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Exact Operating Characteristics for Linear Sum of Envelopes of Narrowband Gaussian Process and Sinewave

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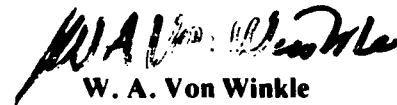
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Preface

This research was conducted under NUSC Project No. A96420, "Combat System Engineering and Analysis Program," Principal Investigator Jeffrey S. Cohen (Code 61), Sponsoring Activity NAVSEA PMS-393, Program Manager S. Goldstein. This research was also conducted under NUSC Project No. A75205, Subproject No. ZR0000101, "Applications of Statistical Communication Theory to Acoustic Signal Processing," Principal Investigator Dr. Albert H. Nuttall (Code 33), Program Manager CAPT Z. L. Newcomb, Naval Material Command (MAT-05B).

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<p>The characteristic function of a linear sum of M independent Rice variates is derived and evaluated exactly and then used in a numerical procedure to determine the exceedance distribution function, as a function of the threshold, the input signal-to-noise ratio, and M. Plots of the detection probability and false alarm probability for a wide range of signal-to-noise ratios are given, for values of M up to 8192. In addition, the required threshold values and input signal-to-ratios are tabulated and plotted for specific values of M, false alarm probability, and</p>		

20. (Cont'd)

detection probability. A program and explanation are included for those users interested in extending results to their particular application.

TABLE OF CONTENTS

	Page
LIST OF ILLUSTRATIONS	ii
LIST OF TABLES	iii
LIST OF SYMBOLS	iv
INTRODUCTION	1
METHOD OF EVALUATION	3
Characteristic Function Details	3
Special Cases	7
Asymptotic Performance for Large M	8
RESULTS	11
SUMMARY	33
APPENDICES	
A. DERIVATION OF RICE CHARACTERISTIC FUNCTION	A-1
B. DESCRIPTION OF PROGRAMS AND LISTINGS	B-1
REFERENCES	R-1



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By _____	
Distribution/	
Availability Codes	
Dist	Avail and/or Special
A-1	

LIST OF ILLUSTRATIONS

FIGURE	PAGE
1 Required Input S/N for $P_D=.5$	13
2 Required Input S/N for $P_D=.9$	14
3 Required Input S/N for $P_D=.95$	15
4 Required Input S/N for $P_D=.99$	16
5 Required Input S/N for $P_D=.999$	17
6 Receiver Operating Characteristics for $M=1$	19
7 Receiver Operating Characteristics for $M=2$	20
8 Receiver Operating Characteristics for $M=4$	21
9 Receiver Operating Characteristics for $M=8$	22
10 Receiver Operating Characteristics for $M=16$	23
11 Receiver Operating Characteristics for $M=32$	24
12 Receiver Operating Characteristics for $M=64$	25
13 Receiver Operating Characteristics for $M=128$	26
14 Receiver Operating Characteristics for $M=256$	27
15 Receiver Operating Characteristics for $M=512$	28
16 Receiver Operating Characteristics for $M=1024$	29
17 Receiver Operating Characteristics for $M=2048$	30
18 Receiver Operating Characteristics for $M=4096$	31
19 Receiver Operating Characteristics for $M=8192$	32

LIST OF TABLES

TABLE	PAGE
1. Normalized Thresholds Required for Specified M and P_{FA}	12
2. Parameter Cards	B-5
3. Table Cards	B-5
4. File Cards	B-5
5. Command Cards	B-5
6. Plot Device Cards	B-6
7. Sample Input Deck for P_D vs. P_{FA}	B-6
8. Sample Input Deck for SNR vs. M	B-7
9. Sample Input Deck for Printing SNR	B-7
10. Print Out of SNR vs. M	B-8
11. Description of Subroutines	B-10

LIST OF SYMBOLS

M	Number of independent envelope samples summed
e_m	m-th envelope sample
x	Decision variable; sum of M envelope samples
P_D	Detection probability
P_{FA}	False alarm probability
A	Sinewave amplitude
σ	Noise standard deviation
α	A/σ , voltage measure of input signal-to-noise ratio
S/N	$\alpha^2/2$, power measure of input signal-to-noise ratio
$p_e(u)$	Probability density function of envelope random variable e_m
$f_e(\xi)$	Characteristic function of envelope random variable e_m
I_0	Modified Bessel function of order zero
${}_1F_1$	Confluent hypergeometric function
μ_x	Mean of random variable x
$f_x(\xi)$	Characteristic function of decision variable x
$Q_x(u)$	Exceedance distribution function of random variable x; (8)
Im, Re	Imaginary part, Real part
Δ	Sampling increment in ξ ; (12)
N	Size of FFT; (13)-(14)
L	Limit employed on integral on ξ in (9)
b	Bias employed to shift random variable x
Φ	Cumulative distribution function of normalized Gaussian random variable; (19)
Φ^{-1}	Inverse function to Φ defined in (19)
σ_x	Standard deviation of random variable x
dB	Required input signal-to-noise ratio per-sample in decibels
β	Parameter incorporating specified P_D and P_{FA} ; (28)

