associated step causes $\beta = 5$; i.e.,

$$\tau_{i} = (5. - \beta_{i}) / [2|R_{i}g_{i}|]$$

b. If such a step causes $\phi_a > 0.2$, choose instead

 $\tau_{i} = (0.2 - \phi_{ai}) / [2|R_{i}g_{i}|]$

- c. If the resulting $\tau_i > 1$, set $\tau_i = 1$.
- d. If the resulting $\tau_i < .0001$, jump to mode B.
- e. Limit $0 \le \beta \le 10$. Compute J_{i+1} .
- f. If $J_{i+1} \ge J_i$, jump to item 4 below.
- 3. Accept θ_{i+1} and reduce λ ; i.e.,
 - a. Set $\theta_i = \theta_{i+1}$, $\lambda_i = \lambda_i/10$.

- 3.5 -

- b. If J has decreased less than 1 percent in more than five iterations (reductions of λ in item 3.a) since passing through mode B, jump to mode B. Otherwise continue in mode A (jump to item 1 above).
- 4. Reduce τ_i by a factor of 10. If the resulting $J_{i+1} < J_i$ jump to item 3 above, otherwise repeat item 4 above up to five times (according to counter INDEX). If J does not decrease with five reductions of τ_i , jump to mode B. If $J_{i+1} < ERR$, terminate.

Mode B

- 5. Fix $\tau = 0.1$, set $\lambda = 0.01$.
- 6. Increase λ by a factor of 10, increment the count ICLAM, determine the corresponding step direction $(H_i + \lambda_i C_i^2)^{-1} g_i$, parameter set θ_{i+1} , and J_{i+1} . If $J_{i+1} \ge J_i$, repeat item 6.
- 7. Accept θ_{i+1} (i.e., set $\theta_i = \theta_{i+1}$) and set $J_{\lambda} = J_{i+1}$.
- 8. Increase τ_i by a factor of 5^{ℓ} , $\ell = \text{ICLAM}$. 9. Try $\vec{\theta}_{i+1} = \vec{\theta}_i \tau_i \left(H_i + \lambda_i c_i^2\right)^{-1} \vec{g}_i$, if $J_{i+1} > J_i$ or if $J_i > J_i$ 50 and $J_{i+1}/J_{\lambda} < 0.9$, accept $\vec{\theta}_i$ and return to item 1 above, otherwise accept $\theta_{i+1} \rightarrow \theta_i$ and repeat item 8.

The algorithm is depicted in the flow diagram of Fig. 6.





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Experience during development of RELY I proved α to be insufficiently independent of the other parameters to warrant a fourth degree of freedom in the search process. The computer program therefore was modified to accept a priori the estimate of α , and to search in three dimensions for the values of β , ϕ_a , and N_a . From these the estimated number of remaining errors,

$$n_a = n_a(k) = N_a - n_c(K), \quad k = 1, 2, 3, \dots K$$

and mean time to failure

$$MTTF = \frac{1}{\phi_a n_a}$$

are computed. The latter two computed quantities, the sought software reliability factors, were found to be essentially insensitive to reasonable a priori estimates of α . Results below include cases of correct and incorrect fixed α .

3.3 Estimation Results

Results are tabulated and displayed in histograms below for several simulated random data examples. Examples I and II differ in the selection of K and T to vary the number of errors, N_R , remaining in the software. Example I uses test interval length T = 1.5 and 60 intervals which leaves about 30 remaining of the initial 200 software errors. Example II uses longer test intervals, T = 8, and fewer intervals, K = 20, to leave about 5 errors remaining. Example III corresponds to a

- 3.8 -
larger software system having a considerably larger number of initial errors (N_a = 1000), smaller error detection rate (ϕ_a = 0.01), but a better correction rate (α = 0.8), and the longer test intervals (T = 8.0). Finally, the fourth example demonstrates the insensitivity to the fixed value of α . Example IV essentially is Example I with α fixed at 0.5 instead of the "true" value, 0.7.

Before examining the results, we anticipate the nature of the distributions by analytically determining the mean and variance for the simple single interval (K = 1) case. Let the observed number of detected errors N_d be Poisson with mean and variance ρN_a , where $\rho = \phi_a T$. We minimize the squared error J,

$$J = \left(\rho N_{a} - N_{d}\right)^{2}$$

$$\frac{\partial J}{\partial N_{a}} = 2\rho \left(\rho N_{a} - N_{d}\right) = 0$$

to obtain an estimate

$$\hat{N}_a = \frac{N_d}{\rho}$$

which has mean

$$E\left\{\hat{N}_{a}\right\} = \frac{E\left\{N_{d}\right\}}{\rho} = N_{a}$$

and variance

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$$E\left\{\left(\hat{N}_{a}-N_{a}\right)^{2}\right\} = E\left\{\hat{N}_{a}^{2}\right\} - 2 E\left\{\hat{N}_{a}N_{a}\right\} + E\left\{N_{a}^{2}\right\}$$
$$= \frac{E\left\{\frac{N_{d}^{2}}{\rho^{2}} - 2N_{a}E\left\{\hat{N}_{a}\right\} + N_{a}^{2}\right\}$$
$$= \frac{\left(\operatorname{var}\left(N_{d}\right) + \overline{N}_{d}^{2}\right)}{\rho^{2}} - N_{a}^{2}$$
$$= \frac{\rho N_{a} + \rho^{2} N_{a}^{2} - \rho^{2} N_{a}^{2}}{\rho^{2}}$$
$$= \frac{\frac{N_{a}}{\rho^{2}}$$

The number of remaining errors,

$$N_R = N_A - N_d$$

is estimated

$$\hat{N}_{R} = \hat{N}_{a} - N_{d} = N_{d} \left(\frac{1-\rho}{\rho}\right)$$

with unbiased mean

$$E\left\{\hat{N}_{R}\right\} = E\left\{\hat{N}_{a} - N_{d}\right\} = N_{a} - N_{d}$$

and with root mean squared difference from its true value,

- 3.10 -

$$E\left\{\left[N_{d}\left(\frac{1-\rho}{\rho}\right) - \left(N_{a} - N_{d}\right)\right]^{2}\right\}^{1/2} = E\left\{\left(N_{a} - \frac{N_{d}}{\rho}\right)^{2}\right\}^{1/2}$$
$$= \left[N_{a}^{2} - \frac{2N_{a}}{\rho} E\left\{N_{d}^{2}\right\} + \frac{E\left\{N_{d}^{2}\right\}}{\rho^{2}}\right]^{1/2} = \left[N_{a}^{2} - 2N_{a}^{2} + \frac{\rho^{2}N_{a}^{2} + \rho N_{a}}{\rho^{2}}\right]^{1/2} = \sqrt{\frac{N_{a}}{\rho}}$$

Notice that the value of the latter quantity corresponding to:

a. Example I: Let $N_R = N_a - \rho N_a$, or ($\rho = 1 - 31/200 = .845$) is

$$=\sqrt{\frac{200}{.845}} = 15$$

b. Example II: $(\rho = 1 - 5/200 = 0.975)$ is

$$=\sqrt{\frac{200}{.975}} = 14$$

c. Example III: ($\rho = 1 - 185/1000 = 0.815$) is

$$\sqrt{\frac{N_a}{\rho}} = \sqrt{\frac{1000}{.815}} = 35$$

One would expect these values to approximate the standard deviations σ (N_R) for the respective multiple-interval cases (though perhaps with less validity when N_R/N_a is small). The σ (N_R) indicated below for the four examples then are of the magnitude to be expected. Table 1 lists numerical information from the four examples. Examination of results from the four simulated examples indicates that the estimator produces reasonable estimates of the reliability parameters N_R and MTTF.

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Item	Example: I	II	III	IV
"True" Values				
α	.700	.700	.800	.700
β	.100	.100	.100	.100
¢a	.020	.020	.010	.020
Na	200	200	1000	200
NR	31.0	4.93	185	30.6
MTTF	1.61	10.1	0.538	1.64
Time Interval				
Т	1.5	8.0	8.0	1.5
Number of Intervals				
К	60.	20.0	20.0	60.0
A Priori α				
α'	.700	.700	.800	.500
Initial Guess				
β	.300	.300	0	0
ф а	.010	.010	.02	.010
Na	300	300	1500	300
Estimated Values				
N _R	26.3	5.03	195	29.7
а (N _R)	11.2	3.07	40.4	10.8
MTTF	2.19	12.6	.543	1.75
o (MTTF)	.65	7.43	.093	.33

Table 1. Estimation results.

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Fig. 7. Histograms of reliability parameters for Example I.

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Fig. 8. Histograms of reliability parameters for Example II.

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Fig. 9. Histograms of reliability parameters for Example III.

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Fig. 10. Histograms of reliability parameters for Example IV.

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4.0 Conclusions and Recommendations

4.1 RELY I

We have developed and displayed a model which we believe more accurately describes an actual testing environment of a large software package. This new generalized model has been incorporated in an estimation algorithm for the purpose of discerning reliability of the software from its test data. The first version of the algorithm RELY I described here converges in a given region of interest of the model parameters. RELY I is applicable to software test cases where tape replacement (software revision) occurs at uniform intervals of time, and where sufficiently reliable information is available concerning the number of errors detected during each of the successive intervals.

4.2 Data Requirements

Data required for RELY I are simple, viz., the time interval T between software revisions (tape replacements), and the sequence $\Delta_d(k)$ during each of the K successive versions of the software, where k = 1, 2, ..., K. Secondly, the data must be from a process of the type upon which the assumptions of the model were based, viz., the testing of large-scale software packages such as that in the F-16 control system.

There must be an identifiable single continuous line of software package identity throughout the test process. The package passes successively through a sequence of versions k, k = 1, 2, ..., K. At any given time during test, the software is in only one of the versions,

- 4.1 -

"all" of which software version is being tested.¹ Each version, k, is identical to the preceding version, k - 1, and succeeding version, k + 1, except for the software corrections "counted" in $\Delta_c(k)$ and $\Delta_c(k + 1)$, respectively. Figure 11 indicates the time relationship of the several sequential quantities. The requirement is that $\Delta_d(k)$ be precisely known for each version k, where all versions are identified and satisfy this single and continuous identity as described. This requirement is violated if a major untested version is suddenly introduced midstream, or if an alternate part of the software package simultaneously being tested suddenly is adopted. The generated error feature can accommodate a minor amount of this kind of violation, but generated errors are modeled as occurring as a constant proportion of the correction rate.

