- 1) Automatic Detection and Recognition of the First Arrival Phase of Seismic Event Signals Contaminated by Noise
- 2) The Curious Case of the Missing Explosion
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- 2) Eugene Herrin

Southern Methodist University Geophysical Laboratory Inst for the Study of Earth & Man Dallas, TX 75275

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of the RANK 2700 detector.

The second paper presents a scenario for a clandestine, decoupled nuclear explosion in southeastern New Mexico based upon data recorded at Lajitas and other stations at regional distances from the hypothetical clandestine test.

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Automatic Detection and Recognition of the First Arrival Phase of Seismic Event Signals Contaminated by Noise.

Kathleen Ann Alden

ABSTRACT

A historical prospective on the methods that have been used to automatically detect event—signals and pick first arrival times is developed. Then the data set used to characterize and test the detection techniques is described. Next the features which discriminate seismic signals effectively from the background noise are characterized.

Two automatic detection methods are investigated. (1) The AR (8)-spectral estimates of the signal and noise are used to develop the AR (p)-spectral estimate for a synthetic waveform.

(2) Several non-parametric tests are employed in a time domain detector to discriminate event signals from background.

The non-parametric detector chosen, RANK 2700, employs a modified rank sum test to locate the seismic event and pick its first arrival time. Errors in the automatic first arrival picks for 152 of the event traces in the data set are used to analyze the performance of the RANK 2700 detector.

INTRODUCTION

Since the advent of the computer, work has been done in many fields to develop techniques to harness the computer to do large tedious jobs more swiftly and with fewer fluctuations in performance than its human counterpart. One of these fields is the identification and classification of seismic signals from digital seismic event records.

There are thousands of seismic events occurring every day and hundreds of seismic stations. It is not practical to evaluate all these seismic records at the speed necessary to keep up with the influx of data. Thus, many studies have been conducted to implement various techniques to (1) determine first arrival times, (2) classify the seismic events correctly and (3) locate the origins of the seismic events.

It is becoming increasingly uneconomical to pick first arrivals by hand and a computer can be used to identify first arrivals more consistently than would be picked by hand on an oscillogram. At the same time, digital recorders are becoming more common on even low-cost seismic systems, and it can be expected that in the future computer techniques will become more attractive.

<u>Automatically Detect Events and</u> <u>Pick First Arrival Times</u>

Many scientists, including R.V. Allen (1978), K.R. Anderson (1978), R. Blandford (1983), K.S.Fu (1982), J.E. Gaby (1983), and H.H. Liu (1981) have tried to devise an effective automatic pattern recognition system for seismic signals during the last twenty years. Computer techniques for picking first events have yet to gain widespread acceptance (P.J. Hatherly, 1982). Effective results have been elusive due to the nature of the seismic signals and the methods used to model and predict the observed values.

Both statistical methods; i.e., maximum likelihood estimator, maximum entropy spectra, etc. (C.H. Chen, 1981), and structural methods; i.e., pattern recognition schemes that use shape features such as slope, radius of curvative, period and amplitude, (J.E. Gaby and K. Anderson, 1983), have been employed to characterize and identify seismic events.

Statistical classification algorithms can be grouped into one of two types, parametric or non-parametric. Parametric algorithms assume a particular class statistical distribution, commonly the normal distribution, and then estimate the parameters of that distribution, such as the mean and variance, to use in algorithm classification. Non-parametric algorithms make

no assumptions about the class distributions. Non-parametric techniques are sometimes termed robust because they work well for a wide variety of class distributions, if the class signatures (mean, variance, etc.) are reasonably distinct. Parametric techniques usually yield good results under the same conditions as the non-parametric techniques if the signatures of the classes are reasonably distinct, even if the assumed class distribution is invalid.

The most effective example of the statistical methods is 94% correct recognition of regional and teleseismic events using maximum entropy spectra and spectral ratio (energy>0.5Hz/energy<0.5Hz), (C.H. Chen, 1982). But even 94% is not a good enough performance record for an automatic pattern recognition system. A system cannot be considered automatic if it requires human watchdogs to monitor its progress and correct mistakes.

Part of the problem of correctly characterizing and classifying seismic signals is in obtaining an accurate first arrival time. There is always ambiguity associated with measuring the first arrival time from seismograms, whether it is done by seismologist or machine, since these signals of finite bandwidth are of unknown shape and contaminated by noise. Such ambiguity can be reduced by combining the processes of picking arrivals in an iterative fashion (C.H. Chen, 1982). Inaccurate first

arrivals distort the structural and statistical characterizations of a seismic event that are used for classification.

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Effective results have been elusive due to the nature of the seismic signal which contains events contaminated by noise. In many fields of signal processing: example, the development of speech recognition systems, the automatic analysis of signals requires the recognition of specific features in the signal. (Gaby and Anderson, Since more a priori knowledge exists about characteristic shapes of words. it is easier automatically identify a word and classify it correctly. The paucity of a priori information available on the morphologies (shapes) of seismic signals inhibits automatic pattern recognition (Gaby and Anderson, 1983).

The lack of a priori knowledge characterizing seismic signals has prompted the recent application of pattern recognition techniques to find and develop methods to discriminate and classify seismic signals. The methods used to find the broad characteristics of a seismic event include storing the trace in a binary tree structure and using affinity techniques, first developed for image processing, to combine small segments of the signal and store the signal within the tree structure at different levels of complexity (Gaby and Anderson, 1983). In other words, this method would use features such as period,

slope and amplitude to divide the event trace into segments and then try to associate those segments with the sequence of arrival phases that make up a seismic event.

Augmented transition networks (ATN), originally created to provide a formal environment to develop grammatical rules describing finite state grammers, have also been used to develop seismic signal structural characteristics interactively (Anderson, 1981).

All the techniqes for seismic pattern recognition schemes require that the first arrival time of the waveform be known. The seismologist uses the morphology (shape) of the seismic signal to identify the first arrival time correctly along with changes in period and amplitude. Determining the correct first arrival time is important in classifying the signal and locating the origin of the seismic event. For an automatic pattern recognition system to replace the seismologist it must identify first arrival times correctly and classify the event with better than the 94% current success rate of techniques previously attempted.

Automatic processes that have been employed to determine first arrival times rely on statistics to distinguish between two populations (signal and noise) occurring on a seismic trace. Since the density distribution functions for both the signal and noise are not known a priori they have traditionally been estimated

by Gaussian density distribution functions. The reason for this is that the statistics (i.e., the divergence and linear discriminant functions) for characterizing Gaussian (normal) distributions are well known and easily available for computation.

The standard statistics for a normal distribution have traditionally been used to characterize the signal and noise making it possible to determine first arrival times by comparing parameters from the two populations. Using the standard statistics for a normal distribution to characterize the noise a predictive model is constructed to predict future seismic noise values.

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There should be a failure of the observed data to match the predicted value at the first arrival because the arrival of the seismic signal is not predictable from the background noise. Robinson(1967) wrote subroutines to pass the proposed first arrival if there was a significant prediction error at the first arrival and for two terms after it. Significance was established by making comparisions with the prediction errors within the noise. A test, which predicts the values at the first arrival from the previous values using the technique of linear least-square prediction, was first used by Wadsworth et al(1953) to identify seismic events on the basis of a prediction failure. (P.J. Hatherly, 1982)

Most techniques to distinguish noise from the first

arrival signal use parametric statistics which assume the signal has a known distribution. This is an incorrect assumption for the background noise preceding a seismic event. In general, background noise is a non-stationary process due to seasonal changes and atmospheric variations, with an unknown distribution. Over a short time interval, less than 100 seconds, background noise can usually be considered stationary except in the case where it precedes a seismic event. When a seismic event occurs the mean of the signal often flucuates while the transient signal is being recorded by the seismic instrument.

Non-stationary processes have time changing means and/or variances. Stationary process implies that the mean, the variance and the autocovariances of the process are invariant under time translations. Thus the mean and variance are constant, and the autocovariances depend only on the lag time.

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Non-stationary time series have been modeled by several processes: (1) Harrison(1964) used an exponentially weighted moving average, EWMA, to forecast seasonal short term sales. (2) A modified autogressive moving average, ARMA, model with time varying coefficients of the form:

$$\sum_{k=0}^{p} b_{kt}^* e_{t-k} = \sum_{j=0}^{r} c_{jt}^* x_{t-j} + d_t , r = \max(p-1, q-m)$$

has been used by P. Whittle (1965), to model non-

stationary time series. Prediction formulas applied in the stationary case are shown by Niemi(1983) to also be valid in the non-stationary case. The effects of the non-stationarity on the estimation of the parameters of the underlying ARMA model are not shown to be significant. (3) The autoregressive integrated moving average, ARIMA, method, based on Gaussian stochastic processes was developed by Box and Jenkins (1970) to model homogeneous non-stationary time series.

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Homogeneous non-stationarity implies the changing mean can be described by a low order polynomial in time. However, the coefficients of the polynomial are not constant but vary with time. The observations are described by random stochastic trends (polynomials). Tintner (1940), Yaglom (1955), and Box and Jenkins (1976) argue that homogeneous non-stationary sequences can be transformed into stationary sequences by taking successive differences of the series.

In practice it is usually the first or second integral of a non-stationary process which is stationary. The ARIMA technique integrates the non-stationary time series until it is stationary and then models the resulting time series as an ARMA process. Under fairly general conditions the prediction interval for a future observation in an ARIMA scheme is robust with respect to symmetric non-normality of the error distribution,

Heuts (1981).

Non-parametric methods that make no assumptions about the density distribution of the time series are outlined by Kassam and Thomas(1975). Cobben(1982) outlines a non-parametric detector based on a sign test that discriminates between:

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$$H_0$$
: $x(t) = n(t)$ //data = noise//

 H_a : $x(t) = s(t) + n(t)$ //data = signal + noise//

included in Kassam and Thomas (1980). The sign test was extended by Cox (1955) to detect steps and ramps in the presence of additive noise; i.e., signals with sharp onset and gradual onset first arrivals. Cobben's method assumes a stationary signal which may be a valid assumption for a time window of less than 100 seconds, (C.H. Chen, 1978).

Although background noise may possibly be considered stationary, it cannot be assummed to be Gaussian. Various known and unknown processes: microseisms which are "non-Gaussian"; thermal noise due to current across resistors which is "Gaussian"; and seismometer noise which is "distributed as 1/f" contribute to form a non-Gaussian density distribution. Although, Heuts (1981), has shown that an ARIMA prediction model (based on an assumption of Gaussian and stochastic data) is robust with respect to symmetric non-normality of the error distribution.

distinguishing noise from the first arrival.

Part of the difficulty in automatically classifying seismic signals and locating their origin is a direct result of incorrect first arrival times picked by methods assuming an underlying Gaussian distribution for the signal. Since the seismic signals have an unknown underlying distribution a method for first arrival detection using non-parametric statistics should be developed in the hopes that it will improve the accuracy of the first arrival time.

Once the first arrival time can be swiftly and accurately located the morphological and statistical methods of classifying seismic signals can be used with greater accuracy since the correct sequence of events (or states) that characterize the seismic signal will be readily available. The first arrival on the seismic trace identified as the first observation is which is statistically different from the observation before. The point chosen is dependent on the signal to noise (S/N) ratio and the amplification of the seismic signal (P.J. Hatherly, 1982).

Some signal detection schemes used to record events in real time use the spectrum found using the Walsh (Goforth and Herrin, 1981) or Fourier (Blandford, 1983) transforms of the signal to distinguish the signal from the background noise. These methods while effective for

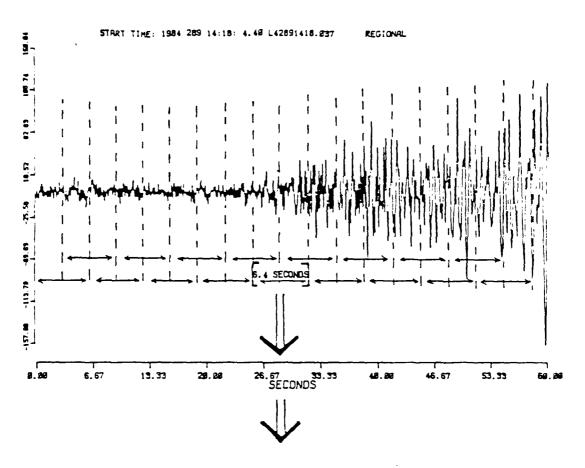
real time detection of events can only locate the first arrival time within one window (a block of observed data values transformed to provide a spectrum for comparision). Figure 1 illustrates the fast Walsh transform.

If this window is large enough to detect the first arrival (greater than two times the longest period of the expected signal, 2.5 seconds in our case) it will not be small enough to accurately determine first arrival times for the purpose of classifying and locating the origin of events. This indicates an effective automatic first arrival picker should be implemented in the time domain.

Methods Attempted in this Study

Since the signal is nonstationary with an unknown distribution function it is logical to develop a detector that does not assume stationarity; i.e., constant mean and constant variance, or a "known" distribution function. There is a broad range of amplitudes and bandwidths which characterize local, regional and teleseismic earthquakes and explosions. Thus, we must develop features to distinguish a broad category of signals from the background noise.

Since the seismic signal we record has an unknown distribution, it is important to determine how valid some of the parametric techniques; i.e., techniques that assume a know distribution, are when applied to a seismic signal.



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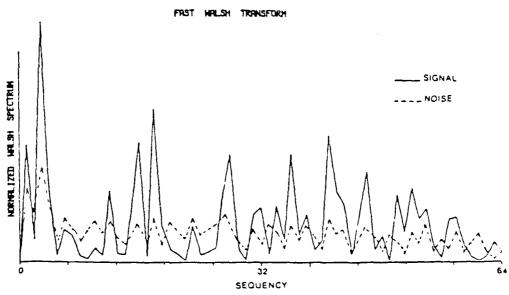


Figure 1. A figurative example of a 6.4 Second Detection Window and its Corresponding Normalized Walsh Spectrum Plotted on Top of the Normalized Walsh Spectrum for the Background Noise.

In particular, Niemi (1983) wrote a paper indicating an ARMA scheme was robust with respect to symmetric nonnormality. The discrete density distributions for the signal and noise were plotted for 100 seismic events recorded at each of the four seismic stations in far west Texas (Lajitas, Marathon, Shafter and Tres Cuevas) to find out if the density distributions of the seismic signals and noise from each station could be characterized as symmetric even though the "true" underlying density distributions were unknown. Histograms of the amplitude of the seismic signal were computed to approximate the discrete density functions for the noise and signal plus noise of each event trace. (See figures 2 through 4) In all cases the discrete density functions for both the signal and noise were symmetric. This assumption allows us to model the non-stationary seismic signal with a simplier model than the ARIMA model for which methods of computation have been more fully explored.

Two methods were attempted to develop an automatic first arrival picker. The first technique uses the AR-spectral estimation of the signal and the background noise to develop a synthetic waveform to cross-correlate with the event trace and pick out the first arrival phase. The second technique employs a non-parametric test within a sliding window detector to identify the seismic event signal and pick the first arrival phase.

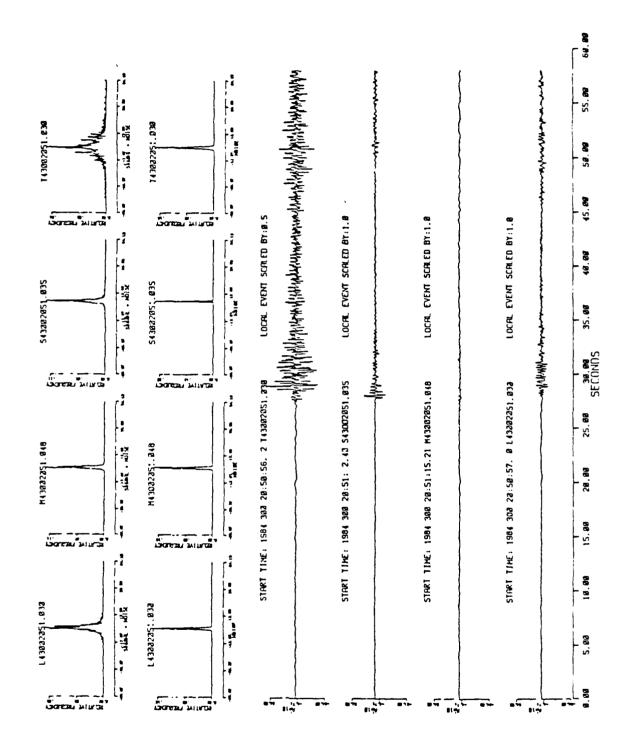
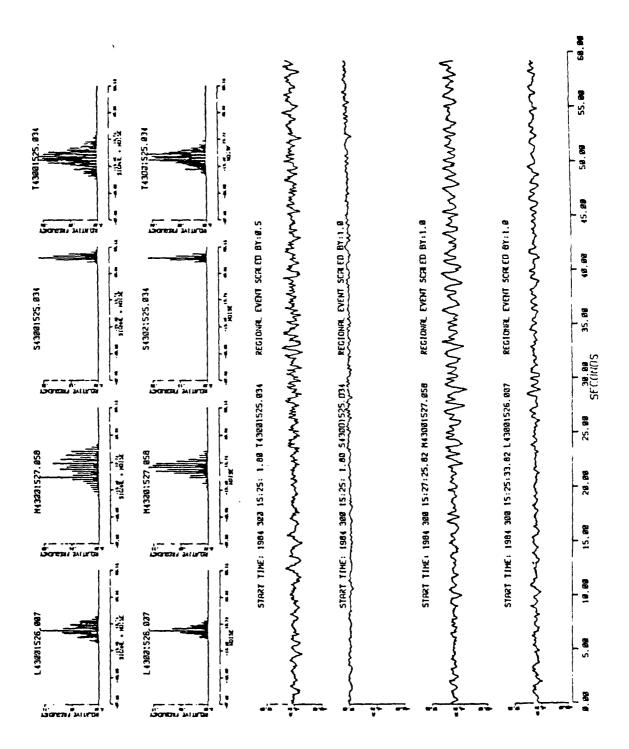
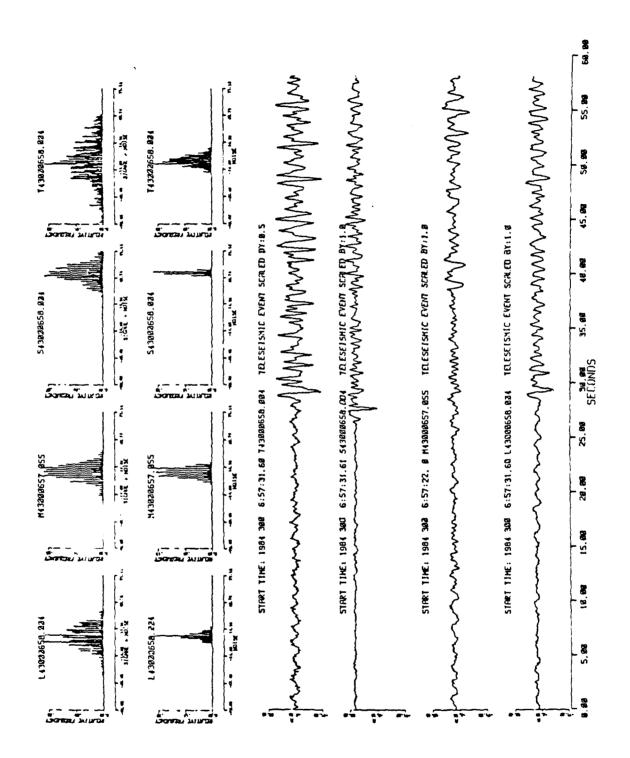


Figure 2. The First 60 Seconds for a Local Event, Recorded at Each of the Four Seismic Stations, Lajitas, Marathon, Shafter and Tres Cuevas, are Plotted from Bottom to Top Respectively. The Corresponding Discrete Density Functions for the First 20 Seconds of Background Noise and the Next 20 Seconds of Signal Plus Noise for Each Trace are Shown from Left to Right.



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Figure Regional З. The First 60 Seconds for Recorded Each of the Four Stations, Event, at Seismic Lajitas, Marathon, Shafter and Tres Cuevas, are Plotted Bottom Top Respectively. to The Corresponding ensity Functions for the First 20 Noise and the Next 20 Seconds of Density Seconds of Background Signal Plus Noise for Each Trace are Shown from Left to Right.



The First 60 Seconds for a Figure Each of the Four Seismic at Recorded Event, are Plotted Shafter and Tres Cuevas, Lajitas, Marathon, Corresponding The Bottom to Top Respectively. Density Functions for the First 20 Seconds Noise and the Next 20 Seconds of Signal Plus Background Noise for Each Trace are Shown from Left to Right.

DESCRIPTION OF THE DATA SET

Instruments Used to Record the Data

Seismic events recorded at four stations in southwest Texas, Lajitas, Maratnon, Shafter and Tres Cuevas, were collected during the Group of Scientific Experts Technical Test (GSETT) in which Southern Methodist University was a participant. GSETT was conducted from October 15, 1984 to December 14, 1984 to test procedures for exchanging seismic data, which include: the extraction of Level I signal parameters (seismic phase identifiers, arrival times, signal amplitudes and periods); the exchange of Global these parameters primarily via the Telecommunications System of the World Meteorological Organization (WMO/GTS); and the collection and assesment of these data at the Data Center in Washington, D.C.

Table 1 gives the locations and descriptions of the instruments used to record the seismic events during the test. The instrument responses are illustrated in figures 5 and 6.

A real time event detector utilizing the fast Walsh transform (Goforth and Herrin, 1981) implemented on a DEC RT/11 micro-computer received the seismic data from the Lajitas, Marathon, Shafter and Tres Cuevas stations via

TABLE 1. Locations and Descriptions of Instruments Used to Record the Event Data

SEISHOHETER DESCRIPTIONS

	NPME	LAJITAS	HARRTHON	SHAFTER	TRES CUEVAS	
	מו	ഥ	на	SH	TR	
	CHENNEL HUMBER	2	. 3	4	1	
	INSTRUMENT TYPE	23988	23960	23900	18388	
	LATITUDE DEG N	29.333	38.326	29.924	29.316	
	LONGITUDE DEG H	103.667	103.255	184.371	183.717	
	ELEVATION METERS	1813	1356	1578	1897	
	_ vca нz	962	1388	1700	2848	
SHAFTER TRES CLE	MARATHON:	5	5		15 \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \	

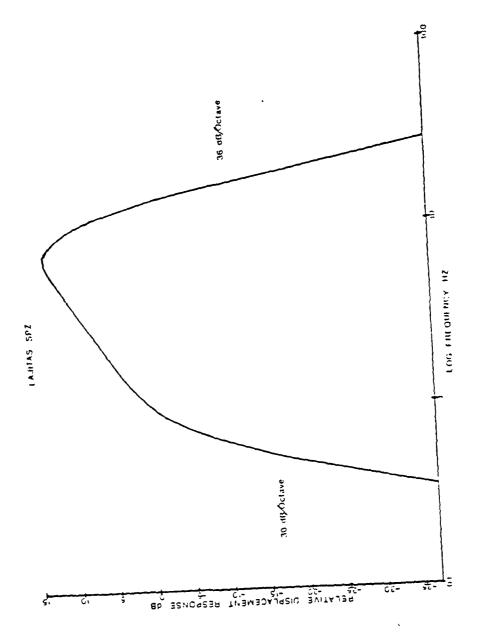
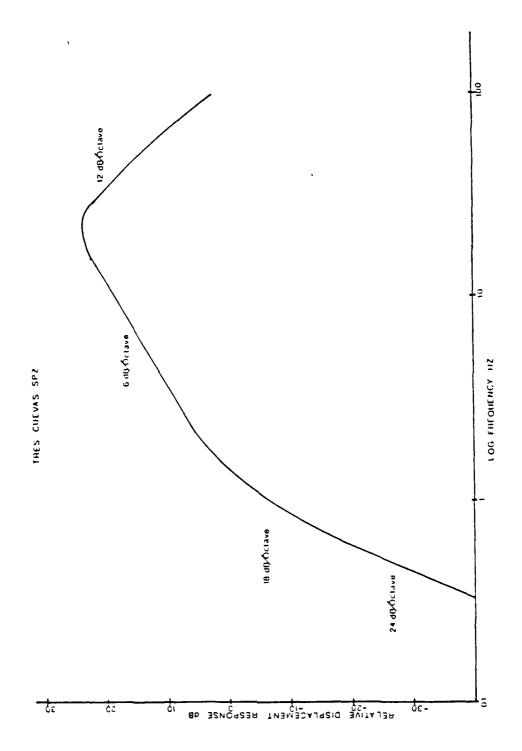


Figure 5. Instrument Response for Lajitas, Marathon and Shafter Recording Instruments



Instrument Response for the Tres Cuevas Figure 6. In Recording Instrument,

telephone line. (See appendix A for a description of the fast Walsh transform detector.) The detected events were recorded on magnetic tape for later graphic display and analysis by a seismologist.

The seismic analyst was responsible for identifying phase (Pn,Pg,P or Lg) and picking first arrival times, amplitudes and periods. The bias of the analyst directly effects the choice of local events v.s. "noise". Discrimination of the phase of the first arrival (P or Lg) is dependent on which phase the analyst's experience indicates the waveform matches. Parameters for the events chosen by the analyst as "true" events, not "glitches" or "background noise", were sent to the U.S. National Data center in Washington using the UNIX-net software of the Eunice operating system.

CONTRACT PROPERTY

GSETT's Final Event List associated all the event information collected from throughout the world and reported the origin time, location, magnitude and number of stations used to define each event which was detected during the two month time period of the test.

The residual travel time (the difference between the arrival time at the station indicated by the best least squares fit of the stations associated with an event and the first arrival time picked by the analyst) was listed for each station used to define an event. The "true" estimates calculated by least squares, assuming the

correct location of the event was found, allow us to determine if the automatic detection method is doing as well, worse, or better than the human analyst in ambiguous cases.

Methods Used to Detect and Record Events in the Event Files

The fast Walsh transform detection window is seconds long with a 50% overlap. (see figure 1) When the spectra of the detection window exceeds the background noise threshold the detection is triggered. seconds of data preceding the detection window is written at the beginning of the event file so that if the Walsh detector triggered on an Lg arrival the P wave first arrival should be present somewhere in the first 30 Thirty seconds was originally seconds of the event file. thought to be adequate to make sure the P arrival was included in the event trace but experience in working with the data collected during the GSETT experiment has shown that 60 seconds is a better predetection interval for insuring the first arrival (Pn) is included in the event file when the detector triggers on Pg.

EVENT SIGNAL CLASSIFICATION

<u>Variation in the Signals to be Distinguished</u> <u>From the Background Noise</u>

223°22327°E25556°E

The collection of seismic events we are attempting to distinguish from the background noise, (recorded on short period seismic instruments with a bandwidth of 0-20 Hz), encompasses a broad range of periods and amplitudes. Local events, occurring within 25 to 150 kilometers of the seismic station, are characterized by amplitudes which may be nearly the same as the noise and by periods of 0.3 to 0.4 seconds. Regional events, occuring on the same continent as the seismic station, are characterized by varying amplitudes and periods of 0.5 to 0.7 seconds. Teleseismic events, events that travel distances greater than about 2000 km to reach the seismic recording station, are characterized by periods of 0.8 to 1.2 seconds. The high frequency component of the background noise similar in spectral content to local seismic event signals while the underlying low frequency component of noise is similar in spectral content background teleseismic signals. Pn can often not be discriminated from the background noise by the analyst, thus Lg which follows Pn is often picked as the first arrival phase for

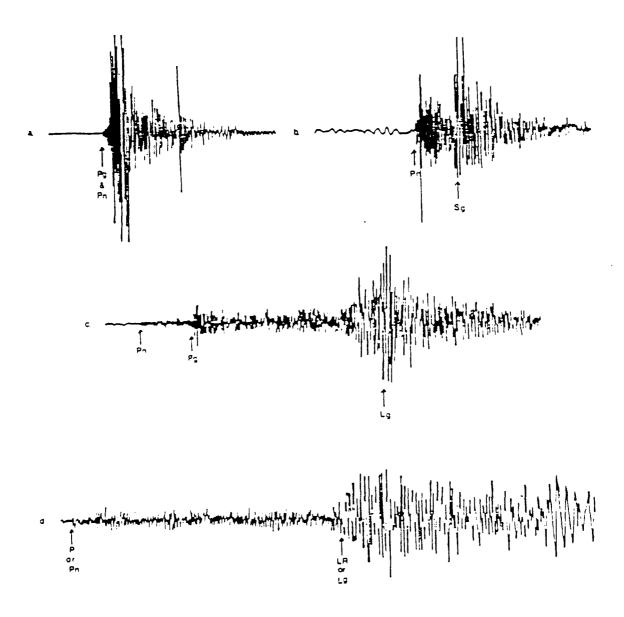




Figure 7. Examples of the Various Types of Seismic Signals Typically Encountered. (a) 33 km - Local Event -SPZ recorded in Berkeley, California. (b) 140 km - Local Event - Closest distance Pn and Pg can be observed separately - SPZ recorded in Colorado. (c) 60 - Regional Earthquake - SPZ recorded in Winner, South Dakota. 130 - Regional to Teleseismic Earthquake California - SPZ recorded in Colorado. (e) Teleseismic Earthquake - Offshore Peru - SPZ recorded Black Rapids, Alaska.

regional events. Figure 7 illustrates typical examples of the types of seismic signals encountered. (R.B. Simon, 1981)

Appropriateness of the Data Set for Testing the Effectiveness of Any Automatic Techniques Devised

The data set collected during GSETT represents collection of event records of different types of events with continuous background noise variations from four different seismic stations. This data set has variations in: signal type (local, regional and teleseismic); background noise (due to changes in wind temperature, humidity, barometric pressure and cultural activity); instrument type and local structure at the seismic station. It is concluded that because the data set includes variations in all the parameters normally expected to change when recording seismic events the data set will be a good test of the robustness of any detection method used to pick first arrival times.

Indentify the Characteristics of the Seismic Signals Which Define a Class Distinctive From the Background Noise

Any classification technique actually consists of a one-time calculation of decision boundaries, followed by a comparison of each sample's feature vector and the location of those boundaries. By incorporating these

facts in classification algorithms, their efficiency can be improved by an order of magnitude or more, with little or no reduction in accuracy. (Schowengerdt, R.A., 1983)

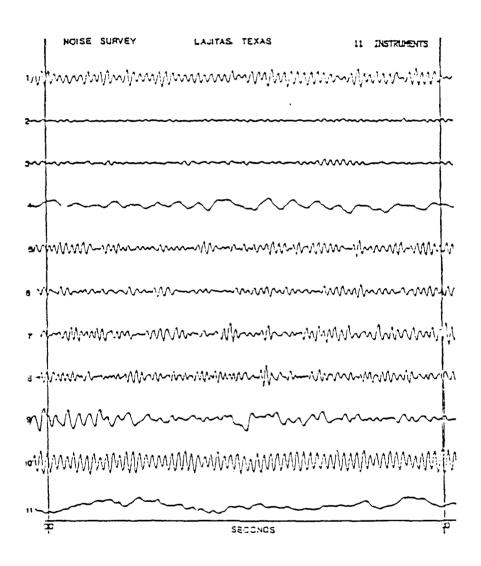
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Easily calculatable features available in the time domain for discriminating seismic signals and background noise are: period estimates, (intervals between zerocrossings and slope direction changes); amplitude; and slope. Figure 8 is a typical example of the background This noise data was recorded during a background noise study February 8, 1985 at the Lajitas test site. the next three sections we determine the effectiveness these features in distinguishing a broad class of seismic signals from the background noise by using: (1) a threshold detector based on an estimate of the mean and variance of the background noise; (2) segmentation of the signal based on the similarity of features to find a pattern; and (3) division of the features based on their ranks into rank quartiles to look for a relative pattern in the feature distribution.

A Threshold Detector to Identify Features

For the threshold detector the mean and variance were calculated for the amplitude and slope (average difference between the ith sample and the samples adjacent to it) of the first 100 observations representating background noise assuming a normal distribution. The threshold for signal

Figure 8. Background Noise Recorded on 11 Different Instruments During a Noise Survey at the Lajitas Test Site, February 8, 1985. Instrument 1 - GS 13 Z (vault), Instrument 2 - GS13N (vault), Instrument 3 - GS13 E (vault), Instrument 4 - 18300 (mine site), Instrument 5 - S750 Z (grouted in), Instrument 6 - S750 Z (vault), Instrument 7 - S750 A (North of mud hut 150m), Instrument 8 - S750A (East of mud hut 150m), Instrument 9 - GS21 (buried 330'), Instrument 10 - GS21 (buried 50'), Instrument 11 - microphones (wind).



detection was set at two standard deviations from estimated mean of the background noise. Thus, given that background noise is all that is present in the seismic signal, there is a 95% probability that the amplitude and slope for any observation will be within two standard deviations of its estimated mean. Figures 9 through illustrate the results of this test on three types seismic events recorded at the four stations Lajitas, Marathon, Shafter and Tres Cuevas. The state table for the color changes occuring along the event traces in figures 9 through 11 is shown in Table 2. The mean and variance estimated for the slope and amplitude were not very effective in clearly distinguishing the background noise from the seismic event signal.

Segmentation Using Affinity Techniques to Identify Features

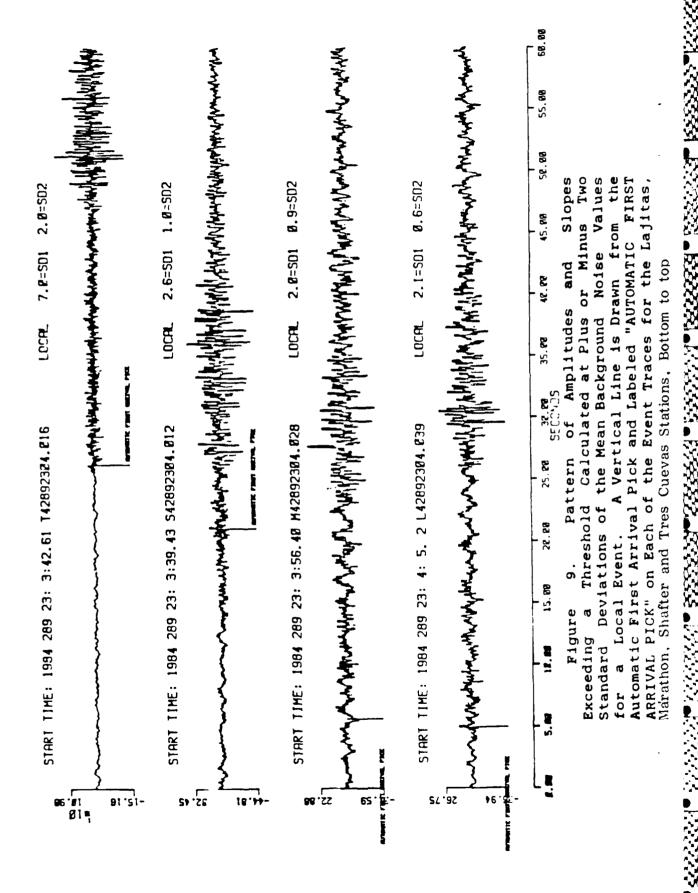
Segmentation of the trace using the estimated mean and variance of the slope and amplitude of the background noise was done using the nearest neighbor decision rule and an affinity algorithm to combine the samples with the most similar features into a finite number of segments. The nearest neighbor decision rule decides which of a segments neighbors is most similar. Then the affinity algorithm combines those segments which mutually consider each other most similar. Figures 12 through 14 illustrate this technique on the three types of seismic signals,

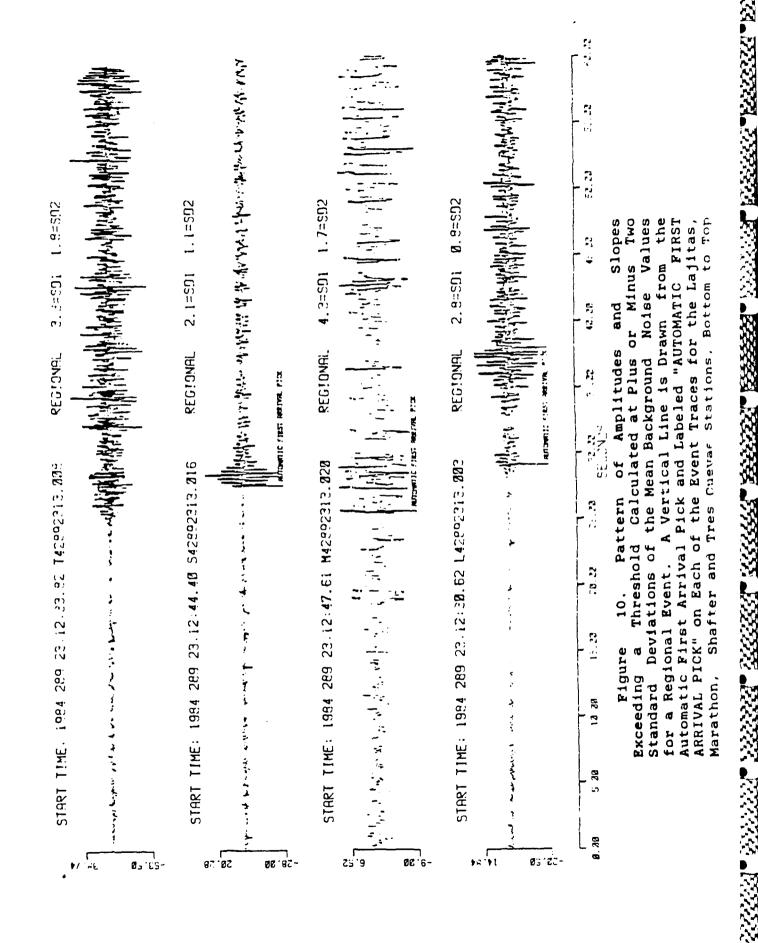
TABLE 2

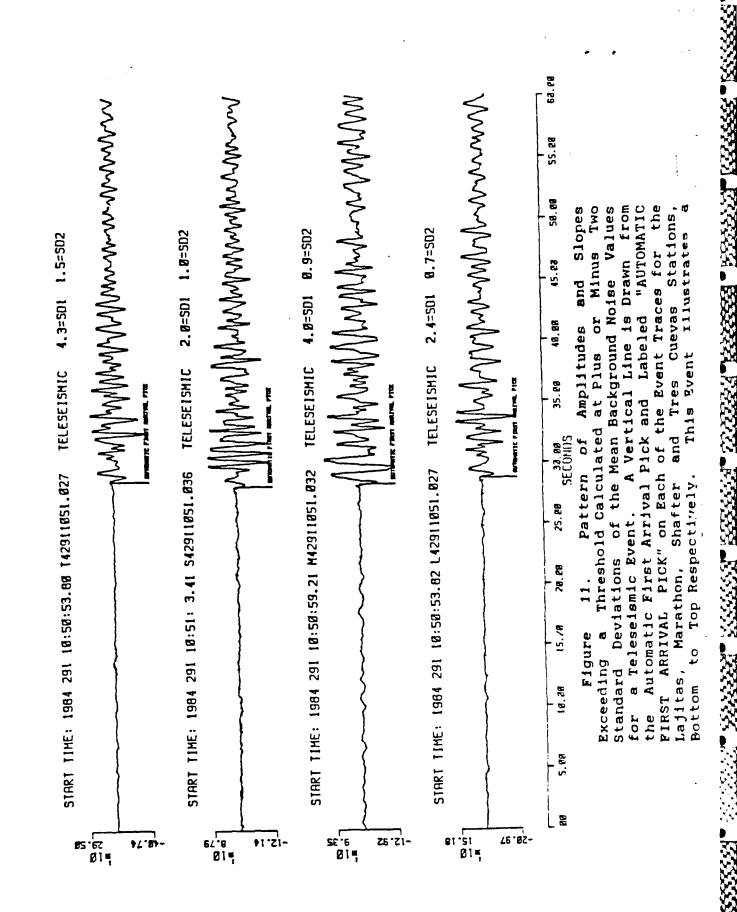
State Table for Color Changes Used to Graphically Depict the Results of the Threshold Detector

	PMPLITUDE	SLOPE
BELON THRESHOLD	8	8
ABOVE THRESHOLD	1	1

COLCR	STATE (AMPLITUDE, SLOPE)		
RED	Ø8		
BLUE	18		
GREEN	81		
BLACK	11		







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Recorded Segments Affinity and Tres Observations for a Local Event, at the Four Stations, Lajitas, Marathon, Shafter and of Combined into a Finite Number Nearest Neighbor Decision Rule £14 (4.0) 西部 別をみ Figure 12. 19. 19. are 4. Using the Techniques. भूगा अभू 1/4 JUE ILB

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Recorded at the Four Stations, Lajitas, Marathon, Shafter Regional Event, Cuevas, are Combined into a Finite Number Rule Using the Nearest Neighbor Decision Observations for a Affinity Techniques. Figure 13. and Tres

START THE: 1984 290 0:19:17. 1 T429,0019.051 TELEBEISHID EVBAT SCALED BY 0.5 المرسي المرك ومرس وي مواده ومدس الموارا المرايا المرايا المرسعم مسترس مرس مرس مسترس مرسوس من المرايات - かんしょうしょう

START TIME. 1984 290 0.20.49.81 S4290021.023 TELESEISMIC EVENT SCREED EVILLE

START TIME: 1964 290 0:19:20.21 M42900019.053 TELEBETSMIC EVENT SCREED BY:1.0

START TIME: 1984 290 0:19:17. 1 L429CO19.CG1 TELESEISHIC EVB4T SCHLED BY-1.0 हिला हिला हिला केला देन केला देन केला हिला हिला हिला हिला हिला हिला हिला Lajitas, Marathon, Shafter Observations for a Teleseismic Event, Cuevas, are Combined into a Finite Number the Nearest Neighbor Decision Recorded at the Four Stations, Figure 14. Tres and

affinity Techniques.

Using

local, regional and teleseismic. The colors represent different segments. No clear pattern was identified in the segmentation to make this a useful method for distinguishing signal from noise.

D. 25.

Both the threshold detection and segmentation methods made the assumption that the noise was Gaussian and stationary when the mean and the variance were estimated. Since this is untrue, the noise could not be effectively characterized by the estimated mean and variance and overlap occured between the boundaries defining the signal and noise classes. The subroutine, "NEIGHBOR", listed in appendix B uses the nearest neighbor rule and affinity techniques to combine the segments.

Using Rank Quartiles to Identify Features

The next attempt to characterize the signal and noise categories was to rank the amplitudes and periods, estimated by the distance between zero crossings, of the event trace. Then the ranks were divided into quartiles and color coded on the event trace so any patterns in the features associated with the signal or noise might be identified. Table 3 illustrates the color code used in figures 15 through 19 to identify which rank quartile an observation belongs.

This technique showed a clear pattern for ranked amplitudes of events with large S/N ratios but was

TABLE 3

State Table for Color Changes Used to Graphically Depict the Rank Quartiles

AMPLITUDE RANK QUARTILES

P		let	2nd	3rd_	4Lh
E R I O	1et	i	2	3	4
R A N K	2nd	5	6	7	8
U R R	3 rd	9	19	11	12
T I L E S	4Lh	19	14	15	8

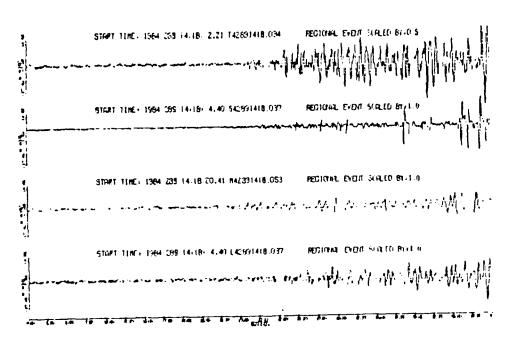


Figure 15.b

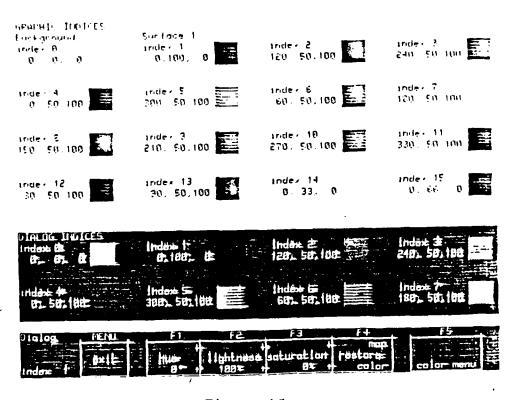


Figure 15.a

Figure 15. The Ranks of Amplitudes and Periods are Divided into Rank Quartiles. The Color Changes in the Event Traces, Recorded at the Stations Lajitas, Marathon, Shafter and Tres Cuevas, Bottom to Top Respectively, Graphically Depict any Patterns that Exist in the Features. (a) The Color Index Corresponding to the 16 Rank Quartiles the Ranks of the Periods and Amplitudes of Each Trace are Divided into as Shown in Table 3. (b) Shows the Rank Quartile Pattern for a Regional Event.

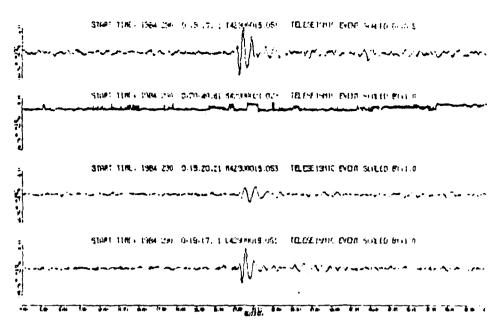


Figure 16.b

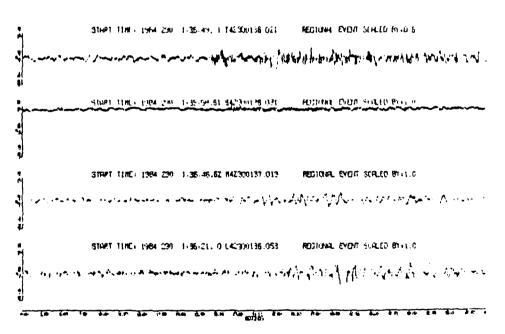


Figure 16.a

Figure 16. The Ranks of Amplitudes and Periods are Divided into Rank Quartiles. The Color Changes in the Event Traces, Recorded at the Stations Lajitas, Marathon, Shafter and Tres Cuevas, Bottom to Top Respectively, Graphically Depict any Patterns that Exist in the Features. (a) Illustrates the Rank Quartile Pattern for a Regional Event. (b) Illustrates the Rank Quartile Pattern for a Regional Event.