EXACT OPERATING CHARACTERISTICS FOR LINEAR SUM OF
ENVELOPES OF NARROWBAND GAUSSIAN PROCESS AND SINEWAVE

INTRODUCTION

The operating characteristics for a linear envelope-detector of a sinewave in narrowband Gaussian noise, followed by summation of M independent envelope samples, were presented in [1] and [2, sect. 8.3]. That approach was based upon evaluation of the first 31 moments of the envelope variate and their use in a type A Gram-Charlier series approximation, or in modified approximations involving averages over different numbers of terms in the series [1, pp. 758-9]. However, there are possible pitfalls to the above approach. First, evaluation of very low exceedance probabilities, like 10^{-10} , may be inaccurate; see [1, Fig. 1]. Second, the effect of a systematic error would be hard to detect, if present, since the method yields only an approximation to the exceedance distribution function, and not its exact value.

We will use an exact approach here, based upon evaluation of the characteristic function of the envelope detector output, from which the exceedance distribution function can be precisely evaluated numerically [3,4]. In this fashion, we avoid moment evaluations altogether; we can evaluate false alarm probabilities in the 10^{-10} range easily (with double precision computer arithmetic); and we can control truncation and aliasing errors to any desired degree; see [3] for details. The results of [4] can not be applied here because each independent envelope sample is the result of a nonlinear operation, namely a square root, applied to a sum of two squares of Gaussian random variables with non-zero means.

In the plots of detection probability vs. false alarm probability to be presented herein, both abscissa and ordinate use the same normal probability scales, regardless of the number of envelope samples M considered. This allows for easier interpolation, and is in distinction to [1], where a different false alarm probability abscissa was used for each M [1, pp. 759-62]. Also, the parameter employed here for indexing the curves is α , a voltage signal-to-noise ratio which is equal to the ratio of the sinewave amplitude to the

rms noise level, rather than the dB parameter employed in [1]. This leads to curves that are more nearly equally spaced, and therefore to easier and finer interpolation capability.

Finally, we present five figures for the required input signal-to-noise ratio per sample required to realize specified false alarm and detection probabilities, as a function of M , the number of envelope samples added. The five figures correspond to detection probability $P_D = .5, .9, .95, .99, \text{ and } .999$ respectively, and each figure contains false alarm probabilities $P_{FA} = 10^{-n}$ for $n=1(1)8$. This total of 40 curves greatly augments the 2 cases presented in [1, Fig. 16] and [2, Fig. 8.18].

A program for the evaluation of the input signal-to-noise ratio required for a specified set of values of M , P_{FA} , and P_D is furnished, along with an explanation of its use. In this fashion, values of M , P_{FA} , and P_D intermediate to those considered here can be easily investigated.

METHOD OF EVALUATION

Characteristic Function Details

In [3,4], a method of calculating the cumulative and exceedance distribution functions directly from a given characteristic function was presented. To utilize those results here, we need the characteristic function of summation random variable

$$x = \sum_{m=1}^M e_m, \quad (1)$$

where e_m is the envelope of a narrowband filter output with a sinewave signal of amplitude A and Gaussian noise of power σ^2 . Through proper normalization, the probability density function of envelope e_m takes the familiar Rice form

$$p_e(u) = u \exp\left(-\frac{u^2 + \alpha^2}{2}\right) I_0(\alpha u) \quad \text{for } u \geq 0, \quad (2)$$

where the single parameter

$$\alpha = \frac{A}{\sigma} \quad (3)$$

is a voltage measure of signal-to-noise ratio per envelope sample. The power measure of signal-to-noise ratio per sample is

$$\frac{S}{N} = \frac{A^2/2}{\sigma^2} = \frac{\alpha^2}{2}. \quad (4)$$

The quantities in (3) and (4) will be referred to as input signal-to-noise ratios, since they are per-sample measures, prior to the summation in (1) which yields the output or decision variable x .

The characteristic function corresponding to random variable e in (2) is given by Fourier transform

$$f_e(\xi) = \int_{-\infty}^{+\infty} du \exp(i\xi u) p_e(u) = \int_0^{+\infty} du u \exp\left(i\xi u - \frac{u^2 + \alpha^2}{2}\right) I_0(\alpha u) \quad , \quad (5)$$

and will be called the Rice characteristic function. A series expansion for (5) is developed in appendix A, and has been programmed in double precision for numerical use here. As a particular special case, for $\alpha=0$, no signal, we have the Rayleigh probability density function and characteristic function:

$$p_e^{(0)}(u) = u \exp(-u^2/2) \quad \text{for } u \geq 0 \quad ,$$

$$f_e^{(0)}(\xi) = \exp(-\xi^2/2) \left[{}_1F_1\left(-\frac{1}{2}; \frac{1}{2}; \frac{\xi^2}{2}\right) + i\left(\frac{\pi}{2}\right)^{1/2} \xi \right] \quad . \quad (6)$$

The latter follows by use of [5, 3.896 3.4] and via manipulation of the hypergeometric function series along with Kummer's transformation [5, 9.212 1]. Formula (6) is particularly attractive numerically, since the series expansion of ${}_1F_1$ contains all positive terms except for one. It should be observed that the imaginary part of Rayleigh characteristic function $f_e^{(0)}(\xi)$ in (6) decays very rapidly with ξ ; this useful feature will also be shared by the Rice characteristic function, $f_e(\xi)$, and is due to the fact that the odd part of the Rice probability density function in (2) is smooth for all u , and is in fact entire in u , for any α . By contrast, the even part of the Rice probability density function in (2) has a discontinuous derivative for real u , thereby leading to slow decay of the real part of $f_e(\xi)$.