Errors are usually classified into certain arbitrary categories, ranging from those obviously to be counted, to those of doubtful pertinence (obviously "repeated" errors, errors associated purely with erroneous test conduct, etc.). Suffice it here to suggest that the criterion for counting a given error or not will be related to its likelihood of occurrence, and its interpretation as a "failure", under operational conditions.

4.3 Recommendations

Recommendations toward improved interfacing with information in a real test process (thus taking greater advantage of inherent features

¹ That is, all the parts of the software package are being exercised in a manner representing that for which the reliability factors, e.g., MTTF, are to be applied later.

Errors, cumulative:	ן א ס ו	(1) 	ที่(2) ∣	Ñ(3) │
Errors, incremental:	0 ₹0	(1)	↓ ≩(2)	↓ Ճ(3)
Version:	(initial)	(1)	(2)	(3)
Error rate:	r (0)	r(1)	r (2)	r (3)
Test time:	0 1	<u>ــــــــــــــــــــــــــــــــــــ</u>	2т	3T
k:	0 1	L	2	3

Fig. 11. Tested software in only one version at a time, identical to neighboring versions except for the changes counted in Δ_{c} at the respective boundaries.

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of the underlying new model) include the following further work:

- 1. Revise the KOST subroutine to accommodate nonuniform test intervals T(k), by using a finite difference technique for the solution of $\vec{\xi}(k, \vec{\theta})$. The increased utility is expected to far outweigh the lesser analytic tractability of the resulting system and the possible increase in required computation time.
- 2. Apply the algorithm to real data. Available data should be gathered, studied, and adapted, by interpretation and transformations, to the requirements of RELY. Residual functions over a variety of cases will indicate how well the model represents the real test process. Experience will lead to further recommendations concerning data requirements, and to possible improvements in RELY such as provisions for using information in real data concerning error correction and error generation.

For example, the quantity $n_a(k)$,

$$n_{a}(k) = N_{a} - n_{c}(k) = \frac{N_{o} - n_{c}(k) + n_{g}(k)}{1 - \gamma} = \frac{n(k)}{1 - \gamma}$$

is the augmented number of errors remaining in the tested software. That is, $n_a(k)$ is the number of errors which would be detected henceforth if the testing process were to continue indefinitely, including those generated after time k. The number of errors remaining in the software, exluding

- 4.4 -

those yet to be generated in the correction process, is

$$n(k) = (1 - \gamma) n_{\alpha}(k)$$

The parameter γ is assumed not observable in the present implementation of the model. If, however, among the detected errors, generated errors are distinguishable from original errors, then the additional quantity $n_{gd}(k)$, the number of generated errors detected, is available. The model state is easily augmented to include $n_{gd}(k)$. The model remains unchanged but, to the extent that the additional information is available in test data, the model parameter γ becomes observable.

Experience in the development of RELY I suggests further investigation of the nature of the $J(\vec{\theta})$ surface. Such investigation should include also the surface associated with the alternative cost functional using cumulative functions N(k), rather than the incremental number of errors $\Delta(k)$. Convergence properties in certain regions of parameter space may be significantly improved using N(k) rather than their derivatives $\Delta(k)$. Indeed, parallel computation using each, respectively, may prove both feasible and advantageous. Another gradient type parameter estimation method, such as the Fletcher-Powell deflected gradient method, may also prove more efficient with the alternative functional.

- 4.5 -

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APPENDIX A

MAIN AND SUBROUTINE DESCRIPTIONS

1. MAIN and Subroutine Diagram

The subroutines of RELY I are indicated in Fig. A.1. Internal



Fig. A.1. RELY I subroutine diagram.

subroutines EIGMIN and MARQ are shown, as well as the UNIVAC MATH-PACK library subroutines RANDEX, TRIDMX, EIGVAL, and EIGVEC. A brief description of MAIN and its subroutines follows.

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2. MAIN

MAIN reads input data, executes the estimation algorithm (see Sec. 3.2), and prints output. It also computes simulated random data $\Delta_{d}(k)$ under the ISIM = 1 option (see Appendix C).

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3. Subroutine RNDTA (A, B, P, NA, T, V, RR, JJJ)

From α , β , ϕ_a , N_a , T, the input initial random numbers and L (corresponding to RNDTA variables A, B, P, NA, T, RR, JJJ, respectively, and corresponding to the MAIN variables ALPH, BETA, PHI, NA, TD, RRR, NN, respectively), RNDTA computes the sequence $\Delta_d(k)$, k = 1, 2, ... L (the RNDTA variable V, and MAIN variable S). The resulting sequence $\Delta_d(k)$ is used as simulated data, the (incremental) number of errors detected successively in each software test interval. The initial random number RRR is passed through POISS to RANDEX.

RNDTA, at each interval k, integrates the system equations:

$$n_{d}(k) = N_{d}(k-1) + \phi_{a}T(N_{a} - n_{c}(k-1))$$

$$n_{c}(k) = n_{c}(k-1) + \alpha \phi_{a}T(N_{a} - n_{c}(k-1)) + \beta T[N_{d}(k-1) - n_{c}(k-1)]$$

 $N_{d}(0) = n_{c}(0) = 0$

using the cumulative (random) number of errors detected $N_d(k-1)$, to obtain $n_c(k)$ for use in POISS. Subroutine POISS generates the random integer $\Delta_d(k)$ according to the mean detection rate $r_d(k) = \phi_a [N_a - n_c(k)]$.

4. Subroutine POISS (DD, ZZ, PP, TT, RRRR, KKK)

From $n_a (=N_a - n_c)$, ϕ_a , T, the input initial random number RRR, and k (POISS variables DD, PP, TT, RRRR, and KKK, respectively, corresponding to RNDTA variables D, RP, RT, RQ, and K), POISS computes the value Δ_d (k), according to

$$\sum_{i=1}^{m} C(i) < T, \quad \Delta_{d}(k) = m$$

The random sequence C(i), i = 1, 2, ... 100, with exponential distribur_d^C tion function $1 - e^{d}$, $r_d = \phi(N_a - n_c)$, is generated by RANDEX. The starting random number required by RANDEX in C(1) is the input initial random number RRR for k = 1, and is the preceding random number $R(2^{25})$ for k > 1, where R is the value C(100) previously computed for the $(k - 1)^{th}$ pass. 5. Subroutine RANDEX (C, 100, U)

Reference: UNIVAC Large Scale System MATH-PACK, Programmer's Reference, UP-7542, Rev. 1.

RANDEX produces a set of 100 pseudo-random numbers C with exponential distribution

$$1 - c^{-UC}$$

by operating on a uniformly distributed variate X, according to the inverse transform method

$$C = \frac{-\ln(1 - X)}{U} .$$

RANDEX uses two other UNIVAC MATH-PACK subrottines RANDU and RANDN. RANDU generates X, $0 \le X \le 1$, for which computation it calls RANDN for random integers $0 \le I \le 2^{35}$. RANDEX requires an initial value, $0 \le C(1) \le 2^{35}$, different integer parts of which produce different output sequences.

6. Subroutine KOST (AA, BB, PP, NNA, SS, NJJ, TT, NMN, DT, H, GJ, DND, COST, DNC, RMTTF, ZNC, RERR)

Given values of the parameters $\vec{\theta}$, time interval T, number of intervals K, test data $\Delta_d(k)$, KOST computes the incremental error sequences (see Sec. 2.2)

$$\delta_{d}(k) = N_{a}\phi_{a}(1 \quad 0)L^{k-1}B$$

$$\delta_{c}(k) = N_{a}\phi_{a}(0 \quad 1)L^{k-1} B$$

where (1 0) and (0 1) are the transposes of the vectors $\begin{pmatrix} 1 \\ 0 \end{pmatrix}$ and $\begin{pmatrix} 0 \\ 1 \end{pmatrix}$, respectively, and

$$L = \begin{pmatrix} 1 & -\phi_{a}T \\ \beta T & 1 - T(\beta + \alpha\phi_{a}) \end{pmatrix}, \qquad B = \begin{pmatrix} T \\ \alpha T \end{pmatrix}$$

KOST further computes the cost scalar (see Sec. 3.1)

--

$$J = \sum_{k=1}^{K} \left[\delta_{d}(k) - \Delta_{d}(k) \right]^{2}$$

the gradient vector components

$$\frac{\partial J}{\partial \theta_{m}} = 2 \sum_{k=1}^{K} \delta_{d}(k) \frac{\partial \delta_{d}(k)}{\partial \theta_{m}}$$

and the Hessian matrix elements

$$\frac{\partial^2 J}{\partial \theta_n \partial \theta_m} = 2 \sum_{k=1}^{K} \left[\delta_d(k) \frac{\partial^2 \delta_d(k)}{\partial \theta_n \partial \theta_m} + \frac{\partial \delta_d(k)}{\partial \theta_n} \frac{\partial \delta_d(k)}{\partial \theta_n} \right]$$

KOST also computes the associated estimate of total errors corrected

$$n_{c} = \sum_{k=1}^{K} \delta_{c}(k)$$

the number of errors remaining

$$n_R = N_a - n_c$$

and the mean time to failure

$$MTTF = \frac{1}{\phi_a n_R}$$

The first derivatives above are given by:

$$\frac{\partial \delta_{\mathbf{d}}(\mathbf{k})}{\partial \theta_{\mathbf{m}}} = N_{\mathbf{a}} \phi_{\mathbf{a}} (1 \quad 0) L^{\mathbf{k}-2} \left[(\mathbf{k}-1) \frac{\partial L}{\partial \theta_{\mathbf{m}}} \mathbf{B} + L \frac{\partial B}{\partial \theta_{\mathbf{m}}} \right]$$
$$+ \delta_{\mathbf{d}}(\mathbf{k}) \left[\frac{1}{\phi_{\mathbf{a}}} \frac{\partial \phi_{\mathbf{a}}}{\partial \theta_{\mathbf{m}}} + \frac{1}{N_{\mathbf{a}}} \frac{\partial N_{\mathbf{a}}}{\partial \theta_{\mathbf{m}}} \right]$$

where

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$$\frac{\partial \theta_{\mathbf{i}}}{\partial \theta_{\mathbf{m}}} = \begin{cases} 1 & \mathbf{i} = \mathbf{m} \\ 0 & \mathbf{i} \neq \mathbf{m} \end{cases}$$

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<u>9</u> Г =	0	0	$\frac{\partial B}{\partial B} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$
a a	0	0)	$\partial N_{a} \left(0 \right)$

