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PEAKEDNESS OF WEIGHTED AVERAGES OF JOINTLY DISTRIBUTED  
RANDOM VARIABLES. (U) FLORIDA STATE UNIV TALLAHASSEE  
DEPT OF STATISTICS W CHAN ET AL. AUG 87

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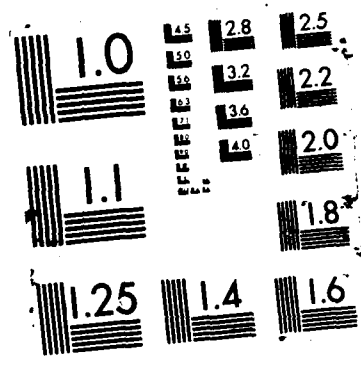
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FSU-STATISTICS-TR-M712R AFOSR-TR-87-1574

F/G 12/3

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AD-A185 611

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TATION PAGE

2

1a. REPORT SECURITY CLASSIFICATION

UNCLASSIFIED

RESTRICTIVE MARKINGS

2a. SECURITY CLASSIFICATION AUTHORITY

NA

3. DISTRIBUTION/AVAILABILITY OF REPORT

Approved for Public Release; Distribution Unlimited.

2b. DECLASSIFICATION/DOWNGRADING SCHEDULE

NA

4. PERFORMING ORGANIZATION REPORT NUMBER(S)

FSU TECHNICAL REPORT NO M712R

5. MONITORING ORGANIZATION REPORT NUMBER(S)

~~AFO SR NM Report No 85-184B~~  
AFO SR NM 87-1574

6a. NAME OF PERFORMING ORGANIZATION

FLORIDA STATE UNIVERSITY

6b. OFFICE SYMBOL (If applicable)

(If applicable)

7a. NAME OF MONITORING ORGANIZATION

AFO SR/NM

6c. ADDRESS (City, State and ZIP Code)

Department of Statistics  
Tallahassee, FL 32306-3033

7b. ADDRESS (City, State and ZIP Code)

Bldg. 410  
Bolling AFB, DC 20332-6448

8a. NAME OF FUNDING/SPONSORING ORGANIZATION

AFO SR

8b. OFFICE SYMBOL (If applicable)

NM

9. PROCUREMENT INSTRUMENT IDENTIFICATION NUMBER

F49620-82-K-0007

8c. ADDRESS (City, State and ZIP Code)

Bldg. 410  
Bolling AFB, DC

10. SOURCE OF FUNDING NOS.

PROGRAM ELEMENT NO.  
6.1102F

PROJECT NO.  
2304

TASK NO.  
A-5

WORK UNIT NO.

11. TITLE (Include Security Classification)

PEAKEDNESS OF WEIGHTED AVERAGES OR JOINTLY DISTRIBUTED RANDOM VARIABLES

12. PERSONAL AUTHOR(S)

Wai Chan, Dong Ho Park, and Frank Proschan

13a. TYPE OF REPORT

~~Technical Report~~ Journal

13b. TIME COVERED

FROM \_\_\_\_\_ TO \_\_\_\_\_

14. DATE OF REPORT (Yr., Mo., Day)

August 1987

15. PAGE COUNT

1-6

16. SUPPLEMENTARY NOTATION

17. COSATI CODES

FIELD GROUP SUB. GR.

18. SUBJECT TERMS (Continue on reverse if necessary and identify by block number)

Key Words - Peakedness, convex combination, majorization, Schur-concave density, Cauchy distribution.

19. ABSTRACT (Continue on reverse if necessary and identify by block number)

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*Cauchy distribution convergence*

*sub 2*

DTIC ELECTED  
OCT 29 1987

20. DISTRIBUTION/AVAILABILITY OF ABSTRACT

UNCLASSIFIED/UNLIMITED  SAME AS RPT.  DTIC USERS

21. ABSTRACT SECURITY CLASSIFICATION

UNCLASSIFIED

CO D

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22c. OFFICE SYMBOL

AFO SR/NM

REVISION

**AFOSR-TR- 87-1574**

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by

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FSU Technical Report No. M 712R  
AFOSR Technical Report No. 85-184R

August, 1987

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AMS (1981) subject classification No. 60E15.

87 10 15 040

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This note extends the Proschan (1965) result on peakedness comparison for convex combinations of i.i.d. random variables from a  $PF_2$  density. Now the underlying random variables are jointly distributed from a Schur-concave density. The result permits a more refined description of convergence in the Law of Large Numbers.