The characteristic function of output variable x in (1), for statistically independent envelope samples $\{e_m\}$, is given by

$$f_x(\xi) = [f_e(\xi)]^M \quad , \quad (7)$$

in terms of the Rice characteristic function (5). This relation could be used directly to find the exceedance distribution function of x according to [3, (5)-(6)]

$$Q_x(u) = \int_u^{+\infty} dt p_x(t) = \frac{1}{2} + \int_{0^+}^{+\infty} d\xi \operatorname{Im} \left\{ \exp(-iu\xi) \frac{f_x(\xi)}{\pi\xi} \right\} \quad . \quad (8)$$

However, the slow decay of $\text{Re}\{f_x(\xi)\}$ prompts us to use a modified version given in [6, (15)]:

$$Q_x(u) = \frac{2}{\pi} \int_{0^+}^{+\infty} \frac{d\xi}{\xi} \cos(u\xi) \text{Im}\{f_x(\xi)\} \quad \text{for } u > 0 \quad (9)$$

This form is applicable to positive random variables, of which x , as given by (1) and (2), is certainly a member.

To see why form (9) is preferred over (8), we develop (7) as

$$f_x(\xi) = [f_r(\xi) + if_i(\xi)]^M = \sum_{m=0}^M \binom{M}{m} i^m [f_i(\xi)]^m [f_r(\xi)]^{M-m} \quad (10)$$

where $f_r(\xi)$ and $f_i(\xi)$ are the real and imaginary parts of Rice characteristic function $f_e(\xi)$. Then

$$\text{Im}\{f_x(\xi)\} = \sum_{\substack{m=1 \\ m \text{ odd}}}^M (-1)^{\frac{m-1}{2}} \binom{M}{m} [f_i(\xi)]^m [f_r(\xi)]^{M-m} \quad (11)$$

contains $f_i(\xi)$ to at least the first power in all terms, thereby yielding a rapid decay with ξ .

Development (11) has been used to show why $\text{Im}\{f_x(\xi)\}$ decays rapidly with ξ . However, when we employ (9) in a numerical evaluation, we simply take the imaginary part of the power in (7), and do not use (11) at all; (11) is an alternating series of large terms for large M .

Actual numerical evaluation of (9) proceeds as follows [3]: for the Trapezoidal rule with sampling increment Δ in ξ ,

$$Q_x(u) \cong \frac{2}{\pi} \left[\frac{1}{2} x^\Delta + \sum_{n=1}^{\infty} \frac{1}{n} \cos(un\Delta) \text{Im}\{f_x(n\Delta)\} \right] \quad (12)$$

where we used $f_x(\xi) \sim 1 + i\mu_x \xi$ as $\xi \rightarrow 0$. Then, restricting the u values to a particular selection,

$$Q_x\left(\frac{2\pi m}{N\Delta}\right) = \frac{2}{\pi} \left[\frac{1}{2} \mu_x \Delta + \sum_{n=1}^{\infty} \frac{1}{n} \cos(2\pi mn/N) \operatorname{Im}\{f_x(n\Delta)\} \right] =$$

$$= \frac{2}{\pi} \operatorname{Re} \sum_{n=0}^{N-1} z_n \exp(-i2\pi mn/N) \quad , \quad (13)$$

where collapsed sequence $\{z_n\}_0^{N-1}$ is defined as

$$z_0 = \frac{1}{2} \mu_x \Delta + \sum_{j=1}^{\infty} \frac{1}{jN} \operatorname{Im}\{f_x(jN\Delta)\} \quad ,$$

$$z_n = \sum_{j=0}^{\infty} \frac{1}{n+jN} \operatorname{Im}\{f_x((n+jN)\Delta)\} \quad \text{for } 1 \leq n \leq N-1 \quad . \quad (14)$$

Form (13) is particularly attractive since it can be accomplished via an N -point FFT. It can be shown that only the values for $0 \leq m \leq N/2$ are useful in (13); the remainder are heavily aliased and must be discarded. Thus there is a trade-off: use of only the imaginary part of $f_x(\xi)$ results in aliasing twice as coarse. However, the rapid decay of the imaginary part far outweighs the aliasing.

The summations in (12) and (14) cannot be conducted to infinity. Rather the integral on ξ in (9) is terminated at limit L , where the truncation error is guaranteed to be sufficiently small. A trial and error procedure [3] yielded the following rules which control the truncation and aliasing errors:

$$L = \min(9, 17/\sqrt{M}),$$

$$\Delta = .12/\sqrt{M},$$

$$b = \min(0, -M\sqrt{\pi/2} + \sqrt{M}6). \quad (15)$$

The inverse \sqrt{M} dependence of L and Δ for large M can be anticipated by observing that the characteristic function of random variable x in (1) then

approaches a Gaussian function with argument proportional to M^2 . The bias (or shift) b is added to random variable x in order to yield a new random variable that remains just positive, even for large M ; this allows us to take maximum advantage of the fundamental aliasing interval $(0, \pi/\Delta)$ in u in (12) and (13). The linear term (in M) of b in (15) is due to the mean of the Rayleigh variate (for $\alpha=0$) which is $\sqrt{\pi/2}$; the algebraic term in \sqrt{M} is due to the fact that the standard deviation of random variable x in (1) increases according to \sqrt{M} .