The second derivatives are:

$$\frac{\partial^{2} \delta_{d}(\mathbf{k})}{\partial \theta_{n} \partial \theta_{m}} = N_{a} \phi_{a} (1 \quad 0) \left\{ (\mathbf{k}-2) \mathbf{L}^{\mathbf{k}-3} \frac{\partial \mathbf{L}}{\partial \theta_{n}} \left[(\mathbf{k}-1) \frac{\partial \mathbf{L}}{\partial \theta_{m}} \mathbf{B} + \mathbf{L} \frac{\partial \mathbf{B}}{\partial \theta_{m}} \right] \right\}$$
$$+ \mathbf{L}^{\mathbf{k}-2} \left[(\mathbf{k}-1) \left(\frac{\partial^{2} \mathbf{L}}{\partial \theta_{n} \partial \theta_{m}} \mathbf{B} + \frac{\partial \mathbf{L}}{\partial \theta_{m}} \frac{\partial \mathbf{B}}{\partial \theta_{n}} \right) + \frac{\partial \mathbf{L}}{\partial \theta_{n}} \frac{\partial \mathbf{B}}{\partial \theta_{m}} \right]$$
$$+ \mathbf{L} \frac{\partial^{2} \mathbf{B}}{\partial \theta_{n} \partial \theta_{m}} \left\} + \frac{\partial}{\partial \theta_{n}} \left\{ \delta_{d}(\mathbf{k}) \left[\frac{1}{\phi_{a}} \frac{\partial \phi_{a}}{\partial \theta_{m}} + \frac{1}{N_{a}} \frac{\partial \mathbf{N}_{a}}{\partial \theta_{m}} \right] \right\}$$

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7. Subroutine EIGMIN (HH, CORR, GGJ, DDAM, EH, EV)

Given the Hessian HH(4, 4) and the gradient GGJ(4) of the cost functional $J(\vec{\theta})$, and the current Marquardt parameter (λ) DDAM, EIGMIN computes the eigenvalues EV(4) and eigenvectors EGV(4, 4), and the outer product EH(4, 4, 4), for J. EIGMIN then computes λ_i such that the Marquardt matrix $(H_i + \lambda_i C_i^2)$ is positive definite, and the corresponding parameter correction factors CORR(4) = $(H_i + \lambda_i C_i^2)^{-1} \vec{g}_i$.

8. Subroutine TRIDMX (N, NM, A, D, B)

Reference: UNIVAC Large-Scale System MATH-PACK, Programmer's Reference, UP-7542, Rev. 1, Sec. 9, p. 1.

TRIDMX transforms a real symmetric matrix, B(4, 4), to tridiagonal form using Householder's method, where D(4) are the resulting diagonal elements and B(4) are the off-diagonal elements. Input integers N and NM are equal to the order, 4, of B.

- A.10 -

9. Subroutine EIGVAL (LP, E, A, B, W, F)

Reference: UNIVAC Large-Scale Systems MATH-PACK, Programmer's Reference, UP-7542, Rev. 1, Sec. 9, p. 8.

EIGVAL evaluates the eigenvalues of a symmetric tridiagonal matrix, using Sturm sequences. A(4) are the diagonal elements and B(4) are the off-diagonal elements of the matrix. The eigenvalues E(4) are stored in descending order of absolute value.

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10. Subroutine EIGVEC (LP, NM, R, A, B, E, V, P, Q)

Reference: UNIVAC Large-Scale Systems MATH-PACK, Programmer's Reference, UP-7542, Rev. 1, Sec. 9, p. 15.

EIGVEC evaluates the eigenvectors of a real symmetric tridiagonal matrix using Wilkinson's method. A(4) are the diagonal elements and B(4) are the off-diagonal elements of the matrix. E(4) are the eigenvalues, and V(4, 4) are the eigenvectors.

11. Subroutine MARQ (EEH, EEV, DDLAM, CCORR, GGGJ, HHH)

Given the outer products EEH(4, 4, 4) of the eigenvectors of the Hessian, the eigenvalues EEV(4), the gradient GGGJ(4), and Marquardt parameter (λ_i) DDLAM, MARQ computes the step CCORR(4), $\left(H_i + \lambda_i C_i^2\right)^{-1} g_i$, in parameter space.

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APPENDIX B

RELY I GLOSSARY AND INDEX

(Library subroutines RANDEX, RANDU, RANDN, TRIDMX, EIGVAL, EIGVEC are not included here. See UNIVAC MATH-PACK references given in Appendix A for detailed information.)

MAIN (including internal subroutines EIGMIN and MARQ)

Variable	Description	Line Number
ALPH	Value for α in simulat	18, 22, 25, 30 tion mode
B(4, 4)	Normalized Hessian mat	162, 167, 170, 172 rix in EIGMIN
BETA	Value for β in simulat	18, 22, 25, 30 tion mode
CNC	Cumulative values for	24, 33, 34 n _c in simulation mode
CND	Cumulative values for	23, 32, 34 n _d in simulation mode
CORR (4) (MARQ: CCORF	Vector of corrections	3, 48, 57, 58, 61-63, 72, 78, 86-88, 95, 114, 115-117, 127, 137, 160, 163, 201 to parameters 205, 207, 209, 229
COST	Most recently computed	25, 30, 43, 46, 49, 50, 56, 65, 67, 69, 70, 73, 76, 91, 93, 94, 100, 120, 122, 124-127, 129, 130, 135, 145, 147 value for the mean-squared error
COSTI	Cost for currently acc to determine whether t reduces the cost	46, 130, 145 epted parameter values used in mode B the new estimate for the parameters
DDD	Denominator used to no EIGMIN	194, 195 rmalize the matrix (H + λ I) ⁻¹ in

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Variable Description Line Number DDDD 223, 224 Denominator used to normalize the matrix $(H + \lambda I)^{-1}$ in MARQ DIA(4) 162, 170-172 Diagonal entries of the tridiagonalized Hessian matrix used in EIGMIN to compute eigenvalues DLAM 45, 48, 72, 74, 78, 95, 103, 111, 114, 128, 137 Value for λ (in EIGMIN: DDAM 160, 188, 205, 217) (in MARQ: DDLAM 182) DND(300) 4, 25, 32, 43, 67, 76, 91, 120, 135 Incremental values in n_d EGV(4, 4)162, 172, 176 Eigenvectors for the Hessian matrix EH(4, 4, 4) 3, 48, 72, 78, 95, 114, 128, 137, 160, 162, 176, 188 Outer products of the eigenvectors for the Hessian matrix used to compute $(H + \lambda I)^{-1}$ (in MARQ: EEH 205, 207, 217) ENC(300) 4, 25, 33, 43, 67, 76, 91, 120, 135 Incremental values for n ERR 5, 7, 56, 93, 126 Value for termination criterion EV(4) 3, 48, 72, 78, 95, 114, 128, 137, 160, 162, 172, 181, 182, 185, 188 Eigenvalues for the Hessian matrix (in MARQ: EEV 205, 207, 217) GJ(4) 3, 25, 43, 48, 67, 72, 76, 91, 95, 114, 120, 127, 135, 137 Gradient vector for the cost functional (in EIGMIN: GGJ 160, 163, 201) (in MARQ: GGGJ 205, 207, 229) H(4, 4) 3, 25, 43, 48, 67, 72, 76, 78, 91, 95, 114, 120, 127, 135, 137 Hessian matrix for the cost functional (in EIGMIN: 160, 162, 167, 194) HH 205, 207, 223) (in MARQ: HHH

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Variable	Description	Line Number
I	Index for various loo	ps
ICLAM	Index which counts th in mode B	104, 113, 144 e number of times that λ is increased
IFLAG	Index that counts the mode A reduces the co	47, 70, 71, 106 number of times that an iteration of st by less than 1 percent
INDEX	Index that counts the has been reduced	51, 80, 81 number of times that the step size
ISIM	Indicates whether the estimation with real of	15, 17 run is a simulation (ISIM = 1) or an data
J	Index for various loo	ps
JDEX	JDEX = 1 indicates th successful in a given	105, 143, 144 e first time that changing λ has been iteration of mode B
KK	Index used for DO loo products of eigenvect	p in EIGMIN for computing outer ors
KL	Index used for DO loo (H + λ I) ⁻¹	184, 185, 188, 214, 217 p in EIGMIN and MARQ for computing
LMBEX	LMBEX = 1 indicates t	112, 122, 123 hat λ was changed in mode B
NA	Value for N in simul	2, 18, 22, 25, 30 ation mode
NJ	Number of test interv	10, 25, 43, 67, 76, 91, 120, 135 als
NN	Number of tape version	5, 7, 10, 11, 22, 25, 31, 43, 67, 76, 91, 120, 135 ns

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Variable	Description Line	Number
OFDI(4)	162, Off-diagonal entries in tr	, 170-172 idiagonalized Hessian matrix
PCOST	50, Current minimum value for	69, 70, 73, 94, 124, 127, 129 the cost functional
PHI	18, Value of ϕ_a in simulation	22, 25, 30 mode
R(4, 4)	Matrix $(H + \lambda I)^{-1}$ computed	166, 188, 195, 201 I in EIGMIN
RR(4, 4)	Matrix $(H + \lambda I)^{-1}$ computed	211, 217, 224, 229 I in MARQ
REFF	26, 121, Estimated number of errors period	30, 44, 49, 68, 77, 92, 100, 122, 125, 136, 147 remaining at the end of test
RMTTF	25, 120, Estimated mean time to fai	30, 43, 49, 67, 76, 91, 100, 122, 125, 135, 147 Llure
RRR	18, Randomization value in sim	19, 22 mulation mode
S(300)	3, 1 120, Error data	2, 22, 25, 38, 43, 67, 76, 91, 135
T(300)	3, 1 Tape version replacement (3, 25, 43, 67, 76, 91, 120, 135 imes
TAU	57-6 Step size	53, 79, 86-88, 102, 115-117, 144
TD	5, 7 120, Length of each test interv	7, 13, 22, 25, 43, 67, 76, 91, , 135 7al
TEMP1(4) TEMP2(4)	163, 163, Vectors which are used ten the eigenvalues and eigenv	, 171, 172 , 171, 172 aporarily in the computation of rectors of Hessian matrix