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## 1. Introduction

Proschan (1965) shows that:

**1.1 Theorem.** Let  $f$  be  $PF_2$ ,  $f(t) = f(-t)$  for all  $t$ ,  $X_1, \dots, X_n$  independently distributed with density  $f$ ,  $\underline{p} \stackrel{m}{\geq} \underline{p}'$ ,  $\underline{p}, \underline{p}'$  not identical,  $\sum_1^n p_i = 1 = \sum_1^n p'_i$ . Then  $\sum_1^n p_i X_i$  is strictly more peaked than  $\sum_1^n p'_i X_i$ .

(Definitions of majorization ( $\underline{p} \stackrel{m}{\geq} \underline{p}'$ ),  $PF_2$  density, and peakedness are presented in Section 2.)

The Law of Large Numbers asserts that the average of a random sample converges to the population mean under certain conditions. Roughly speaking, Theorem 1.1 states that a weighted average of i.i.d. random variables converges more rapidly in the case in which weights are close together as compared with the case in which the weights are diverse.

In the present note, we extend the basic univariate result to the multivariate situation in which the underlying random variables have a joint Schur-concave density. Theorem 2.4 presents the precise statement of the multivariate extension.

## 2. Peakedness comparisons

The theory of majorization is exploited in this section to obtain more general versions of the result of Proschan (1965). We begin with some definitions. The definition of peakedness was given by Birnbaum (1948).

**Definition 2.1.** Let  $X$  and  $Y$  be real valued random variables and  $a$  and  $b$  real constants. We say that  $X$  is more peaked about  $a$  than  $Y$  about  $b$  if

$$P(|X-a| \geq t) \leq P(|Y-b| \geq t)$$

for all  $t \geq 0$ . In the case  $a = 0 = b$ , we simply say that  $X$  is more peaked than  $Y$ .

Next we define the ordering of majorization among vectors.

**Definition 2.2.** Let  $a_1 \geq \dots \geq a_n$  and  $b_1 \geq \dots \geq b_n$  be decreasing rearrangements of the components of the vectors  $a$  and  $b$ . We say that  $a$  majorizes  $b$  (written  $a \stackrel{m}{\geq} b$ ) if

$$\sum_{i=1}^n a_i = \sum_{i=1}^n b_i$$

and

$$\sum_{i=1}^k a_i \geq \sum_{i=1}^k b_i \quad \text{for } k = 1, \dots, n-1.$$

**Definition 2.3.** A real valued function  $f$  defined on  $R^n$  is said to be a Schur-concave function if  $f(a) \leq f(b)$  whenever  $a \geq b$ .

A nonnegative function  $f$  on  $(-\infty, \infty)$  is called a Pólya frequency function of order 2 ( $PF_2$ ) if  $\log f$  is concave. If  $f$  is a  $PF_2$  function then  $\phi(x) = \prod f(x_i)$  is Schur-concave. Thus the random vector  $x = (X_1, \dots, X_n)$  has a Schur-concave density under the conditions of Theorem 1.1. A function  $f$  defined on  $R^n$  is said to be sign-invariant if  $f(x_1, \dots, x_n) = f(|x_1|, \dots, |x_n|)$ . In the following theorem, we give a peakedness comparison for random variables with a sign-invariant Schur-concave density.

**Theorem 2.4.** Suppose the random vector  $X = (X_1, \dots, X_n)$  has a sign invariant Schur-concave density. Then for all  $t \geq 0$ ,

