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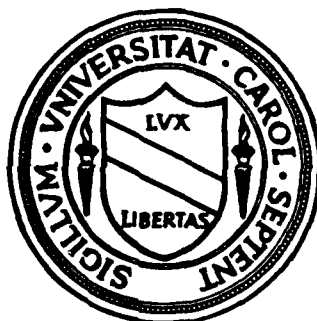
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A COUNTEREXAMPLE CONCERNING THE EXTREMAL INDEX

by

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Technical Report No. 237

August 1988

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2a. SECURITY CLASSIFICATION AUTHORITY <b>NA</b>		3. DISTRIBUTION/AVAILABILITY OF REPORT  Approved for Public Release; Distribution Unlimited	
2b. DECLASSIFICATION/DOWNGRADING SCHEDULE <b>NA</b>			
4. PERFORMING ORGANIZATION REPORT NUMBER(S) Technical Report No. 237		5. MONITORING ORGANIZATION REPORT NUMBER(S)  <del>AFOSR TR 83-1098</del>	
6a. NAME OF PERFORMING ORGANIZATION  University of North Carolina	6b. OFFICE SYMBOL (If applicable)	7a. NAME OF MONITORING ORGANIZATION  AFOSR/NM	
6c. ADDRESS (City, State and ZIP Code) Statistics Department CB #3260, Phillips Hall Chapel Hill, NC 27599-3260		7b. ADDRESS (City, State and ZIP Code) Bldg. 410 Bolling AFB, DC 20332-6448	
8a. NAME OF FUNDING/SPONSORING ORGANIZATION  AFOSR	8b. OFFICE SYMBOL (If applicable)  NM	9. PROCUREMENT INSTRUMENT IDENTIFICATION NUMBER  F49620 85 C 0144	
8c. ADDRESS (City, State and ZIP Code) Bldg. 410 Bolling AFB, DC		10. SOURCE OF FUNDING NOS.	
		PROGRAM ELEMENT NO. 6.1102F	PROJECT NO. 2304
11. TITLE (Include Security Classification) A counterexample concerning the extremal index			
12. PERSONAL AUTHOR(S) Smith, R.L.			
13a. TYPE OF REPORT preprint	13b. TIME COVERED FROM 9/1/87 TO 8/31/88	14. DATE OF REPORT (Yr., Mo., Day) 1988, August	15. PAGE COUNT 6
16. SUPPLEMENTARY NOTATION			
17. COSATI CODES		18. SUBJECT TERMS (Continue on reverse if necessary and identify by block number)  Key words and phrases: N/A	
FIELD	GROUP SUB. GR.		
XXXXXXXXXXXXXX			
19. ABSTRACT (Continue on reverse if necessary and identify by block number)  The concept of an extremal index, which is a measure of local dependence amongst the exceedances over a high threshold by a stationary sequence, has a natural interpretation as the reciprocal of mean cluster size. We exhibit a counterexample which shows that this interpretation is not necessarily correct.			
20. DISTRIBUTION/AVAILABILITY OF ABSTRACT  UNCLASSIFIED/UNLIMITED <input checked="" type="checkbox"/> SAME AS RPT. <input type="checkbox"/> DTIC USERS <input type="checkbox"/>		21. ABSTRACT SECURITY CLASSIFICATION  UNCLASSIFIED	
22a. NAME OF RESPONSIBLE INDIVIDUAL  Major Brian W. Woodruff		22b. TELEPHONE NUMBER (Include Area Code) (202) 767-5026	22c. OFFICE SYMBOL  AFOSR/NM

A COUNTEREXAMPLE CONCERNING THE EXTREMAL INDEX

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Summary

The concept of an extremal index, which is a measure of local dependence amongst the exceedances over a high threshold by a stationary sequence, has a natural interpretation as the reciprocal of mean cluster size. We exhibit a counterexample which shows that this interpretation is not necessarily correct.

December 1987

Supported by Air Force Office of Scientific Research;

Grant Number: F 49620 85C 0144

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Let  $\{\xi_n, n=0,1,2,\dots\}$  denote a stationary sequence and define  $M_n = \max(\xi_1, \dots, \xi_n)$ . Under suitable conditions it is possible to prove results of the form

$$nP(\xi_1 > u_n) \rightarrow \tau \iff P(M_n \leq u_n) \rightarrow e^{-\theta\tau} \quad (0 < \tau < \infty) \quad (1)$$

where  $0 \leq \theta \leq 1$ . The parameter  $\theta$  was termed the extremal index by Leadbetter (1983), though the concept had occurred earlier in papers of Newell (1964), Loynes (1965), O'Brien (1974a, 1974b) and Davis (1982). For a general overview of extremes in stationary sequences, see Leadbetter, Lindgren and Rootzén (1983).

It is possible to define an exceedance point process  $N_n$  on  $(0,1]$ , such that  $N_n(s,t]$  is the number of exceedances of the level  $u_n$  among  $\{\xi_r: ns < r \leq nt\}$ . Convergence of  $\{N_n\}$  as  $n \rightarrow \infty$  is studied by Hsing, Hüsler and Leadbetter (1988). One of their main results is that, if a limiting point process exists, then it must be compound Poisson. The atoms of this limiting process correspond to clusters of exceedances. Somewhat parallel results have also been obtained by Alpuim (1987).

A natural interpretation of  $\theta$  is that  $1/\theta$  is the main cluster size in the limiting point process. Hsing et al. were not, however, able to prove this without making additional assumptions. The following example shows that the result is false without such assumptions.