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Variable	Description	Line Number
ZA		40, 43, 49, 52, 67, 82, 91, 96, 100, 107, 120, 122, 125, 131, 135, 139, 147
	Value for α in simula	tion and estimation mode
ZB		40, 43, 49, 53, 57, 61, 64, 66-67, 83, 86, 89-91, 97, 100, 108, 115, 118-120, 122, 125, 132, 135, 140, 147
	Value for β in simula	tion and estimation mode
ZN		40, 43, 49, 55, 63, 67, 85, 88, 91, 99, 100, 110, 117, 120, 122, 125, 134, 135, 142, 147
	Value for N in simul	ation and estimation mode
ZNC	Estimated value for n	26, 30, 44, 49, 68, 77, 92, 100, 121, 122, 125, 136, 147 at the end of the test period
ZP	Value for A in circul	40, 43, 49, 54, 58, 62, 67, 84, 87, 91, 98, 100, 109, 116, 120, 122, 125, 133, 135, 141, 147
		ation and estimation mode
ZZA	Currently accepted va	52, 76, 82, 96, 107, 137, 139 lue for α
ZZB	Currently accepted va	53, 76, 83, 97, 108, 115, 132, 140 Lue for β
ZZN	Currently accepted va	55, 76, 85, 99, 110, 117, 134, 142 Lue for N _a
ZZP	Currently accepted va	54, 76, 84, 98, 109, 116, 133, 141 lue for ₉ a
	Subrouti	ne RNDTA
A (db1)	Correction rate param	1, 7, 19 eter α
B (dbl)	Correction rate param	1, 8, 19 eter β
D	Estimated number of r	4, 21, 18 emaining errors n _a

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Variable	Description	Line Number
E (dbl)	D	17, 18
JJJ	Number L of intervals	l, 15 (software versions)
К	Index corresponding to	15, 21, 27 p k th interval
NA (dbl)	Initial number N $(\gamma - a)$	1, 3, 10, 16, 17, 19 augmented) of software errors
P (dbl)	Detection rate paramet	1, 9, 16, 19 ^{ter φ} a
RA	A	4, 7
RB	В	4, 8
RNA	NA	4, 10
RP	Р	4, 9, 21
RQ	RR	5, 6, 21
RR	Storage place for MAIN RANDN. Contains first RANDEX upon return.	l, 6 N input initial random number for t exponential random number from
RT	Т	4, 11, 21
RZ	Poisson random A _d (k) 1	4, 21, 22, 27 returned by POISS
T (db1)	Time interval T	1, 11, 16, 19
V(300) (db1)	RZ, Poisson random sec	1, 12, 27 Juence $\Delta_d(k)$ returned by RNDTA

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Variable	Description Line Number
X(2) (db1)	12-14, 16, 17, 19, 20, 22, 25 Temporary memory for current cumulative random \vec{N}
Y(2) (dbl)	12, 16, 19, 20 $\vec{\mathbf{M}}$ Temporary memory for current cumulative estimate $\vec{\mathbf{M}}$
	Subroutine POISS
C(100)	2, 4, 6, 8, 11 Set of uniformly distributed random numbers generaged by RANDEX to be used by POISS as the sequence of times between successive detected errors
DD	1, 7
	Number of remaining errors n a
K	10, 11, 13 Counter of successive detected errors
KKK	1, 3 Number L of test intervals
PP	1, 7 Parameter value ϕ_a
Q	9, 11, 12 Cumulative time $\Sigma_i C_i$ during the test interval, accumulated until it exceeds T
RRRR	1, 4 Initial random number for starting RANDN. Its value for k = 1 is input by MAIN. Subsequent values are set by POISS RRRR = 2^{25} C(100).
TT	1, 12 Test interval T
U	7, 8 Mean frequency $\phi_a n_a$ of the error detection r_d
ZZ	1, 13, 16 Poisson random number of detections $\Delta_d(k)$ generated by POISS for the k th interval. It is Δ_d such that:
	$\sum_{i=1}^{\Delta_{d}} C_{i} \leq T, \Delta_{d} \leq 100$

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APPENDIX C

PROCEDURE FOR OPERATION OF RELY I

The program listing (FORTRAN V, Appendix D) accompanying this report consists of MAIN with internal subroutines EIGMIN and MARQ, and external double-precision subroutines RNDTA, POISS, KOST, TRIDMX, EIGVAL, and EIGVEC. The subroutines which call single-precision library functions have the necessary coding for converting between double- and single-precision variables. Certain double-precision library functions are used, viz., DABS, DEXP, and DSQRT.

The input deck depends on whether data are to be simulated or are to be read from input cards.

INPUT DECK (to simulate data and estimate)

Card 1: TD, NN, ERR, [F5.2, 2X, 14, 2X, F8.6]

Card 2: ISIM [I2] (must be unity)

Card 3: ALPH, BETA, PHI, NA [4(G14.6, 2X)]

Card 4: RRR [16X, G14.1] (input initial random number)

Card 5: ZA, ZB, ZP, ZN [4(G14.6, 2X)]

INPUT DECK (to estimate using punched card data)
Card 1: TD, NN, ERR, [F5.2, 2X, I4, 2X, F8.6]
Card 2: ISIM [I2] (must be zero)
Card 3: ZA, ZB, ZP, ZN [4(G14.6, 2X)]
Card 4-(K+3): S(1), 1 = 1, 2, ... K [G10.1]

The integer part of any real number RRR, $0 \le RRR < 2^{35}$.

Sec. 1

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determines a unique repeatable random sequence $\Delta_{d}(k)$ from the simulator.

Initial guesses $\vec{\theta}_0$ for the model parameters are recommended as follows:

0 < ZA < 1. (typically 0.8)
0 < ZB < 1. (typically 0.0)
0 < ZP < 0.2
0 < ZN</pre>

To produce different simulated random data, the operator must change the input initial random number RRR, $0 < RRR < 2^{35}$.

Though J < ERR is the internal stopping criterion, experience in random cases proved maximum pages of printed output (say, 10) to be as practical a stopping criterion as any.

The program prints out the current accepted parameter estimates $\vec{\theta}_{1} = \alpha$, β , ϕ , N_{a} , together with running estimates of the reliability parameters N_{R} and MTTF, and selected auxiliary quantities, at each step i. However, the label "CHANGING LAMBDA" indicates only tentative parameter values produced in mode B (see Sec. 3.2). Therefore, these tentative values must not be taken as values which minimize the cost functional J. The final estimates of the reliability parameters are those associated with the last accepted iteration i.

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APPENDIX D

RELY I PROGRAM LISTING AND SAMPLE OUTPUT

RELY#RELY(1).	MAIN
1	IMPLICII REAL+8(A-H+0-Z)
2	KEAL+8 HA
3	LIMENSIUN T(300),S(300),H(4,4),GJ(4),CORR(4),EH(4,4,4),EV(4)
4	UIMENSIUM DND (300) +ENC (300)
5	REAU(5+007)TL+NH+ERR
o	667 FURMAT(F5.2/2X/14/2X/F8.6)
7	WRITE(6+886)TU+NN+EKR
5	- 386 FORMAT(3X) INTERVAL LENGTH = 1,2X,F5.2,2X, NUMPER OF INTERVALS = 1,
9	12X+14+2X+TERMINATION CRITERION =++F8.6)
10	NJEMIA
11	UG 14 J=1,NN
ΤŞ	5(J)=0
13	נ(J) = (J−_+) + Tu
14	14 CONTINUE
10	READ(5:150)ISIM
10	150 FURMAT(12)
17	1F(ISIM+NE-1)60 10 10
10	READ (5+668) ALPH+DETA+FHI+NA
19	nĽAU(5+069)KRR
c U	#KITE (6+689)KKR
21	069 FURMAT(16X;614.4)
22	CALL RNUTA (ALPH+EETA+PHI+1,A+TD+5+PPR+NN)
د>	C.,D=0
2 4	C.+C=0
≥5	CALL KOST (ALPH+BETA+PHI+14+S+NJ+T+NN+TD+H+GJ+DNU+COST+ENC+PMTTF+
20	12inC+nERn)
∠7	wkITE(6+169)
έo	169 FORMAT(37%, ALPHA', 5%, HETA ',4%, PHI ',5%, NA ',5%,
29	1' COST'+JX+'EST.NC'+4X+' MTTF '+4X+'NA-NC')
30	WRITE(6+170)ALPH+0ETA+PHI+NA+COST+7NC+RMTTF+PERR
31	LO 89 I=1,NN
32	Cwū=Chū+LND(I)
33	
34	WHITE(6+890)CHD+UNC
ა5	390 FORMAT(1UX+1ND=++2X+F10.3+2X++NC=++2X+F10.3)
36	89 CUNTIFUE
37	GU TO 11
38	10 KEAU(5+151)(S(I)+1=1,NN)
39	151 FORMAT(G10.1)
40	11 REAU(5+088)ZA+ZB+2P+ZN
41	888 FCRMAT(G14.6.2X.G14.6.2X.G14.6.2X.G14.6)
42	#KITE(6,169)
43	CALL KOST (ZA+ZB+ZP+ZN+S+NJ+T+NN+TD+H+GJ+DHD+COST+ENC+RMTTF+
44	12(NC+RERK)
45	ULAME1
46	COSTIECUST
47	
48	CALL EIGMIN(H)CORH, GJIDLAM, EH, EV)
49	72 WRITE (6) 171) ZA, 26, 29, 20, CUST, 2NC, RMTTF, RERP
50	
51	
54	
33	
54	
55	
56	IF (CUSTILIERRIGU TO 73

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37		TAU=(J=LB)/(Z=DAES(CURR(Z)))
58		if(1AU.g1.(.2-ZP)/(2+DABS(CuRR(3))))TAU=(.2-ZP)/(2+DABS(COPR(3)))
59		1F(TAU-GT-1)TAU=1
60		IF(TAU.LT0001)60 TO 74
01		28=28-Curr(2)+TAU
62		2F=2P+Cukr(3)+TAU
63		2H=2N-CURR(4)+TAU
64		IF(28.LT.0)28=0.
05		IF (COST.GT.10)DLAM=1.
0 0		1F(20.67.10.)28=10.
υ7		LALL KOST (ZA+ZB+ZF+ZN+S+NU+T+NN+TD+H+GJ+DND+COST+ENC+RMTTF+
00	1	(21+C+RERK)
69		IF(COST.GE.PCUST)GO TO 76
70		IF (COST/PCOST.GT99) IFLA3=IFLAG+1
71		IF(IFLAG.GT.4) GU TO 74
72		CALL EIGMIN(H,CURK,GJ,DLAM,EH,EV)
75		PCOST=CUST
74		DLAM=DLAM/10
75		60 TO 72
70	76	CALL KOST (ZZA+ZZB+ZZP+ZZN+S+NJ+T+NN+TD+H+GJ+DND+COST+ENC+RMTTF+
77	1	ZNC+RERN)
7o		CALL EIGMIN(H, CURR, GJ, DLAM, CH, EV)
79	71	TAU=TAU/10
60		
61		IF (INDEA.GT.5) GU TO 74
62		2 h = 2 2 h
63		
64		
6 5		20=22N
60 60		2/1-228-CORR(2)+TAU
b7		
6h		ZUEZN-CORR(4)+TAU
1.4		F(2n-L)-0)2(=0-
40		F(28-61-10-)28=10-
41		CALL KOST (7A + 7B + /F + 7H + S + HJ + 1 + HN + TD + H + GJ + DHD + COST + FNC + R ^M TTF +
92	1	
4.5	-	IF (COSTALTARRIGO TO 73
44		1F (COS1-6F-PCOS1)60 TO 71
45		CALL FISHIN(H+CONF+GJ+DLAM+EH+EV)
46		27A 27A
97		
94		//927P
99		22/122N
144		k 1 TE (6+172) 74+2m+2P+7N+COST+7NC+RMTTF+REPR
101		
102	74	
16.3		
104		
145		
166		
147		74=/74
100		762778
1.19		/4272P
110		
111	77	
112		
113		1706778 76184711 AMA1