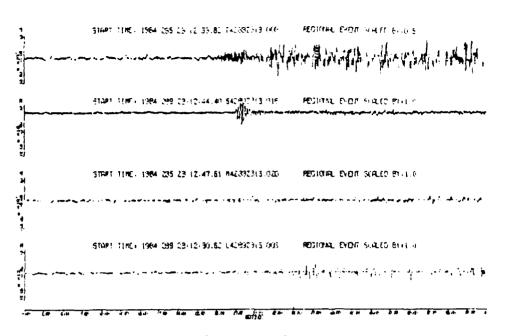


Figure 17.b

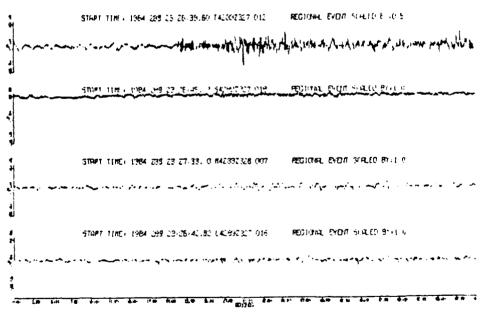


Figure 17.a

Figure 17. The Ranks of Amplitudes and Periods are Divided into Rank Quartiles. The Color Changes in the Event Traces, Recorded at the Stations Lajitas, Marathon, Shafter and Tres Cuevas, Bottom to Top Respectively, Graphically Depict any Patterns that Exist in the Features. (a) Illustrates the Rank Quartile Pattern for a Regional Event. (b) Illustrates the Rank Quartile Pattern for a Regional Event.

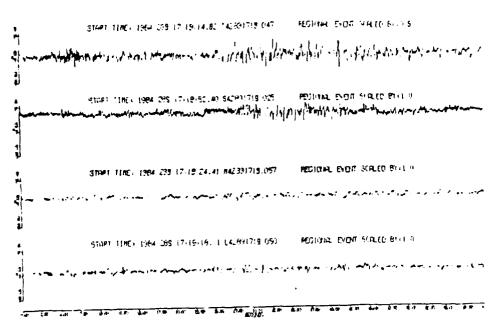


Figure 18.b

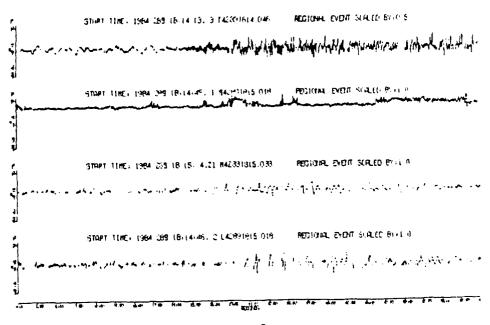


Figure 18.a

The Ranks of Amplitudes and Periods are Figure 18. Divided into Rank Quartiles. The Color Changes in Event Traces, Recorded at the Stations Lajitas, Marathon, Shafter and Tres Cuevas, Bottom to Top Respectively, Graphically Depict any Patterns that Exist in Features. (a) Illustrates the Rank Quartile Pattern for a Regional Event. (b) Illustrates the Rank Quartile Pattern for a Regional Event.

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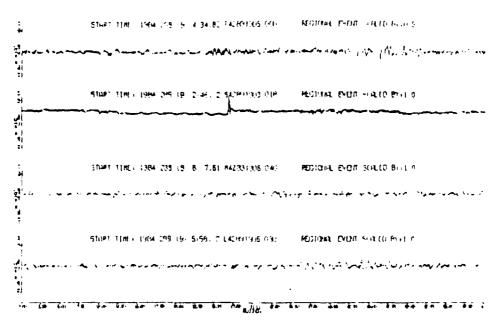


Figure 19.b

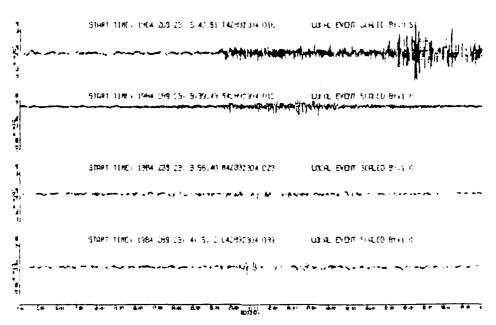


Figure 19.a

Figure 19. The Ranks of Amplitudes and Periods are Divided into Rank Quartiles. The Color Changes in the Event Traces, Recorded at the Stations Lajitas, Marathon, Shafter and Tres Cuevas, Bottom to Top Respectively, Graphically Depict any Patterns that Exist in the Features. (a) Illustrates the Rank Quartile Pattern for a Local Event. (b) Illustrates the Rank Quartile Pattern for a Regional Event.

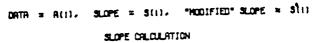
ambiguous for signals buried in noise. The rank pattern for estimated periods showed no strong pattern for distinguishing signal from noise. However, the ranks for the amplitude and estimated periods, which were added together and divided into quartiles, was the the best indicator of a clear pattern in the features to be used to discriminate signal and noise. This is reasonable since the criteria the human analyst uses to distinguish signal from noise especially in ambiguous cases with low signal to noise ratios is based on changes in both period and amplitude from the preceding background. Slope is sensitive to changes in both amplitude and period which makes it the most favorable feature to use in creating an automatic first arrival detection algorithm.

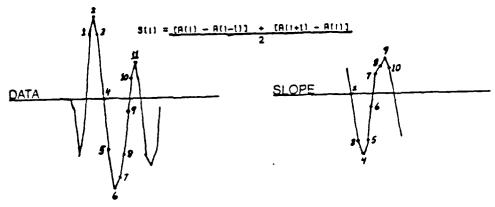
Conclusions

Since parametric techniques appear to allow too much overlap in the classifications to clearly distinguish signal from noise, a non-parametric technique using the slope of the event trace appears to describe the detection method with the best chance of discriminating background noise from a broad class of seismic event signals.

The feature with the least amount of overlap between background noise and seismic event signals, "modified" slope, was calculated through discrete integration of the event trace slope over segments with the same slope

Figure 20 demonstrates this calculation. direction. Various period estimations of the background noise were too similar to one or more of the local, regional or teleseismic signal classes for effective discrimination between signal and noise. Amplitude was discriminator when the S/N ratio was high. However, the detectors failed to trigger on the first arrival for events with low S/N ratios. Both the slope and "modified" slope of the event trace are characterized by a constant mean equal to zero and a variance which changes with time. This makes the "modified" slope of the event trace a more attractive feature for distinguishing between signal and noise than the unmodified event trace whose mean and variance both change with time.





DATA	SLOPE	"MODIFIED" SLOPE
R(1) = 4.	-	
A(2) = 6.	S(2) = 9.	S(2) = 8.
A(3) = 4.	S(3) = -3.	5 13) = -3.
A(4) = 8.	S(4) = -3.75	51(4) = -6.75
A(5) = -3.5	3(5) = -3.	\$15) = -9.75
A(6) = -6	S(6) = -0.75	5161 =-18.5
A(7) = -5	\$(7) = 1.25	\$771 = 1.25
A(8) = -3.5	3(8) = 1.75	5 18) = 3.
R(9) = -1.5	\$(9) = 2.25	\$19) = 5.25
A(18)= 1.	\$(18)= 1.75	\$(18)= 7.
R(11)= 2.	-	-

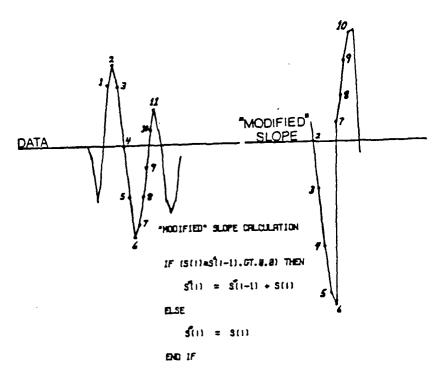


Figure 20. Examples to Graphically Illustrate the Slope and "Modified" Slope Calculations for the Slope and "Modified" Slope Features used to Distinguish Event Signals from Background Noise.

DETECTORS

AR-Spectral Estimation Used to Develop a Synthetic Waveform

Assumptions

The autoregressive (AR) model is correct for modeling minimum phase stationary processes. To estimate the spectral density function we must assume that we can obtain an "adequate approximation" by using a finite order AR model (Priestley, The development of 1981). synthetic waveform to pick out the first arrival phase event trace from the autoregressive spectral estimate that same event trace requires several assumptions. First, in order to use this technique, we must assume the first arrival phase of the seismic signal generated by an earthquake or explosion is minimum phase. Second we assume the corner frequency, or dominate period, of the seismic signal has the same period as the first arrival phase. Third we assume that the background noise does not contain an obvious corner frequency. Forth we assume the seismic signal is stationary or at least can be estimated by a process that assumes the driving function is white noise. Seismic signals contaminated by noise are nonstationary processes with unknown distributions.

Niemi, (1983), studied the effects of nonstationary noise on ARMA models and concluded that the prediction formulas applied in the stationary case were also valid in the nonstationary case provided the underlying density distribution of the nonstationary process was symmetric. Discrete density distributions for 400 seismic signals and their associated background noise were plotted and in all cases the density distributions were symmetric. This allows us to conclude that an AR model is appropriate for estimating the spectra of the event trace given our initial assumptions.

Description of Method

The first step in developing a synthetic waveform was to compute the coefficients for an eighth order AR model using the Burg algorithm. See appendix B for the subroutines, "Polr_ArBURG" and "AR_SPECTRA", to compute the Burg coefficients and the AR-spectra. Coefficients were computed for the first 800 samples of the event file representing the background noise and the next 800 samples of the event file thought to contain the first arrival phase. Then the AR(8) spectra, $\hat{p}(f)$, was computed for both the signal and the noise and plotted (Priestley, 1981).

$$\hat{p}(f) = \frac{2\hat{\sigma}^2}{\left|1 - \hat{\theta}_1 e^{-i2\pi f} - \cdots - \hat{\theta}_8 e^{-i16\pi f}\right|^2}$$

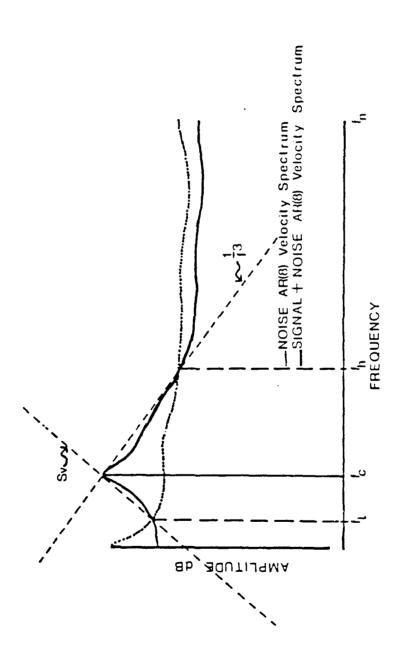
$$(\hat{\theta}_1 \dots \hat{\theta}_8) = AR(8) \text{ coefficients}$$

$$\hat{\sigma}^2 = \text{variance of the white noise}$$

$$f = \text{frequency Hz}$$

To determine the shape of the synthetic waveform's AR spectral density we employ the doctrine parsimony. The minimum number of poles needed to describe the spectral shape with the information we have available is three. The poles are located at the low cut-off, high cut-off and corner frequencies of the signal bandwidth. The corner frequency is assumed to be the frequency with the signal spectra's peak amplitude. The low and high frequency cut-offs define the signal spectra interval where the signal spectra is consistently above the noise spectra. The amplitude of the low frequency cut-off is determined by the roll-off below the corner frequency due to the sensivity of the recording instrument, approximately 6dB per octave for the Lajitas station instrument. The amplitude of the high frequency cut-off is determined by the roll-off above the corner frequency due to attenuation. The expected roll-off is -3dB to -4dB per octave. Figure 21 illustrates the proposed method.

The coefficients of the synthetic AR(p) model are derived from the synthetic AR power spectral density,



the Proposed AR-Spectral Method for Creating 1/f3 Denotes the Sv_Denotes the Spectrum for a Synthetic Waveform. Sensitivity of the Recording Instrument. 1 Roll-off with Prequency Due to Attenuation. Figure 21.

COLLINE RESERVED PRESERVED TO THE PERSON PRESERVED BESSELVED TO THE PERSON OF THE PERS

 $\hat{p}(f)$, described in the preceding paragraph. The AR theoretical power spectral density, p(f), is defined by

$$p(f) = \left| S(f) \right|^2$$

where

$$\left| S(f) \right|^2 = S(f)S^*(f)$$

and

$$S(f) = \frac{\sqrt{2\sigma}}{\left[1 - \theta_1 e^{-i2\pi f} - \cdots - \theta_p e^{-i2\pi fp}\right]}$$

Now by making use of the fact that the theoretical power spectral density, p(f), can also be written in terms of the autocovariance, R(k),

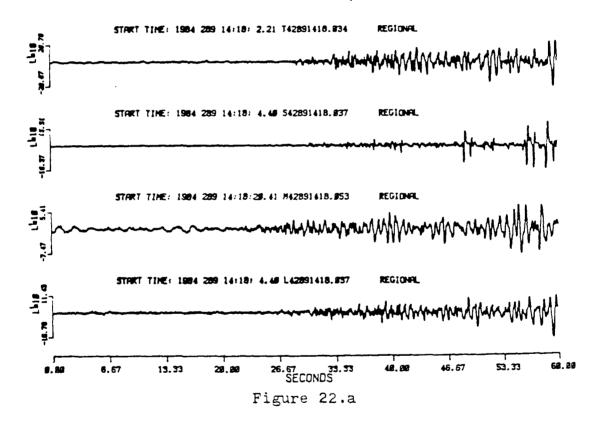
$$p(f) = \int_{0}^{\infty} e^{-i2\pi f k} R(k) dk$$

we can obtain the autocovariance, R(k), of the synthetic AR model by taking the inverse fast fourier transform (IFFT) of the synthetic AR power spectral density, p(f). $p(f) = \int_{-\infty}^{\infty} e^{-i2\pi fk} R(k) dk$ IFFT $\int_{-\infty}^{\infty} x(k) x(k-t) dt = R(k)$ Then the autocovariance, R(k), of the synthetic process, x(t), is used by the Yule-Walker algorithm (Kay and Marple, 1981), (see subroutine "AKAIKE" in appendix B), to obtain the AR(p) coefficients for an AR model of order p(t) (determined by Akaike's information criteria ,AIC, Priestley, 1981). This synthetic AR(p) model is then convolved with a spike plus white noise to create the synthetic waveform. The synthetic waveform is then cross-correlated with the event trace to try and pick out the first arrival phase.

Effectiveness of Method and Conclusions

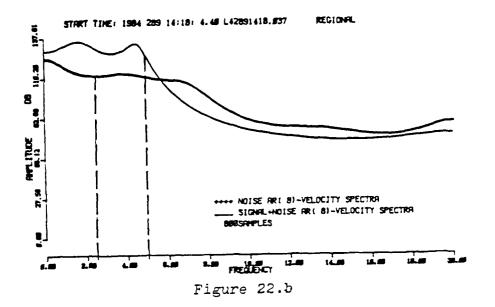
The method hypothesized above started to break down with the assumptions that the first arrival phase has the same period as the corner frequency and the background noise had no obvious corner frequency. The period estimated by the analyst for the first arrival phase of the Lajitas event traces only correlated well with the corner frequencies of the corresponding AR(8) spectral estimates in 25% of the events studied; i.e., teleseismic events with a high signal to noise (S/N) ratio.

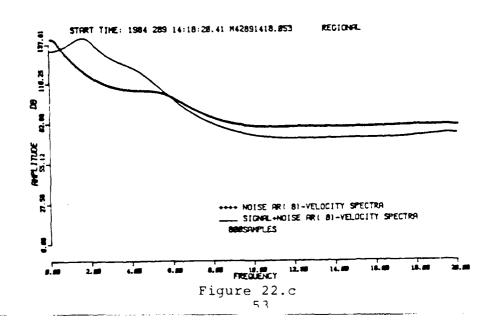
often a corner appeared in the background noise making it difficult to determine the correct spectral bandwidth for the signal automatically. For the above reasons this method for identifying the first arrival phase for a broad category of seismic events was abandoned in favor of a more robust non-parametric technique, the rank sum detector. Representative examples of the AR(8) spectral estimations for several of the events studied are shown in figures 22 through 26 with the analyst's estimations of the dominate periods for the noise and the first arrival phase marked for the Lajitas seismic events.



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AR(8)-Spectral Estimates for the Signal 22. (a) Event traces recorded at each and Noise for Event 1. Lajitas, Marathon, Shafter and Tres of the four stations, (b) Lajitas AR(8)-Spectral for Event 1. Cuevas, The Analyst Estimates for the Dominate Period Estimates. of the Noise and the First Arrival Phase with Dashed Lines Drawn to their Respective (c) Marathon AR(8)-Spectral Estimates. Estimates. Shafter AR(8) Spectral Estimates. (e) Tres Cuevas AR(8) Spectral Estimates.





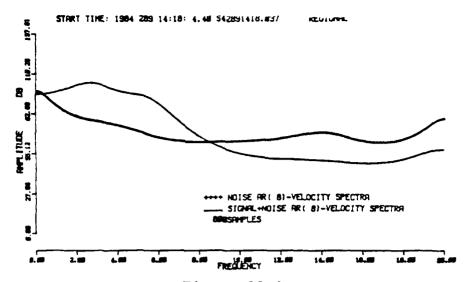


Figure 22.d

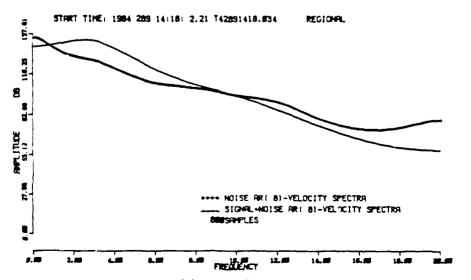


Figure 22.e 54

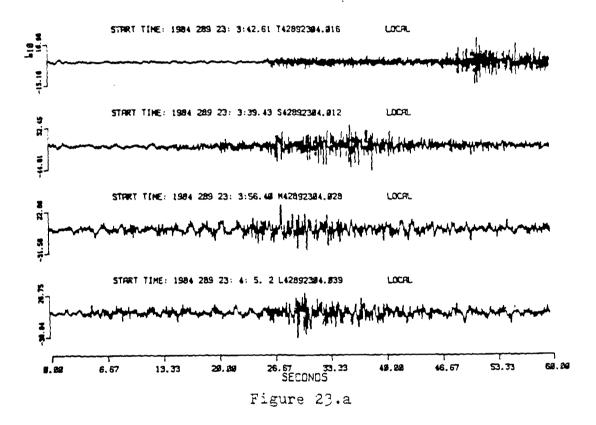
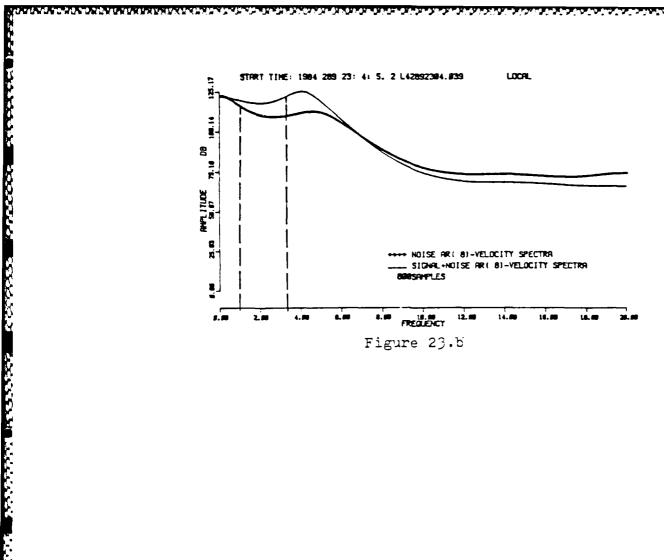
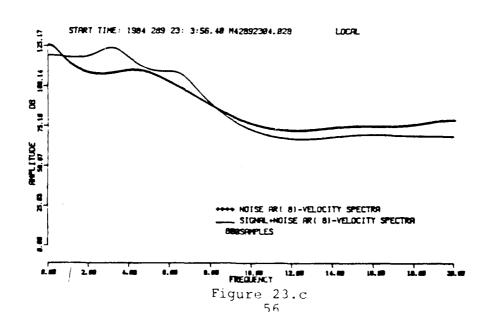
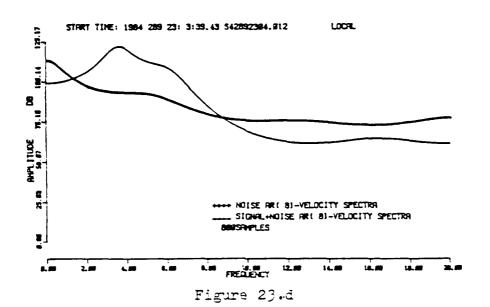


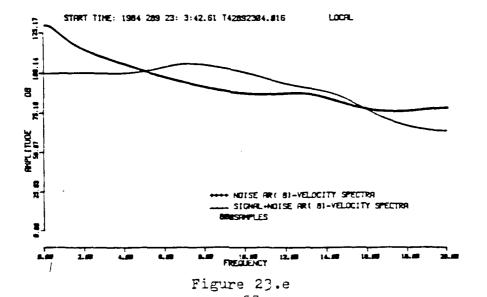
Figure 23. AR(8)-Spectral Estimates for the Signal and Noise for Event 5 are Normalized by the Variance of (a) Event Traces Recorded at Each of the Four the Noise. Stations, Lajitas, Marathon, Shafter and Tres Cuevas, for Event 5. (b) Lajitas AR(8)-Spectral Estimates. Analyst Estimates for the Dominate Period of the Noise and the First Arrival Phase are Indicated with Dashed Lines Drawn to their Respective Spectral Estimates. (c) Marathon AR(8)-Spectral Estimates. (d) Shafter AR(8) Spectral Estimates. (e) Tres Cuevas AR(8) Spectral Estimates.

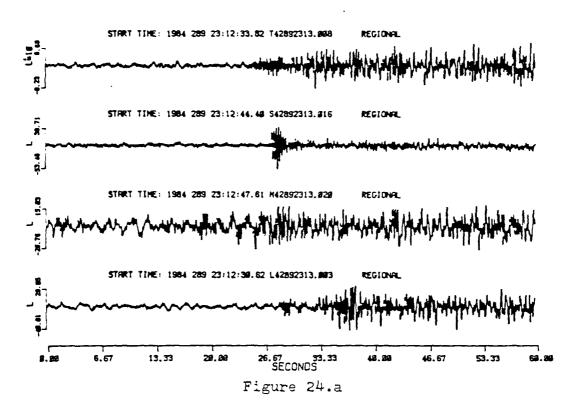


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Figure 24. AR(8)-Spectral Estimates for the Signa. and Noise for Event 6 are Normalized by the Variance of (a) Event Traces Recorded at Each of the Four the Noise. Stations, Lajitas, Marathon, Shafter and Tres Cuevas, for The (b) Lajitas AR(8)-Spectral Estimates. Analyst Estimates for the Dominate Period of the Noise and the First Arrival Phase are Indicated with Dashed Lines Estimates. Drawn to their Respective Spectral Marathon AR(8)-Spectral Estimates. (d) Shafter AR(8) (e) Tres Cuevas AR(8) Spectral Estimates. Estimates.

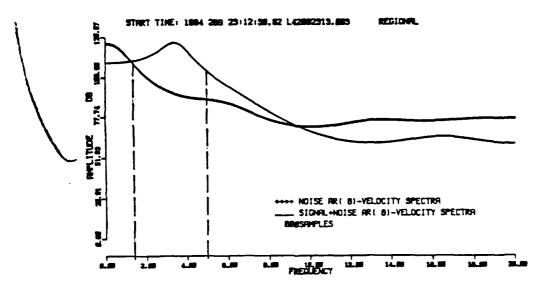


Figure 24.b

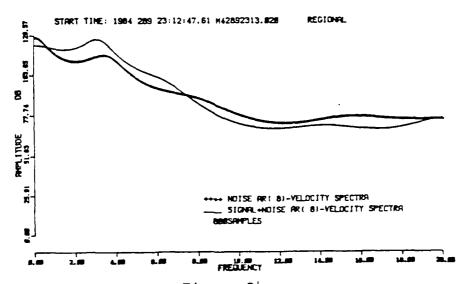
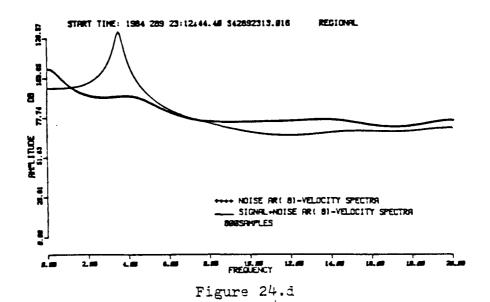
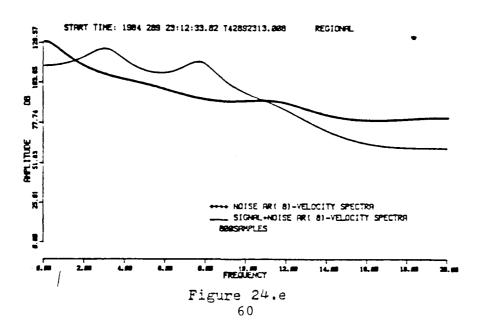


Figure 24.c

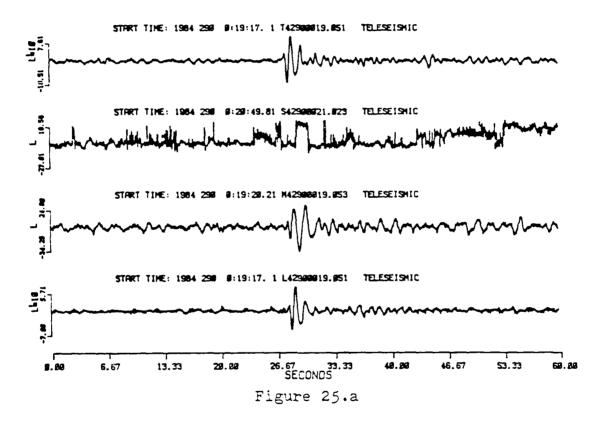


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Figure 25. AR(8)-Spectral Estimates for the Signal and Noise for Event 8 are Normalized by the Variance the Noise. (a) Event Traces Recorded at Each of the Four Stations, Lajitas, Marathon, Shafter and Tres Cuevas, for (b) Lajitas AR(8)-Spectral Estimates. Analyst Estimates for the Dominate Period of the Noise and the First Arrival Phase are Indicated with Dashed Lines Drawn to their Respective Spectral Estimates. Marathon AR(8)-Spectral Estimates. (d) Shafter AR(8) Spectral Estimates. (e) Tres Cuevas AR(8) Spectral Estimates.

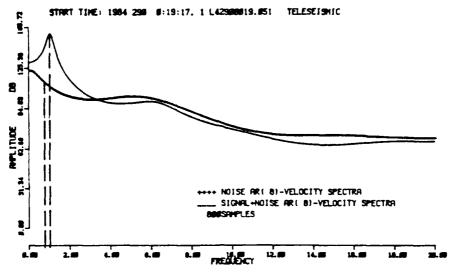


Figure 25.b

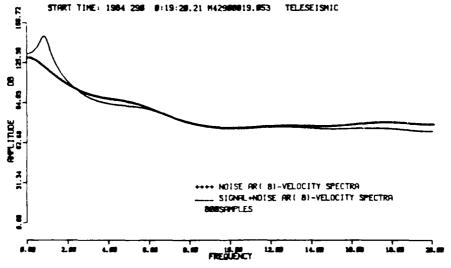
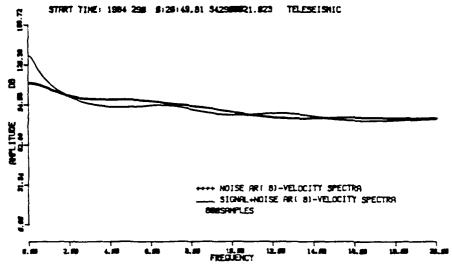
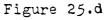
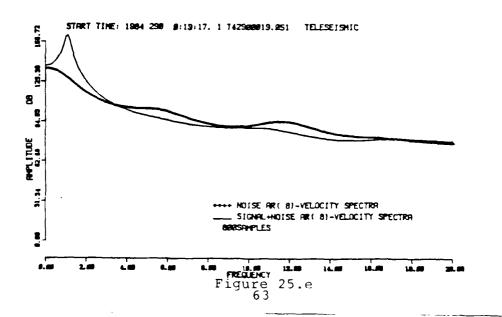
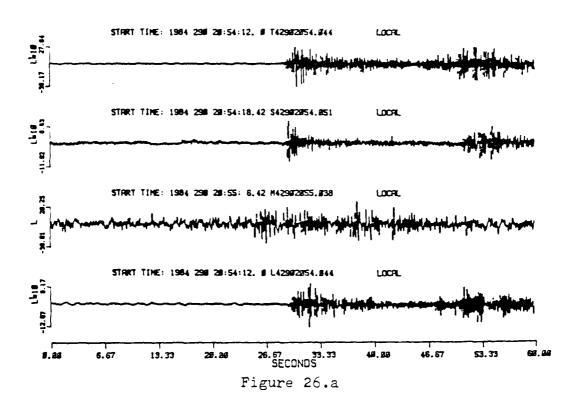


Figure 25.c 62



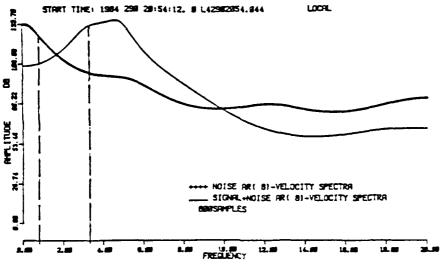


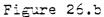


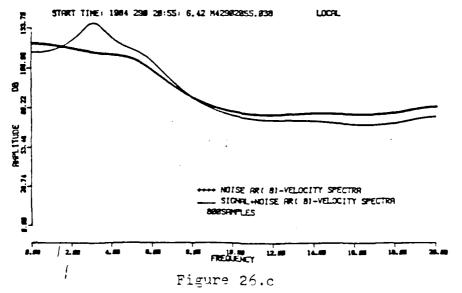


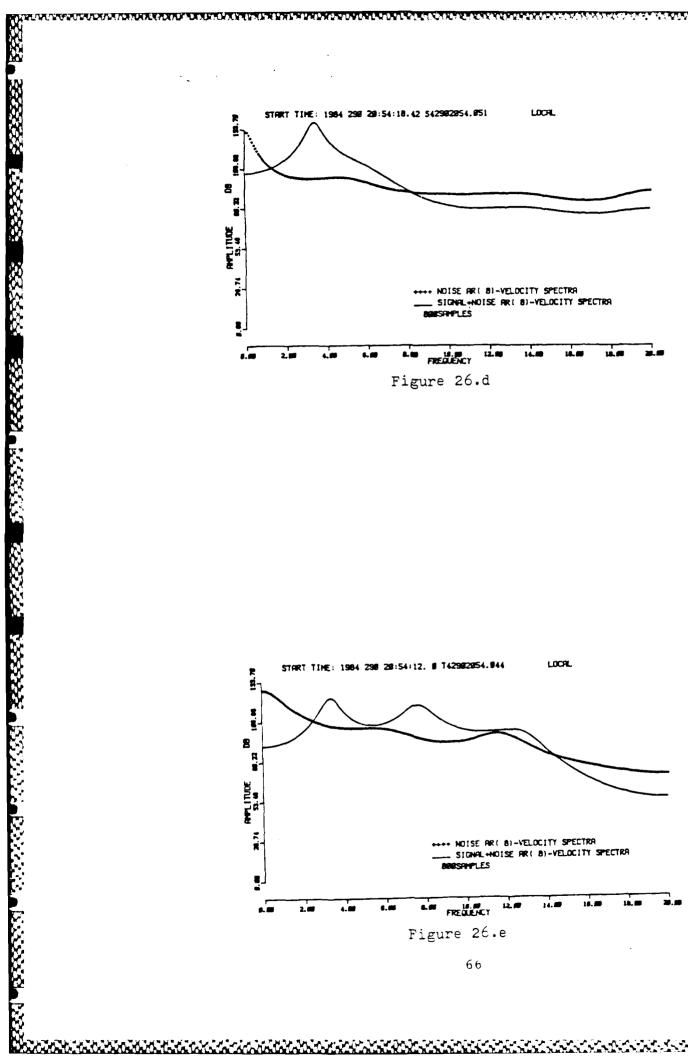
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Figure 26. AR(8)-Spectral Estimates for the Signal and Noise for Event 11 are Normalized by the Variance of the Noise. (a) Event Traces Recorded at Each of the Four Stations, Lajitas, Marathon, Shafter and Tres Cuevas, for Event 11. (b) Lajitas AR(8)-Spectral Estimates. The Analyst Estimates for the Dominate Period of the Noise and the First Arrival Phase are Indicated with Dashed Lines Drawn to their Respective Spectral Estimates. (c) Marathon AR(8)-Spectral Estimates. (d) Shafter AR(8) Spectral Estimates.









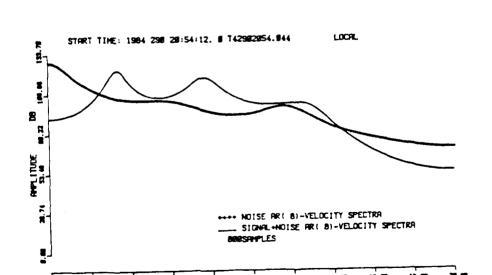


Figure 26.e

Non-parametric Detectors

Description of the Detector Used to Implement the Non-parametric Tests

The non-parametric tests used to distinguish signal from noise in the detector were performed on the "modified" slope of the event trace observation data. This "modified" slope is obtained by discrete integration of the slope of the event trace over segments with the same slope direction. Figure 20 illustrates how the slope and "modified" slope were calculated for each data point in the event traces. The subroutine, "SLOPE", in appendix B calculates the "modified" slope of the event trace.

The number of observations selected for testing the performance of each non-parametric technique as a detector The sample size was chosen so that the time 100. period of each observation window would be greater than largest period expected to occur in the event trace. periods of the various components of the event trace signal and noise range from 0.1 to 1.2 seconds. of 100 observations, taken at a sample rate of 40 samples per second, correspondes to two and one-half seconds. first 100 observations of each event trace is assummed to be a representative example of the background noise for The first the time period of the event trace. observations representing background noise are denoted

X(i), i=1,100, in the descriptions of the non-parametric tests following this description of the detector. The groups of 100 sequential observations taken at later times in the event trace and compared with the background noise, X(i), are denoted by Y(i), i=1,100. The null hypothesis used for each non-parametric test, H_0 , is the assumption that both X(i) and Y(i) are taken from the same population; i.e., that the cummulative distribution functions for X(i) and Y(i), F(x) and F(y), respectively are the same.

Since we know the first 30 seconds into the event trace precedes the 6.4 second window which triggered the fast Walsh detector, we make the assumption that the first arrival phase of the seismic event (P or Lg) is within the first 60 seconds of the event trace. To detect the seismic signal a window of 100 observations is moved in increments of 10 samples, 0.25 seconds, down the event trace. Each window of 100 samples is compared to the first 100 samples of the trace representing background noise using a non-parametric statistical test. Detection occurs when the non-parametric test indicates the two windows of 100 samples are not from the same population; i.e., the underlying distributions, F(x) and F(y), are not the same.

Three non-parametric tests, (1) the two sample sign test, (2) the run test, and (3) the rank sum test, were

tried in the detector to determine the most effective test for discriminating between signal and noise. The following describes each test, the assumptions each test made, the validity of those assumptions with respect to the data set and the effectiveness of each test in discriminating between the background noise and the signal plus noise. The detector subroutine, "DETECTOR", is listed in appendix B.

Two sample sign test

Assumptions

The data for the two sample sign test consists of two observations from the random samples, N population; i.e., background noise, and N observations from the treatment population; i.e., signal plus noise. In our case we choose N observations, $X(1), \ldots, X(N)$, from the beginning of the event trace representative of the background noise and N observations, $Y(1), \ldots, Y(N)$, from the remaining portion of the event trace. Unfortunately, the observations taken from a seismic signal are not Instead there is random variables. independent dependency between the observations. Walsh (1949,1951) has shown that the sign test will have similar results if the observations in the two samples are mildly dependent.

The two sample sign test is sensitive to changes in

the location and spread of the distribution. If a signal is present the spread of the X(i) observations should be significantly different from the spread of the background noise, Y(i) observations. The sign test run on the seismic data assumed that no ties, (i.e., X(i) = Y(i)), were present.

Procedure

To test

$$H_0: F(x) = F(y)$$

1. Define indicator variables

$$Z(i) = \begin{cases} 1 & \text{if } [Y(i) - X(i)] > 0 \\ 0 & \text{if } [Y(i) - X(i)] = 0 \\ -1 & \text{if } [Y(i) - X(i)] < 0 \end{cases}.$$

$$\delta(i) = \begin{cases} 1 & \text{if } Z(i) > 0 \\ 0 & \text{if } Z(i) \leq 0 \end{cases}$$

2. Set

$$\operatorname{Sn} = \sum_{i=1}^{N} Z(i) \delta(i).$$

The statistic Sn is the number of positive Z's.

3. For a one-sided test of HO versus the alternative,

$$H_a: F(x) = F(y),$$

at the alpha level of significance,

Reject H_0 if Sn < [N - b(alpha, N, 1/2)]

Accept H_0 if Sn > [N - b(alpha, N, 1/2)],

where the constant b(alpha, N, 1/2) under the null

hypothesis satisfies P[Sn > b(alpha, N, 1/2)] = alpha. That is, b(alpha, N, 1/2) is the upper alpha percentile point of the binomial distribution with sample size N and p = 1/2.

For large sample approximation under the null hypothesis define:

$$\operatorname{Sn}^* = \frac{\operatorname{Sn} - \operatorname{E}(\operatorname{Sn})}{\operatorname{SQRT[Var}(\operatorname{Sn})]} = \frac{\operatorname{Sn} - \operatorname{N}/2}{\operatorname{SQRT}(\operatorname{N}/4)}$$

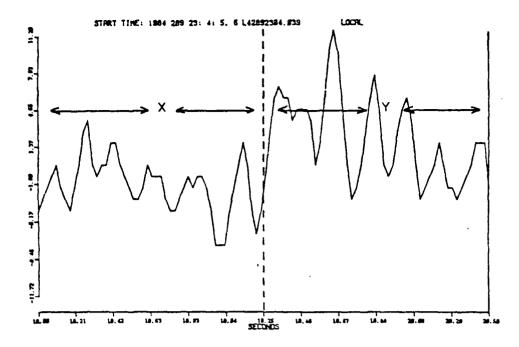
When ${\rm H_0}$ is true, the statistic Sn has an asymptotic (N tending to infinity) standard normal distribution, z(alpha). The normal theory approximation test of ${\rm H_0}$ versus the alternative ${\rm H_a}$, at the alpha level of significance is

Reject
$$H_0$$
 if $Sn^* \ge z(alpha)$
Accept H_0 if $Sn^* < z(alpha)$.

In regard to ties, X(i) = Y(j), if there are zero values among the Z's, discard them and redefine N to be the number of nonzero Z's. Figure 27 is an example of the two sample sign test with N = 30, k = 10 for a confidence interval (0.137-0.583), and alpha = 0.01.

Effectiveness

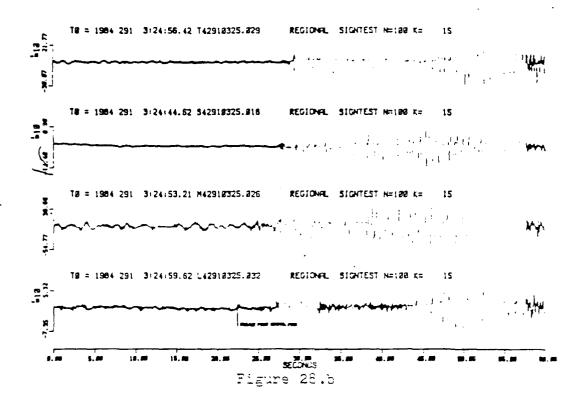
A large number of ties occur when the X(i) and Y(i) observations taken from seismic signals are compared. These ties are due to the limited dynamic range for small values of seismic observations near zero; i.e., the same values occur more often near zero since there is a more limited range of values to choose from. Since the



X	Y	SIGN
-2. Ø	-1.0	-1
4.8	1.0	-1
5.2	2.8	-1
9.3	1.2	1
-8.2	-3.2	-1
-11.0	-7.8	-1
-11.Ø	-13.0	1
-10.2	-15.2	1
-7.Ø	-13.Ø	1
-4.8	-6.2	1
0.0	2.0	1
3.Ø	7.2	1
7.2	11.2	1
6.0	10.0	1
5.0	7.0	1
3.2	1.2	-1
4.8	-4.0	Ø
8.2	9.2	-1
9.0	9.2	8
9.0	20.0	1
10.0	29.2	1
12.0	28.0	1
16.0	18.2] 1
18.0	6.0	-1
20.0	0.0	-1
17.3	-4.2	-1
11.0	1.2	-1
4.0	12.2	1
-1.0	17.8	1
-4.0	19.2	1 1

THO SAMPLE SIGN TEST
H_{O} : $F(x) = F(y)$
Ha: F(x) ≠ F(y)
N = 30
$E(S_n) = 15.0$
$ S_n - E(S_n) = 2.0$
K = 10
REJECT THE NULL HYPOTHESIS

Figure 27. Example of the Two Sample Sign Test. The Sign, (1,0,-1), Indicates the Absolute Value of the X(i) Observation is, Greater Than, Equal, or Less Than, the Absolute Value of the Y(i) Observation. Sn is the Summation of the Positive Signs.



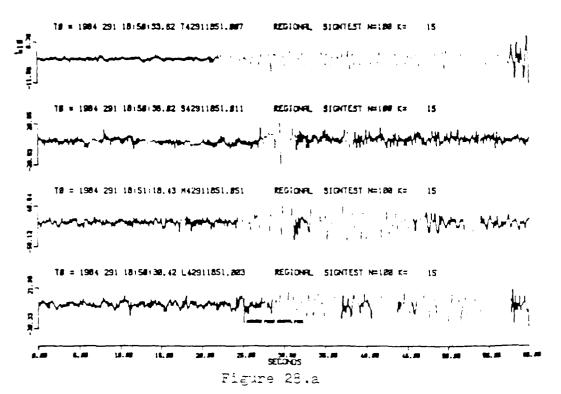
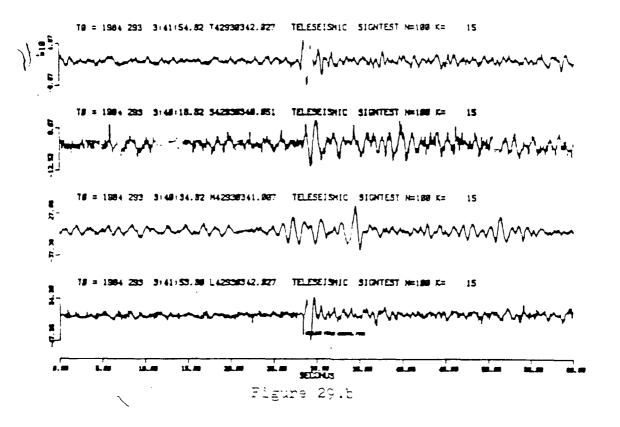


Figure 28. The Performance of the Two Sample Sign Test When Used to Detect Seismic Events. The Null Hypothesis for this Test Assumes all the Observations, X(i) and Y(i), Come from the Same Population, the Background Noise. The Blue Color Indicates a Rejection of the Null Hypothesis, i.e. a Signal is Present. (a) Event 19. (b) Event 14.



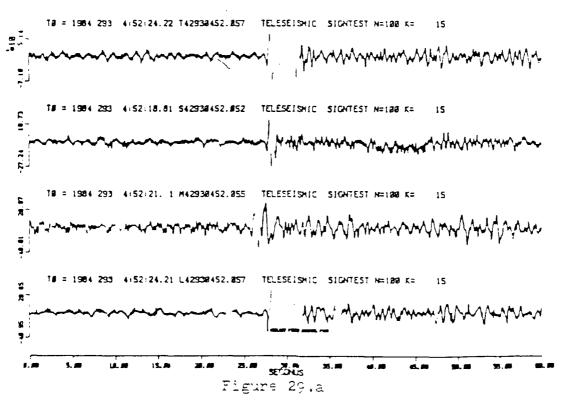


Figure 29. The Performance of the Two Sample Sign Test When Used to Detect Seismic Events. The Null Hypothesis for this Test Assumes all the Observations, X(1) and Y(i), Come from the Same Population, the Background Noise. The Blue Color Indicates a Rejection of the Null Hypothesis, i.e. a Signal is Present.

were not taken into consideration in calculating the value of the sign test they dilute its sensitivity to differences in the spread of the distributions. Ties were not accounted for because of the increase in computation time it would require; i.e., recalculation of the test for the reduced number of untied observations. Figures 28 and 29 illustrate the performance of the two sample sign test when used to detect signals. The sign test subroutine, "SIGNEXP", is listed in appendix B.