In order to use this characteristic function approach, we also need the mean of random variable x in (1). Using (2), this is given by [5, 6.631 1]

$$\begin{aligned} \mu_x = M\mu_e &= M \int_0^{\infty} du u^2 \exp\left(-\frac{u^2 + \alpha^2}{2}\right) I_0(\alpha u) = \\ &= M \left(\frac{\pi}{2}\right)^{1/2} \exp\left(-\frac{\alpha^2}{2}\right) {}_1F_1\left(\frac{3}{2}; 1; \frac{\alpha^2}{2}\right) \end{aligned} \quad (16)$$

This non-alternating series yields accurate values for the mean.

Special Cases

For general M , the characteristic function approach described above must be used. However, for $M = 1$ and 2, closed form expressions for the false alarm and detection probabilities are possible. Specifically, from (1) and (2), for $u \geq 0$,

$$\left. \begin{aligned} P_{FA} &= \int_u^{\infty} dt p_e(t) = \int_u^{\infty} dt t \exp(-t^2/2) = \exp(-u^2/2) \\ P_D &= \int_u^{\infty} dt t \exp\left(-\frac{t^2 + \alpha^2}{2}\right) I_0(\alpha t) = Q(\alpha, u) \end{aligned} \right\} \text{for } M = 1. \quad (17)$$

And for $M = 2$, the false alarm probability can be determined by convolving two Rayleigh probability density functions of the form of (6), to give, for $u \geq 0$,

$$P_{FA} = \exp(-u^2/2) + \sqrt{\pi} u \exp(-u^2/4) \left[\Phi\left(\frac{u}{\sqrt{2}}\right) - \frac{1}{2} \right] \quad \text{for } M = 2. \quad (18)$$

Here, Φ is the cumulative distribution function of a normalized Gaussian random variable:

$$\Phi(u) = \int_{-\infty}^u dt (2\pi)^{-1/2} \exp(-t^2/2) \quad . \quad (19)$$

The detection probability of random variable x in (1) is not available in closed form for $M > 1$.

Asymptotic Performance for Large M

For large M , decision variable x in (1) is approximately Gaussian. The mean of x was given in (16); a similar approach for the mean square of x yields the variance as

$$\sigma_x^2 = M\sigma_e^2 = M(2 + \alpha^2 - u_e^2) \quad . \quad (20)$$

The probability density function of x is then approximately

$$p_x(u) \cong \frac{1}{\sqrt{2\pi} \sigma_x} \exp\left[-\frac{(u - \mu_x)^2}{2\sigma_x^2}\right], \quad (21)$$

with exceedance distribution function

$$Q_x(u) \cong \Phi\left(\frac{\mu_x - u}{\sigma_x}\right) = \Phi\left(\frac{M\mu_e - u}{\sqrt{M}\sigma_e}\right) \quad . \quad (22)$$

For input signal-to-noise ratio $S/N=0$, we have $\alpha=0$ from (4), and (22), (16), and (20) specialize to

$$P_{FA} \cong \Phi\left(\frac{M\sqrt{\pi/2} - u}{\sqrt{M}\sqrt{2 - \frac{\pi}{2}}}\right) \quad . \quad (23)$$

On the other hand, for $S/N > 0$, (22) yields the detection probability P_D . We now use the inverse function $\tilde{\Phi}$ to definition (19) and solve (23) and (22) according to

$$\frac{M\sqrt{\pi/2-u}}{\sqrt{M}\sqrt{2-\frac{\pi}{2}}} = \tilde{\Phi}(P_{FA}) \quad , \quad \frac{M\mu e^{-u}}{\sqrt{M}\sigma_e} = \tilde{\Phi}(P_D) \quad . \quad (24)$$

Eliminating threshold u in (24), we have

$$\sigma_e \tilde{\Phi}(P_D) = \sqrt{M} \left(\mu e^{-\sqrt{\frac{\pi}{2}}} \right) + \sqrt{2-\frac{\pi}{2}} \tilde{\Phi}(P_{FA}) \quad . \quad (25)$$

But also, for large M , the required per-sample input signal-to-noise ratio α will be small, giving

$$\begin{aligned} \mu_e &= \sqrt{\frac{\pi}{2}} {}_1F_1\left(-\frac{1}{2}; 1; -\frac{\alpha^2}{2}\right) \approx \sqrt{\frac{\pi}{2}} \left(1 + \frac{\alpha^2}{4}\right) \quad , \\ \sigma_e^2 &= 2 + \alpha^2 - \mu_e^2 \approx 2 - \frac{\pi}{2} \quad . \end{aligned} \quad (26)$$