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114	CALL MARG(EH:EV:ULAM:CORR:GJ:H)
112	78 20=225-CURR(2)+TAU
116	ZP=ZZP-LURR(3)+TAU
117	ZN=ZZN-LORR(4)+TAU
113	IF(20.GT.10)2B=10.
119	IF(28.L1.0)28=0.
120	CALL FOST (ZA+2B+2P+ZN+S+NJ+T+NN+TD+H+GJ+DND+COST+ENC+RMTTF+
121	
122	IF (LMUEA.EQ.1) WRITE (6, 190) ZA, ZB, ZP, ZN, COST, ZNC, RMTTF, RERR
123	
124	
125	MRIE (D/17)/ZM/ZD/ZF/ZM/CUSI/ZNC/RM/FF/RERP IS/CAST I T SDOVAL 76 78
1.7	$\frac{1}{16} \frac{1}{16} \frac$
120	1. AMENTER VI
129	TELEDET/FORST LT - Q AND POOST AT SAVAR TO 72
150	
1 51	
1.52	
133	//=//P
1.54	
135	CALL KUST (ZA+ZB+ZH+ZH+S+Nu+T+NN+TD+H+GJ+DND+COST+FNC+RMTTF+
100	12ILC+KERK)
1.57	CALL EIGMIN(H,CORR,GJ,DLAR,EH,EV)
84 ل	らい Tひ 7と
139	91 ZZA=ZA
140	22b=2B
141	¿2P=ZP
142	22N=ZN
143	JUEX=JDEX+1
144	IF (JDEX.EG.1) TAU=TAU=5*+ICLAM
145	COSTIECOST
140	GU IV 76 73 - 6175/4 - 73 - 74 - 70 - 74 - 67 - 746 - 67 - 756 - 67 - 67
147	13 WRITE (DT1/3/2A)2D)27/2N/CO31/2N/TM/TF/KERM 370 Schwat(Sta)(Sta)(Atton: VALUES, ADSAV, ADS)
144	110 FORMATION SIMULATION TALUES ARE STEATSTRASTEAT
1.50	171 FCHWATTALING OFFANTARE VALUESTARY 3/FA.5.291.
151	$13(FA,2*23) \cdot F10, b \cdot 23 \cdot FA, 2)$
152	172 FORMAT (5X, *AFTER REDUCING STEP SIZE + 4X, 3(FA, 5, 2X).
153	13(Fp.2+ <x)+f10.6+2x+f8.2)< td=""></x)+f10.6+2x+f8.2)<>
154	173 FORMAT (9X++FINAL VALUES ARE++2X+3(F8.5+2X)+
155	13(Fa.2, zA), F10.6, zX, F8.2)
100	190 FURMAT(5X//CHANGING LAMBDA //2X/3(F8.5/2X)/
157	13(F0.2, LX), F10.6, 2X, F8.2)
158	191 FORMAT(5X; STEEPEST DESCENT VALUES '2X;3(F8.5;2X);
159	13(F8.2, LX), F10.6, 2X, F8.2)
λuŭ	SUBROUTINE EIGMIN(HH,CORR,GGJ,DDAM,EH,EV)
161	IMPLICIT REAL+8(A-H,O-Z)
102	DIME(SIUN_HH(4,4),B(4,4),EV(4),EGV(4,4),EH(4,4,4),DIA(4),OFDI(4),
163	17EMP1(4), TEMP2(4), R(4,4), CORR(4), GGJ(4)
104	
103	U /UI J=1)4
167	
164	0(1/0/-IN(1/0//036K)(0AD3(NN(1/1/)*NN(J)J)) 701 (0AT1ANG
120	
170	CALL TRAINEY (4-4-H-DTA-OFD()
÷, .	

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171		CALL EIGVAL(4,EV,LIA,OFDI,TEMP1,TEMP2)
172		CALL EIGVEC(4,4,6,DIA,OFDI,EV,EGV,TEMP1,TEMP2)
175		LU 708 1=1,4
174		DO 709 J=1,4
175		LG 710 KK=1+4
176		EH(1,J,KK)=EGV(J,1)+EGV(KK,1)
177	710	CONTILUE
176	709	CONTINUE
179	708	LONTINUE
100		10 713 1=1.4
161		IF(FV(1)-6T-0)60 10 713
162		$1F(DDAM.LTEV(I))DDAM=DA_{1}S(EV(i))+.1$
183	713	CONTINUE
104		10 712 N=1+4
165		$1 \in (1 \times 10^{-5} \times 10^{-5}) \to 1 \to 1 \to 0.001 \to V(KL) = 1$
100		
1+7		
1		00 704 0-174
1.0	7.1	
107	2, 3	
141	703	
191	102	
192		
195		DC 706 J=1,4
194		LJD=CSGRI(DABS(HH(I+I)+HH(J+J)))
195	. .	R(I,J)=R(I,J)/DDD
190	706	CONTINUE
197	705	CONTINUE
195		
199		
200		JU 712 J=1,4
201		CORK(I)=CORR(I)+K(I+J)+GGJ(J)
202	712	CUNTINUE
203	711	CONTINUE
204		KETURN
205		SUBROUTINE MARQ(LEH,EEV,GULAM,CCORR,GGGJ,HHH)
200		INPLICIT REAL=8(A-H+0-Z)
207		LIMENSION EEH(4,4,4), EEV(4), CCORR(4), GGGJ(4), RR(4,4), HHH(4,4)
208		UO 714 1=1+4
209		CCORR(I)=0
210		JQ 725 J=1,4
211		R((↓ • J) = J
212	725	CONTINUE
213	714	CONTINUE
214		U0 716 AL=1/4
215		60 717 1=1,4
210		UO 718 J=1,4
217		RR(I,J)=RR(I,J)+EEH(KL,I,J)/(EEV(KL)+DDLAM)
218	718	CONTINUE
219	717	CONTINUE
220	716	CONTINUE
221		40 719 J=1.4
222		00 720 1=1.4
223		UUDDEDSWRT (DABS (HHH (I.I.) +HHH (J.J.))
224		KR(1+J)=KR(1+J)/4L00
225	720	CONTINUE
220	719	CONTINUE
227		NO 721 A=1.4
		** 1 ** * = 4 77

- D.4 -

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260		Lu 722 J=1,4
229		LCORR(I)=CCORR(I)+RR(I+J)+GGGJ(J)
200	722	CONTINUE
231	721	CONTINUE
202		RETURN
233		END

<**>

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WHATIS NELY (1) .. NOST

伝統は彼らい 歩きしょう やらい どちこ

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i.

RELY+RELY(1).	KUST
1	SUBROUTINE KOST (AA+BB+PP+INA+SS+11J)+TT+NMN+DT+H+G1+DND+
2	1CUST, ENL + RMTTF + 2NC + RERR)
ప	IMPLICIT REAL+8(A-H,O-Z)
4	REAL+B INNA
5	UIMEISIUN TT(300),55(300),60(4),H(4,4)
9	DIMENSION DUMD(4), DND(300), UNC(300), ZYTTF(300)
/	LIMENSION DI(4.4)
Ö	COSTED
9	
10	
1.	
14	
14	
19	
10	
17	$a_{\lambda}(1) = 0$
1	90 12 JE1.4
19	
έů	$u_1(x, y) = 0$
<u>_1</u>	12 CONTINUE
ic c	13 LUNFINUE
د ع	MII-AA+PP+DT
£4	Ac=(1-Ax)*DT+6B+xx=+P+0T+xA++2
e J	2L=1-UT+(88+AA+PP)
έo	∧7=P₽≠U「
72	λd=(FP+DT)++2
K 0	A9=AA+Ab
C Y	A1U=A6+AA++2
20	A11=AA+A7
31	
32	
3 5	
35	
30	00 10 REFERENCE 171111111
37	LUC (X) SPPENNAEDTE (71) 21+4+.(1) 221
-18 -18	
39	2(1C=21)C+UNC(K)
40	ZHTTF(K)=1/(PP+(NNA-2NC))
41	IF (K.EG.NMN) HMTTF=ZMTTF(K)
42	IF (K.EG.IVMN)RERETINA-ZNC
43	IF (K. EU. WMN)ZZNA=NNA-ZNC
44 X4	CUST=CO≥T+(SS(K)-LNU(K))++2
45	00 41 I=1+4
40	GJ(I)=-∠#(S5(K)-uhū(K))#UµND(I)+GJ(I)
47	JU 42 L=1,4
48	H(I,L)=2+(DDND(I)+DDND(L)-(SS(K)-DND(K))+D1(I,L))+H(I,L)
49	42 CONTINUE
20	
51	
5.5	£671671471777777777777777777777777777777
34	€►NEI=6₽₽EI=€EEEEEEEEEEEEEEEE //₭/3557/↓316{#8880#11↓7↓↓336
55	##N#E=#################################
56	UUNU (2) IPPANNARDTa (KeA2#ZLL 12)

- D.6 -



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Contraction of the second second