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Run test

Assumptions

The run test selects N observations, X(1), ..., X(N), of representative background noise taken from beginning of the event trace and M observations, $Y(1), \ldots, Y(M)$, at some time later in the event trace. Then the observations are ordered in ascending order the number of runs (groups of X or Y observations) counted. The run test in this application assumes the two samples, X and Y, are independent random variables. Since there is a dependence between the observations in each sample the designated level of significance for the test will not be preserved.

If X(i) and Y(j) are from the same population then X and Y will be well mixed and the number of runs will be large. However, if X(i) and Y(j) are from widely separate

populations then there will only be two runs. The run test is sensitive to differences in both shape and location of the distributions.

Procedure

To test

$$H_0: F(x) = F(y)$$
.

- 1. Order the observations, X(i), i=1,N and Y(i), i=1,M, together in ascending order from least to greatest.
- 2. Set Z equal to the number of distinct groups of Y's. The statistic Z is the number of runs in the ordered array of observations.
- 3. For a one-sided test of H_0 versus the alternative

$$H_a: F(x) = F(y)$$
,

at the alpha level of significance,

Reject
$$H_0$$
 if $Z \ge z0$

Accept
$$H_0$$
 if $Z < z0$,

where the constant z0 is the largest integer which satisfies

$$\sum_{z=0}^{z\phi} P[z = z] = alpha.$$

That is, z0 is the lower percentile point of the distribution of Z for sample sizes N and M.

The following specifies P[Z=z] under the null hypothesis so that we can determine the integer z0 for a given test size.

$$P[Z = z] = P[Z = 2k] = \frac{2\binom{M-1}{k-1}\binom{N-1}{k-1}}{\binom{M+N}{M}}$$

$$P[Z = z] = P[Z = 2k+1] = \frac{\binom{M-1}{k}\binom{N-1}{k-1} + \binom{M-1}{k-1}\binom{N-1}{k}}{\binom{M+N}{M}}$$

For large sample approximation define

$$Z^* = \underbrace{Z - E(Z)}_{SQRT[Var(Z)]}$$

$$E(Z) = \underbrace{2MN}_{M+N} + 1$$

$$Var(Z) = \frac{2MN(2MN-M-N)}{((M+N)^{**}2)(M+N-1)}.$$

When H_0 is true, the statistic z0 has an asymptotic (N tending to infinity) standard normal distribution. The normal theory approximation for the one-sided run test is

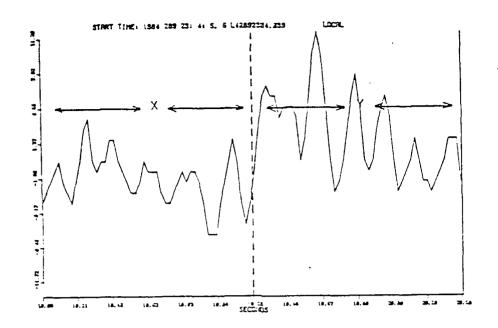
Reject
$$H_0$$
 if $Z^* \ge z(alpha)$
Accept H_0 if $Z^* < z(alpha)$.

For $M \ge 10$ and $N \ge 10$ the asymptotic normal distribution can be used to determine z0 (Mood, Graybill and Boes, 1974). If there are ties, Z(i) = Z(j), we delete the tied observations and recalculate z0 for the new value of N. Figure 30 illustrates an example of the run test for N = 30, M = 30, Z0 = 40, and alpha = 0.01. The run test subroutine, "RUNEXP", is listed in appendix B.

Effectiveness

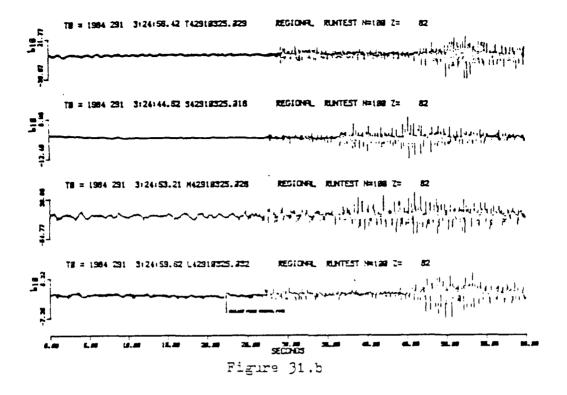
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Due to the way seismic signals are recorded digitally, there are a large number of ties which impede



INCEK	CATA T	YPE	INCEK	RTRC	TYPE	RUN TEST
1	2.3	Х	31	7.3	Y	$H_0\colon F(x)=F(y)$
2	3.3	X	32	7.3	Y	Ha: F(x) ≠ F(y)
3	2.3	Y	33	7.3	Х	N = 30 H = 30
4	8.3	Y	34	8.3	Х	E(Z) = 31.3
5	1.3	Y	35	8.3	X	VAR(Z) = 14.7
6	1.3	Y	35	9.3	X	z = 28
7	1.3	Y	37	9.3	Y	23 = 40
3	1.3	Х	39	9.3	X	REJECT THE NULL HYPOTHESIS
9	1.3	Y	39	13.3	X	
12	1.3	Y	42	12.3	X	
11	2.3	X	41	13.3	Y	
12	2.3	Y	42	11.3	X	
13	2.3	Y	43	11.3	Y	
14	3.3	Y	44	11.3	X	
15	3.3	X	45	11.3		
15	3.3	X	46	12.3	1	
17	4.3	X	47	12.2	Y	
18	4.3	Х	48	13.3		
13	4.3	Y	49	13.3		
23	4.3	Х	53	15.3		
21	4.3	X	51	16.3		
22	4.3	Y	52	17.3	1	
23	4.3		\$3	17.3		
24	5.3	X	54	18.3		
25	5.3	X	55	18.3		
25	6.3	X	56	19.3		
27	6.3	7	57	23.3		
23	6.3	Y	58	22.3	i	
29	7.3	X	59	29.2		
33	7.3	ΙΥ	62	1 29.3	113	

Figure 30. Example of the Run Test. The X(i) and Y(i) Observations are Ordered in Ascending Order. Then the Number of Runs of X and Y Values, Z, are Counted. If the Number of Runs is Less than the Threshold, z0, the Null Hypothesis that X and Y Both Come from the Same Population is Rejected.



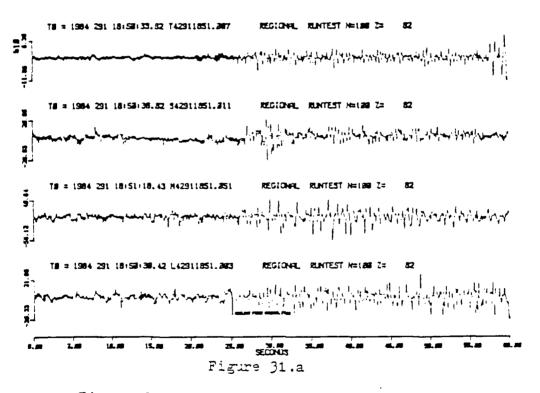
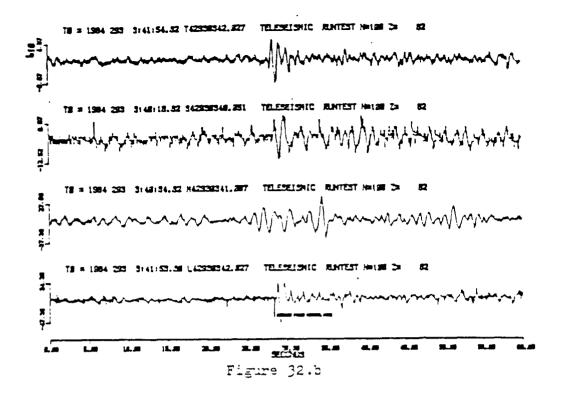


Figure 31. The Performance of the Run Test When Used to Detect Seismic Events. The Null Hypothesis for this Test Assumes all the Observations, X(i) and Y(i), Come from the Same Population, the Background Noise. The Blue Color Indicates a Rejection of the Null Hypothesis, i.e. a Signal is Present. (a) Event 19. (b) Event 14.



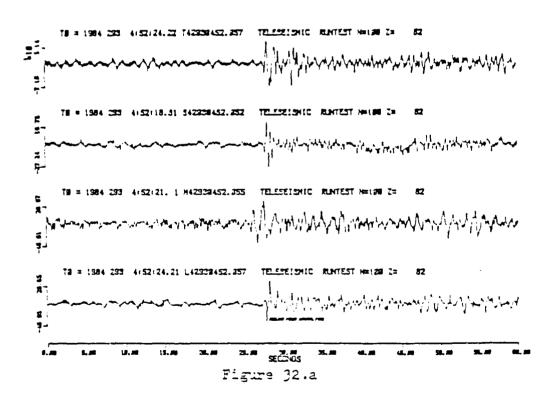


Figure 32. The Performance of the Run Test When Used to Detect Seismic Events. The Null Hypothesis for this Test Assumes all the Observations, X(i) and Y(i). Come from the Same Population, the Background Noise. The Blue Color Indicates a Rejection of the Null Hypothesis, i.e. a Signal is Present.

the effectiveness of the run test in distinguishing the differences in shape between two different populations. Since this non-parametric technique does not have a good way of handling the number of ties naturally occurring in seismic signal comparisons, this technique was rejected in favor of the rank sum test. Figures 31 through 32 illustrate the performance of the run test as a signal detector.

Rank sum test

Assumptions

The rank sum test selects N sequential observations, $X(1), \ldots, X(N)$, of the "modified" slope of the representative background noise from the beginning of the event trace. Then M sequential observations of "modified" slope, $Y(1), \ldots, Y(M)$, are chosen from some time later in the event trace. The N plus M observations are ordered in ascending order and assigned a rank based on their position in the ordered sequence. The ranks for the Y(i) observations, Ry(i), are summed and the absolute value of the difference between that rank sum, Ty, and the estimated mean, E(Ty), is compared with a predetermined threshold value, k.

The rank sum test assumes the X(i) and Y(i) observations are independent random variables. Since background noise is not a purely random process and the

seismic signal can be thought of as the background noise "treated" with the addition of a seismic event, neither the X(i) or the Y(i) observations are independent random Serfling (1968) investigating the robustness variables. of Wilcoxon test, upon which the rank sum test is based, let the two samples, X and Y, be independent of each other but let the random variables within a sample be possible dependent. It is found that the robustness of the test statistic for the Wilcoxon two-sample procedure under the hypothesis with departures from the assumption of random samples depends upon the grade of correlation of the variables X(i) and X(i+1). In other words, similar results for the rank sum test should be obtained when only mild dependence occurs between the observations in the samples. Since we do not know the exact degree of dependence between the observations, there is a probability that the significance level, (alpha), assumed for the rank sum test does not reflect the true significance level.

Procedure

To test

$$H_O: F(x) = F(y)$$
.

- 1. Order the N and M observations from least to greatest and let Ry(i) denote the rank of Y(i) in this ordering.
- 2. Set

$$TY = \sum_{i=0}^{M} RY(i).$$

The statistic Ty is the sum of the ranks assigned to the Y's.

3. For a one-sided test of HO versus the alternative

$$H_a: F(x) = F(y),$$

at the alpha level of significance,

Reject
$$H_0$$
 if $Ty \ge w(alpha, M, N)$

Accept
$$H_0$$
 if Ty < w(alpha,M,N),

where the constant w(alpha,M,N) under the null hypothesis satisfies

$$P[Ty > w(alpha,M,N)] = alpha.$$

Values of w(alpha,M,N) are given in Table C.1, Appendix C (Hollander and Wolfe, 1973).

For large sample approximation under the null hypothesis define

$$Ty^* = Ty - E(Ty)$$
 $SQRT[Var(Ty)]$.

When ${\rm H}_{\rm O}$ is true, the statistic Ty has an asymptotic (minimum of N or M tending to infinity) standard normal distribution. The one-sided normal approximation theory for the test statistic is

Reject
$$H_0$$
 if $Ty^* \ge z(alpha)$
Accept H_0 if $Ty^* < z(alpha)$.

For M \geq 7 and N \geq 7 the asymptotic normal approximation is quite accurate. (Mood, Graybill and Boes, 1974)

If there are ties; i.e., the ith observation in

ascending order is tied with the k observations following the ith observation then the summation of the indices divided by k+1 is the average rank assigned to each of the tied values. For the large sample approximation, compute Ty using average ranks, and replace Var(Ty) by

$$Var(Ty) = \frac{MN}{12} \left[M+N+1 - \sum_{i=p}^{L} f(i)(f(i)**2 - 1) \atop (M+N)(M+N-1) \right],$$

where L is the number of tied groups of ranks and f(i) is the size of the ith tied group. An untied observation is considered to be a tied group of size 1. Hence if there are no tied observations the right hand side of Var(Ty) reduces to

$$\frac{MN (M+N+1)}{12}$$

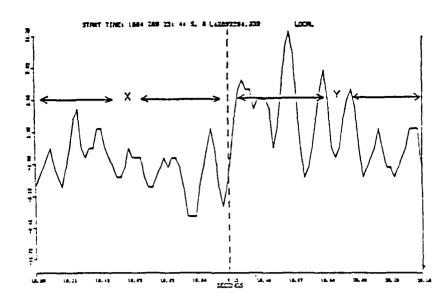
Figure 33 illustrates an example of the rank sum test for N=30, M=30, and alpha = 0.001. The rank sum test subroutine, "RANKTEST", is listed in appendix B.

Effectiveness

Since the ties are handled by averaging the ranks this test is sensitive to the relative changes in the shape of the seismic signal's density distribution. Figures 34 and 35 illustrate the application of the rank sum test to detect a seismic event.

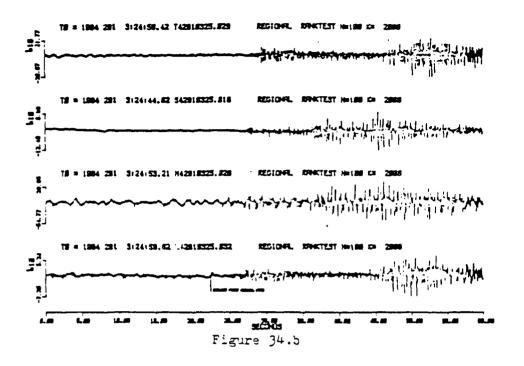
Conclusions

The rank sum test was the most effective of the non-



oata j	RENK	FREG	_ cata (RENK	FRED	RANK TEST
2.3	12.3	3.3	1.3	7.5	6.3	H_{G^*} $F(x) = F(y)$
I	23.3	7.3	1.3	7.5	5.3	Ha, F(xì ≠ F(yì
5.3	24.5	2.3	2.3	12.3	3.3	N = 33 H = 33
2.3	2.5	4.3	1.3	7.5	6. J	E(7q) = 915.3
8.3	34.5	2.3	3.3	15.3	3.3	VAR(Ty) = 4575.3
	43.5	4.3	7.3	31.3	5. J	[Ty - E(Ty) = 7.3
11.3		4.3	13.3	48.5	2.3	K = 222
12.3	40.2	3.3	15.3	53,3	1.3	ACCEPT THE MULL HYPOTHESIS
7.3	21.3	5.3	13.3	48,5	2.3	
	23.3	7.3	6.3	27.3	3.3	
a. 3	2.5	4.2	2. 3	12.3	3.3	
3.3 {	IS. J	3.3	7.3	31.3	5.3	
7.2	31.3	5.3	11.3	43.5	4.3	
	27.3	3.3	13.3	40.3	3.3	
5.3	24.5	2.3	7.2	31.3	5.3	
3.3	15.3	3.3	1.3 (7.5	6.3	
4.3	23.3	7.3	4.3	23.3	7.3	
8.3	34.5	2.3	8.3	2.5	4.3	
9.3	27.3	3.3	9.3	37.3	3.3	
9.3	37.3	3.3	23.3	57.5	2.3	
13.3	43.3	3.3	29.3	63.3	1.3	
12.3	46.5	2.3	28.3	59.3	1.3	
16.3	51.3	1.3	18.3	54.5	2.3	
18. J	54.5	2.3	6.3	27.3	3.3	
23.3		2.3	3.3	2.5	4.3	
17.3	52.5	2.3	4.3	22.3	7.3	
11.3		4.3	1.3	7.5	5.3	
4.3	23.3	7.3	12.3	46.5	2.3	
1.2	7.5	6.3	17.3	52.5		
4.3	23.3	7.3	19.3	56.3	1.3	

Figure 33. Example of the Rank Sum Test. The X(i) and Y(i) Observations are Ordered in Ascending Order and Assigned at Rank Indicating their Position in the Sequence. Then the Ranks, $R\gamma(i)$, for the Y(i) Values are Summed, Ty. If the Absolute Value of the Rank Sum, Ty, minus its Expected Value, $E(T\gamma)$, is Less than the Threshold, K, we Accept the Null Hypothesis that X(i) and Y(i) Come from the Same Population.



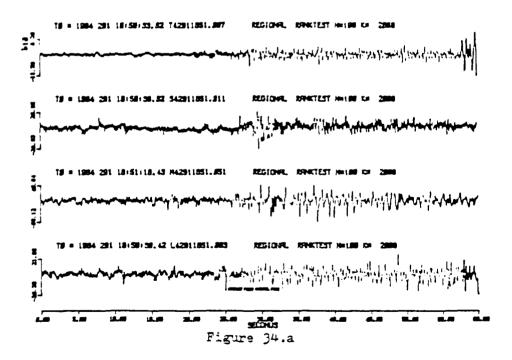
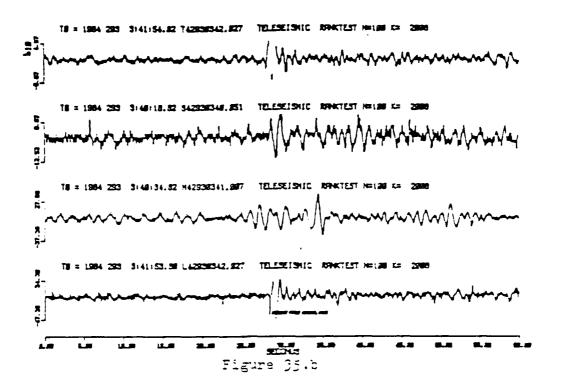


Figure 34. The Performance of the Rank Sum Test When Used to Detect Seismic Events. The Null Hypothesis for this Test Assumes all the Observations, X(i) and Y(i), Come from the Same Population, the Background Noise. The Blue Color Indicates a Rejection of the Null Hypothesis, i.e. a Signal is Present. (a) Event 19. (b) Event 14.



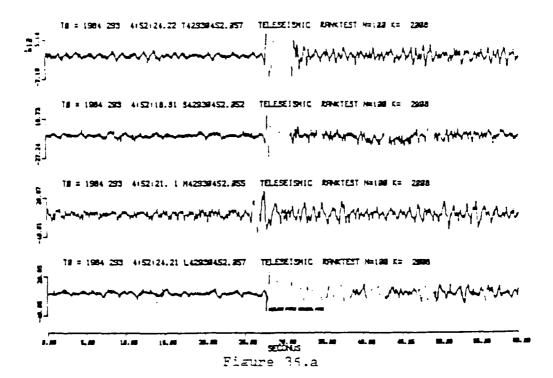


Figure 35. The Performance of the Rank Sum Test When Used to Detect Seismic Events. The Null Hypothesis for this Test Assumes all the Observations, X(i) and Y(i), Come from the Same Population, the Background Noise. The Blue Color Indicates a Rejection of the Null Hypothesis, i.e. a Signal is Present.

parametric tests in discriminating between signal and noise. After studying figures 28,29,31,32,34 and 35 one can see that the rank sum test does the best job of discriminating the event signal from background noise in a trace where an event signal is known to be present. The success of the rank sum test is due to its method of handling the large number of ties indigenous to seismic signals.

DETECTION TECHNIQUE AND FIRST ARRIVAL PICKER EMPLOYED TO FIND THE FIRST ARRIVAL PHASE OF EARTHQUAKES AND EXPLOSIONS IN SEISMIC SIGNALS CONTAMINATED BY NOISE

Description of Non-parametric Detector and Picker

rank sum test applied to the "modified" slope of observation the event trace data most clearly distinguished seismic signals from the background noise of the three non-parametric techniques tested. Several implementations of a detector based on the rank sum test were run on a training set of 48 event traces. The level of significance for the rank sum test was varied to find the threshold for detection which would allow the rank sum test to clearly discriminate between signal and noise. However, the "unmodified" application of the rank sum test between signal and noise was discriminate "consistent" in picking the seismic event due to varying S/N ratios. To allow the threshold for detection change with S/N ratio a "modified" version of the rank sum test was implemented.

The minimum and maximum rank sums for a each trace were determined in this "modified" implementation of the rank sum test. If the range of the rank sums, (i.e., sum of the ranks in each window), for a trace,

maximum rank sum - minimum rank sum,

is greater than 3600, which was observed to constitute the "high" S/N case, then the threshold for signal detection is computed to be the minimum rank sum plus 2700. range of the rank sums is less than 3600, which was observed to constitute the "low" S/N case, then the signal detection threshold is computed to be the minimum rank sum plus 75 percent of the range of the rank sums. Based on a detailed analysis of the rank sums test when applied to traces in the data set, this the event approximately 99.9 percent of the rank sums due to background noise to fall below the threshold in the "high" S/N cases and 99 percent in the "low" S/N cases.

The automatic first arrival pick is defined as the first zero-crossing preceding the first observation in the first detection window whose modified slope is at least 1.05 times greater than the maximum modified slope value of the representative background noise.

The event detector based on the rank sum test takes the first N=100 observations in the event trace to be a representative sample of the background noise present in that event trace. The rank sum test is computed for observation windows, N equal 100 samples in length, moved in increments of 10 samples, 0.25 seconds, down the event trace. The rank sum of each window is computed and compared with the rank sum of the background noise window.

The subroutine to detect and pick the first arrival phase, "RANK2700", is listed in appendix B.

Results

The detector which used the "modified" version of the rank sum test described in the preceding section to discriminate between signal and noise was run on 152 of the event files recorded during the GSETT experiment. These event files were selected to give the broadest representation of the wide variety of seismic signals we would expect to encounter if the detector were run over all events recorded by the fast Walsh detector on a daily basis.

To analyze the performance of the "modified" rank sum detector we looked at 128 of the event traces. Since 24 of the 152 event traces the detector was run on were also included in the training set used to develop the "modified" version of the rank sum detector they were excluded in this analysis to give us an unbiased look at the detector's performance. One hundred and twenty-four automatic picks were compared to their respective analyst's picks. Sixty-five were within 0.4 seconds of the analyst's pick. After analizing the automatic picks differing from the analyst's picks by more than 0.4 seconds, it was concluded that 5 appeared to be more correct than the analyst's original picks. Twenty-four of

the automatic picks were ambiguous, either the automatic pick or the analyst pick could be correct. Thirty-five of the automatic picks had errors larger than 0.4 seconds relative to the analyst picks: 4 local; 28 regional; and 3 teleseismic events. The representative background noise for 22 of the regional and local automatic picks with errors larger than 0.4 seconds contained P arrivals. This indicates that the violation of the assumption that the first 100 samples is background noise causes deterioration in the accuracy of the automatic first arrival pick. The apparently erroneous automatic picks that occur for teleseismic events are caused by emergent precursors to the P-arrival which the analyst picked.

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Occasionally, the assumption that the first 100 observations in the event trace represent background noise is violated. Either because the fast Walsh detector did not detect the P arrival but instead triggered on the Lg arrival or because the signatures from two seismic events occurring near the same time overlap each other on the the event trace. This causes either the P arrivals or the coda from a preceding event to contaminate the first 100 observations assummed to contain only background noise.

When the initial assumption that the first 100 samples of the trace represent background noise is violated the detector discriminates between signal and

TABLE 4. Errors (Automatic Minus Analyst Picks in Seconds), and Ranges of the Rank Sums for the 152 Events Used to Test RANK2700

ALL EVENTS

ALL EVENTS

SERVEY SERVEY TOTAL OF THE STREET PROPERTY STREET STREET STREET INSTRUMENT STREET

event name	RANGE	ERROR	SHAN THEYS	RANGE	ERRCR
	50F7 3	9.0	142921533.236	5751.5	-8. €
542491418.837	5057.8	8.2	M42962838, 881	4733.5	B. 4
L42892313.803	4822.5	8.2	\$42981821.825	4187.8	-3.4
M42911851.851	3742.8		L42911887.323	4572.5	8.4
742932383.819	5323.8	8.3	L42912142.315	5732.2	8.4
142973815.846	5751.5	8.3 8.3	M4231:827.829	4144.5	8.4
L42971753.818	5272.8		142912213.857	4443.5	-2. ↓
\$42971753.818	6174.0	8.2	L42921429.223	2235.2	0.5
T42980019.851	5 \$64.5	-0.1 Ø.1	M42932129.231	5872.2	-∂ .5
\$42980136.831	3581.8	9. i	T42892313.828	5 583.5	8.6
M42982855.838	4294.8	-8.1 8.1	L42902247.213	4145.8	-2. 6
142891418.034	5 349.5	Ø. I	742910325, 229	5643.8	8.6
M42912148.848	5149.0	8.1	542932129.318	6385.5	-2. 5
T42912213.857	4814.5 5145.5	-2.:	M42891815.238	3481.5	8.6
L42920528.008	4184.8	Ø. !	L+231:851.823	4626.5	-2 .6
542928528.881		-8.1	M42512325.225	55+3.3	8.7
142928528.888	5536.2 5469.2	8 .1	#42982218.822	4852.2	8.7
L42921533.036		Ø. 1	L42911959.217	4392.8	-2.7
H42921533.023	5871.8 5261.5	Ø. 1	542962238.214	4264.5	8.8
\$42921533.026		Ø. 1	742911682.883	5197.5	8.8
L42921859.231	3598.2 4275.8	8.1	542892327.218	4095.2	-8.8
742921859.028		-8.1	T42902246.038	3742.8	8.9
L42932129.889	5495.5	Ø. 1	M42921830, 008	3958.5	8.9
142932129.006	5 797.8 4 928.5	-8.1	M42978815.843	6135.5	-2 .9
L42891418.237	5 539.5	8.1	M42911602.038	3788.3	8.9
142951988.257	3728.5	8.1	M42971951.302	2483.5	-1.3
L42962037.316	5929.5	8.1	L42972815.248	6891.8	-1.1
\$42978815.839	5454.2	Ø. 1	\$42951989.222	5251.5	-1.1
\$42892313.016	3311.5	Ø. 1	T42921829.239	4849.5	-1.2
M42892328.007	4579.5	8.1	L42831719.050	4166.5	1.2
M42988819.853	5984.8	8.2	742891814.846	4812.8	1.2
T42982854.844	5848.2	8.2	L42971952.227	4583.8	1.2
L42932303.019	5869.5	8.2	\$42332303.035	5628.5	-1.2
N42932383.235	3994.8	-8.2	L42918325.832	5958.8	1.4
\$42921 438. 238	5725.8	8.2	142962237.216	3655.2	1.5
M42951908.050	4528.5	8.2	M42891719.857	3825.8	1.7
\$42911603.000	4569 S	8.2	L42892327.316	4615.5	-1.8
\$42900021.023	5475.5	8.2	742891719.347	4242.5	1.9
M42891418.053 L42900019.051	4862.8	-8.2	342321529.823	5681.5	-2.2
L42302054.044	5584.2	-8.2	M42912214.852	4867.8	2.1
M42971753.038	5839.5	8.2	\$ 42921858.356	3 923.5	2.3
742892384.816	57 48 . Ø	8.2	\$42911851.811	5646.8	-2. 4
M42981828.847	4772.3	Ø. 3	M42891986.848	334 9.5	2.7
542911807.026	4369.2	Ø. 3	T42912148.815	5 518.5	-2.9
M42921430.014	4358.8	8.3	L42932229.327	2860.8	3, 2
142911807.032	4579.5	-8. 3	T#2971752.852	5585.2	3.3
L42951988,857	5912.5	8.3	#4292Ø528.804	5 799.5	-3, 4
T42921 429. 318	3944.8	-0.3	742971951.000	3419.5	3,9
142892327.212	5168.3	Ø. 3	M42902246.029	4954.5	4. 4
142911851.207	5059.5	8.3	M42892384.828	4193.5	-5.1
T42911958.256	4692.5	8.3	\$42997211.000	2888. 3	5.1
L42921829,839	4112.5	8.3	542902247.016	2987.8	5.4
742891985.888	3880.2	8.3	\$42911958.058	6332.8	-6.3
742980136.821	4554.5	8.3	542892324.212	5496.2	-6.3
T42982209.838	3713.8	Ø.3	L42891815.218	4178.8	6.5

TABLE 4 (Continued)

ALL EVENTS

LOCAL EVENTS

EVENT NAME	RANGE	ERROR	EVENT NAME	RANGE	ERRCR
342891719.025	4825.5	7.4	L42971753.018	5272.8	0.2
L42961821.816	4485.5	7. 5	542971753.218	6174. a	8.8
T42981828.837	3292.8	-9.1	L42962837.216	3722.5	8.1
\$42971950.048	4355.5	18.1	M429@2855.@38	4294.8	8.1
\$42912139.938	3748.8	18.4	M42912148.848	5149.2	0. :
\$42891615.918	3478.5	18.4	T42892384.816	5748.8	8.2
M42980137.019	4228.2	12.6	T42982854.844	5984.8	8.2
M42911959.208	4199.2	13.4	M42971753.238	58 39.5	8.2
542891923.218	335:.2	-15.1	L42587254.844	5524.2	3 . 2
M42892313.322	4762.5	-15.3	M42962838,881	4733.5	8.4
L42891906.230	3868. 5	-17.5	L42912148.815	5732.2	8.4
L42892384.839	55 76.5	-19.9	\$42562838.814	4264.5	0.8
\$42912325.216	5376.2	-22.9	M42971951.208	2423.5	-1.8
M42921859.854	4645.5	~23.6	L42971952.887	4583.2	1.2
\$42917215.201	5845.8	-23.7	T42962237.216	3655.2	1.5
L42900136.053	4227.5	-25.8	T42912148.815	5518.5	-2.9
\$42902054.051	6894.2	-26.8	742971752.852	5585.2	3.3
L42911602.007	4522.2	-33.3	T42971951.202	3419.5	3.9
			M42892324.228	♦193.5	-5. ì
NUMBER OF EVENTS	= 128		542892384.812	5496.2	-6.3
MEDIAN = 0.5			\$42971958.040	4355.5	10.1
MERN = 3.2			542912139.838	37 4E. Z	18.4
YARIANCE = 48.3			L42892384.839	\$5 76.5	-19.9
			\$42972854.851	6894.2	-26.8

NUMBER OF EVENTS = 24
MEDIAN = 8.8
MEPN = 4.8
VARIANCE = 326.3

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L22832131,889 4871,5 8,8 L2381713,858 418,5 1.2 NA2311251,5 15 774,4 8,8 L2381713,858 418,5 1.2 NA2311251,2 18 544,6 8,1 1.4 14291125,2 27 545,3 1.4 1		EVENT NAME	RANGE	ERROR	EVENT NAME	RANGE	ERROR	
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142772737, 81:9								
### MAZESTOR, ### 3211,5		T42932383,819	5323.8	8.2	142921829.239	4849.5		
SACADETINE, 291 2501, 2								
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LC2271573, 355					T42891719.847	4242.5	1.9	
M42211533, 222 5871, 8 8.1 542211554, 255 5846, 2 -2.4 L42211551, 231 3386, 2 8.1 542211554, 255 5846, 2 -2.4 L42211559, 225 1 3386, 2 8.1 44221155, 226, 2 314, 2 526, 2 526, 2 314, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2 526, 2								
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M42931829.247		T42988136,821	4554.5	0.3	L42983136.253	4227.5	-25. a	
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M42918325. #26		M+2891815.238	3481.5	8. 6				
M42922218.822 4852.8 8.7 L42911959.817 4398.3 -8.7 T42911682.883 5197.5 8.8 542892327.318 4896.8 -4.8 T42942246.338 3742.3 8.9 M42911682.838 3788.3 8.9 M42921833.808 3958.5 8.9								
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T42911682.883 5197.5 8.8 542892327.818 4896.8 -8.8 T42982246.838 3742.8 8.9 M42911682.838 3958.5 8.9 M42921838.888 3958.5					11			
T42987246.838		T42911602.003	5 197.5	9. 8				
M42911682.338 3788.3 8.9 M42921833.308 3958.5 8.9								
M42921832.208 3958.5 &.9								
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<u>P'</u> 8					3
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Ž	TABLE 4	(Contin	ued)		3
3	TELESE	ISMIC EVENTS			
	EYENT NAME	RANGE	ERROR		
	T42978815.846 T42988819.851 L42928528.888	5751.5 5564.5	8.8 -8.1		3
2000 Contraction of the Contract	\$42928528.888 \$42928528.888 \$42928528.828 \$42951988.857	5145.5 4184.8 5536.8 5539.5	-0.1 8.1 -3.1 8.1		
	\$42978815.239 #42988819.253 L42988819.251	5989.5 4573.5 4862.8	8.1 8.1 -8.2		
<u> </u>	\$42988821.223 M42951988.252 L42951988.257	4569.5 5725.8 5912.5	8.2 8.2 8.3		3
X X	M42978815.843 L42978815.846 \$42951989.888	6135.5 6891.8 5251.5	-8.9 -1.1 -1.1		
	M42928528.284 MUMBER OF EVENT MEDIAN = 8.1	5 799.5	-3. 4		â
	MEAN = 8.5 VARIANCE = 859.	8			2000
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noise poorly. One way to improve the effectiveness of the detection method is to extend the time preceding the 6.4 second window triggering the fast Walsh detector to 60 seconds. Past experience in picking the first arrival has shown that while the Pn arrival sometimes precedes the Pg arrival triggering the Walsh detector by more than seconds all of the Pn arrivals are included in the first 60 seconds preceding the window triggering the detector. Then only overlapping seismic events will cause performance of the detector to deteriorate. Of the event traces in the data set only four were overlapping events. Table 4 lists the errors relative to analyst's picks in the automatic first arrival picks the range of the rank sums for all the events and for each individual type of event; i.e., local, regional and The illustrations depicting the performance teleseismic. of the automatic detector and first arrival picker on 152 event traces are included in appendix D.

CONCLUSIONS AND SUGGESTIONS FOR FUTURE WORK

rank sum detector is extremely robust and The excellent for distinguishing a broad category of seismic signals from noise. However, the time preceding the window triggering detection in the fast Walsh detector should be increased to 60 seconds. To illustrate this conclusion let us examine an extreme case where we would like the detector to pick either Pn or Lg as the first arrival for later event association and 1 cation. Suppose the fast Walsh detector fails to trigger on the Pn arrival from a regional event 130 from the recording station. maximum time interval between Pn and Pg at 130 is 57 seconds. So if the fast Walsh detector triggered on Pg, the Pn arrival would still be included in the event trace file if the predetection time was increased to 60 seconds. If the fast Walsh detector triggered on Lg instead of Pg, then we conclude the Pn arrival must be too far below the background noise level for even a human analyst to discern otherwise the fast Walsh detector would have triggered on Pg instead of Lg. In this event the human analyst would pick Pg as the first arrival when we would rather have him pick Lg for later association purposes. Since the time interval between Pg and Lg for a regional event 130

the recording station is 3 minutes, the automatic detector would select Lg as the first arrival. The real time fast Walsh detector writes seismic events occurring closer together than 60 seconds into a continuous event file. There should be no significant problem with overlapping events in extending the predetection time to 60 seconds for each event file. Then 99% of the automatic picks should be correct within a median error of plus or minus 0.2 seconds (see Table 5).

The next step is to develop a criteria to break up the class of signals detected into local, regional and teleseismic events and then into earthquakes and explosions. Decision tree or stratified layer classification is designed to take advantage of such situations (one or two features; i.e. slope, period or amplitude) to improve efficiency and, if possible accuracy. The decision tree classifier progresses through a series of stages or layers; at each layer certain classes (local, regional and teleseismic) are separated in the simplest manner possible. It is flexible and permits different features and classifications (boundaries) to be used to separate different classes. (Schowengerdt, R.A., 1983).

TABLZ 5. Errors (Automatic Minus Analyst Picks in Seconds), and Ranges of the Rank Sums of Events Which do not Violate the Initial Assumptions

ALL EYENTS EXCEPT THOSE WHICH VICLATE THE INITIAL ASSUMPTIONS OF THE RANKZTON DETECTOR

ALL EVENTS EXCEPT THOSE WHICH VIOLATE THE INITIAL ASSUMPTIONS OF THE RANKSTOR DETECTOR

EVENT NAME	RENGE	ERRCR	EVENT HAME	RANGE	ercr
			542321430.333	3994.2	-3.2
\$42891418.J37	5857.2	8.2	742892384.316	5748.8	ø. 2
L42892334.339	5576. <i>S</i>	a. a	\$42999821.323	4569.5	8. 2
L12892313.283	4887.5	a. 3	M42931419.253	5475.5	a.2
M42900137.319	4228.3	3.3	L42332333.313	5243.0	7.2
L42981821.318	4425.3	3.3	M42332333.335	5289.5	8.2
T42281923.337	3232.3	3.3	H42371753.038	5039.5	a. 2
L42310325.032	5953. ð	8.3	\$42311683.288	4528. S	a. 2
L42311582.287	4522.3	3.3	L+2582354.844	5284.8	~3. 2
M4291:522.J35	3788.3	8.3	M423S1988.JS3	5725.8	8.2
T42911582.203	5197.5	8.2	L42988819.351	4862.3	-8.2
L12911851.303	4626.5	8.2	\$42911807.025	4369.2	0.3
M42911851.351	3742.3	8.3	142311807.332	4573.5	-∂.3
M42311959.208	4199.J	3.3	142392327.312	5168.J	a.3
T42912148.315	\$518.5	3.3	142930136.321	4554.5	a.3
M42321833.238	3953.5	3.3	M42321438.314	4350.0	a. 3
\$42921858.356	3922.5	8.3	T42911851.J27	5053.5	a. 3
M42932129.331	5070.3	3.3	T4231.429.319	3944.3	-3.3
\$42332129.319	6385.3	2.3	142891985.208	3888.2	a. 3
\$42932333.335	\$628.5	9.3	742911958.356	4692.5	8.3
T42332333. 819	5323.2	3.3	44291838.847	4773.2	9.3
\$42951309.200	5251.5	ø. 3	L429Z1329. J39	4112.5	2.3
542362229.314	4254.5	3.3	L12951388.357	5912.5	9.3
T42362337.316	3655.3	ø. J	142931384.037	3713.2	9.3
142372815.346	5751.5	a. a	#42911897. J29	4144.5	2.4
142371753.318	5272. J	a. 3	L42912213.057	4443.5	-3.4
\$12971753.318	6174.3	8.3	T42921533. J36	5751.5	-3. ↓
T42971752.352	5585.8	8.3	142921333.223 142962233.221	4733.5	8.4
M42971951.300	2+33.5	3.3	542981821.825	4187.3	-3.4
T42371351.888	3413.5	3.3	L12912143.315	5722.3	8.4
\$42571533.326	5251.5	a.:	L12912140.013 L12911897.029	4572.5	9.4
M42892319.287	3311.5	a.:		41.45.J	-3.5
L (2321853.331	3958.3	8.1	L42582247.213 T42892313.208	5583.5	3.5
M42533819.353	4573.5	a.:	742918325.329	5643.2	9.6
T42321853.328	4275. J	8.1		3481.5	8.6
L42932129.289	5455.3	~∂.:	M42891815.838	5540.0	8.7
742999819.351	5 564.5	⊸∂. 1	N+2910325.326	4398.2	-8.7
L42891418.337	4923, 5	-₹. દ	Li2911959.817	4052.3	3.7
742932129.206	5737.3	8.1	H42502210.022	4296. J	-₹.3
\$42323135.331	3531.3	8.1	\$42992327.318	3742.3	3.3
742351418.334	53 +3. 5	-3 .:	742382246.338	6135.5	-8.3
542892313.315	5454.2	8. 1	M42978815.343	6291.3	-1.1
142351938.357	5539.5	ð. i	L12973815.346	1849.5	-1.2
L42962837.316	37 12. E	ð. :	7 (1371929, 339	4583.8	
M42912148.348	51 49. 3	a.:	L12971352.307	4812.8	
M42902255.038	4234.3	3.1	742.91614.346	4166.5	
\$42372815.339	5589.5	3.1	L42891719.353	3826.J	
T42312213.357	4814.5	a. 1	M42891719.357	4615.5	_
L42920513.208	51 45. 3	-3.1	L\$2892327.316	1212 5	
542323523.231	4184.3	a. 1	T42891719.347	3349.5	
142322513.398	5535.2		n42891986.343	2862.2	
L42921533.825	5469.2	_	L42922239.327	5739.5	_
H42921533.323	5871.2		M42923529.204	4133.3	· .
142932354.344	5924.3		H42892334.328	4733	

			GRANDED ROLL OF OFFICE OF OFFICE OF OFFICE OF OFFICE OFFIC		CONTROL OF THE PROPERTY OF	A Children Con Now Now Now No.
		m.nrn 6	/ Committee d \			
		TABLE 5	(Continued)			
	•					
ALL EVENTS E	KCEPT THOSE	HHICH	LOCAL EVENTS I	EXCEPT THOSE	HHICH	
VIOLATE THE II			VIOLATE THE I			
OF THE RANG	K2788 DETECT	TOR	OF THE RAN	OF THE RANKETBB DETECTOR		
EVENT NAME	RANGE	ERROR	EVENT NAKE	RANGE	ERRCR	
42892384.812	5496.8	-6.3	L42892384.839	5576.5	ø. g	
42891719.025	4805.5	7.4	T42912140.815	5518.5	0.8	
			\$42962838.814	4264.5	6.3	
UMBER OF EVENTS	= 188		742962837.816	3655.2	8.8	
EDJAN = 8.2			L42971753.218	5272.2	8.3	
ERN = 8.5			\$42971753.818	6174.2	0.2	
PRIPNCE = 1.3			142971752.852	5525.2	8.2	
			M42971951.888	2483.5	Ø. B	
			142971951.000	3419.5	ø. g	:
			H42982855.838	4294.2	Ø. 1	1
			L42962237.316	3722.5	0. 1	j
			M42912140.840	5149.8	8. i	,
			L42902054.044	5584.8	-2.2	
			T42892304.216	5742.8	Ø. 2	į.
			T42982854.844	5924.2	8.2	•
			M42971753.238	5039.5	8,2	

MUMBER OF EVENTS = 21 MEDIAN = 8.1 MEAN = 8.7 YARIANCE = 11.5

L42971952.887

M42892374.828

542892384.812

4583.2

4193.5

5496.2

1.2

-5.1

-6.3

Substitute(

TABLE 5 (Continued)

REGIONAL EVENTS EXCEPT THOSE WHICH VIOLATE THE INITIAL ASSUMPTIONS OF THE RANKETON DETECTOR

EVENT NAME	RANGE	ERROR
142891418.837	5857.8	8.8
42892313.803	4882.5	3.8
142988137.819	4228.2	8.8
42981821.816	4425.5	8.8
142981828.837	3292.8	8.0
42918325.832	5 958.8	8.8
L42911602.807	4522.0	8.9
M42911692.938	3788.2	8.8
T42911692.983	5 197.5	8.8
142911851.083	4626.5	8.8
M42911851.851	3742.8	8.8
M42911959.008	4199.0	8.2
M42921838.808	3958.5	8.8
\$42921858.856	3923.5	8.3
M42932129.831	5973.3	9.9
\$42932129.818	<i>5</i> 385.5	9. 3
\$42932383.235	5628.5	8.8
T42932383.819	5323.2	8.8
L42921533.836	5469.2	8.1
M42921533,823	5871.2	8.1
\$42921533.226	5261.5	9.1
\$42988136.831	3581.2	8.1
L42921859.231	3938.2	8.1
L42891418.237	4928.5	~8. 1
T42921859.828	4275.8	8.1
L42932129.289	5495.5	-8.:
\$42892313.016	5454.8	8.1
M42892328.887	9311.5	9. 1
T42932129.006	5797.0	Ø.1 -Ø.1
T42891418.834	5349.5	8. 1
T42912213.257	4814.5	8.2
M42891418.853	5475.5 5848.8	8.2
L42932383.819	5889.5	8.2
H42932383.835	3994.8	-8.2
\$42921434.838	4528.5	8.2
342911683.888	3889.8	8.3
742891985.288	5859.5	8.3
742911851.887	4778.8	9.3
M42981828.847	4692.5	8.3
742911858.856 742892327.812	5168.8	8.3
M42921438.814	4358.8	2.3
742982289.838	3713.8	0.3
742921429.018	3944. B	-8.3
342911887.826	4369.8	9.3
742911887.832	45 79.5	-8.3
742980136.221	4554.5	g. 3
L #2921829.839	4112.5	g. 3
542921821.825	4187.8	
142921533.036	5751.5	
L42912213.857	4443.5	
L42911887.829	4572.5	
M42911887.329	4144.5	

REGIONAL EVENTS EXCEPT THOSE WHICH VIOLATE THE INITIAL ASSUMPTIONS OF THE RANK2788 DETECTOR

EVENT NAME	RANGE	ERROR
742919325.929	5643.8	5. 6
T42892313.888	5583 <i>.5</i>	8.6
L42982247.813	4145.8	-3.6
M42891815.838	3481.5	8.6
M42982218.822	4852.3	8.7
M42910325.026	5548.8	8.7
L42911959.817	4398.8	-8. 7
\$42892327.818	4896.2	-3.8
742902246.038	3742.8	9. 9
T42921829.839	4849.5	-1.2
1.2891719.858	4166.5	1.2
742891814.846	4812.8	1.2
M42891719.857	3826. 8	1.7
L42892327.816	4615.5	-1.8
742891719.047	4242.5	1.9
M42891986.848	3349.5	2.7
142982289.827	2868.8	3.8
\$42891719.025	4885.5	7.4

NUMBER OF EVENTS = 71
NEDIAN = 8.2
NEAN = 8.5
VARIANCE = 4.5

TABLE 5 (Continued)

TELESEISMIC EVENTS EXCEPT THOSE WHICH VIOLATE THE INITIAL ASSUMPTIONS OF THE RANK2788 DETECTOR

EVENT NAME	RANGE	ERROR
\$42951989.888	5251.5	g.g g.g
T42978815.846 L42928528.888	5751.5 5145.5	-8. I
542928528.881	4184.2	a.:
T42922528.888	5535.2 4579.5	-3.1 8.1
M42988819.853 T42951988.857	5 539.5	B. 1
\$42978815.839	5909.5	Ø.1 -Ø.1
T42988819.851 54298821.823	5564.5 4569.5	8.2
M42951988.858	5725.8	8.2
L42900019.051	4862.0 5912.5	-0.2 8.3
L42951988.257 M42978815.843	6135.5	~8. 9
L42978815.846	6891.8	-1.1 -3.4
M42922528.284	5 799.5	-3.4

NUMBER OF EVENTS = 16
MEDIAN = 8.1
MEAN = 8.4
VARIANCE = 22.9

THE FOLLOWING PAGES CONTAIN
APPENDIXES AND EVENTS

APPENDIX A FAST WALSH TRANSFORM DETECTOR

The Walsh functions are an ordered set of rectangular waveforms whose amplitudes take the values +1 or -1. They are arranged in order of increasing number of zero-crossings per time interval, so called sequency order.