The example is a regenerative sequence of the form

$$\xi_n = \zeta_j \quad \text{for} \quad \sum_{i=0}^{j-1} N_i \leq n < \sum_{i=0}^j N_i \quad (j \geq 1) \quad (2)$$

where

- (i)  $\zeta_j, j \geq 1$  are independent with a common distribution function  $F$



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satisfying  $P(1) = 0$ ,

(ii) For  $j > 1$ , given  $N_1, \dots, N_{j-1}, \zeta_1, \dots, \zeta_j$  with  $m \leq \zeta_j < m+1$ , the probability of the event  $N_j = i$  is  $q_{mi}$ . Here  $\{q_{mi}, m \geq 1, i \geq 1\}$  is a sequence of probabilities with  $q_{mi} \geq 0, \sum_i q_{mi} = 1$  for each  $m$ .

In words, the process remains in state  $\zeta_i$  for a random number of time epochs determined by the probability distribution  $q_{mi}(i \geq 1)$  with  $m = [\zeta_i]$ , and then moves to a new state which is independently chosen from  $F$ .

Let  $p_m = P(m \leq \zeta_j < m+1)$ ,  $\mu_m = \sum_i i q_{mi}$  and suppose  $\mu = \sum p_m \mu_m < \infty$ . Then  $\mu$  is the mean recurrence time of the process. The process may be made stationary by a suitable choice of distribution of  $N_1$ . It may also be regarded as a function of a Harris chain, and may therefore be treated by extreme value arguments of O'Brien (1987) and Rootzén (1987).

Now let us specify  $\{q_{mi}\}$  to be

$$q_{mi} = \begin{cases} (m-1)/m, & i=1, \\ 1/m, & i=m+1, \\ 0, & \text{otherwise.} \end{cases} \quad (3)$$

Then  $\mu_m = 2$  for all  $m$ , and so  $\mu = 2$  also. Let  $(u_n, n \geq 1)$  be a sequence of thresholds such that  $n P(\zeta_1 > u_n) \rightarrow \tau, 0 < \tau < \infty$ .

**Proposition 1**  $\theta = \frac{1}{2}$ .

**Proof.** By Theorem 3.1 of Rootzén (1987), for  $\delta > 0$  and  $\delta' = \delta + 1/n$  we have

$$|P(M_n \leq x) - P(\zeta_1 \leq x)^{n/\mu}| \leq 2\delta' + P(|\nu_n/n - 1/2| \geq \delta) \quad (4)$$

where  $\nu_n$  is the number of regenerations up to time  $n$ . By choosing  $\delta = \delta_n \rightarrow 0$

appropriately and setting  $x = u_n$ , we deduce

$$P(M_n \leq x) \rightarrow e^{-\tau/2}$$

as required.

Proposition 2. The exceedance point process  $N_n$  converges to a simple Poisson process of intensity  $\frac{1}{2}$ , as  $n \rightarrow \infty$ .

Proof. This follows from Theorem 4.2 of Hsing, Hüsler and Leadbetter (1988). Positive recurrent Harris chains are strong mixing (cf. O'Brien 1987), and hence satisfy the mixing condition  $\Delta(u_n)$  used by Hsing et al. For a suitable sequence  $\{r_n\}$  satisfying  $r_n \rightarrow \infty$ ,  $r_n/n \rightarrow 0$ , let

$$\pi_n(j) = P\left(\sum_{i=1}^{r_n} X_{n,i} = j \mid \sum_{i=1}^{r_n} X_{n,i} > 0\right), \quad j=1,2,\dots \quad (5)$$

where  $X_{n,i}$  is 1 if  $\xi_i > u_n$ , 0 otherwise. Let  $\pi(j) = \lim_{n \rightarrow \infty} \pi_n(j)$  for  $j=1,2,\dots$

The theorem of Hsing, Hüsler and Leadbetter asserts that the point process  $N_n$  converges to a compound Poisson process with compounding distribution  $\pi(\cdot)$ .

However, under (3) it is easy to see that  $\pi(j) = 1$  for  $j=1$ , 0 for  $j>1$ . Hence the limiting process in this case is simple Poisson, with a mean cluster size of 1. This completes our description of the example.

From a statistical point of view, the most natural way to estimate the extremal index is via the point process of high-level exceedances. Such a procedure was in fact proposed by Smith (1984). The example here reveals a possible fallacy in this procedure, though it may not be possible to do much about it in practice.

Remarks.

1. What is going wrong is the lack of tightness of the sequence  $(q_{mi}, i \geq 1)$  as  $m \rightarrow \infty$ . In a similar way it is possible to construct an example of the same phenomenon for any  $\theta < 1$ . Hsing et al. show that the extremal index is in general given by  $\theta = \lim_n (\sum_j \pi_n(j))^{-1}$ , so what is at issue is whether

$$\sum_j \pi_n(j) \rightarrow \sum_j \pi(j) \text{ as } n \rightarrow \infty. \quad (6)$$

This is false for the example considered here.

2. It is possible to exhibit other kinds of extremal behaviour by taking other sequences  $(q_{mi})$ . For instance, if  $\mu_m \rightarrow \infty$  we very easily obtain an example of a process with extremal index 0. Taking this one step further, if  $q_{mi} \rightarrow 0$  as  $m \rightarrow \infty$  for each  $i$  but the distribution  $(q_{mi}, i \geq 1)$  converges under some renormalisation to the distribution of a continuous random variable as  $m \rightarrow \infty$  (example: take  $q_{mi} = 1/m$  for  $i=1, 2, \dots, m$ ), then the point process  $N_n$  does not converge but a suitably renormalised sequence converges to a compound Poisson point process with continuous compounding distribution. Such behaviour is admitted in the general theory of Hsing, Hüsler and Leadbetter, but they do not give any examples.

Acknowledgement. This work was carried out during a visit to the Center for Stochastic Processes, University of North Carolina. I am grateful to Ross Leadbetter, Tailen Hsing and Teresa Alpuim for helpful discussions.



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