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SPATIS HELY (1) . SUB2

X

RELY + RELY (1) . SUB	
1	SUBROUTINE KIETA(A)BIPINAITIVIRRIJJJ)
2	IMPLICII REAL+8(A-H+0-Z)
3	KEAL+6 MA
4	REAL RAIRBIRPIRINA, RTIDIRZ
5	KEAL RQ
D	RQ=SNGL(KR)
7	KAISNGL (A)
5	K6=SNGL(6)
У	KP=SNGL (P)
10	KIA=SNGL (NA)
11	rT=\$14GL(T)
12	#RITE(6+67)
13	UIMENSIUN X(2)+V(300)+Y(2)
14	x(1)=0.
15	x(2)=0.
10	00 100 K=1+JUJ
17	Y(1)=X(1)+P+T+(4A-X(2))
15	LINA-X(Z)
19	J=SNGL(E)
2 0	Y(2)=6+1+X(1)+(1-T+(6+A+P))+X(2)+A+P+T+NA
Z1	x(2)=Y(∠)
<u> </u>	CALL POISS(DIRZIRFIRTIRGIR)
د ے	A(1)=X(1)+DuLL(K2)
<u>د ب</u>	ьйITE(6+88)X(1)+X(2)
25 08	FURMAT(10X++KU++22+614+6+2X++KC++2X+614+6)
20	v (K)=DBLE (RZ)
د7 ۲۵۵	CGNTINUL
20 07	' FORMAT(5x++RANDUM VALUES FOR ND AND NC GENERATED FOR SIMULATION+)
29	KETURN
UL UL	E1.0

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GPATIS HELY (1) . SUBS

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4

RELY+KELY (1) . SUE3	
1	SUBROUTINE PUISS (UD, ZZ, PP, TT, RRR, KKK)
č	DIMENSION C(100)
ن	IF (KKK.GT.1)60 TO 555
4	C(1)=RRKR
5	60 TU 600
ະ ວ55	C(1)=C(_U0)+2++25
7 006	U=00+PP
8	CALL RAHLEX(C+100+0)
9	u=0.
10	UO 100 N=1+100
1 L	G=G+C(K)
14	(F(Q.LT.TT)G0 TO 100
13	ZZ=FLŮAI(K-1)
14	60 TO 200
15 100	CONTINUE
16	22=100
17 == 10	RETURN
10	E.J
A	

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HELY+GAUS	SILEAT(1)	080C	
1		SUBROUTINE THIDHA (NONMOLODOB)	TRID0010
5	<i>c</i>	INPLICIT REAL+8(A-H+O-Z)	
3	ž	TD1. TACANAL T. ATTAN AT DEAL EXAMPTATE MATSTV	
	č	IRIUIAGUNALIZATION UF REAL DIMMETRIG MAIRIA.	
6	•	DIMENSION A (INM NOM) +D (NM) +B (NM)	TRID0020
7	Ç	~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~	
8	ç	SAVE ORIGINAL DIAGON/LS IN ARPAY D.	
	•		TRICOLA
11	14	U(1)=A(1,1)	TR100050
12	Č		
13	ç	FOR N-2 RETURN WITHOUT COMPUTING.	
14	ç		
15	<i>c</i>	11 (11-2)00,55,15	TRIDUUSO
17	č		
18	15	60 86 KE3:N	TR100060
19		KX=X=1	TR100070
20	¢		
42	c	SUM CONTAINS THE SUM OF THE SQUARE ELEMENTS	
22	ç	OF A COLUMN, EXCEPT THE FIRST K+2 ELLMENTS.	
23	c		
24			TRIDUGAD
25	20		Te100100
27	Ē		
28	č	THE & ARRAY LONTAINS THE BETA VALUES.	
29	Č	(1.L., BETA=SUM++(1/2))	
30	c	***************************************	
31	-	b(k-2) = bSIGN(DSGHT(SUM)) - A(K-1)K-2)	
32	ç		
33		IF BETALU NU TRANSPORMATION IS INITIATED.	Tu100120
35	c		IKIUUIEU
30	č	THE LOMPONELITS OF THE COLUMN VECTOR . ANE	
37	č	STURED IN THE PUSITIONS OF THE ANNIHILATED	
38	Ċ	ELEMENTS OF A (I.E., LOWER HALF OF A)	
39	C		
40	24	A(K-1+K-2)SUSQRT(U.5=CA85(A(K-1+K-2)/8(K-2))+0.5)	
41		UENUM=-2.+A(K-1+R-2)+B(K-2)	THI00140
42	•	DO 30 IERIN	TRID0160
43	20	A (] + R - 2) = A (] + R - 2) / DENOM	TRIDU170
	47	JUL SA LIAK.N	Te100140
46	ີ້	••••••••••••••••••••••••••••••••••••••	
47	ē	THE BETAS ARE FORMED ONE BY ONE.	
48	¢	WHEN R OF THEM HAVE BEEN FORMED.	
49	C	P AND & CONTAIN ONLY (N-R) ELFMENTS.	
54	c	THE E ARRAY IS UED TO STORE SUCCESSIVE	
21	L.	PIS AND 015.	
25	C		
22			TRICU190
24		TING CONTINUE TO SOU	18100101
20	يەرونى	n(.)\$\delta + L)\$A(L.\$A/L.	TAINDIGS
			· · · · · · · · · · · · · · · · · · ·

- D.10 -

bu C SCAL=#(TRANSPOSE)*P bl JJ=J TR bl JJ=J TR bl JJ=J TR bl U 40 J=K;N bl U 40 J]=b(J)=SCAL=A(U;K-2) TR bl U 40 U]=b(J)=SCAL=A(U;K-2) TR bl U U J=K;N TR bl U U J]=b(J)=SCAL=A(U;K-2) TR bl U U U TR bl C TRANSFORM ALL ELEMENTS OF A TR bl U TRANSFORM ALL ELEMENTS OF A TR bl U TR TR bl U TR TR bl U TR TR bl U TR TR bl U U TR c TR TR TR c	100200
C TRANSFORM ALL ELEMENTS OF A U7 C EXELPT PIVOTAL ROW AND COLUMY. 09 UC 45 JERKIN 70 UC 45 LEUN 71 UC 45 LEUN 72 40 ALLUJER(LIJ)-2.*(A(LIK-2)*0(J)+A(JIK-2)*P(L)) 73 C 74 C 74 C	100215 100220 100230 100240
09 UC 45 JENKIN TR: 70 UC 45 LEUIN TR: 71 45 ALLIUEN(LIU)=2.*(A(LIK=2)*0(J)+A(JIK=2)*P(L)) TR: 72 46 CONTINUL TR: 73 C TR: 74 C RESIJRE ORIGINAL DIAGONALS OF A. STORE DIAGONAL	
74 C RESIJRE ORIGINAL DIAGONALS OF A. STORE DIAGOANL	100250 100260 100270 100275
TO OF IKANSFORMED MATKIX IN ARRAY D.	
77 $U \cup U = 1 + N$ TRI 78 $1 = A (I + I)$ TRI 79 $A (I + I) = U (I)$ TRI 60 $J = N - I$ TRI 61 $U (J + I) = U (J)$ TRI 62 $5 U U (I) = T$ TRI	D0280 D0290 D0300 C0301 D0302 D0310
b3 b5 b(N)=A((*,*)+1) TKI 64 b0 b(1)=0.0 TRI 85 KLTURI: TRI 80 E(N) TRI	C0320 C0330 C0340

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WHAT+5 GAUSSHENT+SUL7

Alter Sala de Cara an

RELYOUS	51.E.T (1	.SU87		
1		SUBROUTINE EIGV	/AL(LP)L;A,B;W)F)	EVAL
2	c	IMPLICIT REALTO		
5	č	 I P	IS THE STOP OF AFRAY A.	 Fval
5	č	F	IS A VECTUR OF LE FLEMENTS WHICH WILL HOLD	EVAL
	č	ŤH	E ETGENVALUES IN DESCENDING ABSOLUTE ORDER.	EVAL
ž	č	A	IS A VECTOR OF LP ELEMENTS GIVING THE DIAGONAL	EVAL
	č	EL	EMENTS OF THE TRIDIAGONAL MATRIX.	EVAL
5	č	8	IS A VECTOR OF LP ELEMENTS, THE LAST LP + 1	EVAL
10	č	2	IS A VECTOR OF LP ELEMENTS USED FOR TEMPORARY	EVAL
īi	č	ŠT	WRAGE.	EVAL
12	ċ	F.	IS A VECTOR OF LP ELEMENTS USED FOR TEMPORARY	EVAL
د ۱	Ċ	ST	IONAGE.	EVAL
14	ċ			
15		DIMENSION E(LP)	(A(LP)(B(LP)(W(LP)	EVAL
10		UINENSIUN FILP)		
17	c			÷-
18	c	F1	IND ABSOLUTE BOUND FOR THE EIGENVALUES	EVAL
19	ć			
μ.		AMEDALS(A(1))		
4 1		LN=U.		EVAL
-2		LU 1 1=L+LP		EVAL
ڏ <u>ن</u>		AN=DNAX1 (AM+DAB	35(4(1)))	
2 4	1	BH.=UMAXI (BM+LAB	3 > (1 (1))	
2 5		BUZAM+bin+bin		EVAL
20		LU D IFAILP		FVAL
∠7	c		***************************************	
žυ	c	тн	US LOOP FORCES THE EIGENVALUES TO LIE PETWEEN	EVAL
29	č		US AND MINUS CHE. THE E AND & VECTUPS ARE	EVAL
ů.	c	KE	SPECTIVELY LUK AND HIGH ESTIMATES TO ALL	EVAL
21	ć	Th	IL FIGERVALUES.	EVAL
26	C			
دن		A(I)=+(1)/bL		EVAL
54		b(I)=C(i)/ bC		EVAL
55		E(I)=-1.0		EVAL
30	ó	w(1)=1.0		EVAL
37		00 50 K=1+LP		EVAL
30	c			
34	c	FI	IND THE K-TH EIGENVALUE. ALSO LOW AND HIGH	EVAL
40	L	LS	STIMATES FOR THE K+1-ST TO LP-TH EIGENVALUES	EVAL
41	C	AR	RE IMPROVED. THE EIGENVALUES APE FUUND IN	EVAL
42	L	AS	SCENDING ONDER.	EVAL
43	L	TH	E K-TH EIGLAVALUE IS CONSIDERED FOUND IF THE	EVAL
-4-4	C	TH	HE HIGH AND LOW PLACES AGREE TO SEVEN DECIMAL	EVAL
40	L	FL	ALES.	EVAL
40	¢.		***************************************	
47	8	1F((a(K)-E(K))/	/WMAX1(DA85(W(K))+DARS(E(K))/1.E-29)-5.E-8)50+50+1	0
40	10	X=(+(K)+L(K))+0).5	EVAL
49	C			
50	ċ	X	IS A GUESS FOR THE K-TH EIGENVALUE, COMPUTE	EVAL
51	c	NU	JABER OF EIGENVALUES EQUAL OR EXCEEDING X RY	EVAL
24	C	US	SING STURM SEQUENCE (OPTEGASS METHOU).	EVAL
53	Ċ			
24		52=1.0		EVAL
22		F(1)=A(1)=X		EVAL
24		IF(F(1)) 102,10	04+104	EVAL