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Leveline

Sequency is defined as one half the number of zero-crossings per interval. The sequency of Walsh transforms is analogous to the frequency of sines and cosines used in Fourier analysis.

Any time series can be expressed as the weighted sum of a series of Walsh functions. The Walsh transform provides the coefficients for the summation in the same way as done by a Fourier transform. The advantage of the fast Walsh transform is that it can be computed much faster than the Fourier transform because it involves only integer addition and shifting. The macro subroutine for the fast Walsh transform used in the real time event detector is listed in appendix C.

ACCOUNT OF STATE OF S

One complication of the fast Walsh transform is that it does not produce a spectrum in sequency order. Instead, it produces a Paley ordered spectrum (sometimes called natural order) in which the subscripts have been bit-reversed. The seismic event detection program uses lookup table A.1 to convert a sequency subscript from the passband into a bit-reversed Paley subscript in the transformed spectrum. (GSETT Report, Paul Golden, 1985)

TABLE A.1 Sequency Subscript to Bit-Reversed Paley Subscript Conversion

SEQUENCY OFFERE	SUBSCRIPTS	BIT-REVERSED PALEY	ORDERED SUBSCRIPTS
1	1	1	1
2	64	2	33 49
3 4	32 33	3 •	17
5	16	5	25
6	49	6	57
7	17	7	41
8	48	8	9 13
8	8 57	.9 18	45
1 9 11	25	11	61
12	48	12	29
13	8	13	21
14	56	14 15	5 3 37
15 16	24 41	16	5
17	4	17	7
18	61	18	39
19	29	19	55 ~
28	36	28 21	23 31
21 22	13 5 2	22	63
23	28	23	47
24	45	24	15
25	5	25	11 43
26	6 8 28	26 27	5 9
27 28	26 37	28	27
29	12	29	19
38	53	38	51
16	21	31	35 3
32	44	32 33	3 ♦
33 34	63	34	36
35	31	35	52
36	34	36	28
37	15	37 38	28 68
38 39	50 18	39	44
33 48	47	48	12
41	7	41	16
42	58	42	48
43	26	43 44	64 32
44 45	39 18	45	24
46	5S	46	56
47	23	47	48
48	42	48	8 6
49 50	3 62	49 58	38
50 51	36	51	54
52	35	52	22
53	14	53	38
54	51	54 55	62 46
55 5 6	19 4 6	56 56	-14
57	6	5 7	10
58	59	58	42
59	27	59 6Ø	58 26
68 61	38 11	61	18
61 62	54	62	50
63	22	ទ	34
54	43	64	2

The time series data in the detection window is transformed using the fast Walsh transform. The absolute values of the weighted coefficients are summed across the pass band. The sum is then normalized by shifting to the value it would have had if the weights had all been 1. This normalized sum is the magnitude which is compared to the threshold for the channel. A magnitude below the threshold triggers the detection.

When detection occurs the signal is written to an event file. This event file is later graphically displayed to enable a seismic analyst to discriminate between an "event" signal and "glitches" or "noise" and to pick the event parameters. (Goforth and Herrin, 1981)

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APPENDIX B

Subroutines

LIST OF SUBROUTINES

AR_SPECTRA - COMPUTES AR(P) SPECTRAL ESTIMATES

RKRIKE - COMPUTES AR YULE-HALKER COEFFICIENTS

COY - COMPUTES AUTOCORRELATIONS FOR A TIME SERIES

DETECTOR - DETECTOR USED TO TEST NONPARAMETRIC TESTS

IFWH64 - INTEGER FAST WALSH TRANSFORM IN MACRO FOR 64 POINT HINCOW

NEIGHBOR - COMBINES SECHENTS WITH NEPREST NEIGHBOR DECISION RULE

POF - COMPUTES DENSITY DISTRIBUTION HISTORORAMS

Poin_A-BURG - COMPUTES AR BURG COEFFICIENTS

RANKFRED - COMPUTES RANKS AND FREQUENCIES FOR N OBSERVATIONS

RANKTEST - COMPUTES THE RANK SUM SUM TEST FOR NAM OBSERVATIONS

RANK2788 - DETECTOR BASED ON A MODIFIED RANK SUM TEST

RUNEXP - COMPUTES THE RUN TEST FOR N+M OBSERVATIONS

SIGNEXP - COMPUTES THE SIGN TEST FOR N OBSERVATIONS

SLOPE - COMPUTES MODIFIED SLOPE FOR A TIME SERIES

```
SUBROUTINE AR SPECTRAIFHAX, FMIN, COEF, NCDEF,
                            PF, PFLOG . NF . VAR . ARMAX)
           This subroutine computes the RR spectral estimate as
c
      defined by H.S. Priestly in Spectral Analysis and Time Series,
      (paperback), p. 681, 1981.
                                                     IFREQUENCY HAXIMUM
      REAL PHAX
                                                     IFREQUENCY MINIMUM
       REAL FHIN
                                                       IAR COEFFICIENTS
       REAL COEF INCOEF
                                          INPLITUDE OF AR SPECTRA AT F
       REAL PF (251)
                                    IMPLITUDE IN 48 OF AR SPECTRA AT F
       REAL PFLUG(251)
                                    IVARIANCE OF WHITE HOISE OF PROCESS
       REFL YAR
                                             INLIMBER OF AR COSFFICIENTS
       INTEGER NOOEF
       COMPLEXAS CHPLXF, CONST. SUM
       COMPLEX=6 THOP:1/(8.8.6.28318538718)/
       REAL THOP1/6.28318538718/
                                        INUMBER OF POINTS IN AR-SPECTRA
       INTEGER NF
      COMPUTE SPECTRAL ESTIMATE FOR EACH FREQUENCY
        F = FMIN
                                            LOSPUTE FREQUENCY INCREMENT
        SRATE = PLOAT (NF-1) / (FWAX-PHIN)
        DO K = 1.NF
                                                      I (K-1) th FREGUENCY
          F = (FLOAT (K-1) / SRAYE) + FHIN
          SUM = CHPLX(1.8.8.8)
          IF INCOST. DE. 11 THEN
            DO I=1.NODEF
                                                           LOOHPLEX FREQUENCY
               DIPLXF = CHPLX(Fe1,8.8)
              SUM = SUM - CAPLX (COEF (I) .8.8) aCEXP (THOP II aCHPLXF)
             DAG 00
           END IF
           CONST = SUM=CONJG(SUK)
           PF(K) = 1./REAL(CONST)
                                            IVAR = VARIANCE OF HOUSE DATA
           PF(K) = PF(K)=2. SYPR
                                              IAMPLITUDE IN & OF SPECTRA
           PFLOG(K) = 28. =PLOG18(PF(K))
                                              INAXIMUM SPECTRAL PHPLITUDE
           ARHAX = MAX (PFLOCIK) . ARHAX)
         END DO IK
         RETURN
         ĐĐ
```

```
SUBROUTINE AKAIKE (X.N.M. ACF. HL. COEF)
¢
                         Modified from Grou-Hoodward ARMA epectral analysis
               package for the IBM FC by Kathleen A. Alden 1-38-86.
                This exproutine computes the Yule-Halker coefficients
                as described by Kay and Hample: (1982).
                REAL X(1) ISERIES
               REPL FPE (38) INORK ARRAY FOR FINAL PREDICTION ERROR (FPE)
                REAL ACF(1) INUTOCORRELATION
                REAL A(38,38) INORK ARRAY
                REAL COEF(1) IYULE-RALKER COEFFICIENTS
                                                   ILENGTH OF X SERIES
                INTEGER N
                                                   IMPAZIMUM ORDER OF FPE
                INTEGER M
                INTEGER ML
                                                  TORDER SELECTED BY FPE .
 C
                NL8G=138
                ICCEF=1
                SUM=2.8
                DO 51 1=1.N
                SUM = SUM + X(I)
        SI CONTINUE
                XBPR=SUM/N
                XSQ=9.8
               DO 52 1=1.N
                X(I) = X(I) - XBFR
                XSQ = XSQ + X(I) = X(I)
        52 CONTINUE
                PV = XSQ/N
                CALCULATION OF AR COEF'S. USING YULE-HALKER EQUATIONS
                A(1,1) = A(F(2))
                FPE(1) = (1.8-(A(1.1)-A(1.1)))-(N+2)/(N-2)
                H1 = H-1
                DQ 2 I=1.M1
                SN = 8.8
                SD = 8.8
                00 3 J=1.I
                SN = SN + RII, J #ACFII+2-J
                50 = 50 + A(I,J) =ACF(J+1)
            3 CONTINUE
                A(I+I,I+I) = (ACF(I+2) - 5N)/(I.B-SD)
                FPE(I+1) = FPE(I) = (N+i+2) = (N-i-1) = (1.9-A(i+1,i+1) = A(i+1,i+1)) / (1.9-A(i+1,i+1) = A(i+1,i+1) = A(i+
              + [ (N-I-2) #(N+I+1) 1
                1:1=t + 00
                A(I+I,J) = A(I,J) - A(I+I,I+I) = A(I,I-J+I)
            4 CONTINUE
            2 CONTINUE
                 1F(100EF.EQ. 8) 00 TO 15
         16 CONTINUE
         15 CONTINUE
                 f=1
                 يير
            7 IF (FPE(I) .LE. FPE(J)) GO TO S
                I = J
                 J = J + 1
                 GO TO 8
            5 J = J + 1
            8 IF(J .GT. H) 00 TO 6
                 60 TO 7
             & ML=I
                 00 65 1=1,ML
          65 COEF (1) =A (ML. 1)
                 RETURN
                  ENO
```

<mark>átásásás</mark>ásás is tertetetet menyen menyen meg e

```
SUBROUTINE COV (X.L.XB, GV.D.N)
ε
           This subroutines calculates N autocorrelations, the
     easole sean, and the variance equal to the zero lag auto-
     variance for the input time series, x(i), i=1,L.
     Taken from the Gray-Hoodward ARMA spectral estimation
      pockage, 1985.
C
       X=SERIES -IMPUT
       LALENGTH OF SERIES - INPUT
       XB=MEAN - OUTPUT
       GY=YARIANCE - DUTPUT
       D=AUTOCORRELATIONS - OUTPUT
       NEWHORR OF AUTOCORRELATIONS TO CALCULATE -INPUT
       NOTE:0(1)=AUTOCORRELATIONS AT LAG I-1
     REGL X(1)
     REAL D(1)
    COMPUTE THE SAMPLE MEAN FOR XIII. I=1.L
      XB=8.8
     DO 25 1=1.L
 25 XB=XB+X(1)
     XB=XB/FLOAT(L)
    COMPUTE THE AUTOCOVARIANCES AT LAGS = 3, (N-1)
     DO 38 I=1.N
     D(1)≈8.8
     NX=L-I+i
     DO 48 J=1.NK
    D(1)=(X(J)=(BX-(I+J-1)-XB)+D(I)
  38 CONTINUE
    SAMPLE VARIANCE EQUALS THE AUTOCOVARIANCE AT LAG = 0.
      GV=0(1)
C
    NORMALIZE THE AUTOCOVARIANCES TO OBTAIN THE AUTOCORRELATION
      DO 52 I≃1.N
     D(1)=0(1)/GV
```

RETURN ENG

SACON BOSSON PROPERTY CARCAGO BOARD

```
SUBROLITINE DETECTOR (NSAMPLES, N. MEAN, K. INC.
                           TRC. SEGLEN. PATTERN. IST. SN. RANK)
C- SUBROUTINE TO PERFORM THO SAMPLE SIGN TEST
C- HOODY, GRAYBILL & BOES
      REAL TEST
      REAL PATTERN(1) 11.=18,2.=19
      REAL TRC(1) ITRACE CONTAINING SIGNAL
      REAL SEGLEN(1) ISAMPLE LENGTH OF PATTERN
      REAL HEAN
      REAL SH IMAX SIGNAL TO NOISE RATIO FOR TRACE
       INTEGER NSAMPLES
       INTEGER INC LINCREMENT HINDON HOVES DOWN TRACE
       INTEGER IST LISTART IN HINDON
       INTEGER N INDIGER OF SAMPLES IN NOISE PART OF TEST
       INTEGER & INCCEPT OR REJECT CRITERIA
       REAL T(388.3) INORK RERRY FOR ORDERING SAMPLES
       REAL RANK(1) ISUMATION OF RANKS OF SIGNAL
       LOGICAL ACCEPT/. FALSE./
      COMPLITE RANK SUM TEST OVER TRACE
        N2 = 2mN
        SN = HEAN
        ENHAX = 8.8
        DO 1=1.N
          L = 1-1
          PRITERN (NSAMPLES-L) = 1.8
          SEGLEN (NSAMPLES-L) = 1.8
          PRITERN(I) = 1.8
           SEGLENII) = 1.8
           T(1,1) = ABS(TRC(1=2-1))
           T(1+N,1) = ABS(TRC(1=2))
           ENHAX = MAX (T(I.1), ENHAX)
           DIMAX = MAX (T (I+N.1) DIMAX)
         ENO DC
         ENARX = ENMAX=1.05
           IR = 8
           DO I=1, (NSAMPLES-H), INC
              IR = IR + 1
             00 J=1.N
                \mathfrak{J}=\mathfrak{J}+(\mathfrak{I}-\mathfrak{l})
                T(J+N.1) = ABS(TRC(JJ))
              איניבנו מם מאפ
              CALL RANKTEST (N.T (1,1),T (N+1,1),K.TY,ACCEPT)
              SN = MAX (TY, SN)
              RANK(IR) = TY
              IF (ACCEPT) THEN
                DO L=[,([+(INC-1])
                  PATTERNIL) = 2.8 IRE ECT HB
                  SEGLENILI = 1.8
                 540 00
               ELSE
                 DO L=1, (1+(INC-1))
                   PATTERNIL = 1.8 IACCEPT HO
                   SECLENIL) = 1.8
                 END 00
               END IF
             FM 00 11
           SN = SN/RANK(1) IMAX SIGNAL RANK /NOISE RANK
            RETURN
           END
```

```
IFWM64
     .TITLE
      LENGTH 64 INTEGER FRST HALSH TRANSFORM.
      INPLEMENTS FORTRAN YERSION IN IFNEGA.FOR
      LABELS AND VARIABLE NAMES IN COMMENTS ARE
      FROM THE FORTRAN VERSION.
       BY DIVIDENT DELIRING ON 22-NOV-82.
       Modified 12-SEP-84 to handle integer overflow
                INTEGER+2 X164)
               CRLL IFNH64(X)
       PRODUCES TRANSFORM IN-PLACE WITH RESULT IN
       BIT-REVERSED DYROIC (NATURAL) DROFT.
                                              . OF ELEVENTS
       517=64
                                              164 = 2.6
       PHR=6
       L=$17/2
                      IFHM64
       .PSECT
                      IFWM64, 4M(R2,R3,R4,R5,R6,R7)
       .ENTRY
, REGISTER USAGE:
                            ITERATION CONTROL
         RØ
                              L AS HORD INDEX
                   L=2
         R1
         R2
                   POINTS TO X(P)
         R3
                   POINTS TO X(PPL)
         84
         PS
                               K AS HORE INDEX
                    K#2
                    XP
                                                        FITERATION COUNT
                       PHR . RE
         MOVZBL
                                                2عاد
                     #<L#2>,R1
         HOVL
                                                      ٠R
                        01,R2
                                                   ADDRIXIPI). P=1
         HOVZBL
                          4(API .R3
                                                JADOR (X(PPL)) . PPL=L+1
             MOVL
 188:
                       R3.R1,R4
         ADDL3
                                                      1=1
                                                          12K=2 (HORD INDEX)
                        -1.R5
          HOVZBL
                            e2.R6
             HOVZBL
                                                           XP=X (P)
 285:
                          (R3), R7
                                                      X(P)=XP+X(PPL) . P=P+1
             MOVH
 385 t
                       R7, (R4), (R3)+
          EMODA3
                                                  ing overflow, OK
                      4Ø5
                                                   reign changed to negative
          BVC
                                                      , largest negative lies IRBS()
                      353
          BLSS
                       4.X8881,-2(R3)
          HOVH
                      425
                                                           sues largest positive
          888
                           **X7FFF,-2(R3)
                                                          X(PPL)=XP-X(PPL), PPL=PPL
              HOVH
  3531
                            (R4) .R7. (R4)+
               SUBH3
                                                   ino overflow. DK
   ARS:
          BYC
                      585
                                                    setion changed to negative
                                                       plangest negativent (so IRASI)
                       453
           BLSS
                       -- X8881,-2(R4)
           MOYN
                      585
                                                           summe largest positive
           BRB
                            exX7FFF,-21R4)
               HOVH
                                                           ,2X=2X+2
   4551
                             •2,R6
               ADD'-2
                        26,21
                                                     , IF (K.LE.L)
           CHPL
                        385
           BLED
                                                        المحصود
                         R1.R3
           ADDL-2
                                                        ,PPL=PPL+L
                         R1.R4
            ADDL-2
                                                    1+1=20
                        25
            INCL
                         P5.22
            CHPL
                                                     , IF (J.LE.R)
                         2£$
            PL ED
                                                   11=1/2
                         -1.R1.R1
            ASHL
                                                  ,R=R=2
                         #1.R2.R2
            ASHL.
                           28,185
             SOBGIR
             RET
             .ENG
```

SUBROLITINE NEIGHBOR!HAXLEN, SEG, NSEGMENTS, NHIHI

UND BEINDE STEINE STEIN

```
The nearest neighbor decision rule first compares the
      feature vector for each measure of an arrow of observations.
      Each segment is tagged with a 2 or 1 to indicate its affinity
      to either the previous or next sequent respectively. Then
     the segments which have an affinity for each other are ocebined
     and their features are averaged. This continues until the
     number of segments becomes less than or equal to the minimum
     number of seconds desired.
C---
     Code written by Kathieen R. Aiden. March 14,1984.
     SEGII.JI - INPUT VECTOR FOR ITH SECHENT, Jai. 8
                SEGII.1) = AMPLITUDE
                 SEG(1.2) = ESTIMATED PERIOD
                 SEG(1.3) = NUMBER OF DBSERVATIONS IN 1TH SEGMENT
C
                 SEG(1.4) = ROSOLLITE DIFFERENCE BETWEEN
2
                           1-1 AND I SECHENT AMPLITUDES
Ĉ
                 SEG(1.5) = ABSOLUTE DIFFERENCE BETWEEN
                          I-1 AND I SECHENT PERIODS
                 SEG(1.6) = PASOLUTE DIFFERENCE BETHEEN
                           I+1 AND I SECHENT AMPLITUDES
C
                 SEC(1,7) = ABSOLUTE DIFFERENCE BETWEEN
ε
                           I+1 AND I SEGMENT PERIODS
                 SEC(1.8) = 8 INDICATES AFFINITY WITH NEXT SECRENT
                           I INDICATES AFFINITY WITH PREVIOUS SEGMENT
                 NSECHENTSHENTS = NUMBER OFSECHENTS IN TRACE
                 NMIN = MINIMUM NUMBER OF SEGMENTS IN TRACE
      REAL SEGUMANLEN. 81
      INTEGER INSEGMENTS ITOTAL NUMBER OF SECHENTS
      INTEGER WHIN IMINIMUM NUMBER OF SECHENTS
C- DETERMINE REFINITY OF EACH SECHENT
      DO MILLEINSEGMENTS. GT. NMIN)
                                 11ST SAMPLE AFFINITY MANDET SECREPIT
       SEG(1.8) = 8.
        SEGINSECMENTS.81 = 1.8
                                     ILAST SAMPLE AFFINITY NUPREVIOUS
        DO 1=2.NSEGMENTS-1
          SEG(1.4) = 985(SEG(1-1.1)-SEG(1.1))
                                                      IPREV AMPLT DIF
          SEG(1.5) = ABS(SEG(1-1.2)-SEG(1.2))
                                                      IPREV PER DIF
          SEG(1.6) = RBS(SEG(1+1.1)-SEG(1.1))
                                                      INEXT AMPLT DIF
          SEG(1,7) = PBS(SEG(1+1,2)-SEG(1,2))
                                                      INEXT PER DIF
          IF (SEG(1,4).GT.SEG(1,6)) THEN
            IF (SEG(1,5), GT. SEG(1,7)) THEN
              SEG(1.8) = 8.8 IPERIOD & AMPLITUDE AFFINITY MANEXT SEG
            FL.SF
             SEG(1.8) = 8.8
                                      IAMPLITUDE AFFINITY WINDST SEG
            DO IF
          ELSE
            IF (SEG(1.5).LE.SEG(1.7)) THEN
              SEG(1.8) = 1.8 IPERIOD & AMPLITUDE AFFINITY MYPREV SEG
            E SE
             SEG(1.8) = 1.0
                                       I PHPLITUDE REVINITY HUPREY SEG
            END IF
          END IF
        EN0 00
     COMBINE SECRENTS WITH AN AFFINITY FOR EACH OTHER
        X = 8
                             LINITIAL NUMBER OF NEW COMBINED SEGMENTS
```

- CONTINUED -SUBROUTINE NEIGHBOR (MAXLEN. SEG, NSEGMENTS, NMIN)

```
IF (SEG(2.8).LT.8.5) THEN
 K = K + 1
end if
DO I = 2, NSEDMENTS
    IF (SEG(1-1.8).LT.8.5) THEN (COMBINE FEATURES FOR NEW SEG
  IF (SEG(1.8) GT.8.5) THEN
                                                    IMPLITUDE
       K = K + 1
       SEG(K.1) = (SEG(I-1.1) + SEG(I.1))/2.
                                                       (PERICO
       SEG(K,2) = (SEG(I-1,2) + SEG(I,2))/2.
                                                       (LENSTH
       SEG(K,3) = (SEG(1-1,3) + SEG(1,3))
     عكلة
                                                     IPHPL:TUCE
        K = K + 1
        SEG(K,1) = SEG(1,1)
                                                       IPERICO
        SEG(K,2) = SEG(1,2)
                                                        ILENGTH
        SEG(K.3) = SEG(1.3)
      ENO IF
    ELSE
      1F (SEG([+1,8].LT.8.5) THEN
                                                      IRMPLITUDE
        K = K+1
         SEG(K.1) = SEG(1.1)
                                                         IPERIOD
         SEG(K.2) = SEG(1,2)
                                                         ILENGTH
         SEG(K.3) = SEG(1,3)
       END IF
     DIG IF
                                   ICOMBINE FEATURES FOR NEW SEC
    540 00
    IF (SEG(1-1.8).LT.8.5) THEN
                                                       I PHPL ITUDE
      SEG(X,1) = (SEG(1-1,1) + SEG(1,1))/2.
                                                          IPERIOD
      SEG(K,2) = (SEG(1-1,2) + SEG(1,2))/2.
                                                          ILENGTH
      SEG(K.3) = (SEG(1-1.3) + SEG(1.3))
     ENO IF
     NSECHENTS = K
   END DO INHILE INSECUENTS, CT, NAINI
   RETURN
   END
```

SSS A LAGGER REPRESENT OF LAGGERT

```
PROGRAM POF
    This program calculates the density distribution
histograms for the background naise, first 1800 points,
and the signal + naise, second 1808 paints, of the selenic
event troce.
PORRMETER (MOXIEN=3888)
PARAMETER (MAXNTRC=4)
PARAMETER (IBUFSIZ=256)
PARAMETER (NINTERV=1888)
CHARACTER #88 BLANKS
CHARACTER#16 FLNAME (MAXNTRC)
CHARACTER=12 TRCTYPE
CHRRACTER MOS TOHAR
LOGICAL SAMESCALE/, FALSE. / INDUST ALL TRACES TO SAME SCALE IF . TRUE.
LOGICAL CLASS
REAL CLASSINT ICLASS INTERVAL FOR PROB.DENS.CALCULATION
REAL TRACE (MAXLEN, MAXNTRC) ITRACES FOR SEISHIC EVENTS
REAL PONDISE (NINTERY, MAXNIRC) | IPROBABILITY DENSITY DIST. NOISE
REAL POSIGNALININTERV. HAXNTRC1 IPROBABILITY DENSITY DIST. SIGNAL+NOISE
REAL SAMPLERATE/48./ ISAMPLES PER SECOND
REAL SCALETRE (MAXNTRC) / 1.8, 1.8, 1.9, 1.8 / ISCALE TRACES FOR PLOTTING
REAL FSVALUE/8.8/
                      IFIRST SAMPLE VALUE IN SECONDS
REGL XINPLOT/24./
                    IX LENGTH OF PLOT IN INCHES
REAL YIMPLOT/28./ IY LENGTH OF PLOT IN INCHES
REAL AMIN
REAL AMAX
REAL CHIN/188./
REAL CHAX/-188./
REAL FAIN/8.8/
                   IFREQUENCY MINIMUM
REAL FMAX/8.8/
                   IFREQUENCY MAXIMUM
REAL AHT/8.21/
                    ISIZE OF CHARACTERS IN INCHES
REAL XIN IX-AXIS LENGTH
REAL YIN IY-AXIS LENGTH
REAL EVTXIN/24./ IEVENT X AXIS IN INCHES
REAL EVTYIN/12./ IEVENT Y RXIS IN INCHES
REAL PROBXIN/S./ IPROBABILITY DENS X AXIS IN INCHES
REAL PROBYIN/3./ IPROBABILITY DENS Y AXIS IN INCHES
REAL EYTHAX/-18.8/ IEVENT MAXIMUM PLOTTED
REAL EVININATOR OF TEVENT HINIMUM PLOTTED
REAL FREQUIN/8.8/ IRELATIVE FREQUENCY MINIMUM PLOTTED
REAL FREGMAX/1.8/ IRELATIVE FREGUENCY MAXIMUM PLOTTED
REAL IXOFF/1.8/
REAL IYOFF/1.8/
REAL FAC/1,8/ IFRCTOR TO SCALE PLOT
REAL XORIGIN/8.8/ HINITIAL PLOT ORIGIN
REAL YORIGIN/8.8/ IINITIAL PLCT CRIGIN
REAL AFACTOR/8.8/ IY-AXIS TRACES OVERLAPP IF.LE.8.
 INTEGER ITIME(8) IARRAY TO COMPUTE CORRECT TIME FOR 1ST SAMPLE
 INTEGER APOFSAPS/1888/ INUMBER OF SAMPLES IN PROSABILITY DISTR.
INTEGER NSAMPLES/2400/ INUMBER OF SAMPLES IN EACH TRACE
 INTEGER NOLASSINT/188/ INCHBER OF CLASS INTERVALS
 INTEGER NPLOTS/1/ LINITIPLIZE THE NUMBER OF PLOTS PLOTTED TO 1
 INTEGER STRTSAMP (MAXNTRC) /1,1,1,1/ IF IRST SAMPLE TO BE PLOTTED
 INTEGER XLABEL (28) ILABEL FOR X-AXIS
 INTEGER YLABEL (28) ILABEL FOR Y-AXIS
 INTEGER PLABEL (28) ILABEL FOR PLOT
 INTEGER NCHARSX/B/
 INTEGER NCHARSY/0/
 INTEGER NOHORSP/8/
 INTEGER NTRACES/4/ INUMBER OF TRACES
 Integer Ibuff | Ibuffer size for standard colooms
 integer no
                 inumber standard calcomp
 Integer igu
                 Hogical unit number for PLOTSIN nameliet
```

CONTROL CONTROL

```
PROGRAM POF
      Integer 11/28/ Hagical unit number for plotting device input integer 1s/21/ Hagical unit number for plotting device output
      integer te/1888/irecolution of plotter inc/inch
      Integer 14/8/ iplotter width in Inches
                       ichart type 8 = 1/2 from plot edge
      Integer It/8/
                                   1 = 8 " from plot edge
                       12-multipen plotter with 1" offeet
      Integer Ir/2/
      Integer 12/8/ 18≃run in immediate mode
      INTEGER NEWPEN/999/ 1999 FOR PLOT18/-3 FOR CRLCCHP
      INTEGER=2 BUFFER(IBUFSIZ) (ELEMENTS 9-16 OF FIRST RECORD CONTRIN
                        HIYR, IHO, IDAY, IHR, IMIN, ISEC, IHSEC, JDAY OF FIRST
                        ISAMPLE OF EVENT TRACE
      INTEGER IYR (HAXNTRC)
      INTEGER IND (MAXNTRC)
      INTEGER IDAY (HAXNTRC)
      INTEGER IHR (HAXNTRC)
      INTEGER IMIN(HAXNTRC)
      INTEGER ISEC (MAXNTRC)
      INTEGER INSEC (MAXNTRC)
      INTEGER JORY (HAXNTRC)
      INTEGER IDEC/2/ INUMBER OF DECIMAL PLACES
      BYTE BLK (2= IBUFSIZ) IBUFFER EQUIVALENT IN BYTES
      EQUIVALENCE (BUFFER(1), BLK(1))
      COMMON /CHARSIZ/PHT, IDEC, AFRCTOR
      NAMELIST/LISTIN/FLNAME.XINPLOT, YINPLOT, NSAMPLES, NTRACES,
                         STRTSAMP, FSVALUE, SAMPLERATE, 11, 10, 10,
                          IN. IL. IF. IZ. NEWPEN, EVTXIN, EVTYIN, PROBXIN,
                         PROBYIN, FREQUEX, FREQUIN, EVTHAX, EVTHIN, FAC.
                         SAMESCALE, NOLASSINT, IDEC, NPDFSMPS, AHT,
                         SCALETRC AFACTOR
   - READ IN NAMELIST
      HRITE(6, =1 'TYPE IN FILEWAHE CONTRINING POF NAMELIST'
      REPO (5,384) FLNAME (1)
      OPEN (UNIT=22, NAME=FLNAME (1).TYPE='OLD')
      ISTAT = 8
      READ (22, NML = ISTIN, IOSTRT=ISTRT)
C
      DO WHILE (ISTAT, EQ. 8)
C- READ IN TRACES
      CMIN = 188.8
      GMAX ≈ -188.2
      FMIN = 188.8
      FHAX ≈ -188,8
      IDEY ≈ 78
      DO I=1.NTRACES
        OPENI UNIT = IDEY.
                 NAME = FLNAME(I),
                 BLOCKSIZE = 2=18UFSIZ,
                 RECL = IBUFSIZ/2,
                 FORM = 'UNFORMATTED',
                 ACCESS = 'DIRECT',
                 TYPE = 'QLD')
         ISTAT = 0
         18LK = 1
         READ(IDEV'IBLK.IDSTAT = ISTAT) BLK
         IF (ISTAT.EQ. 8) THEN
           IYR(I) = BUFFER(9)
```

- CONTINUED -

TOTAL TENSOR SERVER LEGISLES TOTAL T

```
- CONTINUED -
PROGRAM PDF
```

```
INO([] = BUFFER([8])
         IDAY(I) = SUFFER(11)
         IHR(I) = BUFFER(12)
         IMIN(I) = BUFFER(13)
         ISEC(I) = BUFFER(14)
         IHSEC(I) = BUFFER(15)
         JOAY(I) = BUFFER(16)
         J = 8
         KK = 8
         IBLK = IBLK + 1
         REPO(IDEY'IBLK, IOSTAT = ISTAT) BLK
         DO WHILE (ISTAT. ED. 8. AND. J. LT. MAXLEN)
           DO L=1.IBUFSIZ
             K = J + L
             IF (K. GE. STRTSAMP (I) . AND . KK. LT. MAXLEN) THEN
               KK = KK + 1
                TRACE(KK.I) = BUFFER(L) = SCALETRC(I)
                IF (KK.LE.NPDFSHPS2) THEN
                 GMIN = ABS(TRACE(KK.I))
                 CHAX = MAXIGNAX, CMINI
                ENO IF
             END IF
            EN0 00
            J = J + IBUFSIZ
            IBLK = IBLK + 1
            READ(IDEY'IBLK. IDSTAT = ISTAT) BLK
          END DO INHILE (ISTAT. EQ. 8. RND. KK. LT. MAXLEN)
        DO IF
      CLOSE (UNIT=IDEV)
      END DO
C- DEFINE LIMITS
      GHIN = -GHAX
      IF (EVTHAX.GT.EVTHIN) THEN
         CHIN = EVTHIN
         SHAX = EVTHAX
      ENO IF
      IF (SAMESCALE) THEN
        EVTHAX = CHAX
        EVTHIN = CHIN
      ENO IF
    INITIALIZE PROBABILITY DENSITY CLASS INTERVALS TO ZERO
      DO I=1.NCLASSINT
         DO J=1.NTRACES
          PONOISE(I,J) = 8.3
          POSIGNAL(I.J) = 8.8
         END 00
       EN0 00
 C- CALCULATE CLASS INTERVAL USING CREATEST MINIMUM AND
 C- GREATEST MAXIMUM VELOCITY AMPLITUDES
 C- FOR THE FOUR EVENT TRACES PLOTTED
       CLASSINT = (CHAX - CHIN)/FLORT(NCLASSINT)
       FCYALUE = CHIN + CLASSINT/2.
     DRICULATE PROBABILITY DENSITY DISTRIBUTION FOR NOISE SAMPLE 1-MPDFSMPS
```

```
- CONTINUED -
      PROGRAM POF
      DO Jal.NTRACES
        DO 1=1.NPDFSMPS
          K=1
          CLASS = .FRLSE.
          DO HHILET. NOT. CLASS. AND. K. LE. NOLASSINT)
             IF (TRACE(I.J. GT. (K=CLASSINT+CMIN)) THEN
               K = K + 1
             ELSE
               PENDISE(K.J) = PONDISE(K.J) + 1.8
               CLASS = .TRUE.
             END IF
           END DO (MHILE (.NOT. CLASS)
         ENO 000
         DO I=: . NOLASSINT
           PONOISE(I.J) = PONOISE(I.J)/FLOAT(NPDFSHPS) | RELATIVE FREQUENCY
           FHAX = MAX (FMAX, PONDISE (I, J))
         END CO
       END 00
 C- CALCULATE PROBABLITY DENSITY DISTRIBUTION FOR SIGNAL-HIGISE
 C- NPDFSMPS+1 TO 2MMPDFSMPS
  ε
        NPDFSHPS2 = 2mmPDFSHPS
        DO J=1.NTRACES
          DO I=NPOFSHPS+1,NPDFSHP52
            K=1
            CLASS = .FALSE.
             DO WHILE ( . NOT . CLASS . AND . K . LE . NCLASS INT )
              IF (TRACE(I.J.GT. (K=CLASSINT+GM(NI) THEN
                K = K + 1
               ELSE
                 POSIGNAL(K.J) = POSIGNAL(K.J) + 1.8
                 CLASS = . TRUE.
               ENO !F
             END DO INHILE (. NOT. CLASS)
           EMB 200
           DC I = 1.HCLASSINT
             POSIGNAL (I.J) = POSIGNAL (I.J) /FLOAT (NPDFSMPS) | RELATIVE FREQUENCY
             PHRX = MAX (FMAX , POSTONAL (I , JT)
            ENG 00
          DAG 00/G
          IF (FREDMAX.LT.FREDMIN) THEN
            FREGMAX = FMAX
            FRESMIN = FHIN
          DID IF
    C- SET STATUS TO B AND THEN TRY TO READ ANOTHER NAME ST
          ISTAT = 8
          RESO (22, NML=LISTIN, IOSTAT=ISTAT)
          END DO INHILE (ISTAT.ED. 8)
          FORMAT (A16)
           ENO
```

```
SUBROUTINE POLR_A-BURG(X.N.COEF.IP)
          Modified from Group-Hoodward RRMA spectral analysis
C--
c—
      package for the IBM PC by a hayward 1-28-86
         This subroutine openutes the Burg coefficients os
      described by Koy and Marpie, (1982).
        Parameter (MaxOnder=58)
        Parameter (HaxPte=18888)
        Integer
                         N
                                             inumber of input pte
        Real
                                              Itime series inputs
        Real
                         R(MaxOnder)
        Rea!
                         C(MaxOrder+1)
        Real
                         $(MaxOrder+1)
        Rec
                         Coef (IP)
        Regi
                         AT (MaxUnder+1,2)
C
        DIMENSION Q1 (accepta) ,Q2 (accepta) ,P1 (accepta) ,P2 (accepta)
        If (N.gt. HaxPte) then
                Write(6,1) N
                Forest (1h8, 'Number of Input points exceeds soximus')
1
                 Heltolli, II N
                Stop 'Poir_ArBURG - error 1'
        End If
С
        CREATE A ZERO HEAN
C--
        SUM=8.8
        DO 1=1.N
                SUM=SUM+X ( I )
        End Do
        XBFR=SUM/N
C
        10=8
        INITIALIZE EQUATIONS-
C-
        IENC=+1
        00 J=1.1END
                       P1 (J) =X (J+1) -XBAR
                    SHEX LUIS LUID
        End Do
               THP=0
               IEND=N-1
               00 ≥2.1END
                    THP=THP+(X(J)-XBAR)=#2
        End Co
               5(1+8) = (X(1) - XBAR) mm2+(X(N) - XBAR) mm2+2+THP
               THP=0
               IENC=N-1
               DO J=1.IENO
                  THP=THP+PI(J) =QI(J)
        End Do
               R(1+8)=TMP
               C(1+8)=2=(R(1+8)/3(1+8))
               IENO=N-2
               DO JELIEND
                      P2(J)=P1(J+1)-C(1+8)=Q1(J+1)
                    92(J)=Q1(J)-C(1+8)≠P1(J)
        End do
               T=(1+C(1+8) mm2) mS(1+8) -4mC(1+8) mR(1+8)
                5(1+1)=T-(Q1(N-\partial-1)-C(1+\partial))=2(N-\partial-1))=2-(P1(1)-C(1+\partial))=2(1))=2
               DO N1=1.1P
                       IEND⇒+-2
```

```
- CONTINUED -
      SUBROUTINE POLE_P-BURGIX.N.COEF.IPI
                      DO IK=1. IEND
                             P1(IK)=2(IK)
                            Q1(IK)=Q2(IK)
                End do
                      THP=3
                       IEN0≈N-N1-1
                       DO K=1.IEND
                          THP=THP+P1(K)=Q1(K)
                 End do
                       R(1+N1)=TMP
                       THP=3
                       IENC=N-NI-1
                       DO K=1.IENO
                          THP=THP+(P1 (K) ==2+Q1 (K) ==2)
                 End Do
                 S(1+N1)=TMP
                       C(1+N1)=2#(R(1+N1)/S(1+N1))
                       IENC=N-H1-2
                       DO JEI, IENO
                               P2(J)=P1(J+11-C(1+N1)=Q1(J+1)
                                   あいまいい-((1+47)も)()
                        T=(1+C(1+N1)=+2)=S(1+N1)-4=C(1+N1)=R(1+N1)
                        T1=T-(Q1(N-N1-1)-C(1+N1)=P1(N-N1-1))==2
                    S(1+N1+1)=T1-(P1(1)-C(1+N1)=Q1(1))==2
         End Co
 C
                 --- NOH CALCULATE THE A'S---
                DO KI=1.IP
                     AT (K1.2)=1
                DO IOROER=1.IP
                        DO KI=1.10RDER
                           AT (KI.1) =AT (KI.2)
                  End do
                        NI=IORCER-1
                        IENC=N1+1
                         00 K=1.IEND
                                 IT=N1+1
                                 IF (IT. ED. K) AF IRST=8
                                 IF (IT. NE. K) AFTRST=AT (1+K,1)
                                 IT1=+1+1-K
                                 IF (ITI.ES. IT) ASECOND=8
                                 IF (IT1. ME. IT) ASECOND=AT (1+N1+1-K,1)
                            AT (1+K,2) =AFIRST-C(1+N1) =ASECOND
                   End do
           End do
                 DO IK=1.IP
                          DEF (IK) =-AT (IK+1,2)
            End do
                 RETURN
                 ENG
```

```
SUBROUTINE RANKFRED (N. DATH, RANK, FRED, RANKHAX)
    This submoutine computes the rank and frequency
of occurrence for each data point in an array of N
data values.
REAL DATA(H) ICONTRINS DATA
REAL RANK (N) ICONTAINS RANK
REAL FRED (N) ICONTAINS FREDUENCY
REAL RANKHAX
INITIALLY GIVE ALL N DATA POINTS A RANK OF 1.8
                        AND A FREELENCY OF 8.8
N.1=1 00
                     INITIALIZE RANK
  RRNK(1) = 1.8
                     HINITIALIZE FREQUENCY
  FREQ(1) = 0.8
END 00
CRLCULATE RANK OF DATA VALUE
           ANC
CHLCULATE FREQUENCY OF DATA VALUE
DO 1=1.N
  DO J=1.N
     IF (DATA(I).GT.DATA(J)) THEN
                                     LINCREASE RANK OF DATA
      RPNK(I) = RPNK(I) + 1.8
     DAD IF
     IF (DATA(I).EQ.DATA(J)) THEN
                                     LINCREASE FREQUENCY OF DATA
      FREQ(1) = FREQ(1) + 1.8
     END IF
   END DC
 END 30
 ADJUST RANK OF DATA ACCORDING TO FREQUENCY
   RNK(I) = RNK(I) + (FREQ(I)-1.8)/2.8
   RANKHAX = MAX (RANK(I) , RANKHAX)
  RETURN RANKS AND FREQUENCIES
  RETURN
  ĐΦ
```

```
SUBROLITINE KANKTEST (N. XDATA, YDATA, K. TEST, ACCEPT)
C- SUBROUTINE TO PERFORM RANK SUM TEST
C- MOODY GRAYBILL & BOES
C
      REAL XORTA(N)
      REAL YDATA(N)
      INTEGER N INLABER OF SAMPLES IN RUN TEST
      INTEGER & INCCEPT OR REJECT CRITERIA
      REAL ZIZBE.31 INORK ARRAY FOR ORDERING SAMPLES
      LOGICAL ACCEPT
     COMPUTE RPAK SUM TEST FOR WINDOWS OF SIZE N
       N2 = 28N
       N.1=1 00
        Z(I.1) = ABS(XDATA(II)
         Z(I+N,1) = ABS(YCATA(1))
       ENO CO
       CALL RANKFRED (N2. Z(1.1). Z(1.2), Z(1.3), RANKMAX)
       TY = 0.0
       DO 1=1.N
          J = 1 + N
          TY = TY + Z(J,2)
       END 00
 C
          C1 = FLOAT( N2 + 1)
          CZ = FLOAT (NaN)
  C
                                               IEXPECTED VALUE OF TH
          EXV = FLORT (N) =C1/2.
  ٤
                                                     IVARIANCE OF TU
          VAR = C2=C1/12.
  ¢
          TEST = ROS(TY - EXY)
  c
           IF (TEST. GE. FLORT (K)) THEN
                                         TREJECT THE NULL HYPOTHESIS
            ACCEPT = .FALSE.
           ELSE
                                         INCCEPT THE MULL HYPOTHESIS
            ACCEPT = .TRUE.
           END IF
         RETURN
         Đ.Q
```

```
SUBROUTINE RANK2788 (NSAMPLES, AMPTRO, TRO, N, MEAN, INC. TEST, SECLEN,
                     PATTERN, IST, RANGE, RHAX, RHIN, NRANK, RANKI
      This subroutine operates the rank sum test as described
 In the Introduction to the Theory of Statistics, Mood, Graybill
 and Bose, HaGrae-Hill, 3rd Edition, 1974. The observations,
 x(1), I=1.N, are taken from the representative background naise
 assumed to be present in the first 2.5 seconds of the event
 troos. The second group of observations, y(i), i=1,N, is
 chosen from a window of M samples moved incrementally along
 the event trace for each excution of the rank sum test.
      The rank our test statistic is computed for each window
 of observations, y(1), 1=1,N, and the representative noise
 observations, x(1), 1=1.N. Detection occurs when the rank
 eum of the y(11 ranks exceeds the threehold for probable
 notes. The automatic first arrival plak to the first
 zero procesing preceding the first modified slope value that
 exceeds the adxisus sadified slape value of the representative
 noise ecopies, x(i), i=1.N. occurring in the first detection
 REN TEST
 REAL RHAX
 REAL RHIN
 REAL PATTERN(1)
                    INULL HYPOTHESIS = 1.8 // ALTERNATIVE = 2.8
                     IARRAY CONTAINING FEATURE OF SIGNAL TO RANK
 REAL TRO(1)
 REAL AMPTRO(1)
                   TARRAY CONTAINING AMPLITUDES OF SIGNAL POINTS
 REAL SEGLEN(1)
                                      ILLENGTH OF PATTERN SECHENT
                                 LEXPECTED VALUE OF RANK SUM TEST
 REAL MEAN
 REGIL RANGE
                                        IRANK SUM RANGE OF VALUES
 INTEGER HSAMPLES
                          INUMBER OF OBSERVATIONS IN EVENT TRACE
  INTEGER HRANK
                   INUMBER OF RANK SUM VALUES COMPUTED FOR TRACE
  INTEGER INC
                              I INCREMENT WINDOW MOVES DOWN TRACE
 INTEGER 1ST
                                       IFIRST ARRIVAL PICK INDEX
 INTEGER H INLIMBER OF SAMPLES REFERENCE BACKGROUND NOISE HINDOM
 REAL T(388,3)
                     INORK ARRAY TO RANK HOISE & SIGNAL HINDOHS
 REAL RANK (1)
                                 ISUNATION TY OF RANKS OF SIGNAL
 LOGICAL FOUND/. FALSE./
 LOGICAL SECONO/. FALSE./
 LOGICAL FIRST/.TRUE./
COMPUTE RANK SUM TEST OVER TRACE
 N2 = 2=N
 RHAX = HERN
 RMIN = MEAN
 FLANCY = R. R
                   TEVENT NOISE MAXIMUM MODIFIED SLOPE MACHITUDE
STORE REFERENCE BACKGROUND NOISE IN HORK VECTOR
FIND EVENT HOISE MAXIMUM
 DO I=1.N
   L = I-1
    PRITTERN (NSRMPLES-L) = 1.8
    SECLEN (NSAMPLES-L) = 1.8
    PATTERN(I) = 1.8
    SEGLEN(1) = 1.8
    T(I_*1) = ABS(TRC(I=2-1))
    T(I+N.1) = RB5(TRC(I=2))
    ENHAX = MAX(T(I,1),ENMAX)
    ENHAX = HAX(T(I+N.1), ENHAX)
  END DO
  ENMAX = ENMAX = 1.81
                              ITHRESHOLD FOR FIRST ARRIVAL PICKER
```