Substituting these results in (25) and solving for α , we have the required per-sample input signal-to-noise ratio measures for large M in the alternative forms

$$\begin{aligned} \alpha &\approx 2 \left(\frac{4-\pi}{\pi}\right)^{1/4} \frac{\beta^{1/2}}{M^{1/4}} = 1.446 \frac{\beta^{1/2}}{M^{1/4}} \quad , \\ \frac{S}{N} &= \frac{\alpha^2}{2} \approx 2 \left(\frac{4-\pi}{\pi}\right)^{1/2} \frac{\beta}{M^{1/2}} = 1.045 \frac{\beta}{M^{1/2}} \quad , \end{aligned}$$

$$\text{dB} = 10 \log \frac{S}{N} \approx 10 \log \left(2 \sqrt{\frac{4-\pi}{\pi}}\right) + 10 \log(\beta) - 5 \log(M) = .193 + 10 \log(\beta) - 5 \log(M), \quad (27)$$

where the single parameter

$$\beta = \tilde{\Phi}(P_D) - \tilde{\Phi}(P_{FA}) \quad (28)$$

incorporates the specified false alarm and detection probabilities. (27) displays the familiar $5 \log M$ decibel decay for large M associated with the incoherent addition in (1); see also [2, p. 279, Ex. 8.8].

RESULTS

For a given value of M , the output variable in (1),

$$x = \sum_{m=1}^M e_m, \quad (29)$$

will exceed threshold u with false alarm probability P_{FA} when signal-to-noise ratio α is zero. That is

$$P_{FA} = \text{Prob}(x > u \mid \alpha = 0; M). \quad (30)$$

For specified values of M and P_{FA} , this relation can be solved numerically for u ; the values of normalized threshold u/M are listed in table 1 for $M=2^n$, $n=0(1)13$ and for $P_{FA}=10^{-n}$, $n=1(1)8$.

The detection probability depends on threshold u , M , and signal-to-noise ratio $\alpha(>0)$:

$$P_D = \text{Prob}(x > u \mid \alpha; M). \quad (31)$$

For specified values of M , P_D , and u , this relation can be solved numerically for the required input signal-to-noise ratio α . When the threshold results in Table 1 are employed, the results yield the required input signal-to-noise ratio for specified false alarm probability and detection probability at a particular M . These are plotted in figures 1-5 for

$$P_D = .5, .9, .95, .99, .999, \quad (32)$$

respectively. The abscissa is $\log_2 M$, and the ordinate is in decibels, as defined in (27). The fit of (27) is very good for large M , especially for the larger P_{FA} values. These results in figures 1-5 greatly extend the one in [1, Fig. 16] and [2, Fig. 8.18].

Table 1. Normalized Thresholds Required for Specified M and P_{FA}

$M \backslash P_{FA}$	1E-1	1E-2	1E-3	1E-4
1	2.14596603	3.03485426	3.71692219	4.29193205
2	1.87154046	2.46578168	2.92459903	3.31372579
4	1.68491649	2.08494224	2.39281962	2.65432267
8	1.55592564	1.82779134	2.03544098	2.21134522
16	1.46605729	1.65246898	1.79362769	1.91266565
32	1.40314416	1.53192213	1.62866385	1.70984877
64	1.35896377	1.44846093	1.51524477	1.57104117
128	1.32787317	1.39035933	1.43673968	1.47534630
256	1.30596258	1.34974198	1.38210498	1.40896493
512	1.29050601	1.32125803	1.34392160	1.36268942
1024	1.27959472	1.30123656	1.31715039	1.33030674
2048	1.27188832	1.28713956	1.29833595	1.30758095
4096	1.26644357	1.27720181	1.28509047	1.29159844
8192	1.26259580	1.27018998	1.27575385	1.28034098

$M \backslash P_{FA}$	1E-5	1E-6	1E-7	1E-8
1	4.79852591	5.25652177	5.67769243	6.06970852
2	3.65817649	3.97074674	4.25904998	4.52806135
4	2.88639585	3.09755766	3.29282208	3.47544423
8	2.36734857	2.50933650	2.64073862	2.76376208
16	2.01795589	2.11363367	2.20209577	2.28487698
32	1.78141625	1.84629005	1.90615996	1.96210527
64	1.62006566	1.66438962	1.70520835	1.74328423
128	1.50917003	1.53967893	1.56771937	1.59383081
256	1.43244246	1.45357776	1.47297026	1.49100181
512	1.37906412	1.39378248	1.40726893	1.41979378
1024	1.34176981	1.35206123	1.36148146	1.37022180
2048	1.31562790	1.32284604	1.32944798	1.33556910
4096	1.29725887	1.30233301	1.30697131	1.31126956
8192	1.28432858	1.28790149	1.29116614	1.29419029