- D.12 -

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57	102	51=-1.0		EVAL
50		N=0		EVAL
23		GU TO 105		EVAL
b Ú	104	51=1.0		EVAL
¢1		h=1		EVAL
04	105	DO 120 1=2.L		EVAL
UJ		IF(8(1)) 106	113/100	EVAL
64	The	IF (B(1-1)) 10	J7+114+107	EVAL
63	1.17	IF (DABS(F(I~)	[))+UAB5(F(1-2))-1.E-1~)111,112,112	
υç	Č.		TE THE ORDER TO THE TEORE OF THE STILL IS A CONFICT	Even I
6/	Ļ		IF THE PREVIOUS TWO TERMS OF THE STORM SECONDICE	EVAL
69	ç		WERE VERT SMALLE THET ARE FURCED TO BE CLOSER	EVAL
20	, c		TO ONE IN MAGNITODE TO AVOID ONDERFLOW PROBLETS.	EVAL
13	C .			Eval
11	111	- F (1-1)=F (1-1)	1 1 1 1 1 1	EVAL
74	• • •	F(1=2/=F(1=2)	/#1+610	EVAL
73	115	- F(1)-(A(1)-A)	/#F (1=1/=D(1/#3(1/#F(1=2)	EVAL.
74	1.4	5(1)=(A())=)	1451	EVAL
75	112		1+51	EVAL
22	5.14	GU 10 115	AF(1-1)-051((())11)-0(1)-52)	LVAC
1.	114	- F (1/4/4/1/=A)	1+P11=11=0310((1011)+5(1))/321	EVAL
10	110	- 7257 1876/11116.1	17.116	EVAL
12	116	61-1516-151.6		
	110	15/6146.1117	- 120.117	FUAL
01	117	- 1F (31+32/11/1		FVA1
52	1.1	CONTINE		EVAL
	<u> </u>	CONTRICT		
	č		NON LET N HE THE MUMPER OF	EVAL
00	ž		FTGELVALUES SMULLEN THAN X.	EVAL
	2			
		N=1 Bet		EVAL
20		TENSITIK) 60	0 Tu 20	EVAL
40	Ĺ			
ý,	ĩ		A DELOMES AL UPPER FOUND FUR THE	FVAL
<u>.</u>	ū		K-1H TO N-TH EIGENVALVES.	FVAL
4.5	č			
44	1	00.15 326-0		EVAL
45	15	L (J) = X		EVAL
50	<u> </u>	h=0+1		EVAL
97	ĩ			
90	Ĺ.		IF ALL THE EIGENVALUES ARE SMALLER THAT: A.	EVAL
44	<u>ر</u>		TEST AHETHER WE HAVE CONVERGED TO THE	EVAL
100	Č		K-TH EIGENVALUE.	EVAL
101	Ĺ.		***************************************	
102		IF (LP.LI.N) (50 TU 6	EVAL
103	24	UU 20 JEINLY		EVAL
104	Ļ			
100	c		IF A IS LARGER THAN PREVIOUS LOWER	EVAL
100	Ĺ		BOUND INCHEASE THE LOWER BOUND.	EVAL
107	c			
100		If(X = E(J))	818120	_
109	<0	L(J)=X		EVAL
110		60 TO 8		EVAL
111	50	CUNTINUE		EVAL
1	Ĺ		~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~	
113	Ĺ		RESTUPE INPUT AND SCALE EIGENVALUES.	EVAL

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- D.13 -

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C		
6(D0 60 I=1;LP A(1)=A(1)=BD B(1)=B(1)=BD W(1)=C(1)=F(1))=FD=0.5	EV Ev Ev
Č-		٤¥
c	SORT EIGENVALUES IN ARSOLUTE DESCENDING ORDER	E١
	J2LP	 F\
	K=1	F
	D0 80 J=1/LF	-
	1F(DABS(h(K))-DABS(W(J)))03,63,65	•
63	E(])=w(J)	E
	<u>1-0-1</u>	Ē
	GU TO BU	Ē
65	E(I)=N(N)	Ē
	N=K+1	F
80	CUNTIFUL	Ē
	RETURN	Ē
	E offa	-

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SPATIS GAUSSNEWT.SUDB

- D.14 -

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RELY+GAUSS	NE#1(1)	• • SUBB	
1		SUBROUTINE EIGVEL (LP+NM+R+A+B+E+V+P+Q)	EVEC
2	c	IMPLICIT MEAL+8(A-H)(-2)	
3	č	P IS THE GIVEN MATERY.	FVEC
2	č	LP IS DIMENSION OF THE MATRIX P.	EVEC
5	č	NM IS THE MAXIMUM DIMENSION OF B AND V.	FVEC
7	č	A ANT THE DIAGONAL ELEMENTS OF THE THIDIAGONALIZED	R.EVEC
	č	A ALE THE OFF-UIAGUNAL ELEMENTS OF TRIDIAGONALIZED	R.FVEC
ų	č	E ARE THE EIGENVECTORS OF R.	EVEC
10	č	V WILL HOLD THE EIGENVECTORS, STOPED COLUMNWISE.	EVEC
11	č	P ANL & ARE VECTORS FOR TEMPORARY STORAGE TO	EVEC
12	č	HOLL CUEFFICIENTS OF THE LINEAR EQUATIONS WHICH	FVEC
13	Ċ	LETERMINE THE EIGENVECTORS.	EVEC
14	C		
15		JIMENSION R(NM+NM)+A(LP)+6(LP)+6(LP)+V(NM+NM)+P(LP)+Q(LP)	EVEC
16		LP1=LP-1	EVEC
17		50 IA=1+LP	EVEC
16		$\lambda = \Delta(1) - L(1 \times)$	EVEC
19		YEB(2)	EVEC
20	С		
21	C	GENERATE COEFFICIENTS TO COMPUTE THE IX-TH	EVEC
44	C	EIGENVECTOR, FIRST PICK BEST PIVOT ELEMENT.	EVEC
25	C		* •
24		LU 10 I=1+LP1	EVEC
∠ 5		1F (UABS(X)-DABS(B(1+1)))4,6,8	
ĉυ	4	P(1)=0(1+1)	EVEC
∠7		u(I)=A(I+1)→C(IX)	EVEC
2d		V(1+)X)=D(1+2)	
٤4		<u> </u>	EVEC
30		X=Z=Q(I)+Y	EVEC
JL		1F(EF1*ME*1) X=5+/(1*1X)	EVEC
32		GU TO 10	EVEC
ذ ذ	D	1F(X) 8+7+8	EVEC
34	7	A=1.0E=10	FVEC
35	ь	P(I)=X	EVEC
30		0(I)=X	EVEC
57		v(I,IX)=0.0	
30		x=A(I+1)-(B(I+1)/x+Y+E(Ix))	EVEC
37		A=B(1+5)	EVEC
40	10	CUNTINUL	EVEC
41	C		
42	ç	NON SOLVE THE ABOVE EQUATIONS FOR THE EIGENVECTOR.	EVEC
45	C	TEST LAST PLVOT ELLMENT.	EVEC
44	L.	**************************************	
40	20	IF (X)21+28+21	EVEC
40	3 7	V(LP,IX)=1.0/X	MAFC
47	L C		
40	C C	CONTINUE WITH THE BACK SOLUTION.	
49	é		EVER
20	"		EVEC
5.		V(1/1/1/1/1/1/1/1/1////////////////////	FVFC
52		A#¥16714AJ4+&+\$1818AJ#+6 19191	EVEC
53	40	4=4=4	FUEC
34	э.	17 (1/20/00/20	FUEC
22	∠ U	1	EVEC
		V#4.1.91801	

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- D.15 -

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57		GU TO 25	
50	28	V(LP,IX)=1.0E10	EVEC
59		GO TO 2∠	EVEC
60	50	X=DSQRT(X)	
61		D0 311=1+LP	FVEC
62	31	A(I+1X)=A(I+1X)\Y	EVEC
ъĴ	C	**************************************	
1 14	۲.	TRANSFORM EIGENVECTOR FOR THE TRIDIAGONAL	EVEC
60	ç	MATRIX TO AN EIGENVECTOR OF THE ORIGINAL MATRIX.	FVEC
60	c	+=++++++++++++++++++++++++++++++++++++	
07	Ĺ	######################################	
60		IF (LP.EU.2) GU IV 50	EVEC
67			EVEC
70		$K \equiv LP = KK + 1$	5
71		1=0.0	LVEL
12		00 35 IEA, LP	FVEC
73	30	A=A+A(I+TX)+K(I+K-I)	FVEC
74		DO 40 I-RILP	FVEC
75	40	<i>↓</i> (],] <i>X</i> }=V(],] <i>X</i> }=2.0#Y+R(],K=1)	
70	46	CON INC	EVEC
77	50		EVEC
76		KÉTURI.	EVEC
79			

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WAWT RELT(1)-MOS Interval Lebuth = 1.50 NUMLER OF INTERVALS = 60 TERVINATION CRITERION = .100060

5 x 31 2 3

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UPLUS MINIMIZATION ATTH ALPHA FINED