ᡊᢛᡊᢛᡊᢛᡊᢛᡊᢛᡊᢛᢗᢛᢗᢛᢗᢛᢗᢘᢗᢘᢗᢘᢗᢘᢗᢘᢗᡑᢗᡑᢗᡑᠪᡑᡚᢘᢗᡑᠪᢘᡧᢘᢐᢘᢐᢘᢐᢐᢘᢐᠸᡊᡊᡑᠸᢐᠸᡑᡎᡑᢖᢘᡠᢘᠸᢘᠸᢘᠸᢘᠸᢘᡓᢘᡓᢘᡓᢘᡓᢘᡓᡑᡓᡑᡓᡑᡓᡑ᠘

C- COMPUTE RANK SUM FOR A HINDOH OF LENGHT N COMPARED

```
- CONTINUED -
SUBROUTINE RANK2788 (KSAMPLES, AMPTRC, TRC, N. HEAN, INC, TEST, SEGLEN,

C— NITH THE REFERENCE BACKGRUND NOISE MINDON OF LENGHT N

IR = 8

C— SLIDE THE MINDON DOWN THE TRRCE IN INCREMENTS = INC

D I=1. (KSAMPLES-H), INC

IR = IR + 1

D J=1, N

J = J + (I=1)

T(J+N, I) = BBS(TRC(JJ))

EDUD DO 151:N

CPUL RANKFREI(R2,T(1,1),T(1,2),T(1,3), RANKCHEX)

TY = 6.8

D J=1.N

TY = TY + T(J+N,2)

EDUD DO

RMAX = MAX(TY,RMAX)

RMAX SUM MAXIMUM VALUE

RMIN = MINITY,RMIN

RMAX SUM NINTHUM WELLE

RMAX SUM OF THE IRUN MINDON

DO DO 11

MARKE IR

RMANC = IR RANK SUM OF THE IRUN MINDON

C — MIN DETERMINE TARRESHIN ISTONEL AND NOISE RANK SUM INTERVAL

C — MIN DETERMINE TARRESHIN ISTONEL AND NOISE RANK SUM INTERVAL

C — MIN DETERMINE TARRESHIN ISTONEL AND NOISE RANK SUM INTERVAL

C — MIN DETERMINE TARRESHIN ISTONEL AND NOISE RANK SUM INTERVAL

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C — MIN DETERMINE TARRESHIN ISTONEL AND NOISE RANK SUM INTERVAL

C — MIN DETERMINE TARRESHIN ISTONEL AND NOISE RANK SUM INTERVAL

C — MIN DETERMINE TARRESHI
                                       RANGE = RMAX - RMIN
                                                                                                                          ISIGNAL AND NOISE RANK SUM INTERVAL
                            - NON DETERMINE THRESHOLD = 2788 + RMIN.OR.8.75=(RMRX-RMIN) + RMIN
                                        IF ((RMAX-RMIN).LT.3688.) THEN (LOW SIGNAL TO NOISE RANK SUMS
                                             TEST = (RMRX-RMIN) =2.75 + RMIN
                                       ELSE
                                             TEST = 2780. + RMIN
                                       ENO IF
                     C- FIND SIGNAL
                                       FIRST = .TRUE.
                                                                                                                          ILCOKING FOR FIRST DETECTION WINDOW
                                       DO I=1.NRANK
                                             K = (I-1) = INC + 1
                                              IF (RANK(I), GT. TEST) THEN IPROBABLE SIGNAL PRESENT IN HINDON
                                                   DO L= (K+(INC-1))
                                                         PATTERN(L) = 2.8
                                                                                                                                                                                                     IREJECT HE
                                                         SEGLEN(L) = 1.8
                                                   END 00
                                                  L = K
                                                   DO HHILE (FIRST. ANG. (L.LT. (K+(N-1))))
                                                         IST = L
                                                         IF (RBS(TRC(L)), GT, EWAX) THEN
                                                               MM = L - 1
                                                               DO WHILE (FIRST. PND. (MM. CE. 1))
                                                                     IST = MM
                                                                     IF (AMPTRO (MM+1) = AMPTRO (MM) . LT. 8.8) THEN
                                                                          FIRST = .FALSE.
                                                                                                                                                           IFOUND IST=ZERO CROSSING
                                                                     ELSE
                                                                                                                          ICONTINUE LOOKING FOR ZERO CROSSING
                                                                          MM = MM - 1
                                                                     END IF
                                                               END DO INHILE
                                                          FL.SE
                                                               L=L+1
                                                                END IF
                                                    ENG OO INHILE
                                                    IF (FIRST) THEN
                                                                                                                  TOBSERVATIONS IN HINDON NOT .GT. NOISE
                                                                MM = L - 1
```

```
- CONTINUED -
      SUBROUTINE RANK2788 (NSRMPLES.AMPTRC.TRC.N.HERN.INC.TEST.SEGLEN.
              DO WHILE (FIRST. AND. (MM. GE. 1)) | IPICK 1ST ZERO CROSSING
                                                 IBEFORE END OF HINDOH
                IF (AMPTRC (MH+1) =AMPTRC (MH) .LT.8.9) THEN
                  FIRST = .FRLSE.
                                              IFOUND IST=ZERO CROSSING
                ELSE
                  HM = HM - 1
                                  ICONTINUE LOOKING FOR ZERO CROSSING
                END IF
              END DO INHILE
          END IF
        ELSE
          DO L= (K+(INC-1))
            PATTERN(L) = 1.8
                                                            IACCEPT HE
            SEGLEN(L) = 1.8
          ENG DO
        ENG IF
      END DO 11
      RETURN
      END
```

```
SUBROUTINE RUNEXP(N.XDATA, YDATA, IZB, NRUNS, EXV, VAR)
C
           This subroutine opioulates the run test statistic
      for two arrays of data, x(i), i=1.N and y(i), i=1.N.
      The threshold, zil, for the number of runs below which we
      reject the null hypothesis, F(x) = F(y), le provided os
      input to the subroutine. It is calculated outside the
      subroutine using the normal distribution and a specified
      aipha level. The algorithm is based on the description of
      the run test given in the Introduction to the Theory of
      Statistics, Mood, Graybill and Boss, MaGraw-Hill, 3rd Edition,
      1974.
      REAL XDATA(N)
      REAL YDATA (N)
      REAL EXV
                                   HOSPECTED VALUE OF NUMBER OF RUNS
      REAL VAR
                                         IVARIANCE OF NUMBER OF RUNS
      INTEGER N
                                       INUMBER OF SAMPLES IN RUN TEST
      INTEGER IZE
                                          INCCEPT OR REJECT CRITERIA
      REAL Z(288.2)
                                    HORK ARRAY FOR ORDERING SAMPLES
     COMPUTE RUN TEST OVER TRACE
      N2 = 2mN
      DO 1=1.N
       Z(I_2I) = ABS(XDATA(I))
        Z(1.2) = 1.8
        Z(I+N-1) = RBS(YDATA(I))
       Z(1+N,2) = -1.8
      END 00
     DRDER COMBINED, X(I) AND Y(I), DATA INTO ASCENDING ORDER
      DO I=1.N2-1
        DO J=1+1.N2
          IF (Z(J,1).LT.Z(I,1)) THEN
            TEMP = Z(1,1)
            Z(1,1) = Z(J,1)
            Z(J,1) = TEMP
            TEMP = Z(1,2)
            Z(1,2) = Z(J,2)
            Z(J,2) = TEMP
          ENO IF
        END DO
      END D0
      COUNT THE HUMBER OF RUNS
      ZSIGN = 2(1,2)
      NRUNS = 1
      DO 1=2,N2
          IF ((ZSION=Z(1,2)),LT.8.8) THEN
            ZSIGN = Z(I,2)
            NRUNS = NRUNS + 1
          END IF
      END 00 11
     COMPUTE EXPECTED VALUE AND VARIANCE OF THE NUMBER OF RUNS
     ASSUMING THE NULL HYPOTHESIS IS TRUE
        CI = FLORT (2mm)
        C2 = FLOAT (N2)
        EXV = C1/C2 + 1.
        VAR = (C1=(C1-C2))/(C2==2=(C2-1.))
      RETURN
      ENO.
```

```
"JERGUTINE SLOPE (NSAMPLES, TRC. S)
          This subroutine occourse the integrated slope of
     a selecte eignal for discrete segments of the eignal
     with continuous slope direction.
     Kathleen R. Alden, March 16, 1986.
     REAL S(1)
                                        IMODIFIED SLOPE OF OBSERVATION
                      ITRACE CONTRINING SEISHIC SIGNAL OBSERVATIONS
     REAL TRC(1)
     INTEGER NSAMPLES
                                      INUMBER OF OBSERVATIONS IN TRACE
C- ASSUME SAMPLE MEAN OF TRACE INPUT IS ZERO
  - COMPUTE SEISHIC TRACE HODIFIED SLOPE
     $(1) = TRC(2) - TRC(1)
                                                  ISLOPE OF 1ST SAMPLE
     DO I =2.NSRMPLES - 1
       S(I) = ((TRC(I+1)-TRC(I))+(TRC(I)-TRC(I-1)))/2.
                                                            II Lh SLOPE
        IF (S(I) =S(I-1),GT.8.8) THEN
                                       ISLOPE DIRECTION IS CONTINUOUS
         S(I) = S(I) + S(I-I)
                                                      I INTEGRATE SLOPE
       ENO IF
     ENO 00
     $(NSAMPLES) = TRC(NSAMPLES) - TRC(NSAMPLES-1) ISLOPE LAST SAMPLE
     RETURN
     Đ.O
```

SUBROUTINE SIGNED (N. XDATA, YDATA, K. SN. FXV. TEET)

```
Ç
           This subroutine calculates the sign test statistic
      for the absolute value of two arrays of data, x(1), 1=1,N and
      gill, I=L/N. The threshold, k, which sust be less than the
      test statistic to reject the null hypothesis, F(x) = F(y), is
      provided as input to the subroutine. It is calculated outside
      the subroutine using the normal distribution and a specified
      sighs level. The signrithm is based on the description of
      the eign test given in the Introduction to the Theory of
      Statistice, Mood, Graybili and Base, McGrae-Hill, 3rd Edition,
      1974.
       REFL XCATAIN)
       REAL YDATA (N)
                                     ITHRESHOLD FOR THE TEST STATISTIC
       INTEGER K
                                      INUMBER OF SAMPLES IN SIGN TEST
       INTEGER N
                     ISUMPTION OF THE POSITIVE SIGNS OF (Y(I) - X(I))
       REAL SN
                                PEXPECTED VALUE OF SUMMATION OF SIGNS
       REAL EXY
                                         IABSOLUTE VALUE OF (SN - EXV)
       REAL TEST
 C- COMPUTE SIGN TEST OVER TRACE
         SN = 8.8
         DO I=L.N
           IF (ABS(XDATA(I)).LT.ABS(YDATA(I))) THEN
             SN = SN + 1.8
           ENG IF
          ENO 00
  C- COMPUTE EXPECTED VALUE OF SUMMATION
      AND THE TEST STRTISTIC
          EXY = FLOAT (N) /2.
         TEST = ABS(SN - EXV)
        RETURN
        ENO
```

SSENSE TO SERVICE SERVICE SERVICE DESCRIPTION OF SERVICE SERVICES
APPENDIX C

Values of w(alpha,M,N)

TABLE C.1 (Continued)

n = 2

r	m = 3	m = 4	m = 5	m = 6	m = 7	m = 8	m = 9	m = 10	m = 11
6	.600								
7	.400	00ک							
8	.200	.400	.571						
9	.100	.267	.429	.571					
10		.133	.286	.429	.556				
11		.067	.190	.321	.444	.556			
12			.095	.214	.333	.444	.545		
13			.048	.143	.250	.356	.455	.545	
14				.071	.167	.267	.364	.455	.538
15				.036	.111	.200	.291	.379	.462
16					.056	.133	.218	.303	.385
17					.028	.089	.164	.242	.321
18						.044	.109	.182	.256
19						.022	.073	.136	.205
20							.036	.091	.154
21							.018	.061	.115
22								.030	.077
23								.015	.051
24									.026
25									.013

n = 2

I	m = 12	m = 13	m = 14	m = 15	m = 16	m = 17	m = 18	m = 19	m = 20
15	.538								
16	.462	333 گ							
17	.396	.467	.533						
18	.330	.400	.467	.529					
19	.275	.343	.408	.471	.529				
20	.220	.286	.350	.412	.471	.526			
21	.176	.238	.300	.360	.418	.474	.526		
22	.132	.190	.250	.309	.366	.421	.474	.524	
23	.099	.152	.208	.265	.320	.374	.426	.476	.524
24	.066	.114	.167	.221	.275	.327	.379	.429	.476
25	.044	.086	.133	.184	.235	.287	.337	.386	.433
26	.022	.057	.100	.147	.196	.246	.295	.343	.390
27	.011	.038	.075	.118	.163	.211	.258	.305	.351
28		.019	.050	.088	.131	.175	.221	.267	.312
29		.010	.033	.066	.105	.146	.189	.233	.277
30			.017	.044	.078	.117	.158	.200	.242
31			.008	.029	.059	.094	.132	.171	.212
32				.015	.039	.070	.105	.143	.182

TABLE C.1 (Continued)

n = 2

x	m = 12	m = 13	m = 14	m = 15	m = 16	m = 17	m = 18	m = 19	m = 20
33 34 35 36 37				.007	.026 .013 .007	.053 .035 .023 .012 .006	.084 .063 .047 .032	.119 .095 .076 .057	.156 .130 .108 .087
38 39 40 41 42 43							.011 .005	.029 .019 .010 .005	.052 .039 .026 .017 .009
				п	= 3				
x	m = 3	m = 4	m = 5	m = 6	m = 7	m = 8	m = 9	m = 10	m = 11
11 12 13 14 15 16 17 18 19 20 21 22 23 24 25 26 27 28 29 30 31 32 33 34 35 36 37 38 39	.500 .350 .200 .100 .050	.571 .429 .314 .200 .114 .057 .029	.500 .393 .286 .196 .125 .071 .036 .018	.548 .452 .357 .274 .190 .131 .083 .048 .024	.500 .417 .333 .258 .192 .133 .092 .058 .033 .017	.539 .461 .388 .315 .248 .188 .139 .097 .067 .042 .024 .012	.500 .432 .364 .300 .241 .186 .141 .105 .073 .050 .032 .018	.531 .469 .406 .346 .287 .234 .185 .143 .108 .080 .056 .038 .024 .014 .007	.500 .442 .385 .330 .277 .228 .184 .146 .113 .085 .063 .044 .030 .019 .011

TABLE C.1 (Continued)

n = 3

	η = 3									
*	m = 12	m = 13	m = 14	m = 15	m = 16	m = 17	m = 18	m = 19	m = 20	
24	.527									
25	.473									
26	.420	00ک								
27	.367	.450	.524							
28	.316	.400	.476							
29	.268	.352	.429	.500	•					
30	.224	.305	.384	.456	.521					
31	.182	.261	.338	.412	.479					
32	.147	.220	.296	.369	.438	.500				
33	.116	.182	.254	.327	.396	.461	.519			
34	.090	.148	.216	.287	.356	.421	.481			
35	.068	.120	.181	.249	.317	.382	.444	.500		
36	.051	.095	.150	.213	.280	.345	.407	.464	.517	
37	.035	.073	.122	.180	.244	.308	.370	.429	.483	
38	.024	.055	.099	.151	.211	.273	.335	.394	.449	
39	.015	.041	.078	.125	.180	.239	.300	.359	.415	
40	.009	.029	.060	.102	.152	.208	.267	.325	.382	
41	.004	.920	.046	.082	.127	.179	.235	.293	.349	
42	.002	.012	.034	.065	.105	.153	.206	.262	.317	
43		.007	.024	.050	.086	.129	.178	.232	.286	
44		.004	.016	.038	.069	.108	.153	.204	.257	
45		.002	.010	.028	5دَ٥.	.089	.131	.178	.229	
46			.006	.020	.042	.073	.111	.154	.202	
47			.003	.013	.032	.059	.092	.132	.177	
48			.001	.009	.024	.046	.077	.113	.155	
49				.005	.017	.036	.062	.095	.134	
50				.002	.011	.027	.050	.080	.115	
51				.001	.007	.020	.040	.066	.098	
52					.004	.014	.031	.054	.083	
53					.002	.010	.023	.044	.069	
54					.001	.006	.017	.034	.058	
55						.004	.012	.027	.047	
56						.002	.008	.020	.038	
57						.001	.005	٥١٥.	.030	
58							.003	.010	.023	
59							.002	.007	.018	
60							.001	.005	.013	
16								.003	.009	
62								.001	.006	
63								.001	.004	
64									.002	
65						••			.001	
66									.001	

TABLE C.1 (Continued)

x	m = 4	m = 5	m = 6	m = 1	m = 8	m = 9	m = 10	m = 11
18	.557							
19	.443							
20	.343	.548						
21	.243	.452						
22	.171	.365	.543					
23	.100	.278	.457					
24	.057	.206	.381	.536				
25	.029	.143	.305	.464				
26	.014	.095	.238	.394	.\$33			
27		.056	.176	.324	.467			
28		.032	.129	.264	.4()4	.530		
29		.016	.086	.206	.341	.470		
30		.008	.057	.158	.285	.413	.527	
31			.033	.115	.230	.355	.473	
32			.019	.082	.184	.302	.420	.525
33			.010	.055	.141	.252	.367	.475
34			.005	.036	.107	.207	.318	425
35				.021	77ن.	.165	.270	.377
36				.012	.055	.130	.227	.330
37				.006	.036	.099	.187	.236
38				.003	.024	.074	.152	.245
39					.014	.053	.120	.206
40					800.	.038	.094	.171
41					.004	.025	.071	.140
42					.002	.017	.053	113
43						.010	.038	.089
44						.006	.027	.069
45						.003	.018	.05 2
46						.001	.012	.039
47							.007	.028
48							.00 4	.020
49							.002	.013
50							.001	.009
51								003
52								.00.1
53								.00.
54								00.

TABLE C.1 (Continued)

n = 4

x	m = 12	m = 13	m = 14	m = 15	m = 16	m = 17	m = 18	m = 19	m = 20
34	.524								
35	.476								
36	.431	.522							
37	.385	.478							
38	.342	.435	.521						
39	.299	.392	.479						
40	.260	.352	.439	.519					
41	.223	.312	.399	.481	• • •				
42	.190	.274	.360	.443 -	.518				
43	.158	.239	.323	.405	.482	617			
44	.131	.206	.287	.368	.446	.517			
46	.106	.175 .148	.253 .221	. 3 32 . 2 98	.410	.483	6 16		
47	.085 .066	.123	.191	.265	.375 .341	.449 .415	.516 .484		
48	.052	.101	.164	.235	.308	.381	.451	.516	
49	.032	.082	.139	.205	.277	.349	.419	.484	
50	.029	.062	.116	.179	.247	.318	.387	.453	.\$15
51	.021	.051	.096	.154	.219	.287	.356	.422	4×5
52	.015	.039	.079	.131	.192	.258	.326	.392	455
53	.010	.030	.063	.110	.168	.231	.297	.363	426
54	.007	.022	.051	.092	.145	.205	.269	.334	.347
55	.004	.016	.040	.076	.124	.181	.242	.306	368
56	.002	.011	.031	.062	.106	.158	.217	.279	.341
57	100.	.008	.023	.050	.089	.138	.193	.253	.314
58	.001	.005	.017	.040	.074	.119	.171	.228	288
59		.003	.012	.031	.061	.101	.150	.205	.262
60		.002	.009	.024	.050	.086	.131	.183	.239
61		.001	.006	.018	.040	.072	.113	.162	.216
62		.000	.004	.014	.032	.060	.098	.143	144
63			.002	.010	.025	.049	.083	.125	.174
64			.001	.007	.019	.040	.070	.109	.155
65			.001	.nus	.015	.032	.059	.094	.137
66			.000	.003	.011	.026	.(149	.081	.120
67				.002	800.	.020	.040	.069	105
68				.001	.006	.016	.033	.058	.1191
69				100.	.004	.012	.027	.049	.079
70				.000	.002	.009	.021	.041	.067
71					.001	.006	.017	.033	.057
72					.001	.005	.013	.027	,048
73					.000	.003	.010	.022	.041
74					.000	.002	.007	.018	.034
75						.001	.005	.014	.028
76						100.	.004	.011	.023
77						.000	.002	.008	810,

TABLE C.1 (Continued)

a	4

<u> </u>	m = 12	m = 13	m = 14	m = 15	m = 16	m = 17	m = 18	m = 19	m = 20
78						.000	.002	.006	.015
79							.001	.004	.011
80							.001	.003	.009
81							.000	.002	.007
82							.000	.001	200.
83								.001	.004
84								on o.	.003
85								.000	.002
86								.000	.001
87									100.
88									.000
89									oun.
90									.000

1 * 5

X	m = 5	m * 6	m = 7	m = 8	m = 9	m =)(
28	.500					
29	.421					
30	.345	.235				
31	.274	.465				
32	.210	.396				
33	.155	.331	.\$00			
34	.111	.268	.438			
35	.075	.214	.378	.\$28		
36	.048	.165	.319	.472		
37	.028	.123	.265	.416		
38	.016	.089	.216	.362	.500	
39	.008	.063	.172	.311	.449	
40	.004	.041	.134	.262	.399	.523
41		.026	.101	.218	.350	.477
42		.015	.074	.177	.303	.430
43		.003	.053	.142	.259	.384
44		.DO4	.037	.111	.219	.339
45		.002	.024	.085	.182	.297
46			.015	.064	.149	.257
47			.009	.047	.120	.220
48			.005	.033	.095	.185
49			.003	.023	.073	.155
50			.001	:015	.056	127
51				e00.	3140.	.103
52				.005	.030	.082

CONTRACTOR PROSPERSON PROPERTY PROPERTY.

TABLE C.1 (Continued)

			n = 5			
x	m =5	m = 6	m = 7	m = 8	m = 9	m = 10
53				.003	.021	.065
54				.002	.014	.050
55				100.	.009	.038
56					.006	.028
57			,		.003	.020
58					.002	.014
59					.001	.010
60					.000	.006
61						.004
62						.002
63						.001
64						.001

.000

		n	= 6		
*	m = 6	m = 7	m = 8	m = 9	m = 10
39	.531				
40	.469				
41	.409				
42	.350	.527			
43	.294	.473			
44	.242	.418			
45	.197	.365	.525		
46	.155	.314	.475		
47	.120	.267	.426		
48	.090	.223	.377	:523	
49	.066	.183	.331	.477	
. 50	.047	.147	.286	.432	
51	.032	.117	.245	.388	.521
52	.021	.090	.207	.344	.479
53	.013	.069	.172	.303	.437
54	.008	.051	.141	.264	.396
55	.004	.037	.114	.228	.356
56	.002	.026	.091	.194	.318
57	100.	.017	.071	.164	.281
58		110.	.054	.136	.246
59		.007	.041	.112	.214
60		.004	.030	.091	.184
61		.002	.021	.072	.157
62		.001	.015	.057	.132
63		.001	.010	.044	.110

TABLE C.1 (Continued)

n = 6

•	m = 6	m = 7	m = 8	m = 9	m = 10
64			.006	.033	.090
65			.004	.025	.074
66			.002	.018	.059
67			.001	.013	.047
68			.001	.009	.036
69			.000	.006	.028
70				.004	.021
71				.002	.016
72				.001	.011
73				.001	.008
74				.000	.005
75				.000	.004
76					.002
77					.001
78					.001
79					.000
80					.000
81					.000

n = 7

" - 1				
x	m = 7	m = 8	m = 9	m = 10
53	.500			
54	.451			
55	.402			
56	.355	.522		
57	.310	.478		
58	.267	.433		
59	.228	.389		
60	.191	.347	.500	
61	.159	.306	.459	
62	.130	.268	.419	
63	.104	.232	.379	.519
64	.082	.198	.340	.481
65	.064	.168	.303	.443
66	.049	.140	.268	.406
67	.036	.116	.235	.370
68	.027	.095	.204	.335
69	.019	.076	.176	.300
70	.013	.060	.150	.268
71	.009	.047	.126	.237
72	.006	.036	.105	.209

TABLE C.1 (Continued)

		n = 7				n •
x	m = 7	m = 8	m = 9	m = 10	x	m = 8
73	.003	.027	.087	.182	80	.117
74	.002	.020	.071	.157	81	.097
75	.001	.014	.057	.135	82	030.
76	.001	.010	.045	.115	83	.065
77	.000	.007	.036	.097	84	.052
78		.005	.027	.081	85	.041
79		.003	.021	.067	86	.032
80		.002	.016	.054	87	.025
81		.001	.011	.044	88	.019
82		100.	.008	.035	89	.014
83		.000	.006	.028	90	.010
84		.000	.004	.022	9 l	.007
85			.003	.017	92	.005
86			.002	.012	93	. 0 03
87			.001	.009	94	.002
88			.001	.007	95	100.
89			.000	.005	96	.001
90			.000	.003	97	.001
91			.000	.002	98	.000
92				.002	99	.000
93				.001	100	.000
94				.001	101	
95				.000	102	
96				.000	103	
97				.000	104	
98				.000	105	
					106	
	-				107	

87	.025	.084	.180
88	.019	.069	.158
89	.014	.057	.137
90	.010	.046	.118
91	.007	.037	.102
92	.005	.030	.086
93	.003	.023	.073
94	.002	810.	1 60.
95	.001	.014	.051
96	.001	.010	.042
97	.001	.008	.034
98	.000	.006	.027
99	.000	.004	.022
100	.000	.003	17ن.
101		.002	.013
102		.001	010.
103		100.	800.
104		.000	.006
105		.000	<i>-</i> 00∙
106		.000	.003
107		.000	.002
108		.000	.002
109			.00≀
110			.001
111			.000
112			.000
113			- .000
114			.000
115			.000
116			.000
			 -

m = 10

.381 .348

.317

.286 .257 .230 .204

m = 9

.212

.161

n = 8			
<i>x</i>	m = 8	m = 9	m = 10
68	.520		
69	.480		
70	.439		
71	.399		
72	.360	.519	
73	.323	.481	
74	.287	.444	
75	.253	.407	
76	.221	.371	.517
77	.191	.336	.483
78	.164	.303	.448
79	.139	.271	.414

TABLE C.1 (Continued)

	n = 9	,		п = 9		n	- 10
	m = 9	m = 10	x	m = 9	m = 10	<u> </u>	m = 1
86	.500	·	122	.000	.004	121	.12
87	.466		123	.000	£00£	122	.10
88	.432		124	.000	.002	123	.09
89	.398		125	.000	.001	124	.08
90	.365	.516	126	.000	.001	125	.07
91	.333	.484	127	•	.001	126	.06
92	.302	.452	128		.000	127	.05
93	.273	.421	129		.000	128	.04
94	.245	.390	130		. 0 00	129	.03
95	.218	.360	131		. 0 00	130	.03
96	.193	.330	132		.000	131	.02
97	.170	.302	133		. 0 00	132	.02
98	.149	.274	134		. 00 0	133	.01
99	.129	.248	135		.000	134	.01
00	.111	.223				135	.01
01	.095	.200				136	.00
02	.081	.178		n = 10	1	137	.00
03	.068	.158	-			138	.00
04	.057	.139	1		• 10	139	.00
05	.047	.121	-			140	.00
06	.039	.106	1	105 .5	15	141	.00
07	.031	.091			85	142	.00
08	025	.078			56	143	.00
09	.020	.067			27	144	.00
10	.016	.056			98	145	.00
11	.012	.047			70	146	.00
12	.009	.039			42	147	.00
13	.007	.033			15	148	.00
14	.005	.027			89	149	.00
15	.004	.022			64	150	.00
16	.003	.017			141	151	.00
17	.002	.014			218	152	.00
18	.001	.011			97	153	.00
19	.001	.009			76	154	٠.00
20	.001	.007			57	155	.00
121	.000	.005			40		

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Adapted from Table B of A Nonparametric Introduction to Statistics, by C. H. Kraft and C. van Eeden, Macmillan, New York, 1968, with the permission of the authors and the publisher. Copyright © 1968, by the Macmillan Company.

APPENDIX D

POSSESSE BOSSESSE BOSSESSE BOSSESSE BOSSESSE TO

Illustration of the Results of the Automatic Detector and Picker When Run on the Event Traces Recorded by the Four Seismic Stations, Lajitas, Marathon, Shafter and Tres Cuevas for 38 Seismic events

TABLE D.1 Event File Names, Errors (Automatic Minus Analyst Picks in Seconds), and the Range of Rank Sums for the 38 Events Used to Test the RANK2700 Detector

FILE NAME	ERROR IN AUTOMATIC PICK	RANGE OF RANK SUMS
EVENT 1		
L42891418.237	-8.1	4928.5
M42891418.853	8,2	5475.5
342891418.837	8.8	5057.0
742891418.834	-8.1	5348.5
EVENT 2		
L42891719.858	1.2	4166.5
M42891719.857	1.7	3826.2
342891719.825	7.4	4885.5
T42891719.847	1.9	4242.5
EVENT 3		
L42891815.818	5.5	4178.8
M42891815.838	9.6	3481.5
342891815.818	18.4	3472.5
742891814.846	1.2	4612.8
EVENT 4		****
L 42891986, 232	-17.5	3968.5
M42891986,848	2.7	3349.5
542891983.818	-15.1	3361.8
T42891985,888	6, 3	3888.2
EVENT 5		
L42892384,839	-19.9 -5.1	5576.5 4193.5
M42892384.828	-6.3	4193.3 5496. <i>8</i>
342892384.812	0. 2	5742.8
742892384.816 EVENT 6	0 ,	3/90.2
L42892313.883	8.2	4882.5
M42892313.828	-15.3	4768.5
542892313.016	2.1	5454.8
742892313.008	8.6	\$583.5
EVENT 7	2.0	44414
L42892327.016	-1.8	4615.5
N42892328,887	9,1	\$311.5
342892327.818	-8.8	4896.2
142892327.912	8.3	5168.8
EVENT 8		
1.42900019.851	-8.2	4862.0
M42988819.853	9.1	45 79.5
\$42900021.023	8.2	4569.5
T42900019.051	-8. 1	5564.5
EVENT 9		
L42988136.253	-25.8	4227.5
H42900137.819	12.8	4286.8
542900136.031	9. 1	3581.8
142988136.821	8.3	4554.5
EVENT 18	* *	110F F
1.42981821.816 M42981828.847	7.5 a .3	4485.5
542901821,025	-8.4	4778.8 4187.8
T42901828.837	-9.1	3232.0
EVENT 11	2	J
L42992854.844	-8.2	5584.8
M42982855.838	9.1	4294.0
342982854.851	-26. <i>8</i>	6894.8
742982854.844	8.2	5984.8
EVENT 12		•
L42902209.027	3. 9	2868.8
M42902218.022	2. 7	4852.8
\$42902211.000	5.1	2888. 2
142982789.838	8.3	3713.8

AND DEFENSE SERVICES TO SERVICES AND SERVICE

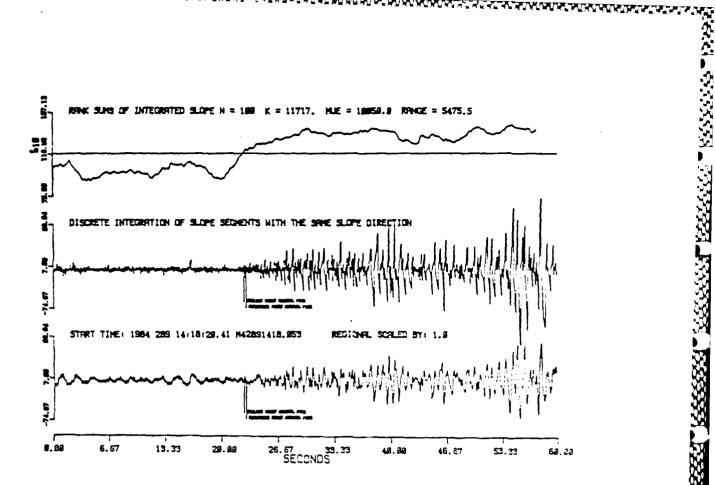
TABLE D.1 (Continued)

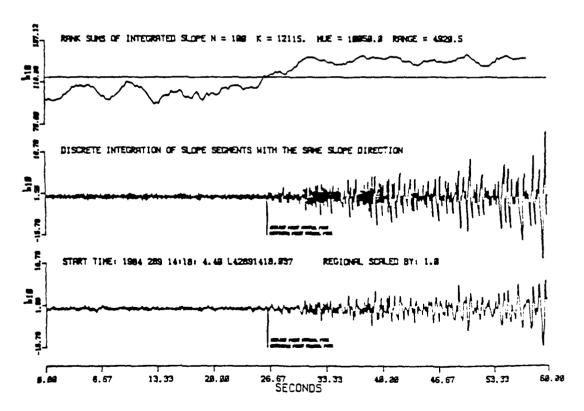
FILE NAME	ERROR IN AUTOMATIC PICK	RANGE OF RANK SUMS
EVENT 13		
L42982247.813	-3.6	4145.8
M42982246.829	4.4	4954.5 2907.0
542982247.816	5.4	2367.6 3742.6
T42982246.838	5.9	3/42.0
EVENT 14		5958.2
L42918325.832	1.4	5548.8
M42918325.826	8. 7	5376.2
\$42918325.816	-22.9	5643.2
T42910325.029	₽.5	30.00.2
■ EVENT 15	-1.5	5457.2
L42918922.842	8.2	4927.B
M42918922.848	8. 1	5262.5
542918922.835	9.1	5462.8
742919922.848	9. •	
EVENT 16	-33.3	4522.8
L42911682.887	8.9	3788.2
M42911682.838	6.2	4538.5
542911603.000	8. B	5197.5
T42911682.883 EVENT 17	2	
L42911887.829	8.4	45 72.5
M42911807.829	8.4	4144.5
542911807.025	8.3	4369.2
T42911807.032	-8. 3	4579.5
EVENT 18		
L42911819.839	-5.5	3839.5
M42911819.835	-23.9	3895.5
542911819.838	8.1	3995.5
742911819.839	8. 2	4139.5
EVENT 19		100C E
L42911851.803	8. 6	4626.5
M42911851.851	6.2	3742.3
542911851.811	-2.4	564 6.8 58 59.5
742911851.807	0. 3	2023.3
EVENT 28		4398.8
L42911959.817	-3. 7	4199.0
M42911959.808	13.4	6332.8
\$42911958.258	-6.3	4692.5
T42911958.256	9.3	4002.0
EVENT 21	8.4	5738.8
L42912148.815	8.1	5149.8
M42912148.848	18.4	3748.8
\$42912139.030	-2.9	5518.5
T42912148.815	- 2.0	
EVENT 22	-2.4	4443.5
L42912213.857 M42912214.852	2.1	4267.3
542912215.281	-23.7	5845.8
T42912213.857	e. 1	4814.5
EVENT 23		
L42928528.888	-0. 1	5145.5
M42928528.884	-3. 4	5799.5
542923528.881	9. 1	4184.8
T42928528.888	-1.2	5536.8
EVENT 24	= -	22,95.8
L42921429.022	9.5	4357.8
M42921430.014	0.3	3994.8
\$42921438.838	-8.2	3944.8
742921429.818	-0.3	₩ 41. p
EVENT 25		5469.8
L42921533.036		5871.2
M42921533.023	8.1 8.1	5261.5
542921533.82	•	5751.5
742921533.83	-w. w	

TABLE D.1 (Continued)

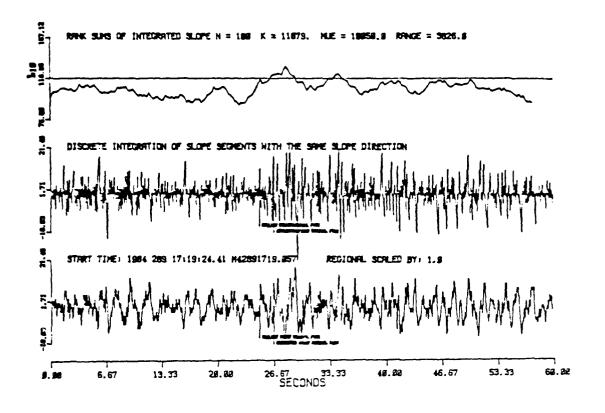
FILE NAME	ERROR IN AUTOMATIC PICK	rance of rank sums
5/5/5 35		
EVENT 26 L42921829.839	9.3	4112.5
M42921832.888	8.8	39 58.5
542921829.823	-2.3	5681.5
T42921829.839	-1.2	4849.5
EVENT 27		T
L42921859.83!	g. i	3998.8 4645.5
M42921859.854	-23.6	90 93.5 3923.5
542921858.856	2.3	4275.6
T42921859.028	8.1	
■ EVENT 28	-2.8	5 797.5
L42922157.814	-4.9	5198.5
M62522157.814	4,8	3101.5
\$42322157.814	9,6	5378.9
T#2922:57.817	• • •	
L42922384.842	-2. 8	5178.8
M42922384.835	-1.6	5616.5
542922384.839	13.1	4682.5
742522384.835	1.9	4453.£
EVENT 38		
L429381 48.826	0. i	4853.8
M429381 48. 229	9.4	5276.5
542939148.223	8.3	5258. 5
T429301 48.025	8.1	46 48 . 8
■ EVENT 31		5375.5
L42931458.814	-21.4	5161.8
M42931449.847	-0.1	7672.8
\$42931450.010	-8.3	5668.9
T42931458.814	-1.5	•
EVENT 32	-8.1	5495.5
L42932129,889	-0.1 -0.5	5878.8
M42932129.031	8, 5	6385.5
\$42932129,818 742932123,886	9. 1	5797.9
142932129.000 EVENT 33		
L42932383.219	8.2	5848.8
M42932303.235	8.2	588 9.5
\$42932303.035	-1.2	5628.5
742932303.019	9.9	5 323. <i>8</i>
EVENT 34		
142951988.857	8. 3	5912.5
M42951988.852	9. 2	5725.8 5251.5
542951989.882	-1.1	5231.5 5539.5
T42951998.857	1.8	2243.4
EVENT 35		3729.5
L42962837.216	6.1 6.4	4733.5
M42962838.881	6. 8	4264.5
542962238.214	1.5	3655. 3
T42962237.216 EVENT 36		
L42978815.846	-1.1	6.1889
M42978815.843	-8 .9	6135.5
542978815.839	6. 1	5989.5
142978815.846	8.9	5751.5
EVENT 37		**
L42971753.818	8.2	5272.0 5639.5
#42971753. # 30	8.2	5039.5 6174.8
542971753.818	0.8	5174. E 5585. B
T42971752.252	3.3	guest v
EVENT 38	1.2	4583.₽
L42971952.007		2483.5
M42971951.888	1	4355.5
\$42971958.848	9.9	3419.5
142971951.000	nto lieed as Pa	rt of the Training

T42371351.887 3.3 3.413.3 * Denotes the Events Used as Part of the Training Set.