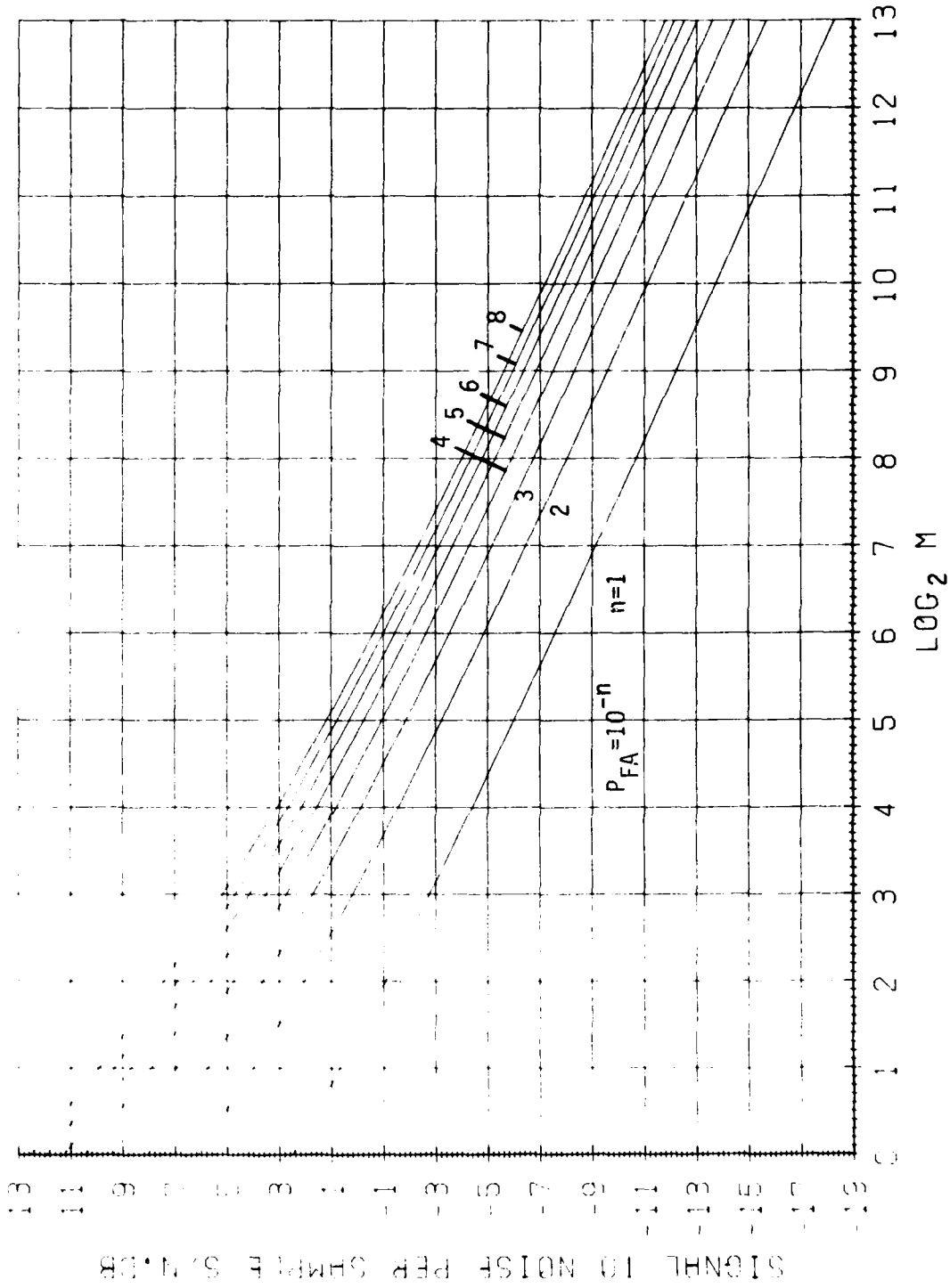


Figure 1. Required Input S/N for P_n=.5

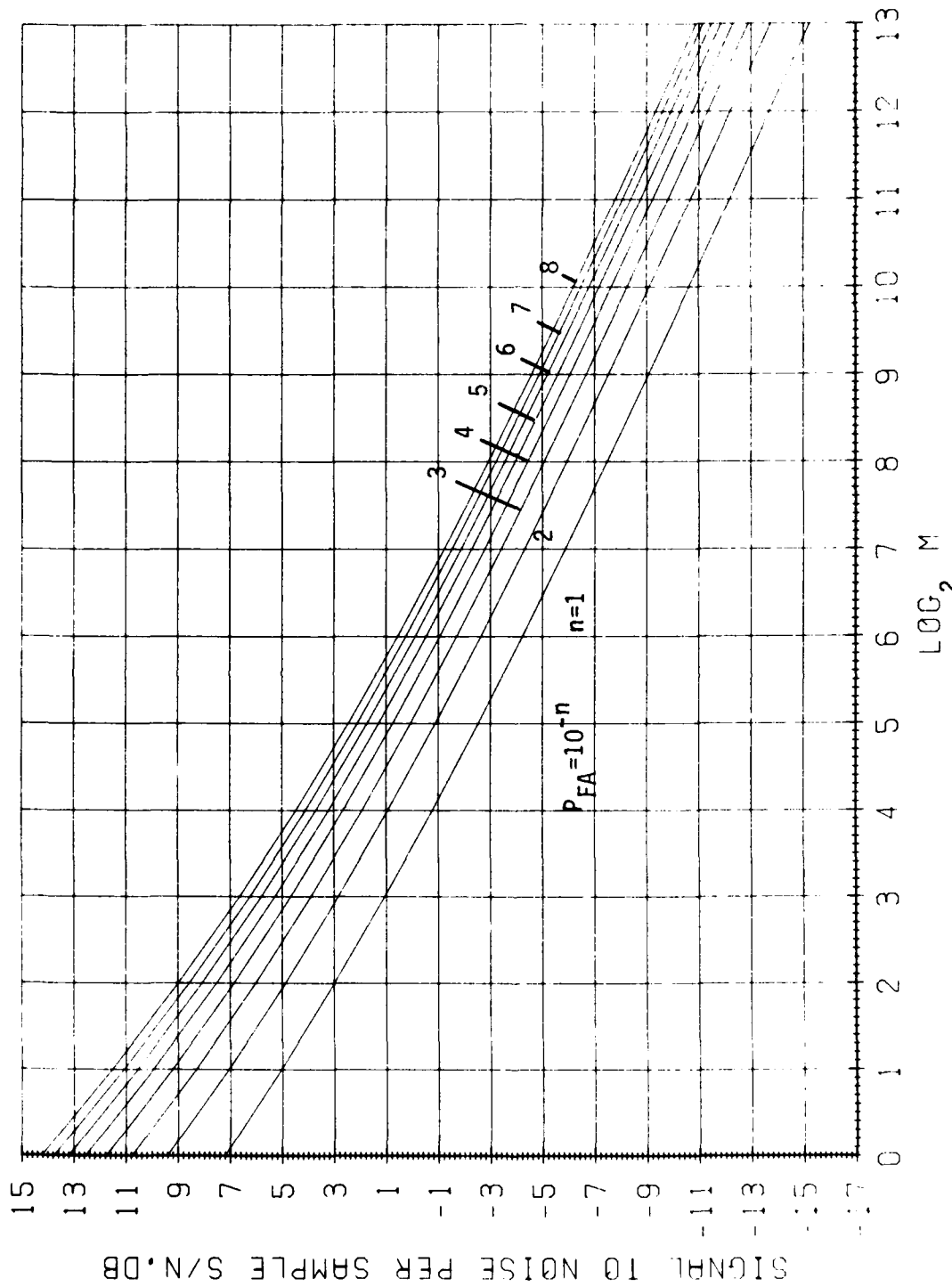