$$\psi(a_1, \dots, a_n) = P(\sum a_i X_i \leq t)$$

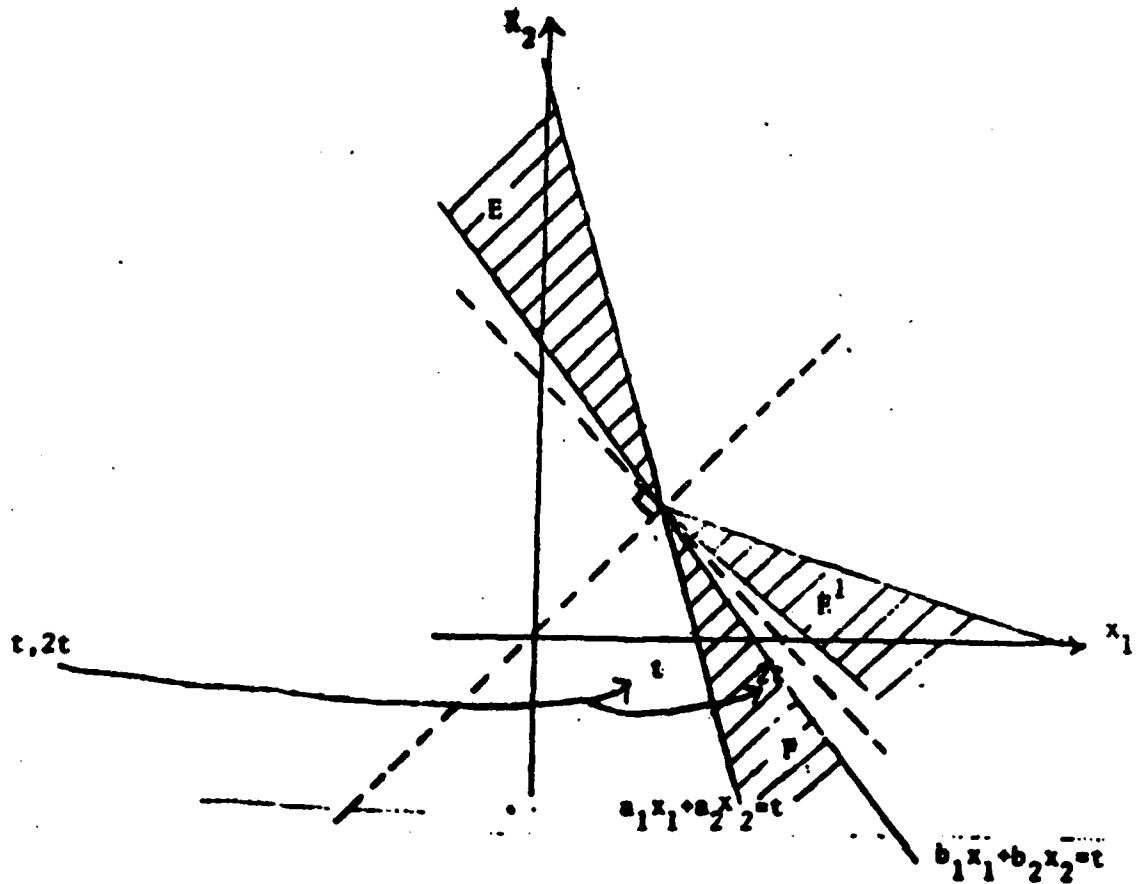
is a Schur-concave function of  $a = (a_1, \dots, a_n)$ ,  $a_i \geq 0$  for all  $i$ . Equivalently,  $\sum b_i X_i$  is more peaked than  $\sum a_i X_i$  whenever  $a \stackrel{m}{\geq} b$ .

**Proof.**

Without loss of generality, we may assume that  $\sum a_i = 1$ . We first consider the case  $n = 2$ .

Let  $\underline{a} = (a_1, a_2)$ ,  $\underline{b} = (b_1, b_2)$ ,  $\underline{a} \stackrel{m}{\geq} \underline{b}$ . Since  $X_1, X_2$  are exchangeable, we may further assume that

$a_1 > b_1 \geq 1/2 \geq b_2 > a_2$ . To show that  $P(a_1 X_1 + a_2 X_2 \leq t) \leq P(b_1 X_1 + b_2 X_2 \leq t)$  for  $t \geq 0$ , consider the following diagram:





Since  $a_1 > b_1 \geq 1/2$ , both lines intersect the  $x_1$ -axis in the interval  $[t, 2t]$  and they intersect the 45 degree line at the point  $(t, t)$  ( $a_1 + a_2 = b_1 + b_2 = 1$ ). We must show that  $P(E) \leq P(F)$ . Now reflect  $E$  across the 45 degree line to form the wedge  $E'$ . Then  $P(E) = P(E')$  because the joint density  $f$  is invariant under permutation. For  $k \geq 0$ , the line  $x_1 - x_2 = k$  intersects  $E'$  at the line segment joining  $(t + b_1 k, t - b_2 k)$  and  $(t + a_1 k, t - a_2 k)$ , and it intersects  $F$  at the line segment joining  $(t + a_2 k, t - a_1 k)$  and  $(t + b_2 k, t - b_1 k)$ . Note that both segments are of equal length. But  $f$  sign-invariant and Schur-concave implies that

$$\begin{aligned} f(t + b_1 k, t - b_2 k) &= f(t + b_1 k, b_2 k - t) \\ &\leq f(t + b_2 k, b_1 k - t) \\ &= f(t + b_2 k, t - b_1 k). \end{aligned}$$

This last fact then clearly implies that  $P(E') \leq P(F)$  by conditioning on  $X_1 - X_2$ .