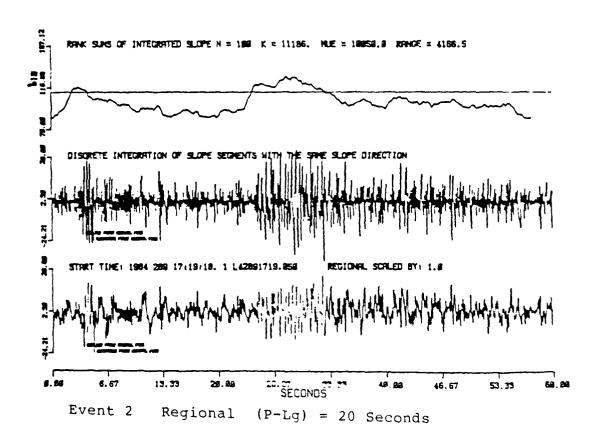




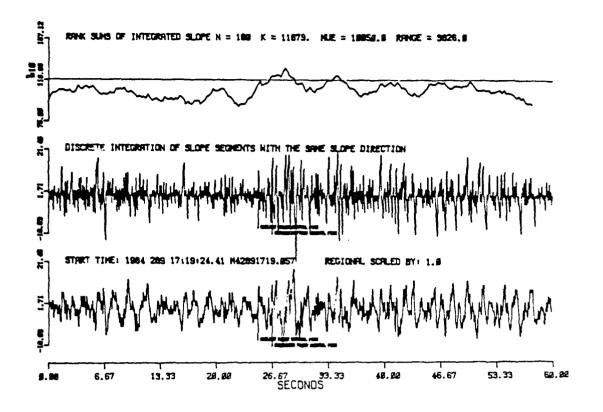
EVENT 1 - Regional (P-Lg) = 2 Minutes

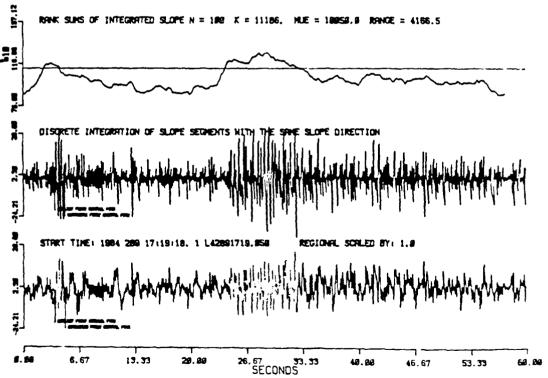


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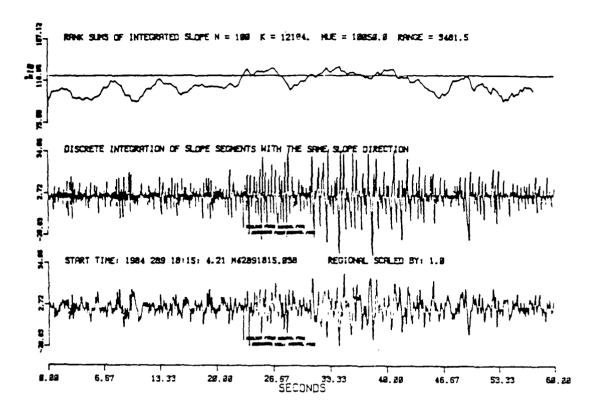


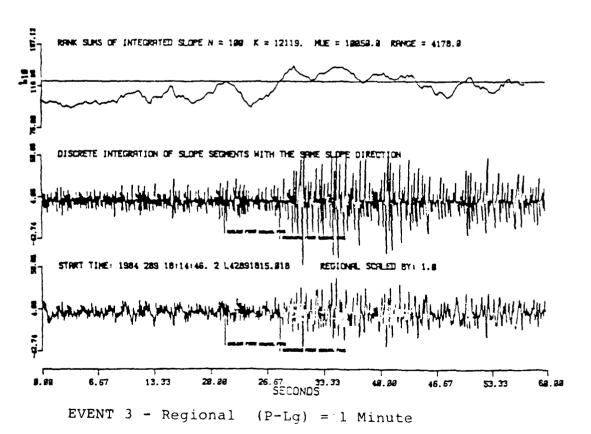
1.48

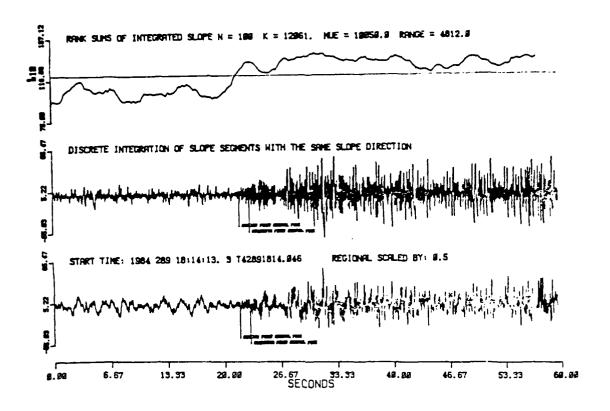


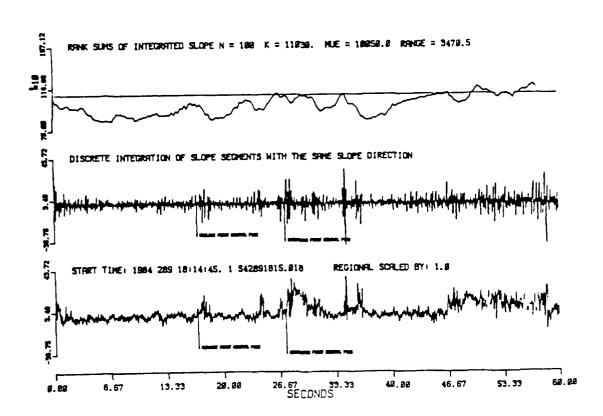


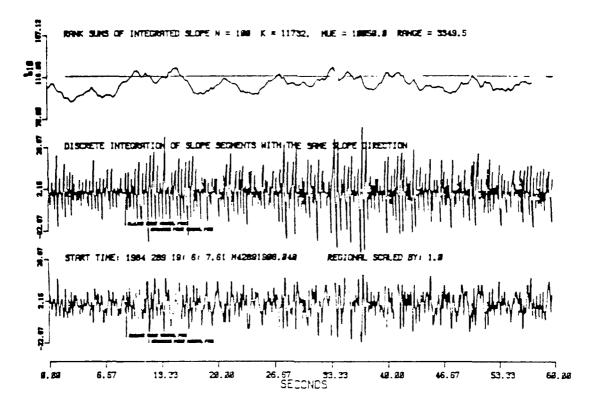
EVENT 2 - Regional (P-Lg) = 20 Seconds

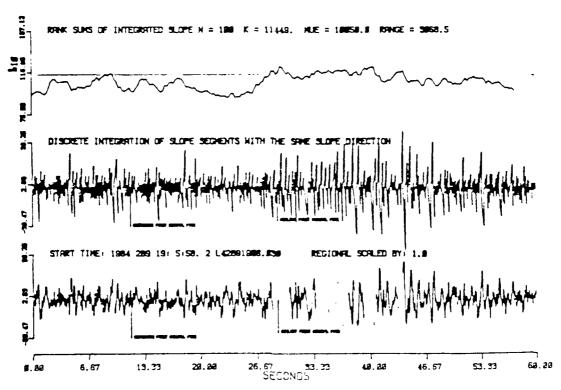




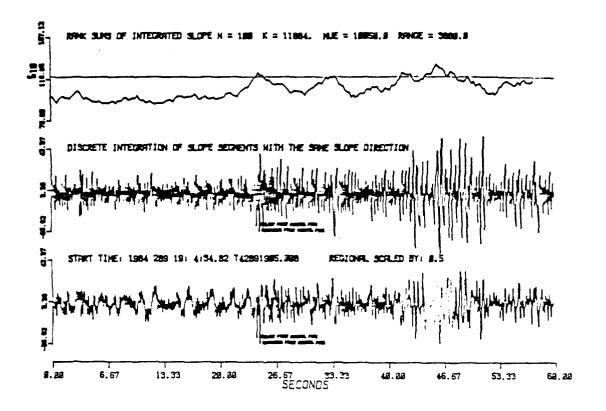


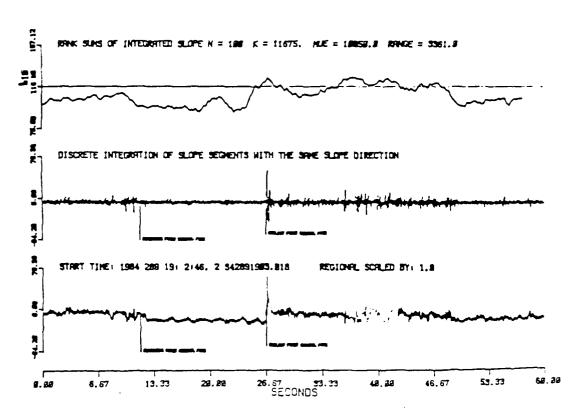


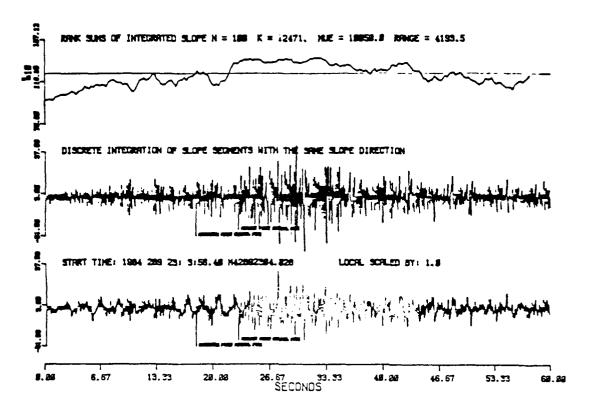


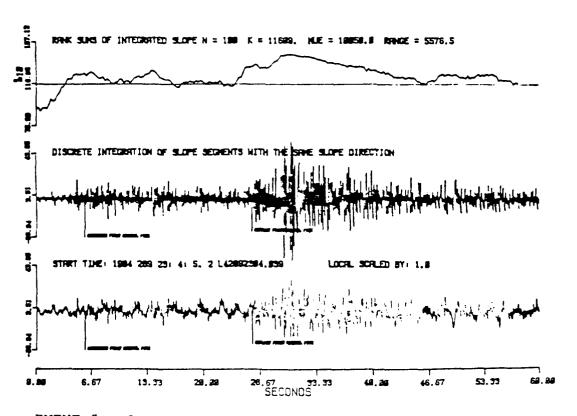


Event 4 - Regional (F - Lg) = 90 Seconds

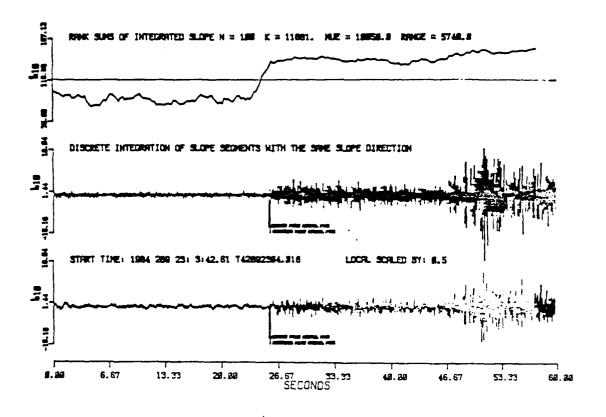


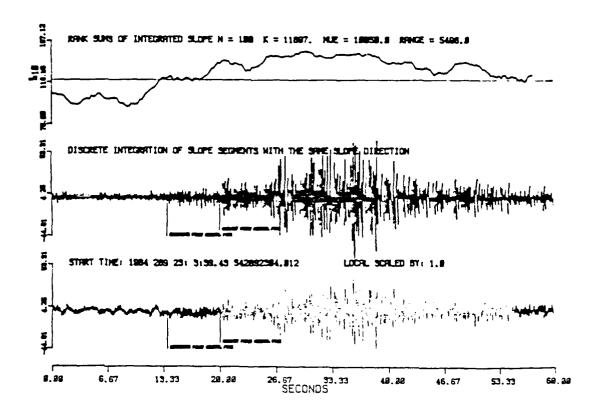


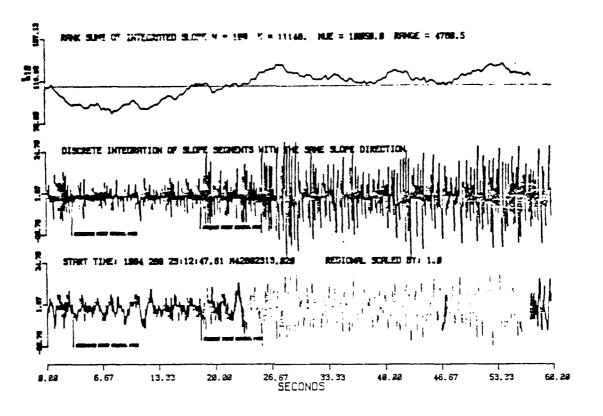


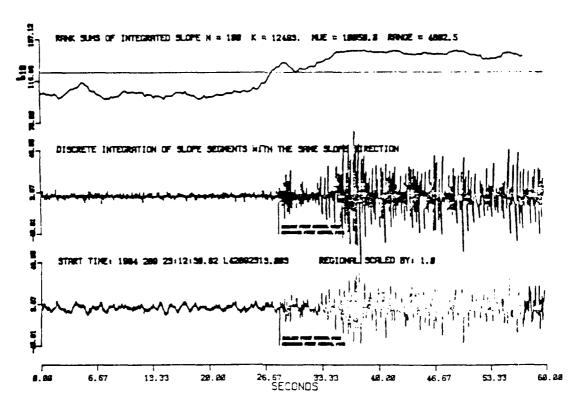


EVENT 5 - Local P Arrival

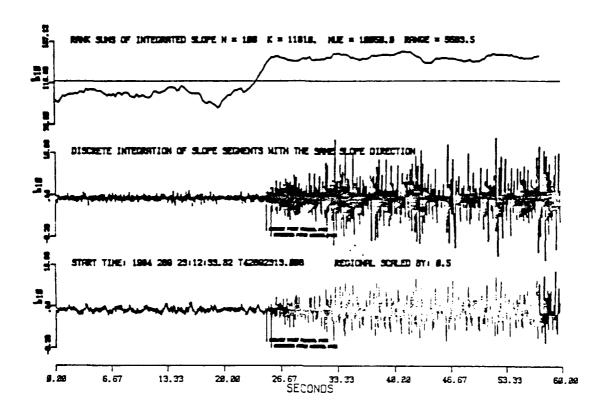


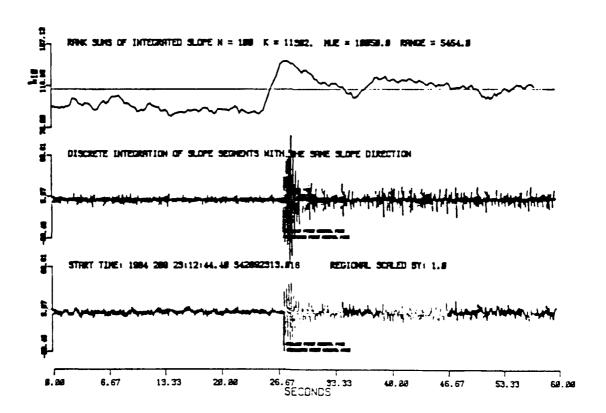


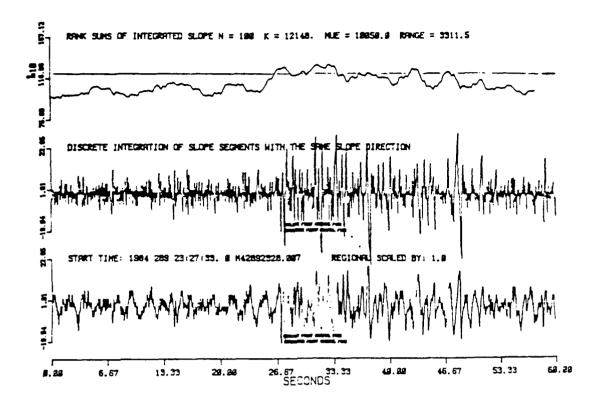


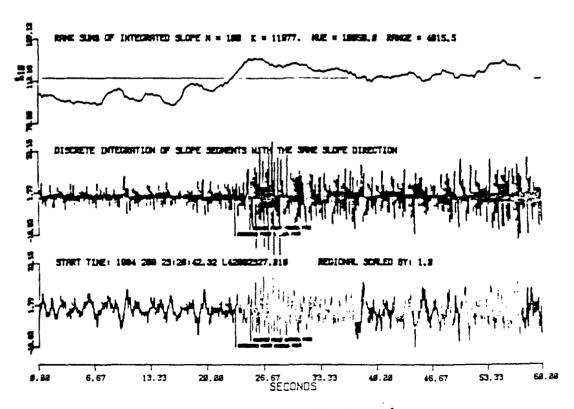


Event 6 - Regional (P - Lg) = 36 Seconds

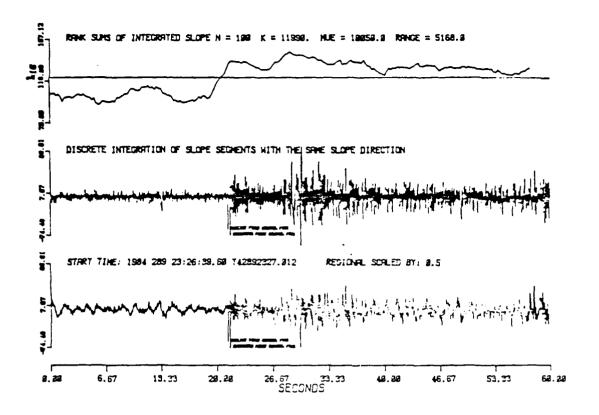


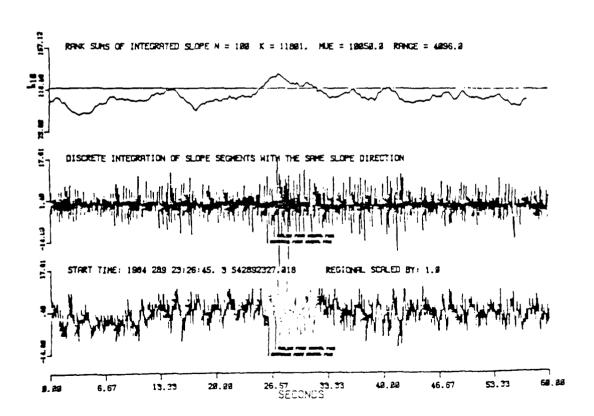


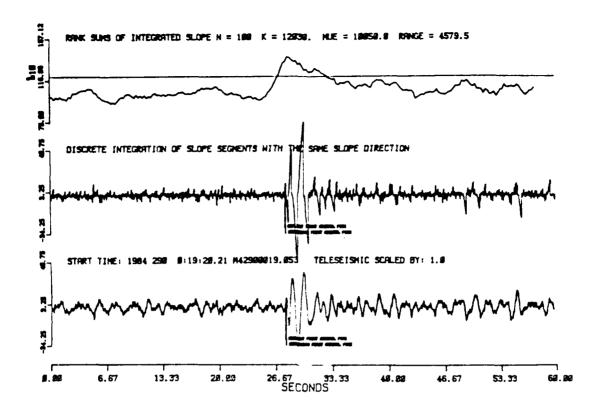


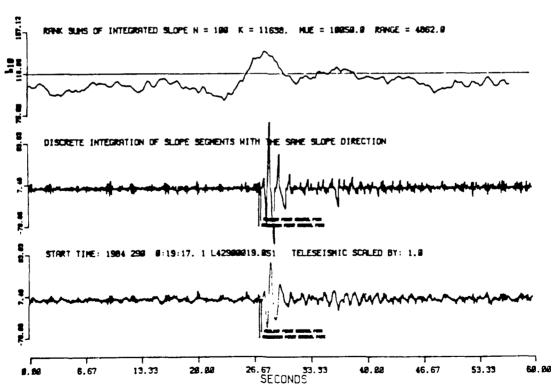


E ent 7 - Regional (P - Lg) = 40 Seconds

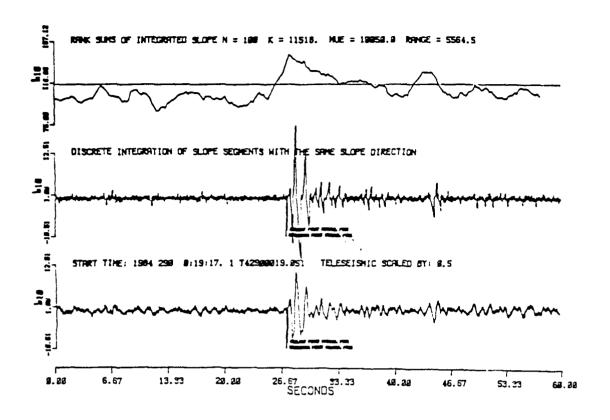




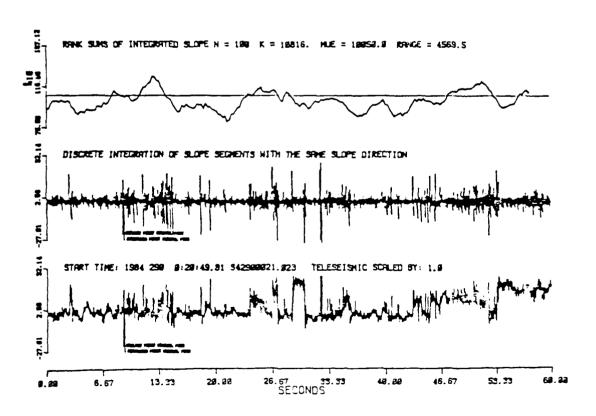


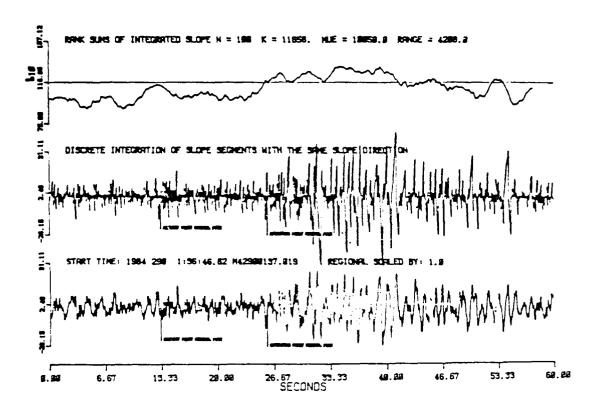


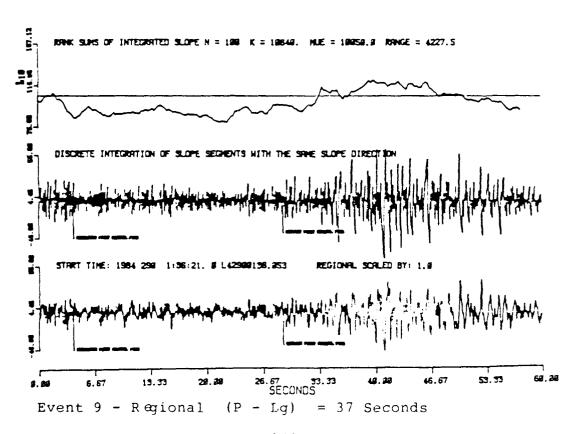
Event 8 - Fiji Islands Region - 19.1S 177.9W - 10/16/84 Origin Time: 00:07:51.3 - Depth: $449 \text{ km} + 89 \text{ M}_{\text{b}}$: 4.5 Residual Error: -0.0

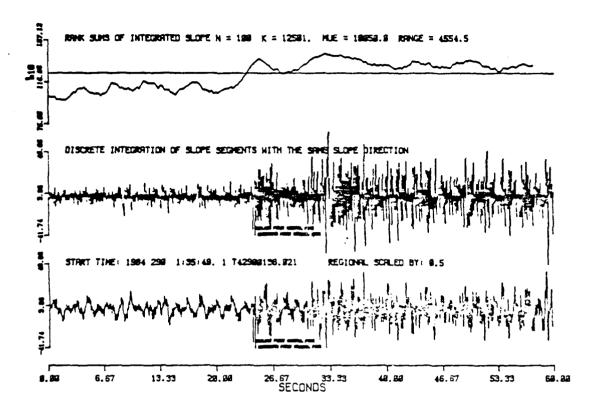


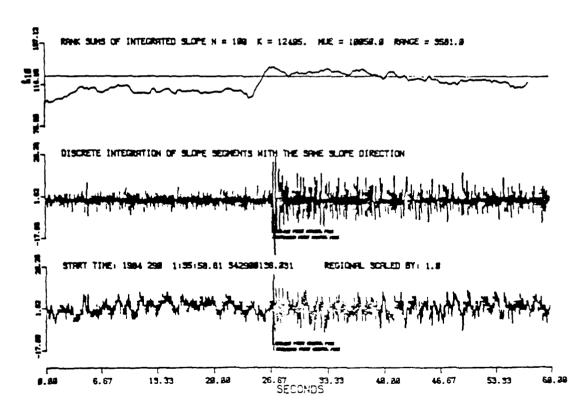
WALL BESKESS - KARKEST ISSESSES - ISSESSES - EDAS

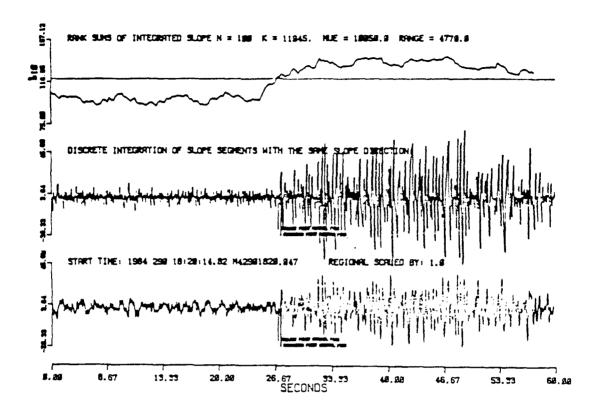








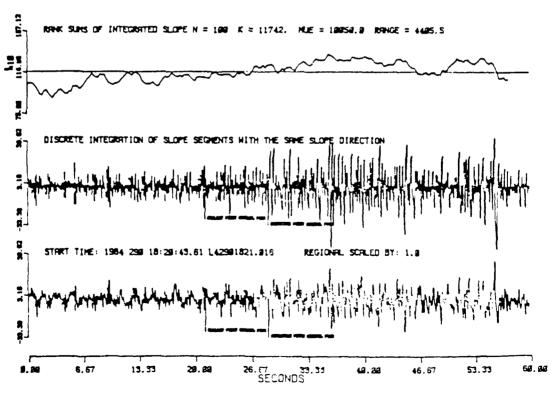




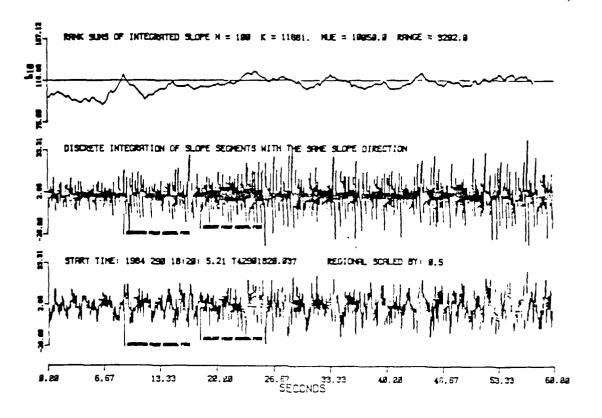
Keekken Reeker

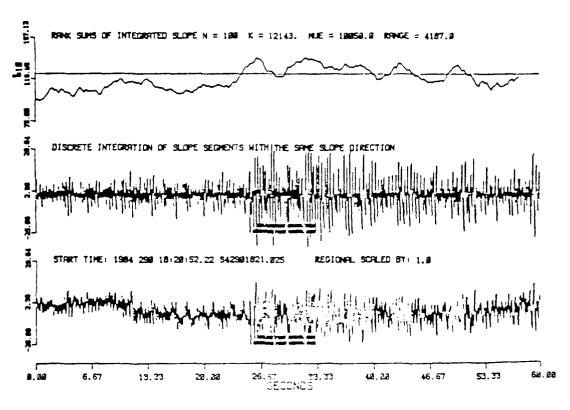
MINITAL SOMETH DECESSES MANNES

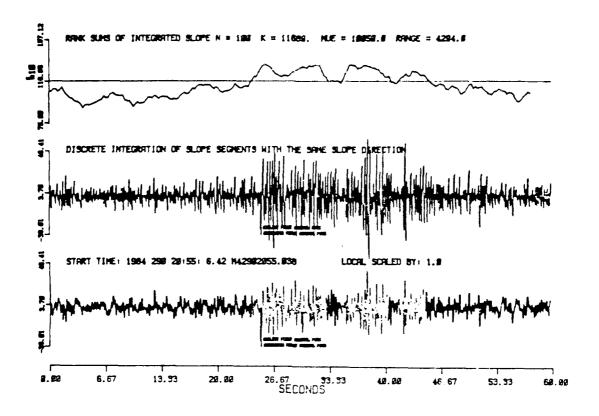
PATOTON RECEDENCE POPULATION WISSON

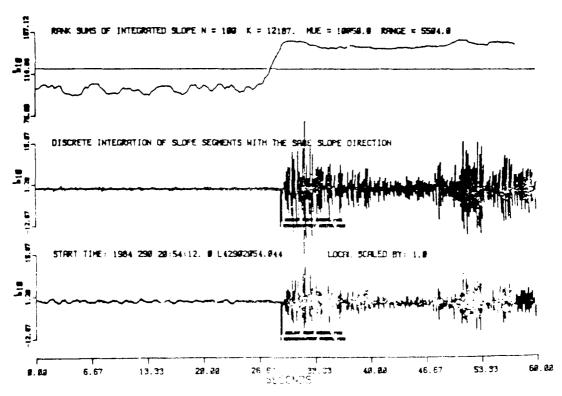


Event 10 - Regional (P - Lg) - 11 Seconds

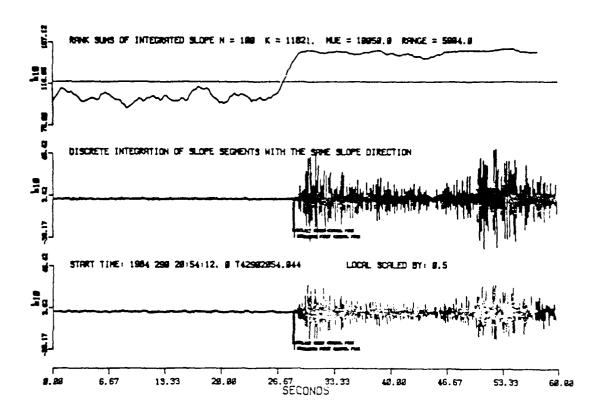


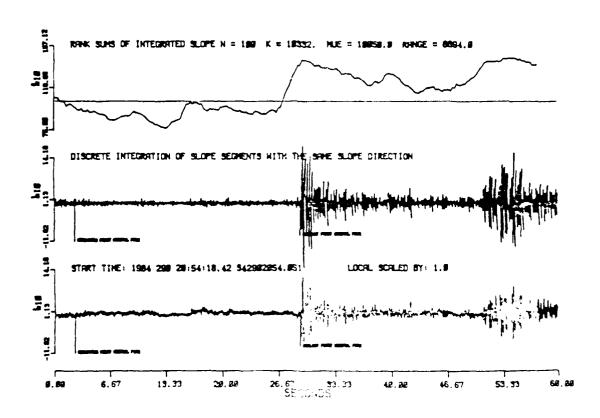


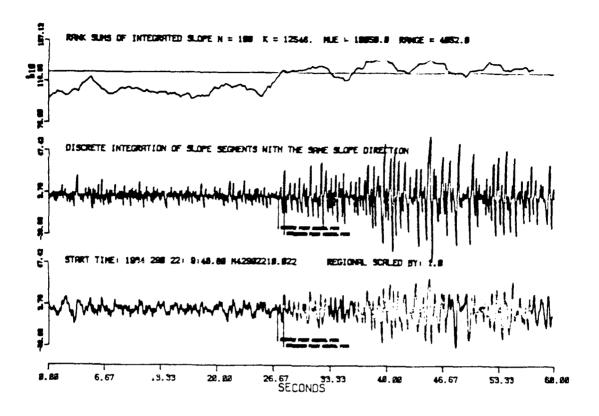


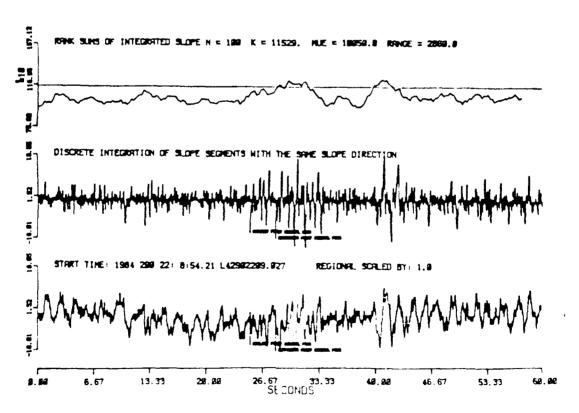


Event 11 - Mexico - October 16, 1984 - Quarry Biast

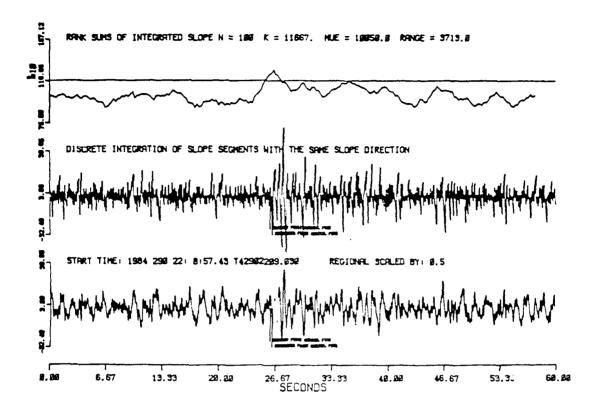




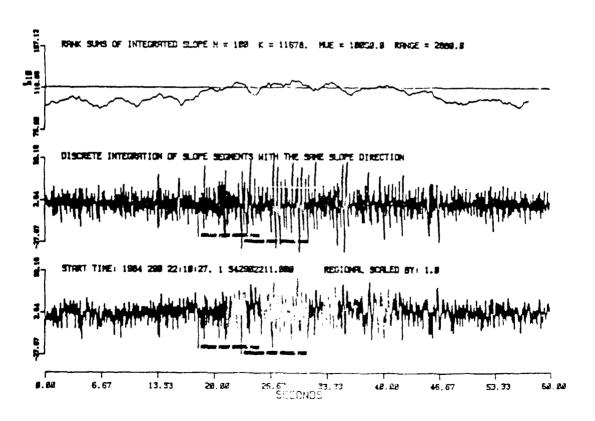


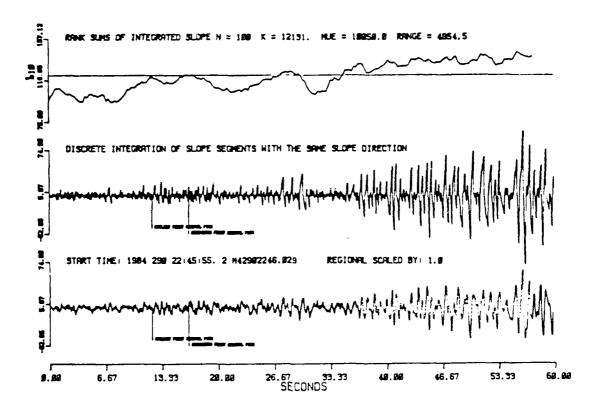


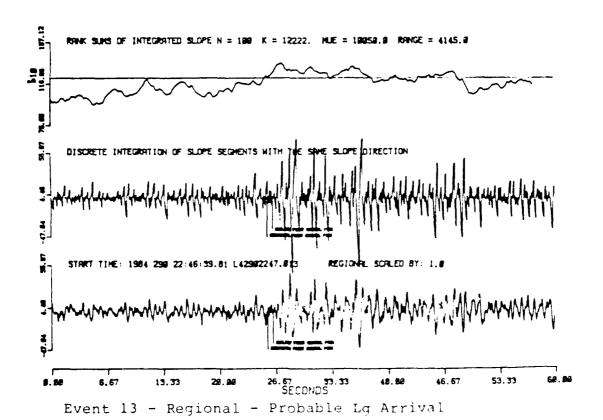
Event 12 - Regional - Probable Lg Arrival

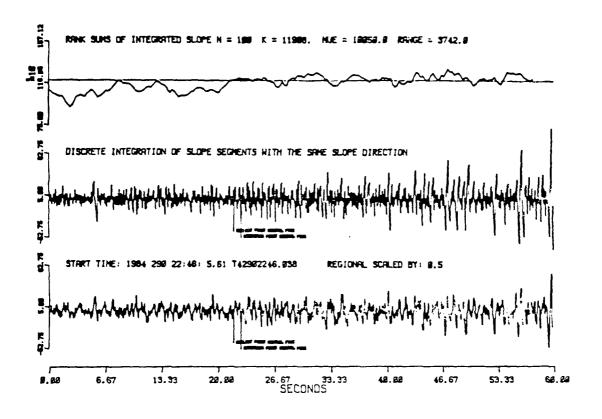


MASSELL BROKKI BROKKI

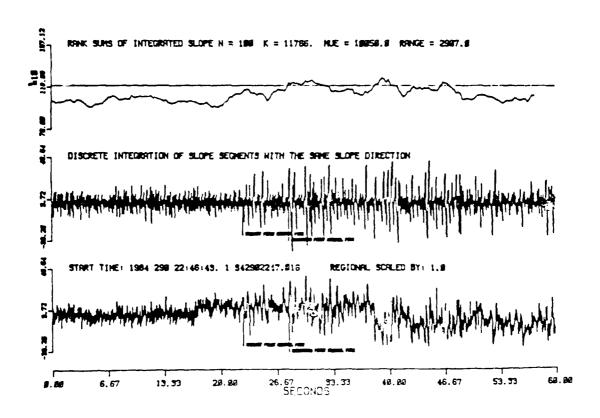


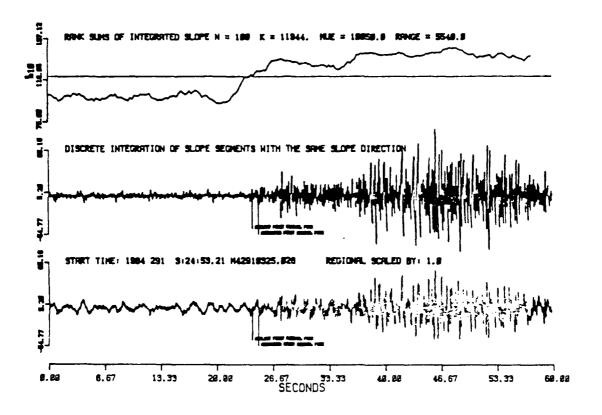




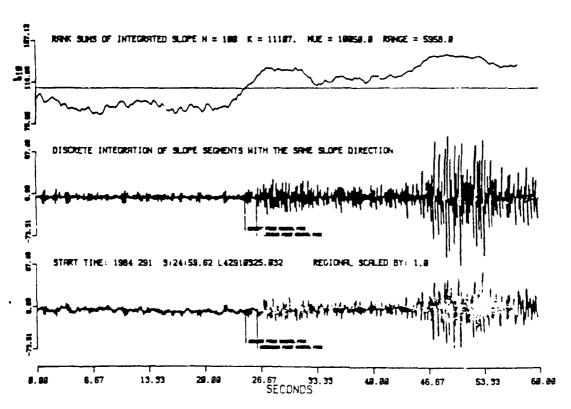


BOSSER BOSSER DEPOSE BOSSER PROTECTION PRODUCT

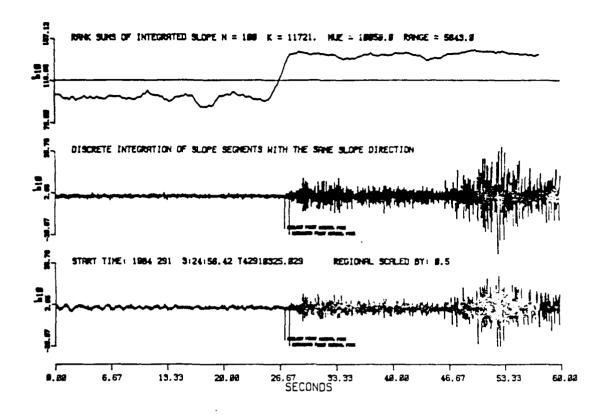


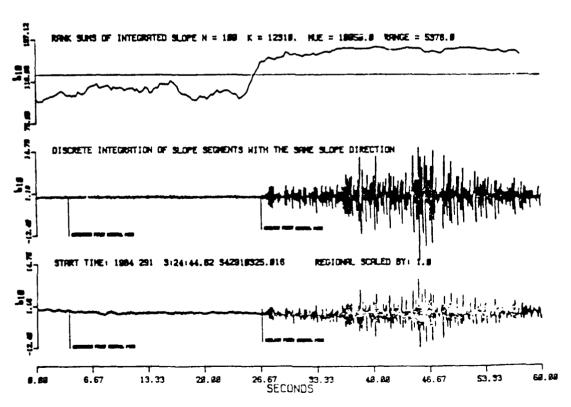


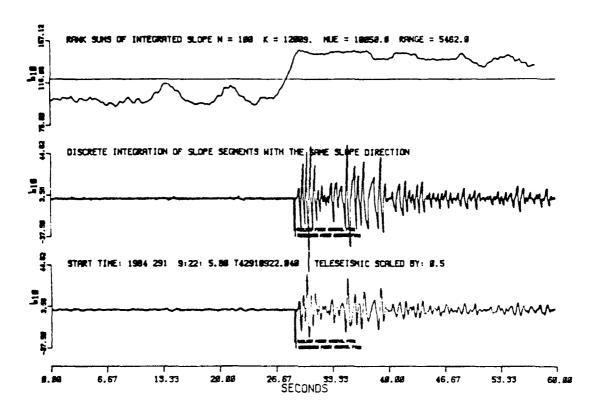
WWW.KKKKKI.LLCCCCCC.KCT.KKKKKKT.KRRRKKT.KRRRRKT.HTKKKKKT.HTKKKKKT.

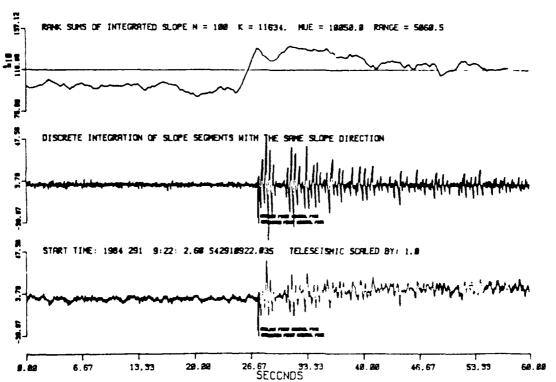


Event 14 - Regional - (P - Lg) = 24 Seconds 172

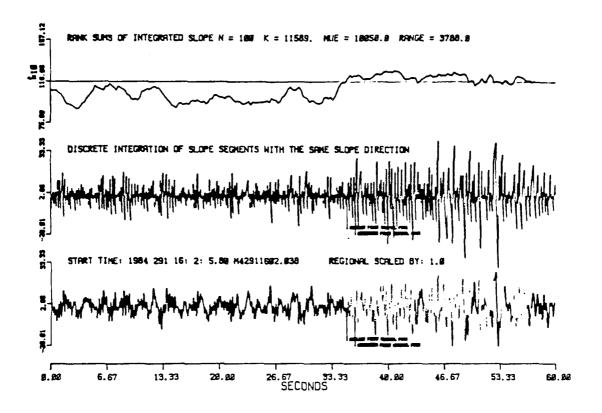




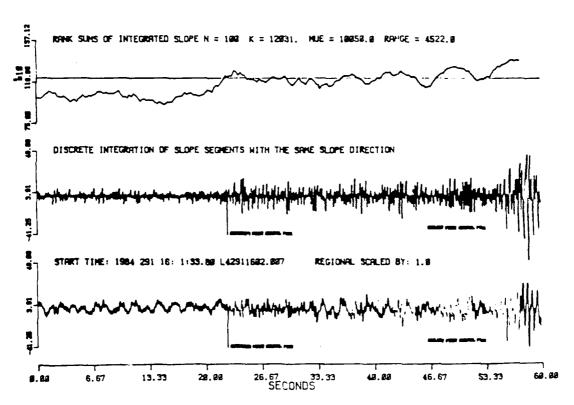




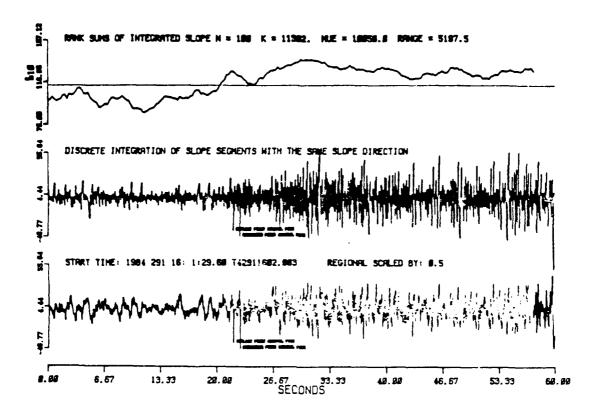
Kurile Islands 50.3N 153.N 153.5E - 10/17/84 Origin Time: 09:11:04.8
Origin Time: 09:11:04.8 - Depth 134km+3 M_b: 5.0
Residual Error: - 0.2

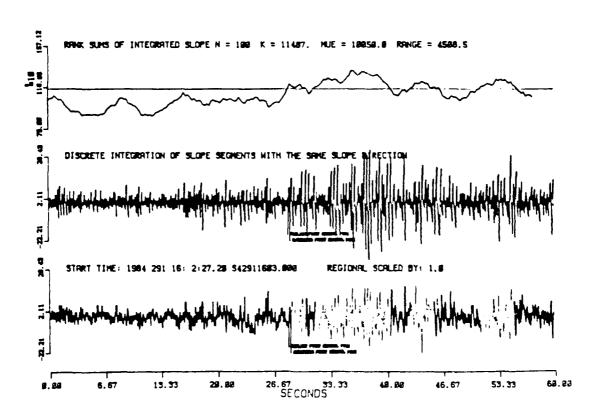


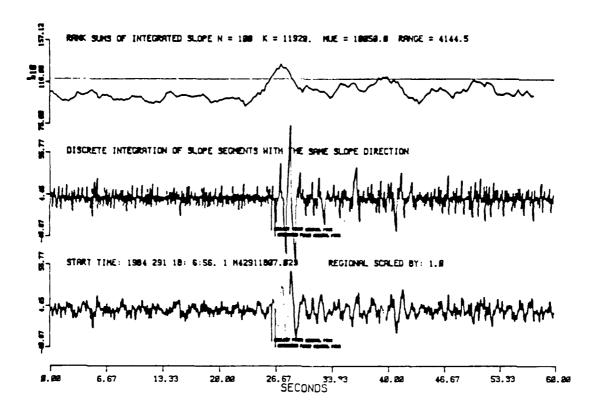
BEREIT RIGHT RECEIVE BREEKE STEELEN BEREIT. BESTEET BESTEET BESTEET

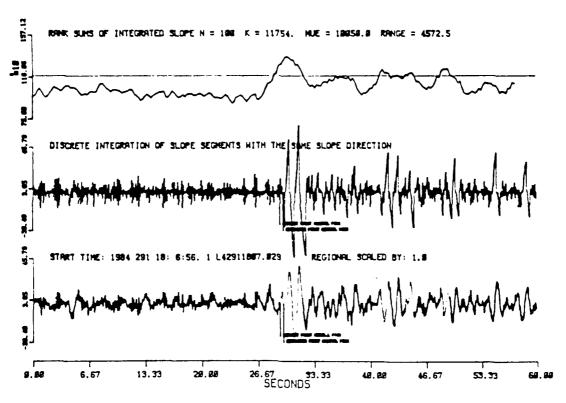


Event 16 - Regional - Probable Lg Arrival

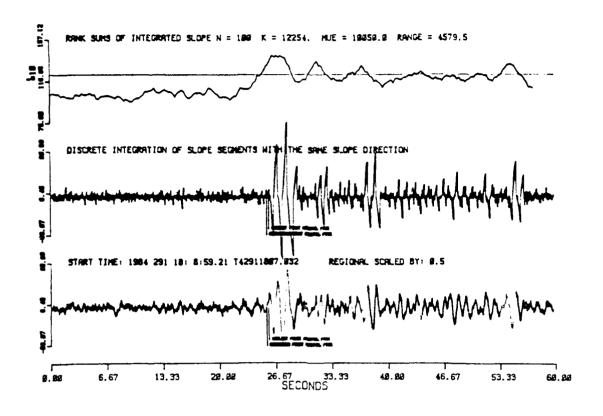


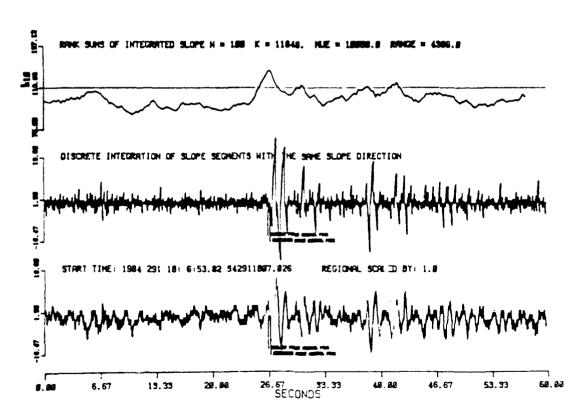


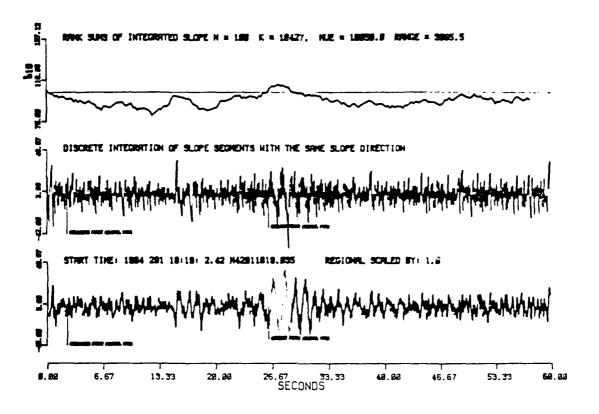


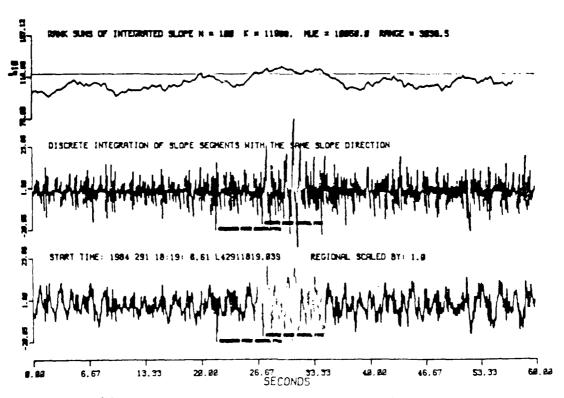


Event 17 - Regional - Probable Lg Arrival

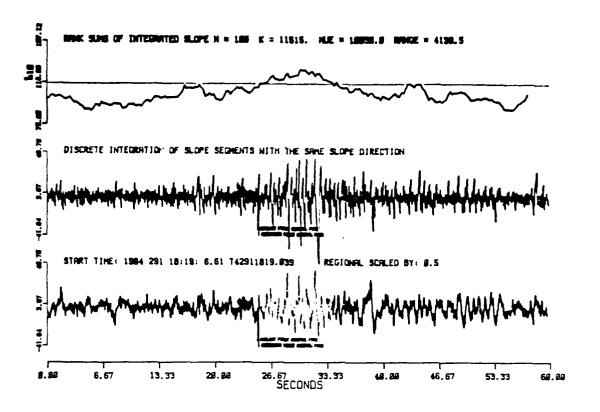


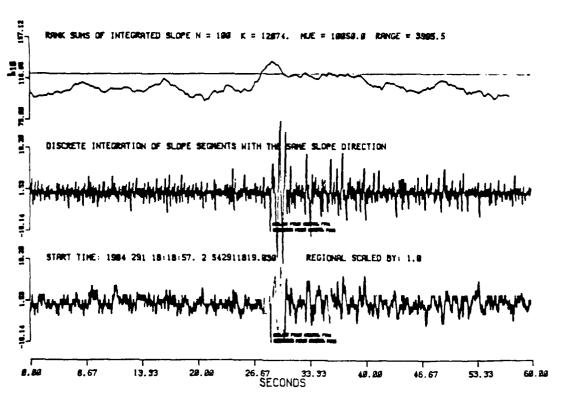




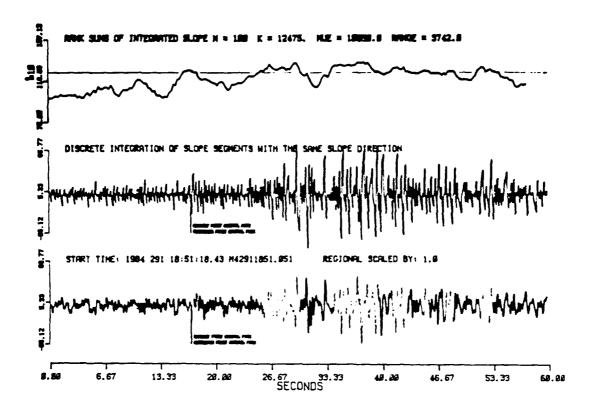


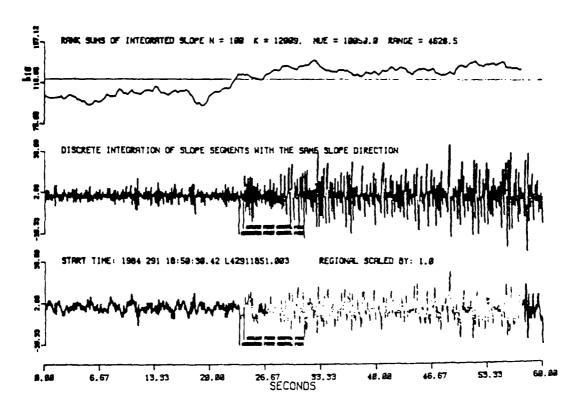
Event 18 - Regional - Probable Lg Arrival