Figure 2. Required Input S/N for $P_D=0.9$

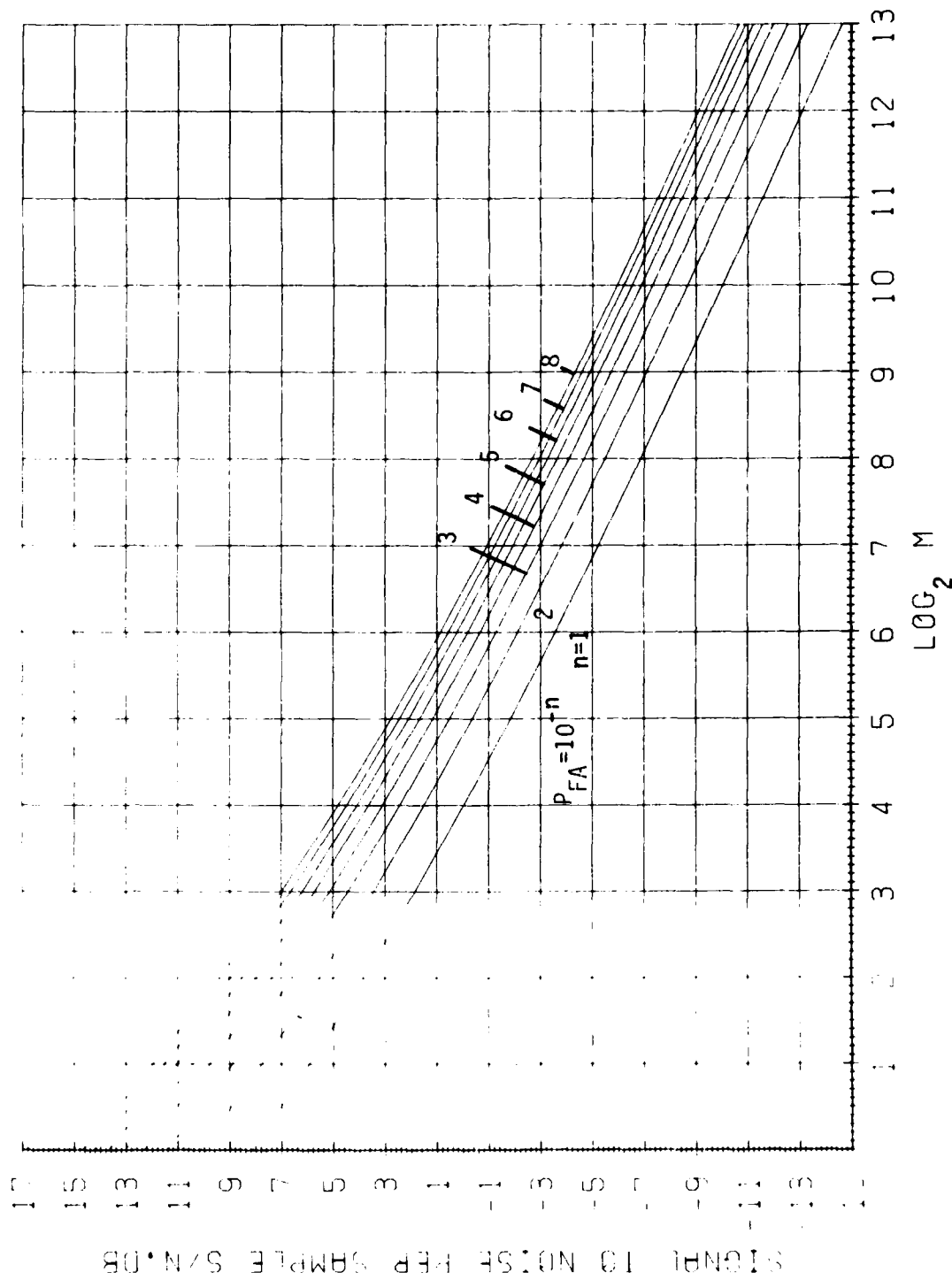


Figure 3. Required