BAUT	אברץ (ג) יאט זאויבהעאר כו	S LN6TH = 1.50	HHUS	רא סב זויונייגערי = פיו	TEPHINATICN CRIFFION = .100000
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				.420004+001	
	27	200004Lu2	22	.873164+001	
	2	.170000+042	U X	.13538/+002	
	2	.220000+002	ъ К	.1797.Jo+UN2	
	2	20000+UU2	ž	*2240U1+U02	
	Š	.330000+002	¥	.265197+002	
	Ş	.350000+002	ŭ	-JIIJ46+002	
	2		ų i	• 35260 / +UR2	
	2	200+00094.	<u></u> צ		
	2:	200+0000+C*	ž i		
	23	5001000000°	ן ר ג ג		
	23	200000+002	202	561431+002	
	22	.770400+062	22	.629067+002	
	2	.620000+UU2	ų ž	.678940+UD2	
	3	SU00004602	¥	•727847+U0ž	
	2	2JU4004078.	ų Š	.769885+002	
	Š	24000046.	ÿ	.8107 35+002	
	ş	.960000+462	ž	#55099+002	
	л Э	.101000+003	ž	.894877+002	
	ž	.101000+003	ЧU	•935353+U02	
	Š	500+000£01°	¥	• 968908+UN2	
	2	.106000+003	¥	200+#2/666*	
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	23	.1120U0+UU0	y ¥	100+919CN1*	
	53	CUU+UU4211.	ž	CUO+/CCOAT•	
	2 3	2004000/77*	, , , ,		
	23	CUU+000211.	د د		
	22	124000+603	ر د ۲		
	2	.147000+003	J ¥	.121530+003	
	NIN N	.128000+003	U Y	.123996+UnJ	
	ž	EUN+JOUEC1.	Ŷ	.126194+U03	
	Ş	.134004+403	¥	 128765+U03 	
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3	14400444	3	J.	.15315-0-03							
2	-1280000	2	J Z								
2	-1000801-	2	y Y	.156033+003							
3	12+099651.	3	ĩ	.15725±+un3							
2	.102000+u(3	ž	. I 584 L < + UOJ							
2	· 102001 • 0	2	y :	500+C29651.							
2	.1000001.	2	ž	.161443+003							
7:	-166400+4	3	ų.	.162937+003							
3		2	ž	707++/1861+							
2	-166000+0	2	y Y	.165201+003							
2	• 10000000	2	Y	CON+755997 .							
3		3	¥	•167303+003		100		1991	- MC		
Simh Allen	VALUES ARE				00.005	160.14	169.45	1.436535			
	6.000										
17	11.674	ÿ		8 582							
17	17.617			12.095							
	23.22			17.699							
Ĩ	28.693	ž		22.356							
17	34.022	ž		27.037							
5 M	39.211	ÿ		31.717							
I QN	962.44	ž		36.375							
H (M	49.168	"X		*0°.994							
II N	53.938	ΰ¥		45.559							
ЦИ	54.572	ÿ		50.059							
	63.070	빙		54.485							
1 1 1	67.435	ÿ		56.828							
17	71.670	ÿ		63.084							
7	75.778	ÿ		67.247							
17	79.761	ÿ		71.315							
	43.621	ÿ		75.284							
II (III	67.460	ų		79.154							
1177	996 95	"¥		62.923							
11)N		ÿ		86.591							
	51.903	ÿ		90.159							
ווש	101.190	ÿ		93.627							
102 102	697.977	ÿ		96.997							
1	107.479	ÿ		100.269							
H J	110.471	Ë Ë		101.444.001							
Ĩ	113.368	ÿ		106.526							
77	116.172	ÿ		109.515							
12	118.876	Ë		112.414							
Я <u>.</u> .	12.121	ÿ		115.224							
1) 1)	144-057	ÿ		117.946							
	126.519	ÿ		120.547							
1.3	106-8-1	¥		23.145 Cal							
12	131.207	ÿ		125.622							
	1-1	2		126.022							
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	1.27.067			132.596							
1	501 - 507										
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2		23									
1	144-400			144 67.							
	1.10.540										
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DATE 06/11/79 MTF 5224558 5527558 5527558 5527558 5527558 5527558 55295197 5529519 5529519 1.0299139 1.029719 1.2256410 1.2256410 1.4510930 1.45109365 1.5539255 1.5539255 1.5539255 1.6223145 1.6223135 1.6223145 1.6223135 1.6223135 1.6223135 1.6223135 1.6223135 1.6223135 1.6223135 1.6223135 1.6223145 1.6223145 1.6223145 1.6223145 1.6223145 1.6223145 1.6223145 1.6223145 1.6223135 1.622315 1.6 EST.NC 1103.04 1103.04 1103.04 1103.04 1103.04 1103.04 1103.04 1103.04 1103.04 1103.04 1103.04 1103.04 1103.04 1103.04 1105.04 105.04 100.04 10 C057 2270.52 2270.52 2270.52 2274.67 2274.67 2274.67 2274.67 2274.67 2274.67 254.67 1559.56 1559.55 15 PHI 0100 01053 01119 01053 01137 01119 01245 01255 01255 01255 01255 01271 01275 02055 02055 02055 02055 02055 02055 02055 02055 02055 02055 02055 02055 02055 02055 02055 02055 02055 151.44.200 151.49.401 151.49.401 151.49.401 151.49.401 151.49.401 151. WE PANAME TEN VALUES ME PANAME TEN VALUES ALPHA FAED 152.146 155.700 155.700 155.700 155.700 155.700 159.720 165.700 165.700 165.700 165.700 165.700 165.700 165.700 1770.801 165.700 1770.801 1770.805 HITE NULTALMINIM

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MANINAZATAUN WATH ALPHA FAAED							DATE 06/	11/79	PAGE
STEEPEST DESULIST VALUES	000045.	00111.	.02068	193.01	159.52	166.52	1.825685	26.49	
CIMMAD.NO LAMDUA	.5000	67111.	.02056	192.91	159.52	166.38	1.824744	26.52	
STEEPEST ULSULNT VALUES	.5uuno	.11119	.02066	19.91	159.52	166.38	1.824744	26.52	
MLA PANAMETEN VALUES	.5u000	.11100	.0205A	193.01	159.52	166.52	1.825485	26.49	
CHANGING LANUDA	.50000	nc111.	.02064	192.96	159.53	166.36	1.021065	26.61	
CHANNESING LAMOUA	00095	.11111.	.02067	192 .9 9	159.52	166.50	1.825411	26.50	
CHANNEL LANBUA	.50000	.11101	.0206R	193.01	159.52	166.52	1.825662	26.49	
CHANULING LANGLA CHANLING LANGLA	10005.	00111.	.02064	193.01	159.52	166.52	1.825483 1.925485	26.49	
CHAMPLING CATCUA CHAMPLINE LANGER	000051	01111	10000-	19.001	150.52	166.52	1.425645	54.69	
CHANNELING LANNADA	.5000	11100	02068	10.51	159.52	166.52	1.825685	26.49	
CHANGLING LANGUA	.50000	.1110	.02068	193.01	159.52	166.52	1.825685	26.49	
CHANGING LANDUA	.50000	.11100	.02068	193.01	159.52	166.52	1.825685	26.49	
CHANELINE LANGUA	• 50000	.11100	.0206R	193.01	159.52	166.52	1.825685	26.49	
CHANNEALING LANGUA	.50000	.11100	.02068	193.01	159.52	166.52	1.825685	26.49	
CHANGING LANGUA	-50000	00111.	.02068	193.01	159.52	166.52	1.625565	26.49	
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			00020°	102.01	150.52	100.32	1.0053000	26.40	
CHANEIRE LANILLA	.50000	.11100	.02068	193.01	159.52	166.52		26.40	
STLEPLET DESCENT VALUES	50000	.11100	.02068	193.01	159.52	166.52	1.825685	26.49	
MEN PARAMETER VALUES	.50000	11100	.02068	193.01	159.52	166.52	1.825685	26.49	
CHANGING LANDA	50000	.11150	.02064	19.49	159.53	166.36	1.421065	26.61	
CHANGING LANGUA	.5000	.1111	.02067	1999	159.52	166.50	1.825411	26.50	
CHANELNE LANEDA	.50000	.11101	.0206P	19.461	159.52	166.52	1.825662	26.49	
CHANGING LAMBUA	.50000	.11100	.0205A	193.01	159.52	166.52	1.825683	26.49	
CHANGING LANUUA	.50000	.11100	.0206A	193.01	159.52	166.52	1.825685	26.49	
CHANGING LANDLA	.50000	.11100	.02068	193.01	159.52	166.52	1.825685	26.49	
CHANGLING LANDUA	.50000	.11100	.02058	193.01	159.52	166.52	1.825685	26.49	
CHANGING LANUUA	.50000	00111.	.02058	193.01	159.52	166.52	1.825685	26.49	
CHANDAN LANDUA CHANDAN SALULA	20000	untit.	.02068	193.01	159.52	166.52	1.825685	26.49	
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CHANNEINE LANNUA				10.01	20°671	100.02	C80C28•1	04 90	
Chanesto Chanesto	00000-		10020.		150.52	166.52	1. 1255600	24.40	
CHARGING LAMEUA	.50000	01111.	.0206A	19.551	159.52	166.52	1.425685		
CHANGING LAMULA	-50000	.11100	02068	193.01	159.52	166.52	1.825665	26.40	
CHANGING LAMUUA	.54000	.11100	.02048	193.01	159.52	166.52	1.825685	26.49	
CHANGING LAMUUA	• 500n0	.11100	.02040	193.01	159.52	166.52	1.825685	26.49	
STLEPLST DESCENT VALUES	.50000	.11100	.0206R	193.01	159.52	166.52	1.825685	26.49	
MEA PARAMETER VALUES	.50000	.11100	.02068	193.01	159.52	166.52	1.825685	26.49	
CHAMPANE LAMBUA	• 50000	.11150	.020ft	192.96	159.53	166.36	1.821065	26.61	
	50000°	.1111	.02067	10,.99	159.52	166.50	1.825411	26.50	
CHANGANG LAPLUA	00005.	10111.	.0206	193.01	159.52	166.52	1.825662	26.49	
	00005*		.02064	193.01	159.52	166.52	1.825683	26.49	
	00005.		.0206	1901	159.52	166.52	1.825685	26.49	
	000000	00111.	-02046	19.01	159.52	166.52	1.825685	26.49	
	00000.	00171.	46020.	193.01	1.9.02	166.32	1.825685	20.44	
ChidNuthias CANUDA	00004	01111	44020°	10.041	20.901	166.52	1. 822556	24.49	
Chimber de Laberda	50000			10.071	179.00	20.001	1000000 I		
Litutivality LAMULA	00005.	.11100	02068	19.001		166.52	1.90556.1	54.9C	
CHAPULNU LAMUUN	.5u000	.11100	.0206	193.01	159.52	166.52	1.825685	26.49	
CHATWELTU LAMBUR	. 5v0n0	.11100	. 0206P	19.691	159.52	166.52	1.825485	26.49	
CHANNELL LAPPLUM	.50010	.11100	.0∠05A	193.01	159.52	166.52	1.825685	26.49	
CHEHLLING CAPPUN	100NS.	.11100	,0205P	193.01	159.52	166.52	1.825685	26.49	

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The above sample output listing is given here only to illustrate the format as programmed in RELY I at the time of delivery. The values shown are of no significance relative to the content of the report (though they are those of a sample from Example IV, Sec. 3.3), nor are they expected to be repeatable exactly with implementation of RELY I at a different computer facility. The sample is offered as an aid to users, showing format and exemplary behavior in a given random case.

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