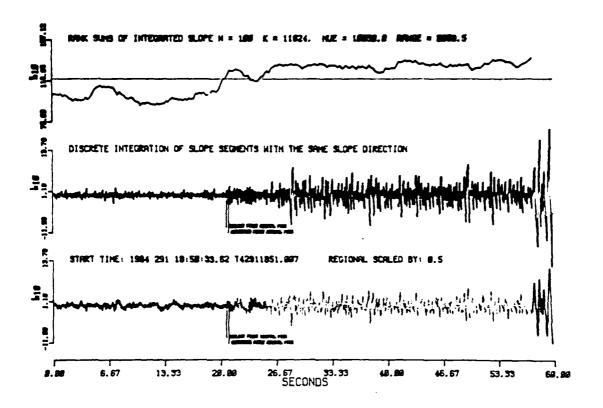


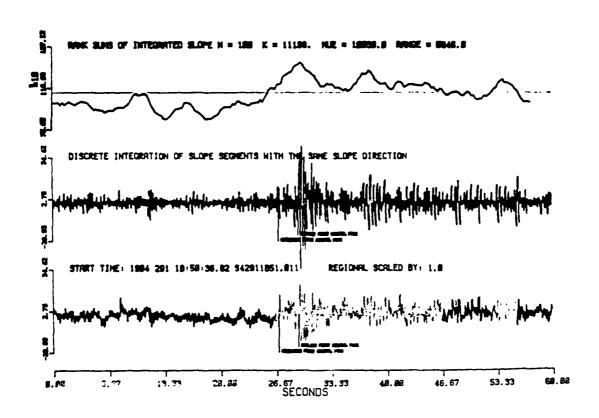


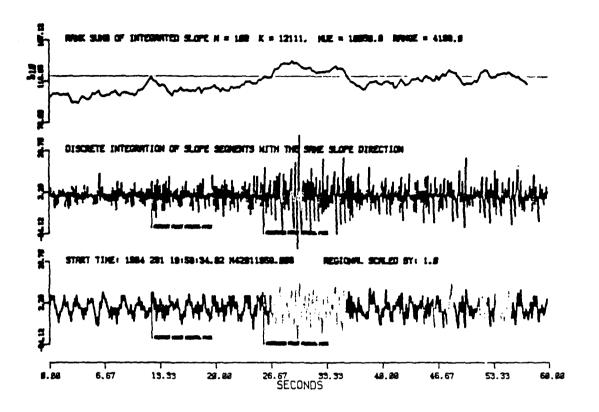
PERSONAL PROPERTY AND PROPERTY OF THE PROPERTY

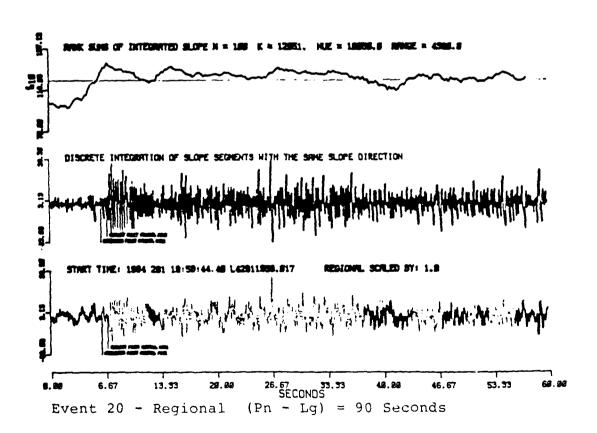


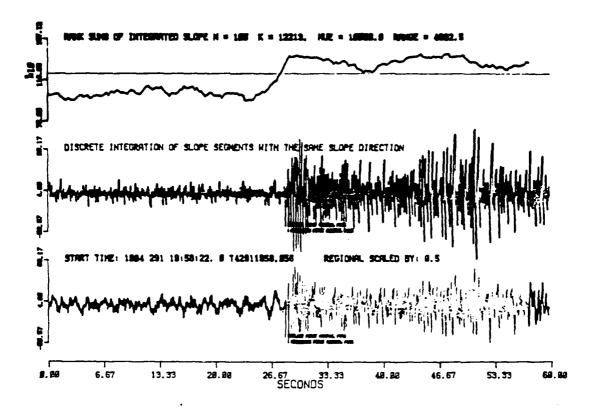




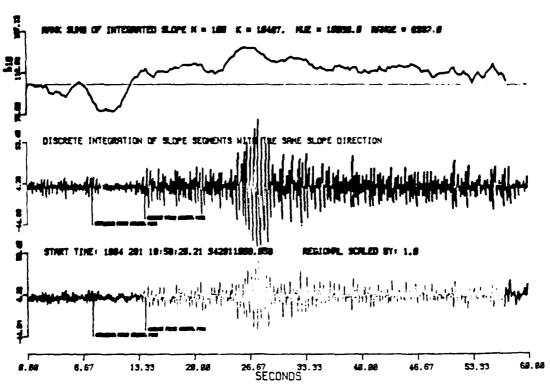




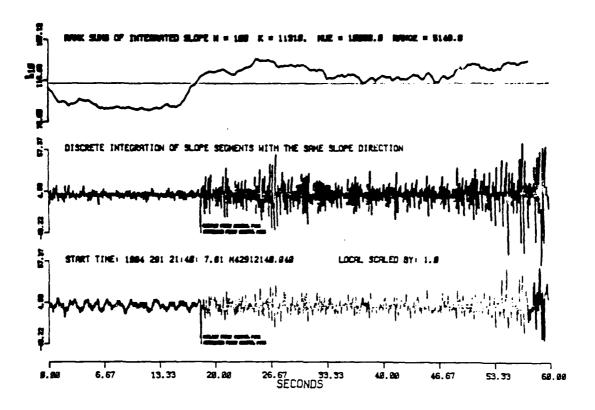


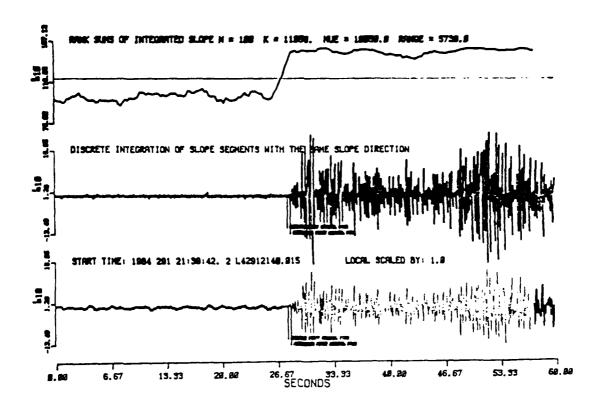


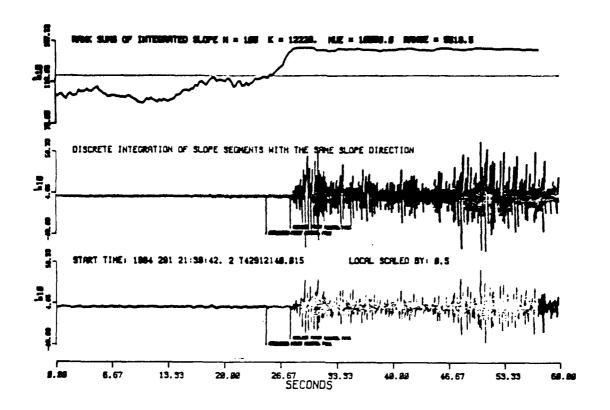
TONGOLD TO SOME SANCOST BOOMS TO SOME SANCOST

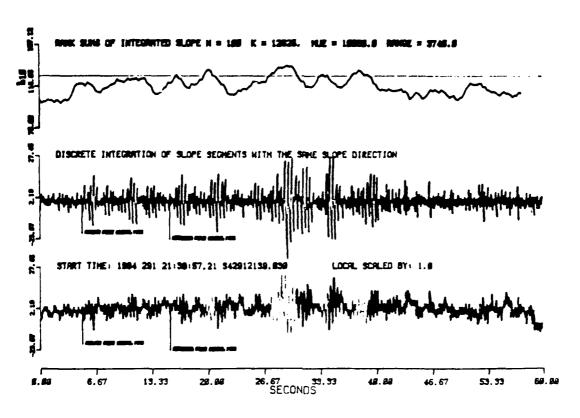


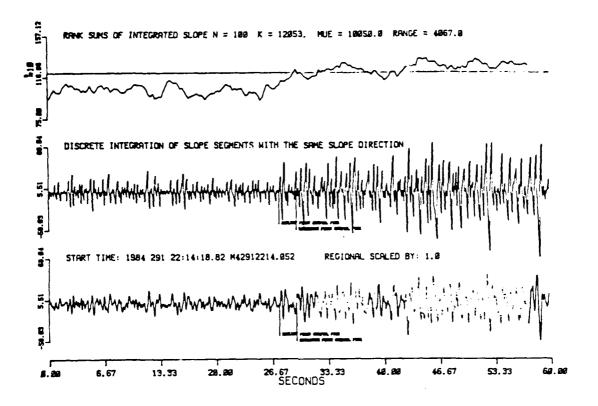
Event 21 - Mexico - October 17, 1984 - Quarry Blast

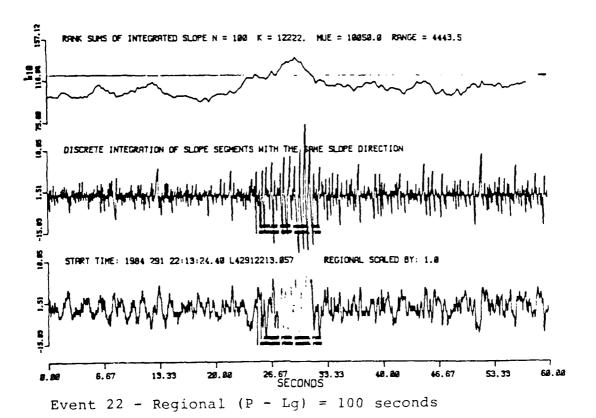


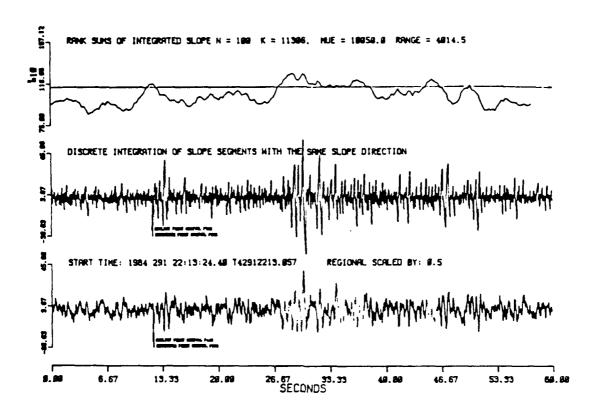


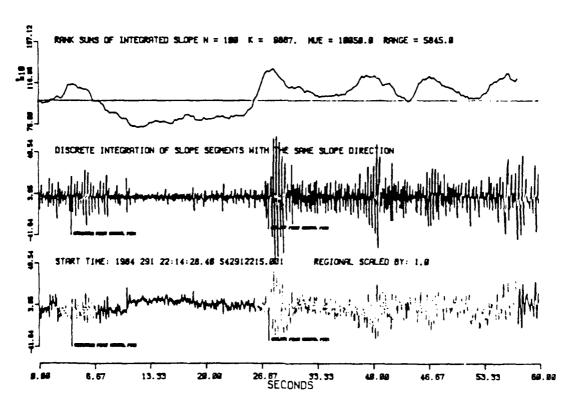


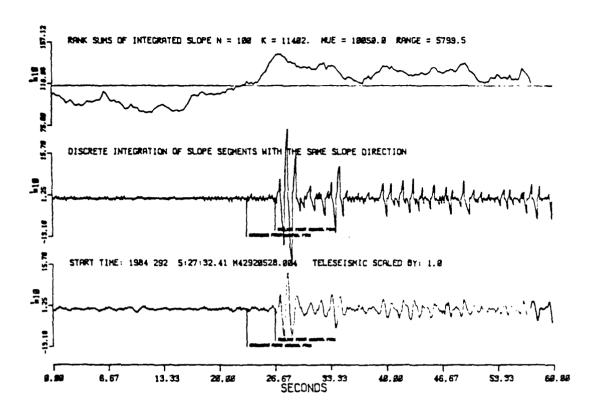


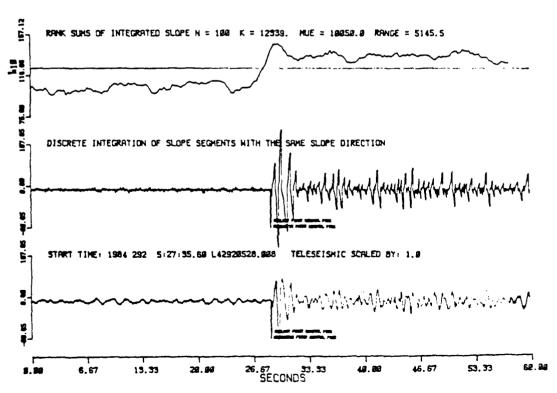




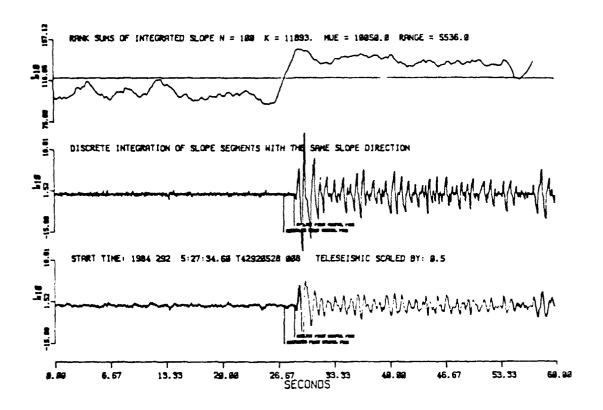


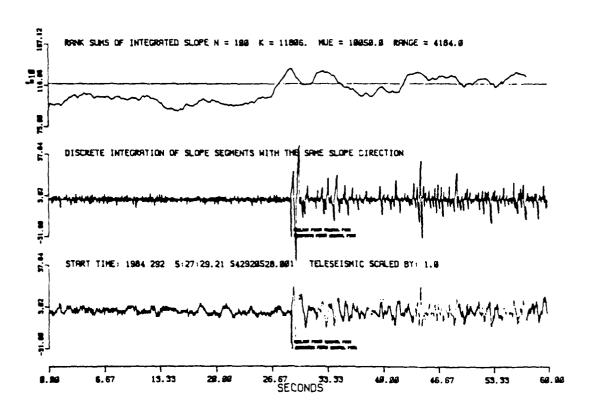


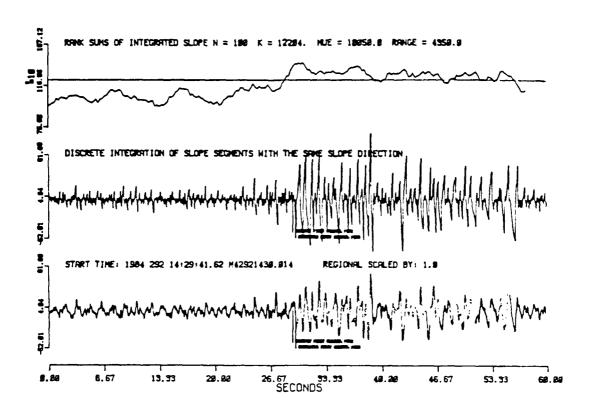


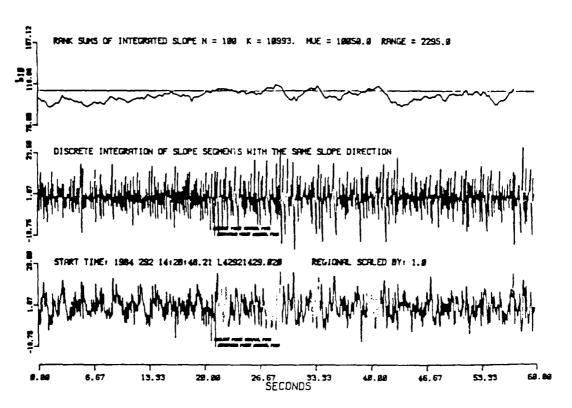


Event 23 - Central Alaska - 63.4N 151.3W 10/18/84 Origin Time: 05:19:53.1 Depth 105 km +7 $M_{\rm b}$: 4.3 Residual error: -0.6



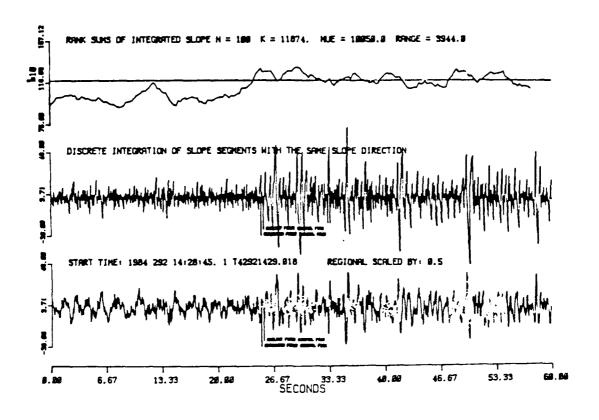


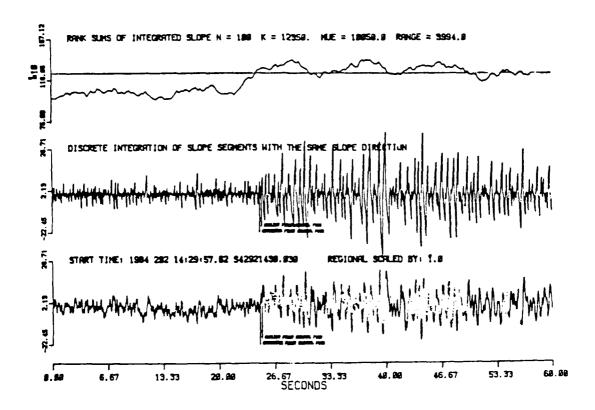


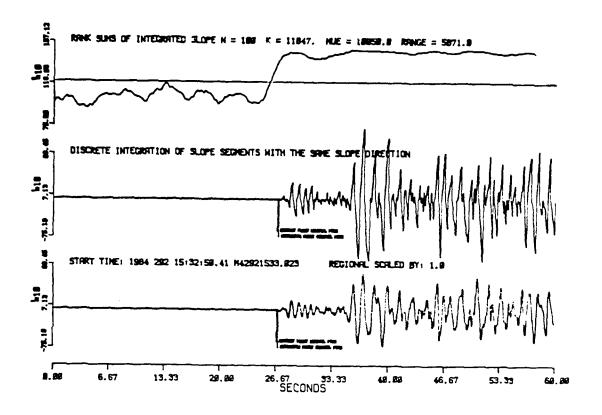


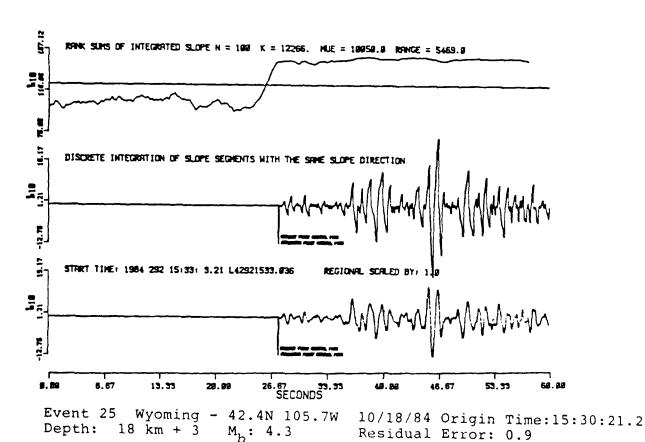
Event 24 - Regional Probable Lg Arrival

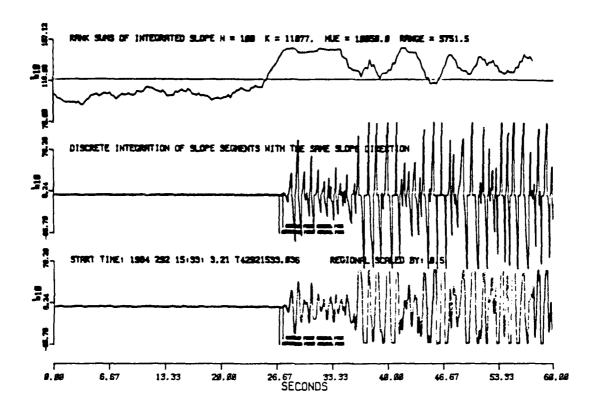
PRODUCE NEW PROPERTY.









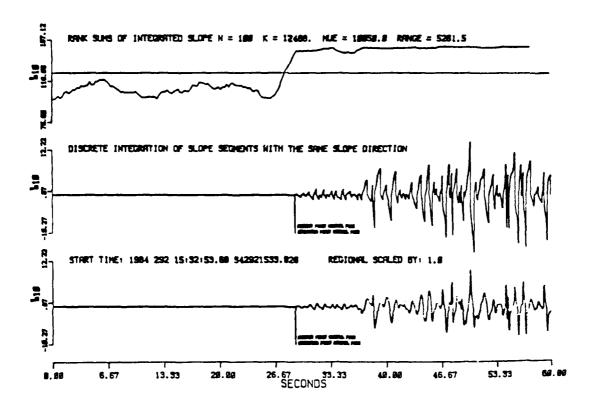


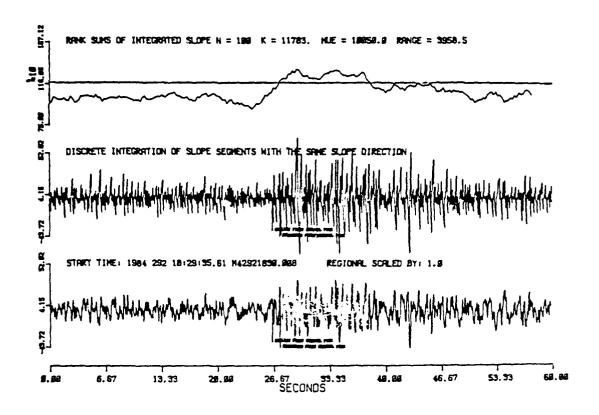
STATES AND SOCIETY STATES

kinitania permana pasasan pasa

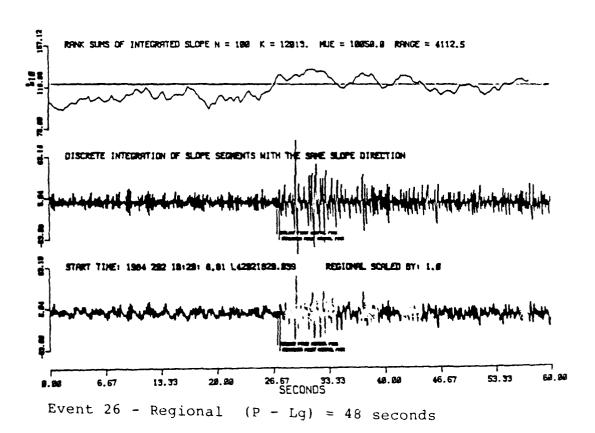
CULTURE BELLING MARCHER BLANCH

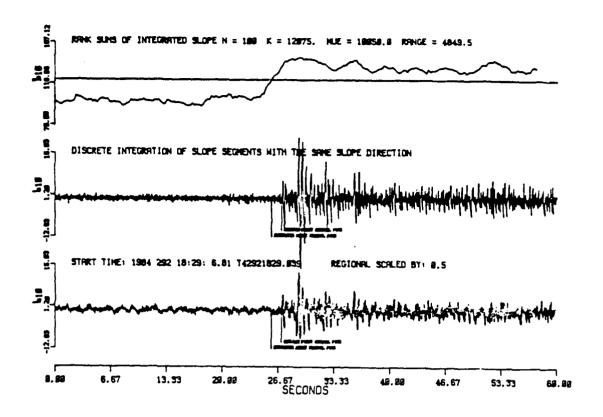
COLORED PROPERTY COLORED

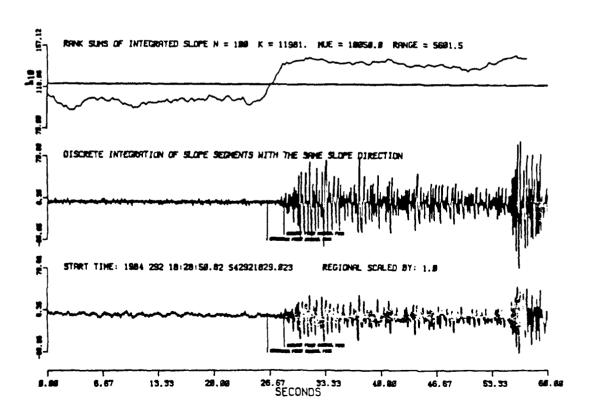




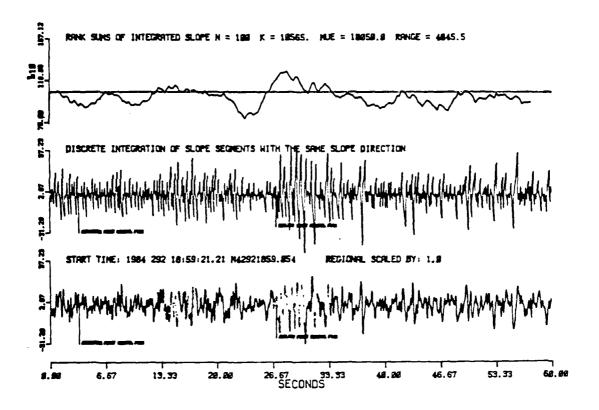
PROGRAM HYDROGON MARKETAN DAWNERS

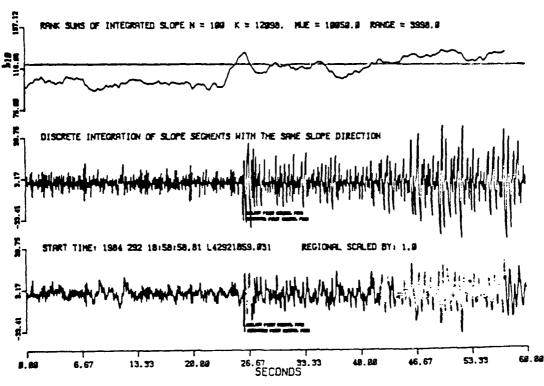




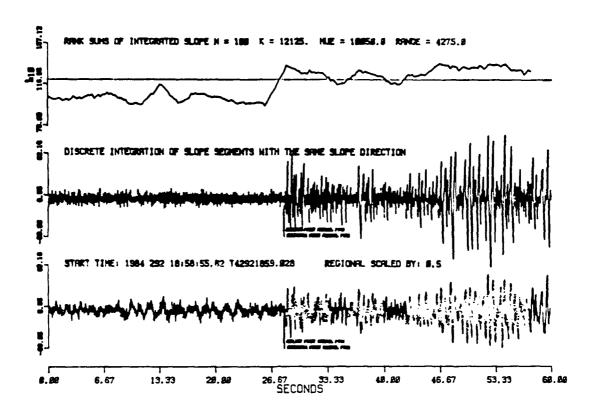


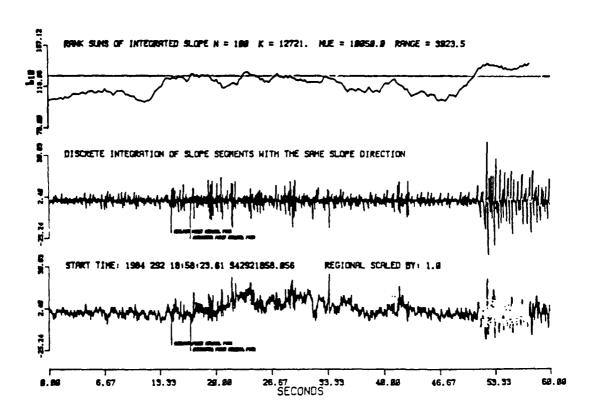
Secondary Additional Industrial Proposed in

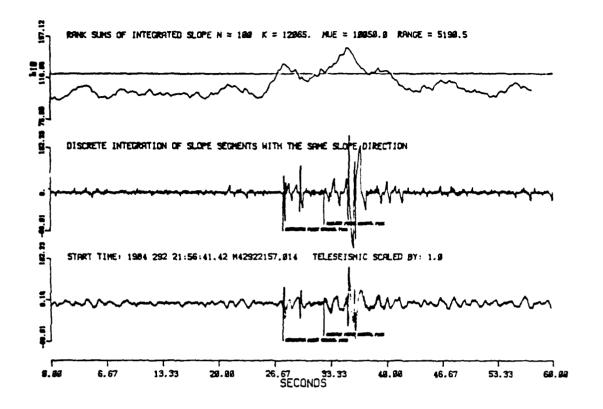


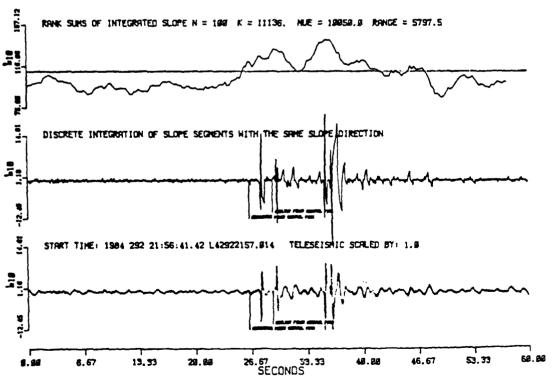


Event 27 - Regional (Pn - Lg) = 95 Seconds

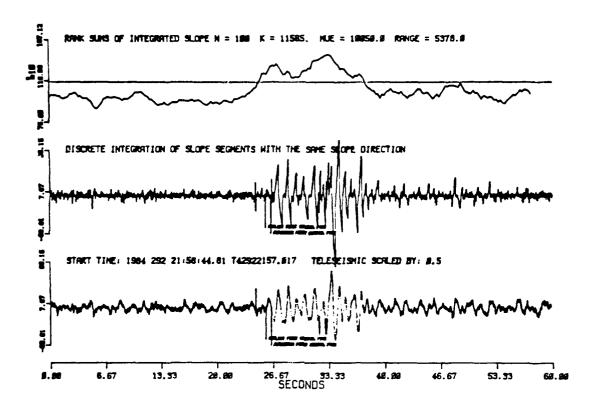


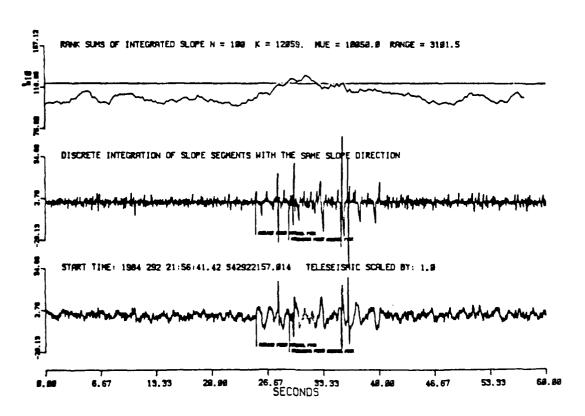
MATTER SECTION OF THE PROPERTY


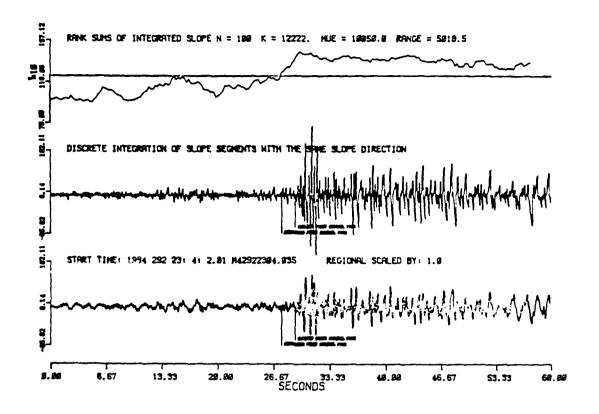


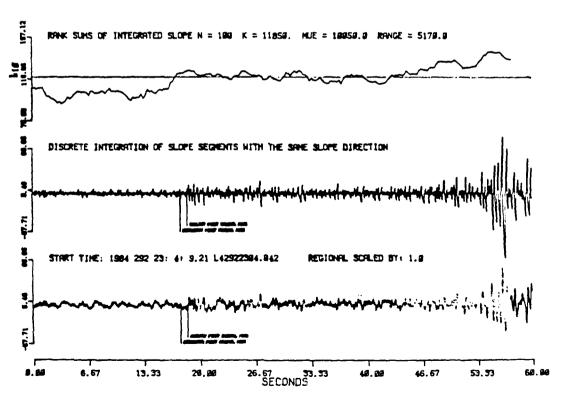


Event 28 Bonin Islands Region 28.1N 139.6E 10/18/84 Origin Time:21:42:38.8 Depth: 529 km + 8 Mb: 4.3 Residual Error: 0.5

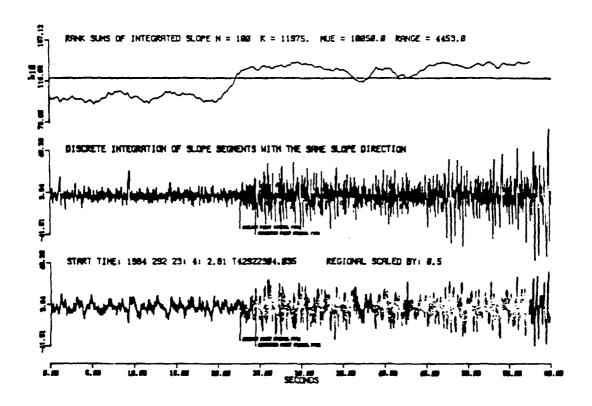


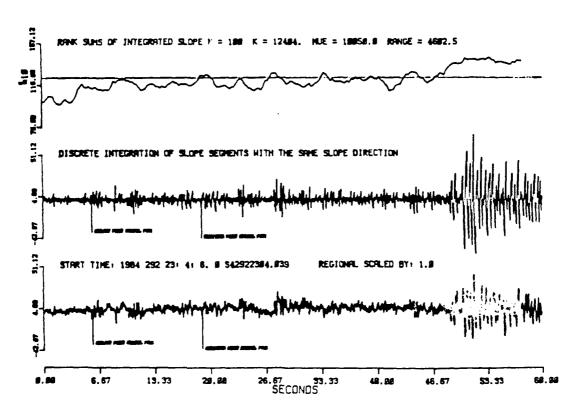


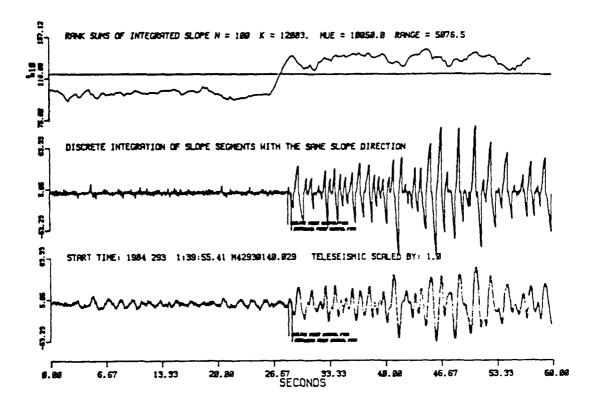


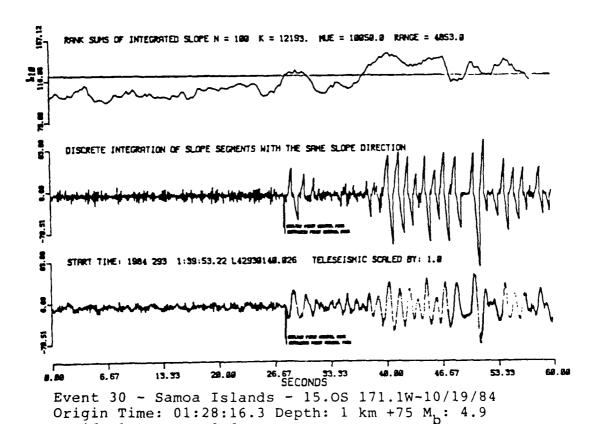


Event 29 - Regional (Pn - Lg) = 37 seconds

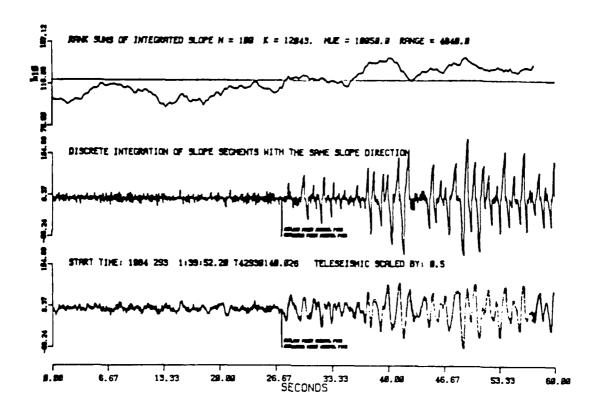


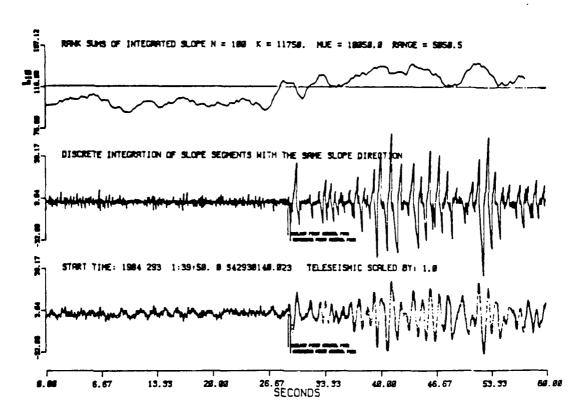


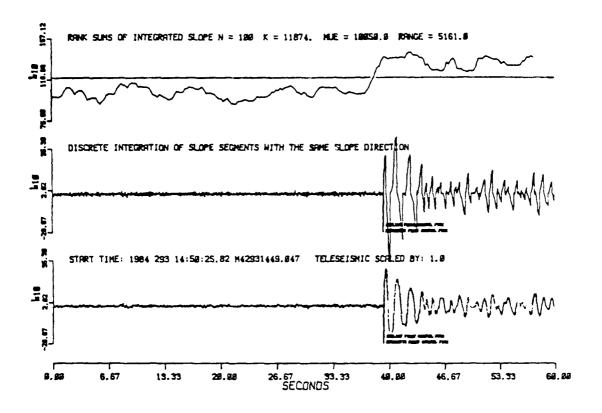


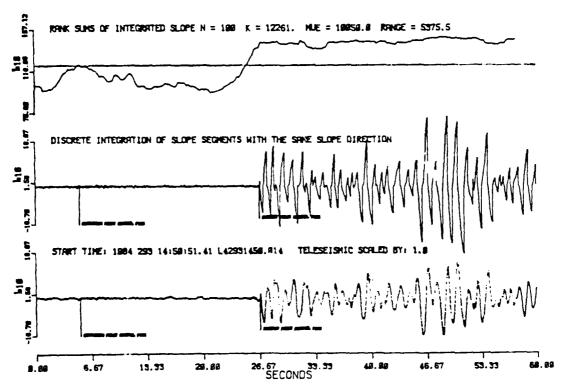


Residual Error: -0.6

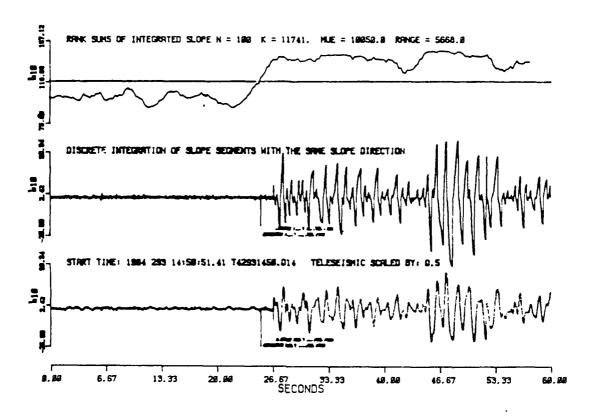






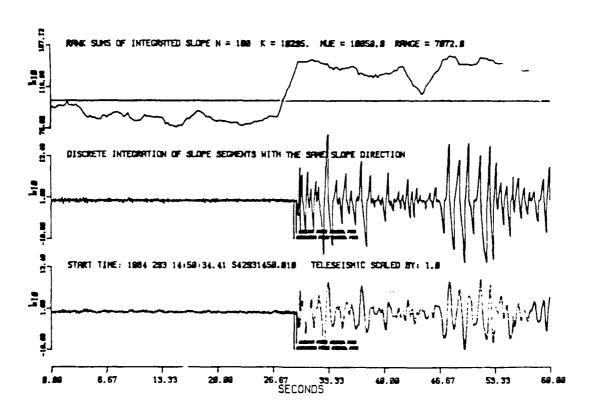


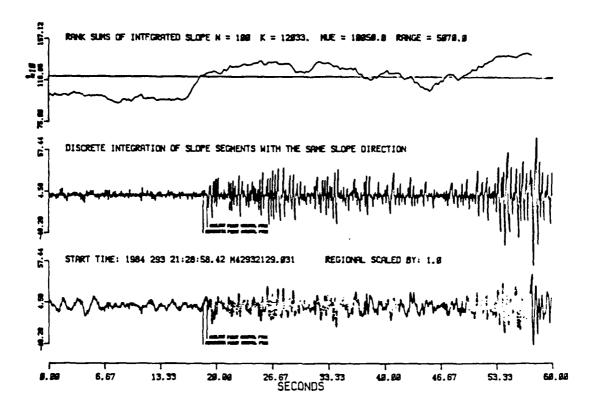
Event 31 - Tonga Islands - 15.7S 173.8W 10/19/84 Origin Time: 14:37:35.6 Depth: 1 km +2 M_b:5.7 Residual Error:1.9



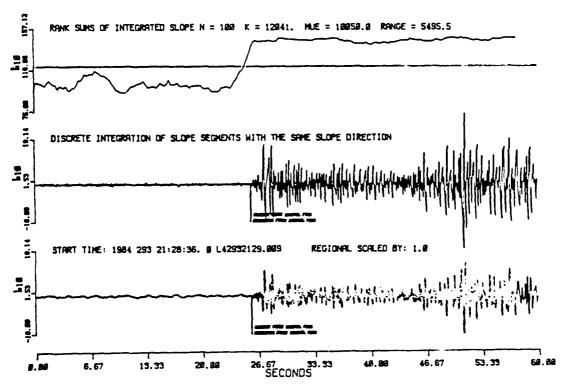
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2)27532E_EEE45252_EE55252.