The result for  $n \geq 3$  now follows since

$$\begin{aligned} P(\sum a_i X_i \leq t) \\ = E [P(a_1 X_1 + a_2 X_2 \leq t - \sum_3^n a_i X_i \mid X_3, \dots, X_n)] \end{aligned}$$

and the conditional density  $f(x_1, x_2 \mid x_3, \dots, x_n)$  is also Schur-concave and sign invariant.

For an example of a Schur-concave density function that is also sign-invariant, consider the multivariate Cauchy density:

$$f(x_1, \dots, x_n) = \pi^{-(n+1)/2} \Gamma((n+1)/2) (1 + \sum_{i=1}^n x_i^2)^{-(n+1)/2}$$

The following result is an immediate consequence of Theorem 2.4.

**Corollary 2.5.** Let  $X_1, \dots, X_n$  be random variables with joint Schur-concave sign-invariant density  $f$ . Then  $\frac{1}{k} \sum_{i=1}^k X_i$  is increasing in peakedness as  $k$  increases from 1 to  $n$ .

**Proof.**

Let  $\underline{a}_1 = (1, 0, \dots, 0)$ ,  $\underline{a}_2 = (\frac{1}{2}, \frac{1}{2}, 0, \dots, 0)$ , ..., and  $\underline{a}_n = (\frac{1}{n}, \dots, \frac{1}{n})$ , where each vector contains  $n$  components. Then  $\underline{a}_1 \succeq \underline{a}_2 \succeq \dots \succeq \underline{a}_n$ . The result follows from Theorem 2.4.

Suppose  $\underline{X} = (X_1, \dots, X_n)$  and  $\underline{Y} = (Y_1, \dots, Y_n)$  are independently distributed with respective densities  $f$  and  $g$  where both  $f$  and  $g$  are Schur-concave and sign invariant. Then Theorem 2.4 implies that  $\sum b_i (X_i + Y_i)$  is more peaked than  $\sum a_i (X_i + Y_i)$  whenever  $\underline{a} \succeq \underline{b}$ . This is true because the convolution of Schur-concave functions is Schur-concave. However, if  $Y_1, \dots, Y_n$  are i.i.d. Cauchy, then the joint density  $g$  given by

$$(2.1) \quad g(y_1, \dots, y_n) = \left(\frac{a}{\pi}\right)^n \prod_{i=1}^n (1 + a^2 y_i^2)^{-1}, \quad a > 0,$$

is not Schur-concave. In Theorem 2.6 below, we show that  $\sum b_i (X_i + Y_i)$  is more peaked than  $\sum a_i (X_i + Y_i)$  whenever  $\underline{a} \succeq \underline{b}$ . This result identifies a different class of densities for which the conclusion of Theorem 2.4 holds.

**Theorem 2.6.** Suppose that the random vector  $\underline{X} = (X_1, \dots, X_n)$  has a sign-invariant Schur-concave density  $f$ . Let  $Y = Y_1, \dots, Y_n$  be i.i.d. Cauchy with joint density  $g$  as given in (2.1). Let  $\underline{X}$  and  $\underline{Y}$  be independent, and  $\underline{a} \succeq \underline{b}$  where  $a_i \geq 0, b_i \geq 0$  for all  $i$  and  $1 = \sum_1^n a_i = \sum_1^n b_i$ . Then  $\sum_1^n b_i (X_i + Y_i)$  is more peaked than  $\sum_1^n a_i (X_i + Y_i)$ .

**Proof.**

Since  $f$  is sign invariant both  $\sum_1^n a_i X_i$  and  $\sum_1^n b_i X_i$  are symmetric random variables. We use the

fact that  $\sum_1^n a_i Y_i, \sum_1^n b_i Y_i$  have the same distribution as does  $Y_1$ . The result now follows

Theorem 2.4 and the Lemma of Birnbaum (1948) by noting that  $Y_1$  has a symmetric and unimodal density.

## References

- [1] Birnbaum, Z.W. (1948). On random variables with comparable peakedness. Ann. Math. Statist. 19 76-81.
- [2] Proschan, F. (1965). Peakedness of distributions of convex combinations. Ann. Math. Statist. 36 1703-1706.

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