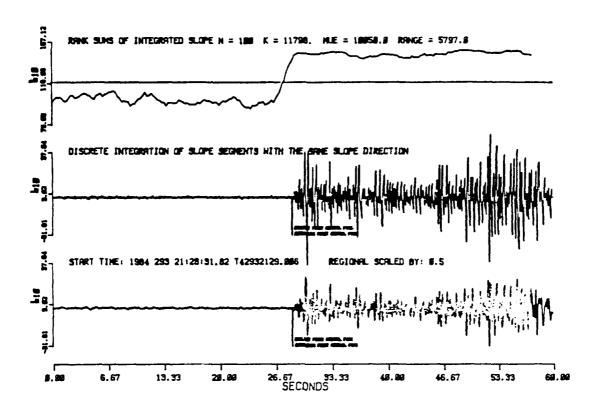


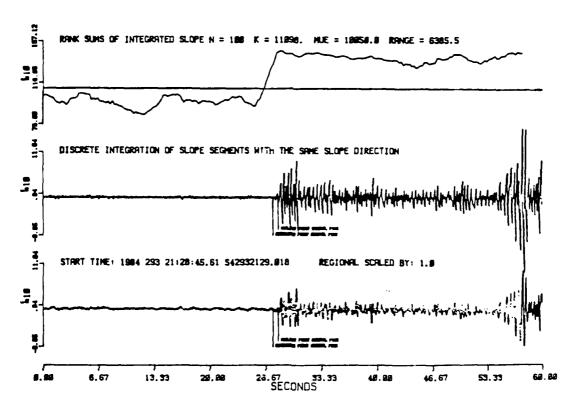


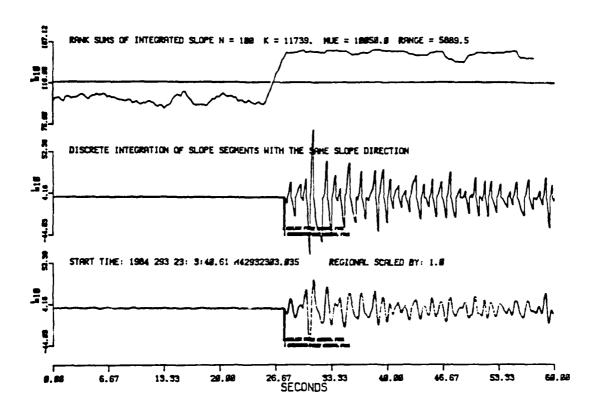
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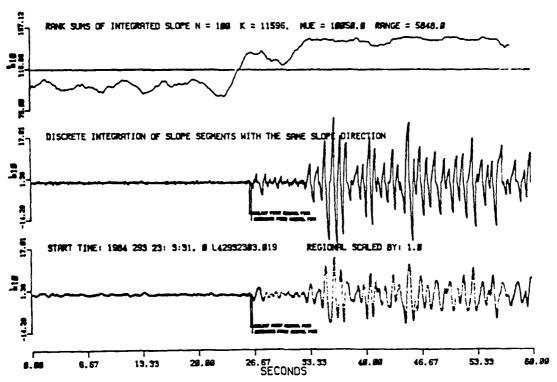


Event 32 - Regional (Pn - Lg) = 21 Seconds

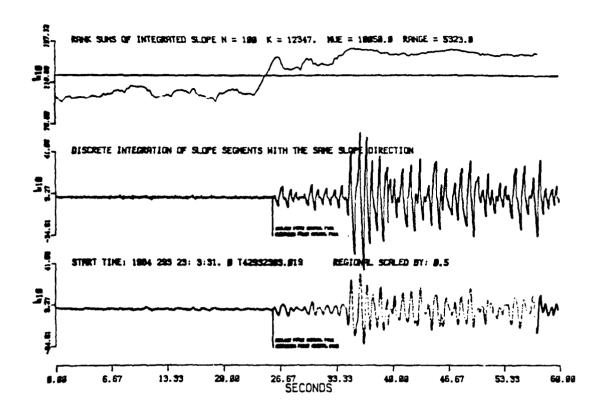


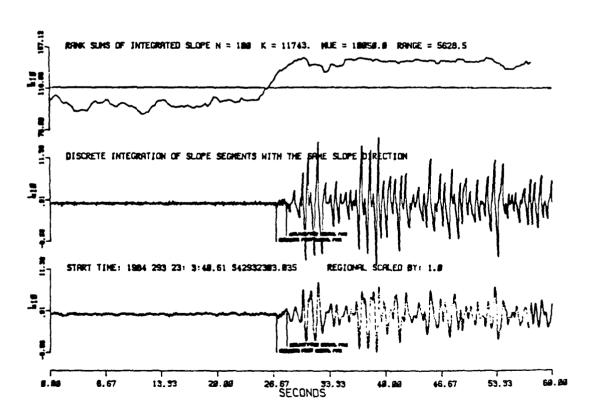


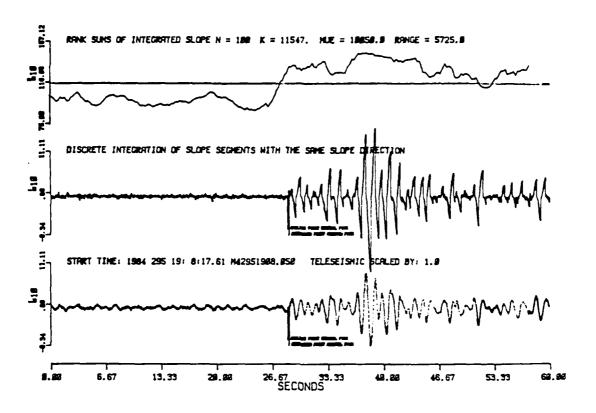


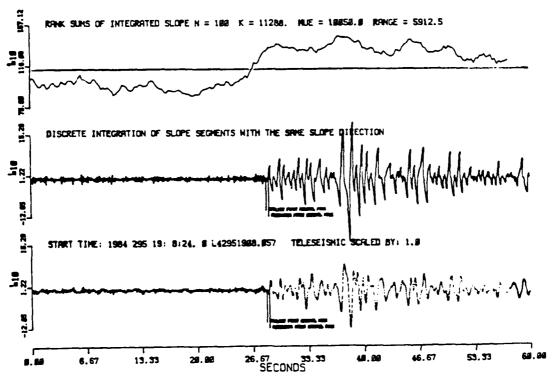


Event 33 - Near Coast of Guerro, Mexiso 16.6N 98.5W 10/19/84 Origin Time: 22:59:57.8 Depth: 25 km + 3 $M_{\rm b}$: 4.9 Residual Error: 3.0

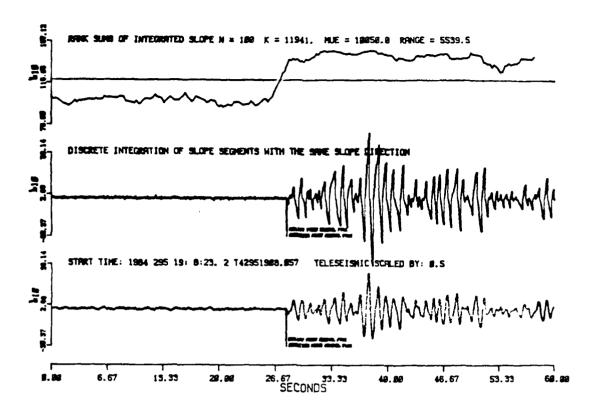


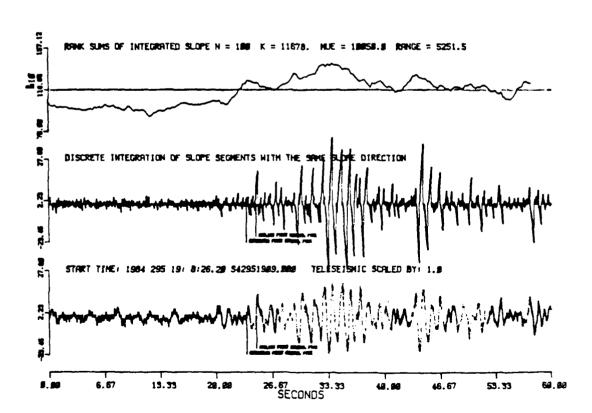


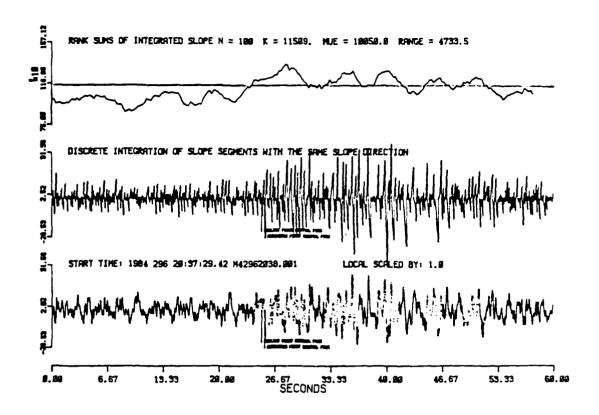


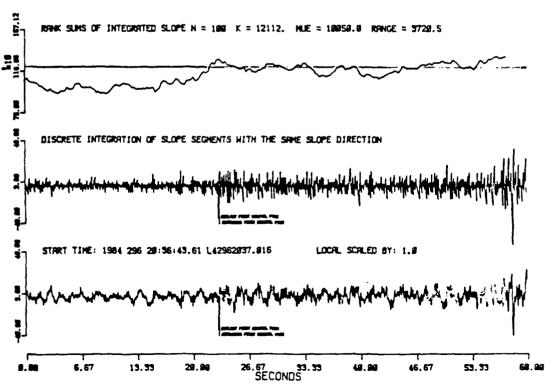


Event 34 - East of Severnoya Zemlya 82.0N 114.2E 10/21/84 Origin Time: 18:57:55.6 Depth: 14km + 4 Mb:4.3 Residual Error: -0.3

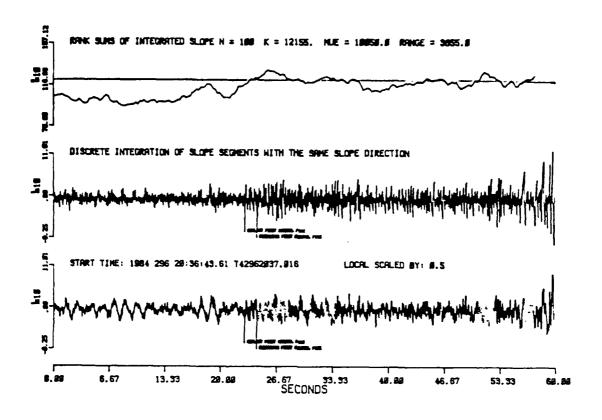


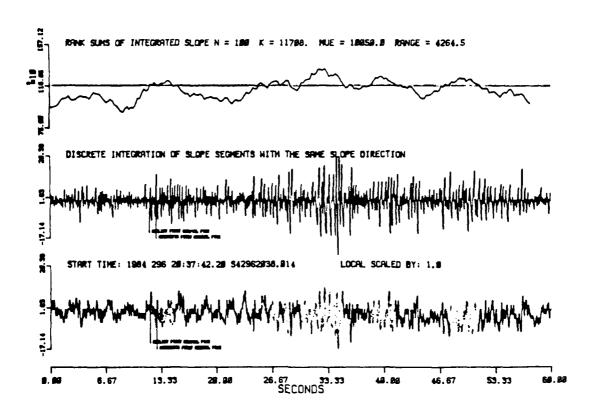


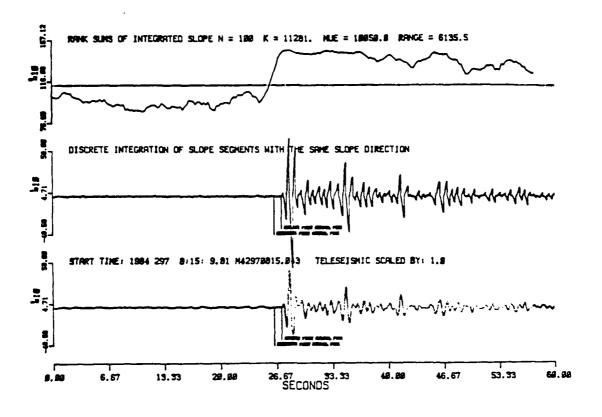


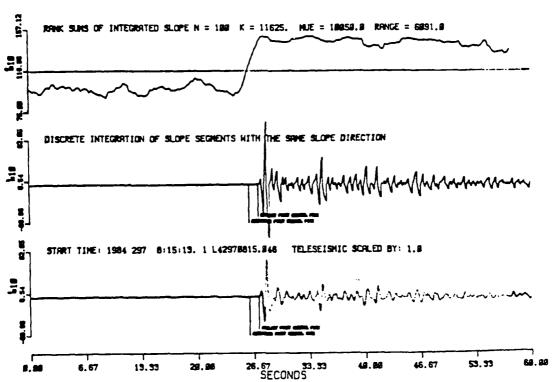


Event 35 Local P Arrival

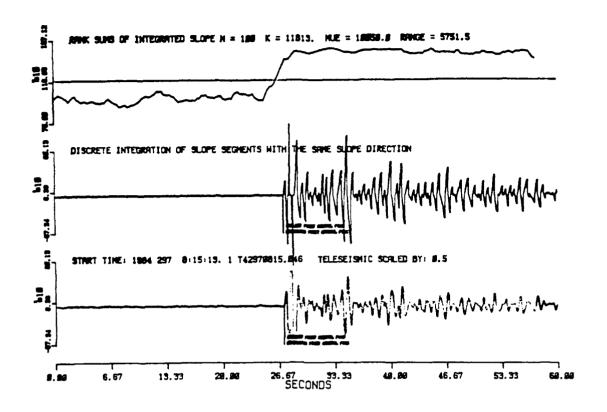


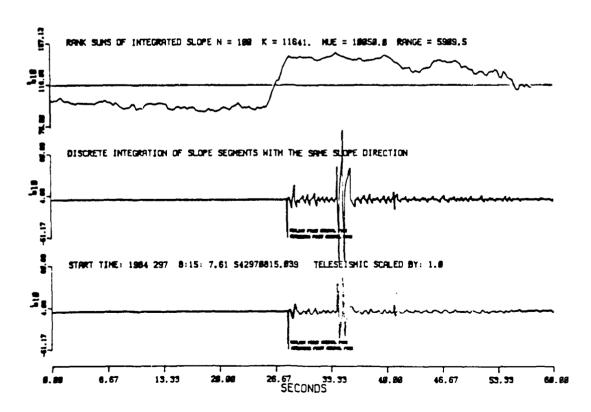


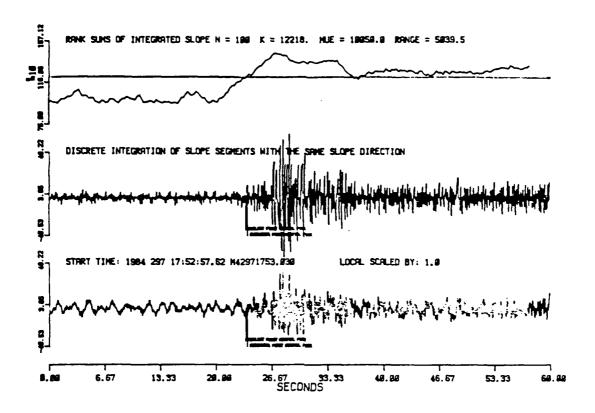


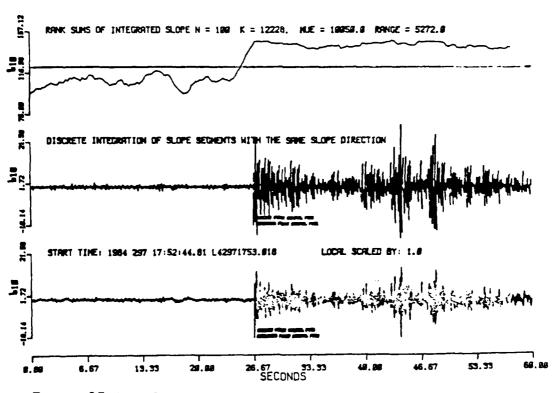


Event 36 Komandorsky Islands Region 55.7N 165.OE 10/23/84 Origin Time: 08:04:46.1 Depth: 20 km +3 M_b:4.8 Residual Error: 0.3

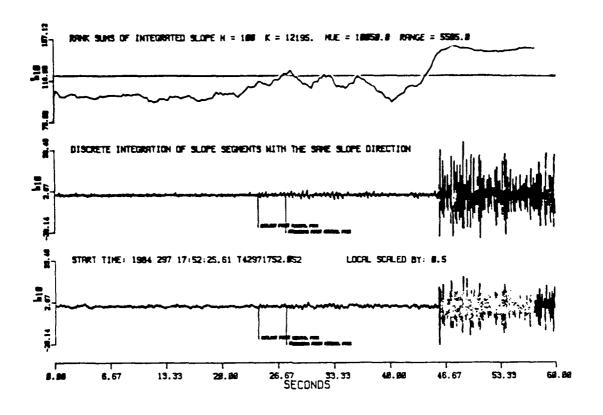


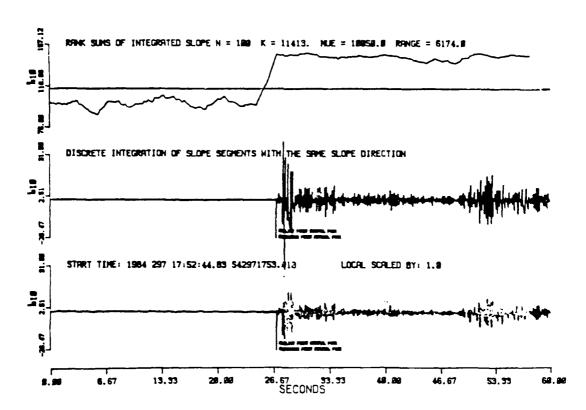


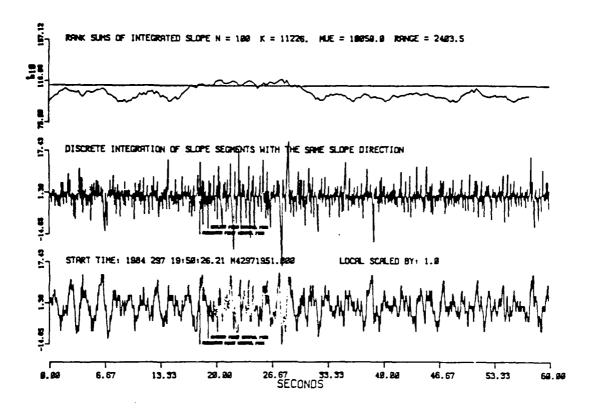


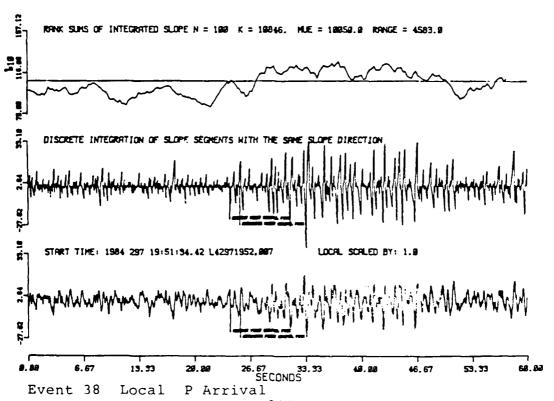


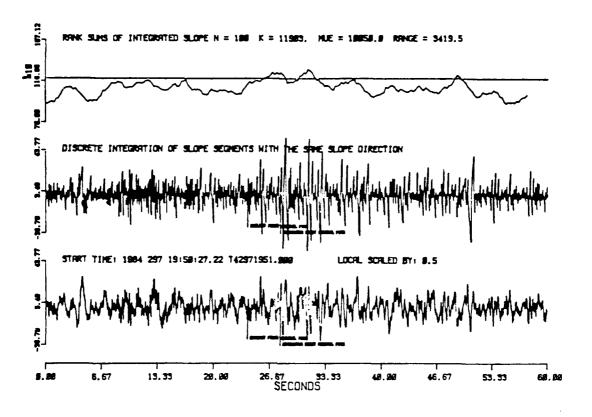
Event 37 Local P Arrival

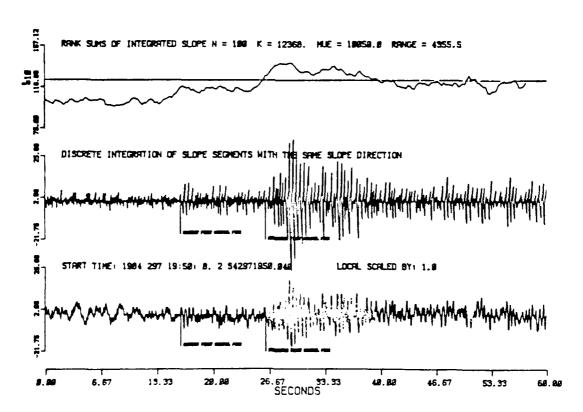












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CONTROLL RECOGNESS SERVICES DESCRIPTION

THE CURIOUS CASE OF THE MISSING EXPLOSION

Eugene Herrin Geophysical Laboratory Southern Methodist University

"Is there any point to which you would wish to draw my attention?"

This narrative could properly be classified as historical Most of what is reported here actually happened, some of the evenus occurred only in the imagination of writer. We begin by considering the capabilities of a regional network designed to monitor an area of thick salt deposits in the western portion of the Permian Basin of Texas and New Mexico. The stations in the network are at Lajitas, Texas; Hobart, Oklahoma; and Winnemucca, Nevada (see Figure 3). Noise levels at the three stations are based upon actual noise observed at these sites and at similar sites. The minimum background noise Lajitas is the lowest ever observed in the frequency band of 5 to This minimum noise level is shown in Figure 40 Hz. minimum levels reported for NORSAR (NORESS site) are somewhat higher, by about 10 dB or more. Measurements at Hobart in the frequency band 1 to 4 Hz show noise levels similar to NORSAR. The background level at Winnemucca based on early measurements is The Lajitas and between the Lajitas and the Hobart values. Winnemucca stations have state-of-the-art, three-component, short

[&]quot;To the curious incident of the dog in the night-time."

[&]quot;The dog did nothing in the night-time."

[&]quot;That was the curious incident," remarked Sherlock Holmes.
("Silver Blaze", A. Conan Doyle)

period instruments in 50 ft. boreholes. At Hobart, a 15 element, short period array of surface instruments is located on basement rock. This array is identical to the NORESS array without the outer ring.

Background levels shown in Figure 1 are the minimum noise levels under ideal conditions which exist only a small fraction of the time. The major source of background noise is the effect of wind at the site as is shown in Figure 2. NORSAR has noise levels similar to those at Lajitas under high-wind conditions. We expect to observe the same effect at the Hobart array.

Figure 3 shows the location of a number of events as well as the locations of the stations. GNOME was a 3 kt explosion tamped in salt which was actually observed at Lajitas, Hobart and Winnemucca. SLEUTH is a planned 3 kt decoupled shot in the Salado formation in the same general area as the GNOME event. Figure 4 shows the stratigraphic units in the area. The Salado salt provides the depth and thickness needed to decouple a 3 kt nuclear explosion.

The southeastern corner of New Mexico is an area dotted with potash mines as shown in Figure 5. Mining potassium bearing minerals from the salt and anhydrite units is accomplished using the room-and-pillar method with the separation of the ore being done on the surface near the working shafts. The area is almost a wasteland covered with mounds of discarded evaporites and dessication ponds. Once the mining has proceeded as far away from the working shaft as is practical, the pillars are systematically removed allowing the mine to subside. This

collapse leads to obvious surface effects over the area which has been mined. After this procedure is completed, a new shaft is dug and the operation is repeated. Travellers crossing this region must beware of dangerous scarps which develop in the highways above collapsing mines.

In the midst of this region under one of the mined areas a cavity with a 40 meter radius has been constructed at a depth of 1000 meters in the Salado salt. The salt removed in this process represents only a small addition to the wastes already present on the surface. A 3 kt nuclear device is placed in the cavity ready for the decoupled test which has been code-named SLEUTH.

The Pn signal levels from a 3 kt shot fully tamped in salt (GNOME) observed at the three stations in the monitoring network are given in Table 1, along with the distances to the stations. The very low signal level at Winnemucca resulted from the high attenuation of Pn across the Basin and Range province. Propagation to Lajitas and Hobart, however, is as expected in the mid-continent. We assume a decoupling ratio of 100 at 1 to 2 Hz for SLEUTH compared to GNOME, thus we can accurately predict the Pn amplitude levels (1-2 Hz) at the three stations for the decoupled shot.

On Tuesday, 25 June 1985, preparations were being made for a 4 kt HE shot (MINOR SCALE) at the White Sands test site (see Figure 3). A cold front was crossing Colorado and Utah at that time as shown in Figure 6. The pattern of fronts moving west to east shown in this map is typical of the weather pattern in this region in the spring and early summer and again in the fall. The frontal movements and wind patterns are highly predictable. On

Wednesday, 26 June (Julian day 177), the cold front had passed into New Mexico, but had not yet begun to affect the wind patterns at Lajitas. That afternoon an earthquake of about magnitude 3 occurred west of Amarillo, Texas, and was recorded on the high-frequency (sample rate 250 per sec) system operating at the Lajitas station. The distance to Lajitas from the epicenter is 670 km. Figure 8 shows the signal and the signal-to-noise spectrum for this event. There is good signal-to-noise-ratio to frequencies greater than 15 Hz. This event was located within the 16-element seismic network operated by Stone and Webster Engineering as part of the nuclear waste disposal survey in the Texas Panhandle; therefore, we were able to compute an accurate epicenter for the event using the network records. Digital data were available for a station 48 km from the epicenter so that a good displacement spectrum could be computed. This spectrum showed a clear corner at 6 Hz, a constant level at lower frequencies and a roll-off above the corner frequency of dB/decade (f $^{-3}$). Using this spectrum and the digital record at Lajitas, we were able to produce a good estimate of the apparent Q for Fn along this path. The value of Q was 246 which is consistent with Q(Fn) reported from northeast of Moscow along a from the Volga River to Vorkuta (Yegorkin and Kun, Izvestiya, 1978, Vol. 14, No. 4, 262-269).

On Thursday, 27 June (Day 178) 1985, the cold front had moved through Oklahoma and much of Texas (Figure 9). Winds at Hobart and Lajitas were 20 to 30 mph from the north. That morning MINOR SCALE was fired at White Sands, and thirty seconds

later the decoupled nuclear shot, SLEUTH, was fired. Figure 10 shows the signal from MINOR SCALE. The amplitude of Pn from this event was about 4 times as large as from the Amarillo earthquake the previous day, as would be expected based on the yield of MINOR SCALE. The wind at Lajitas was 20-30 mph at the time of the event (Figure 10), whereas conditions had been nearly calm during the recording of the Amarillo earthquake on the previous day. Even though the vertical instrument was located at a depth of 100 meters, wind-noise was a major problem on day 178. This effect is clearly seen in Figure 11, where the signal-to-noise spectrum falls to zero dB at 6 Hz for MINOR SCALE. This result can be contrasted with the effective bandwidth of 16 Hz seen for the more distant, significantly smaller event recorded on the previous day when the wind was nearly calm.

SLEUTH was not detected at Lajitas. Pn was below the noise level and the Lg wave train was swamped by the Lg signal from MINOR SCALE. Figure 12 shows the predicted displacement spectra for GNOME and SLEUTH based on the corner frequencies for 3 kt tamped and decoupled events given by Archambeau and the Q-value obtained from the Amarillo earthquake. The displacement spectra of the background noise on days 177 and 178 are also shown. From this figure we see that if SLEUTH had been fired on day 177, Pn would have been detected at Lajitas, in agreement with Archambeau's predictions. By picking the right time to fire SLEUTH, based on the weather patterns and the known time of MINOR SCALE, the most sensitive high-frequency station in the network was made incapable of detecting the event. At Hobart, Oklahoma, the wind was 15 to 20 knots (around 25 mph) from the north

(Figure 9). The surface array there could be expected to produce signal—to—noise improvement over a single surface instrument of about 3.5 at 1 to 2 Hz and about 2 at 10 Hz. The background noise; however, could be expected to be higher by at least a factor of 5 because of the high winds. Thus the Hobart array would fail to detect Pn from SLEUTH. Again, Lg would be lost in the coda of MINOR SCALE, and could not be pulled out by array processing because of the similarity of azimuths for the two events relative to the array. Because of poor propagation across the Basin and Range Province, Pn from SLEUTH could not be detected at Winnemucca no matter what the noise conditions were at that station. Thus we see that an excellent reginal network designed to monitor an area with salt deposits failed to detect a 3 kt decoupled shot.

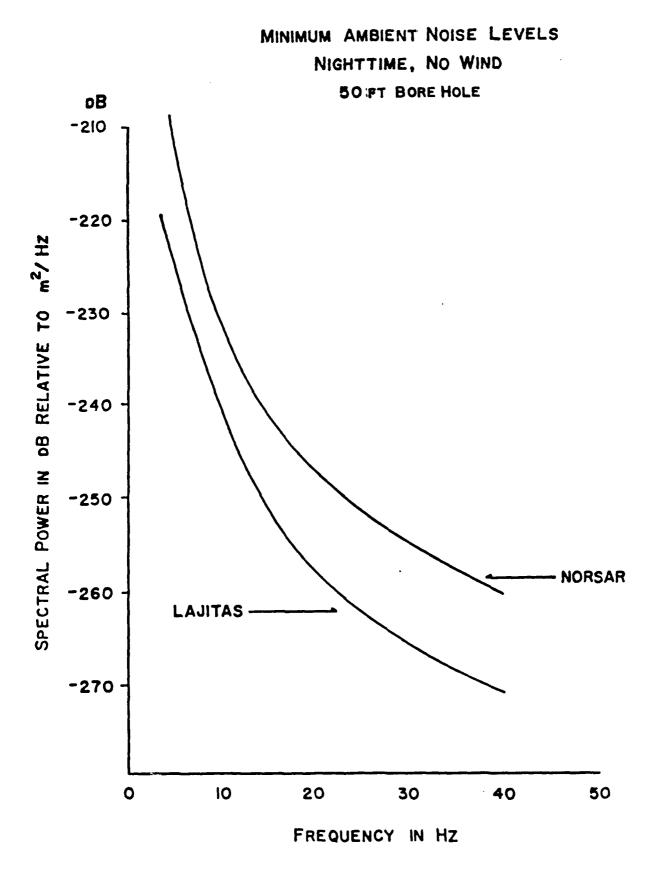
"The dog did nothing in the night-time."

The failure of the network occurred because the evader could pick the most advantageous time to fire the clandestine test. SLEUTH is not one-of-a-kind. Fredictable weather conditions similar to those on 27 June 1985, occur in West Texas and Oklahoma several times each year. Shots of 1/2 kt of HE or larger are not uncommon at White Sands Test Site. Cavity mining could go on year after year in southeastern New Mexico completely masked by the potash mining in the area. By waiting for the

same sequence of events that occurred on 27 June, several decoupled nuclear tests a year could be carried out with virtually no chance of detection.

Two factors are required for a successful clandestine test of the kind described in this paper. The first is a degradation of detection capability at the critical stations because of high wind. The second is a legitimate HE test under the control of the evader which can be used to mask the decoupled shot. The second factor must be regulated by legal means. Control of the first factor depends upon our ability to protect instruments from the effect of wind-induced seismic noise. Until these problems are solved, a clever evader can pick the right time to fire a decoupled shot with little risk of detection.

Acknowledgements. Dick Cromer (Sandia) and Faul Golden (SMU) computed the spectra used in this paper. Stone and Webster Engineering provided digital data from their network in the Texas Panhandle. This work was supported by DARPA/AFGL under contract No. F19628-85-K-0032 and by Sandia National Laboratories. John W. Harrington suggested the use of the quotation from A. Conan Doyle.



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Figure 1. Minimum Seismic Noise Levels at Lajitas and Norsar.

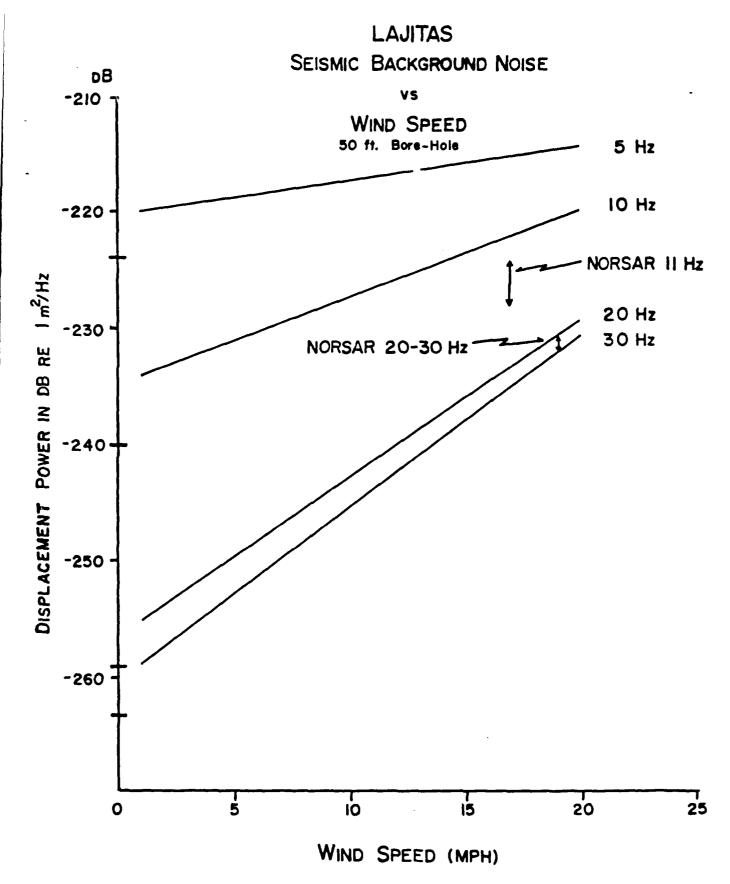


Figure 2. Seismic background versus wind speed at Lajitas.

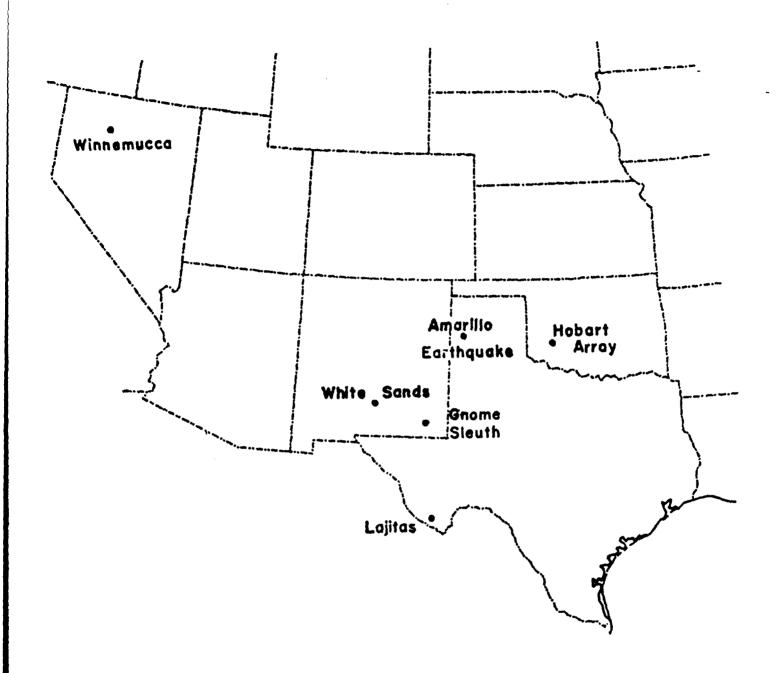


Figure 3. Map showing locations of seismic stations and events.

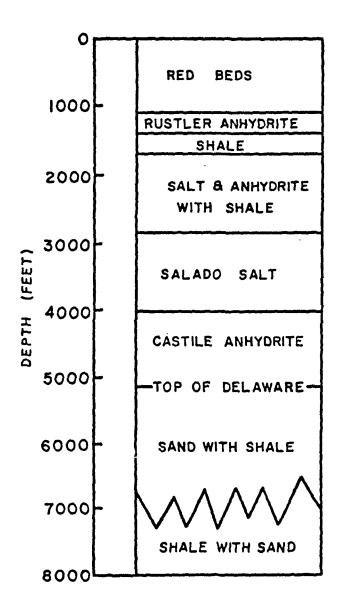


Figure 4. Stratigraphic section of GNOME/SLEUTH site.

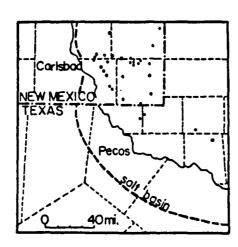


Figure 5. Map showing potash mining activity in the GNOME/SLEUTH vicinity. Dots indicate location of potash mines.

Table 1.

SIGNAL LEVELS

Station	(km) Distance	Pn Amplitudes (millimcrons)	
		GNOME	SLEUTH
Lajitas	378	42	0.4
Hobart	561	35	0.3
Winnemucca	1574	1	0.01

GNOME	3 Kt	tamped in sait
SLEUTH	3 kt	in salt cavity, radius 40 meters
		depth 1000 meters in Salado formation.

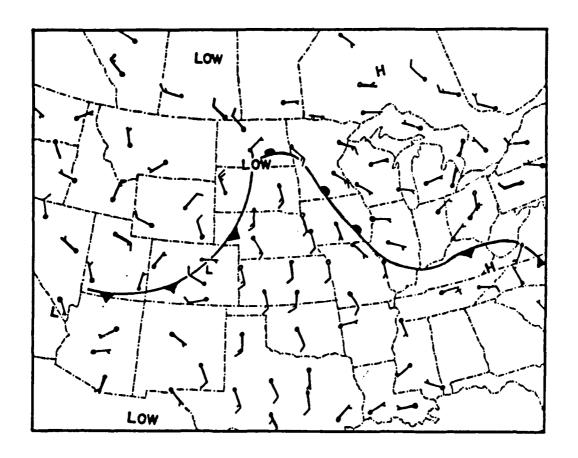


Figure 6. Weather map for Tuesday, 25 June 1985.

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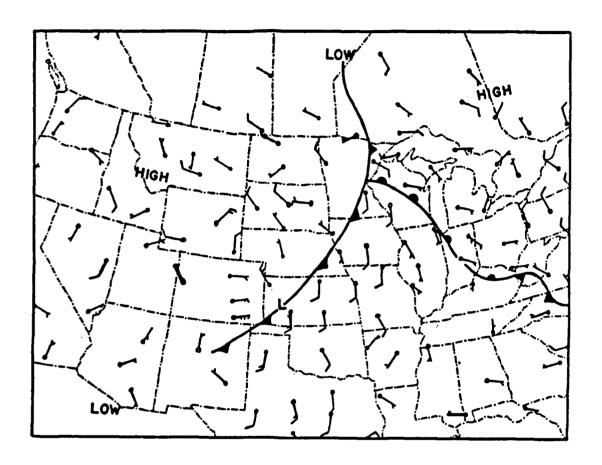
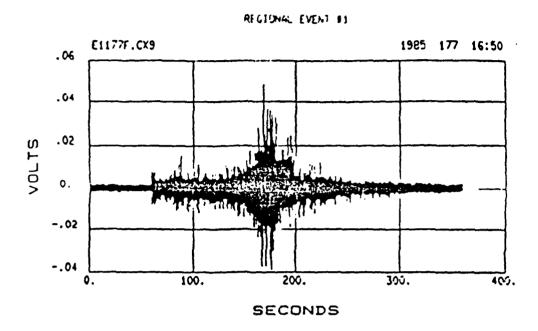


Figure 7. Weather map for Wednesday, 26 June 1985.



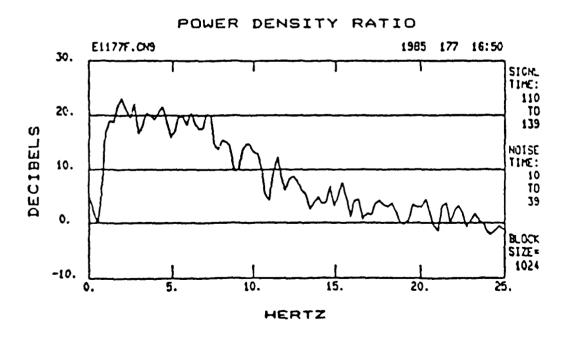


Figure 8. Amarillo earthquake, High-frequency recording (250 samples per second) from 330 ft. Z component at Lajitas and signal-to-noise spectrum for this record.

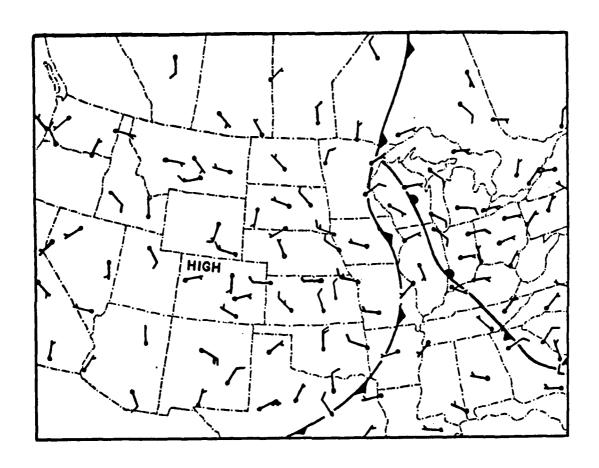
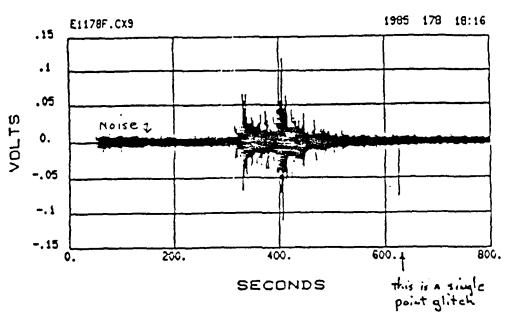


Figure 9. Weather map for Thursday, 27 June 1985.

WSMR EVENT 'MINOR SCALE'



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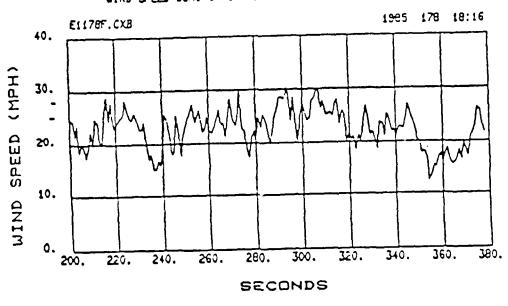


Figure 10. Minor Scale: Record from 330 ft. borehole instrument and plot of wind speed during P-wave arrival time.

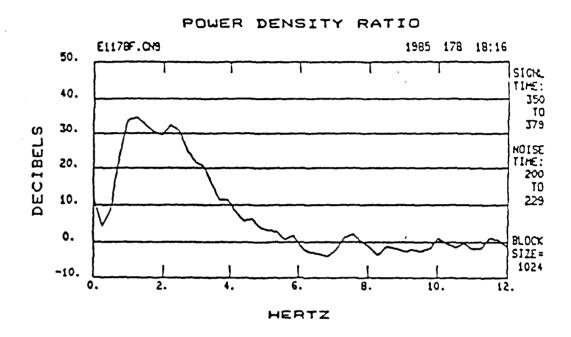


Figure 11. Minor Scale: Signal-to-noise spectrum from P-wave, 330 ft. borehole instrument.

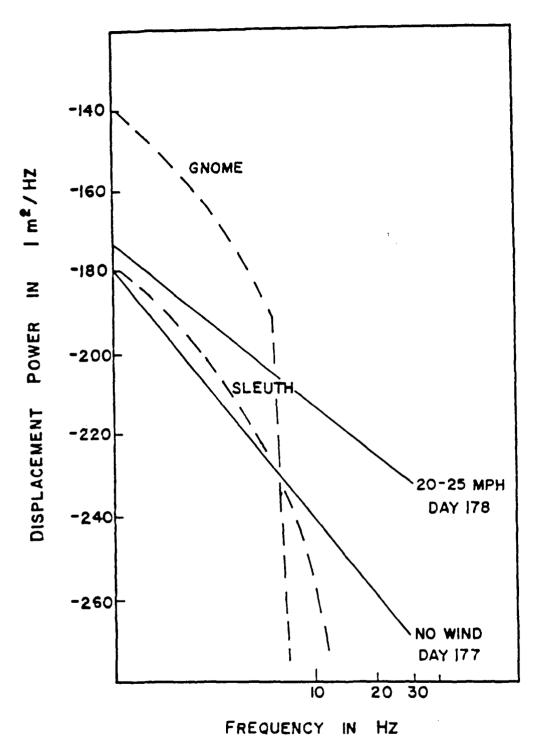


Figure 12. Displacement spectra for GNOME and SLEUTH at Lajitas and the background noise at Lajitas.

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