Input S/N for $P_D = .95$

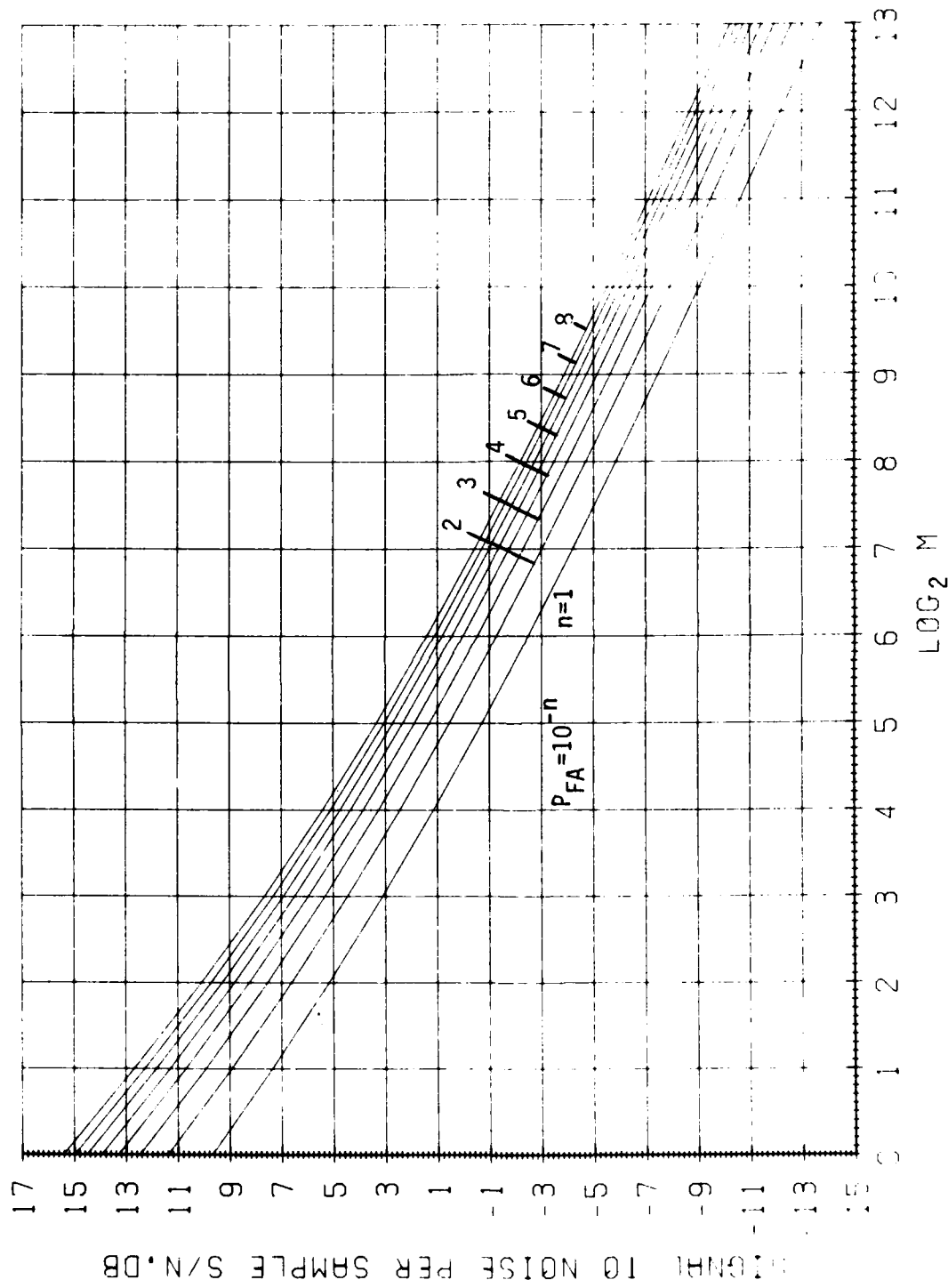


Figure 4. Required Input S/N for $P_D = .99$

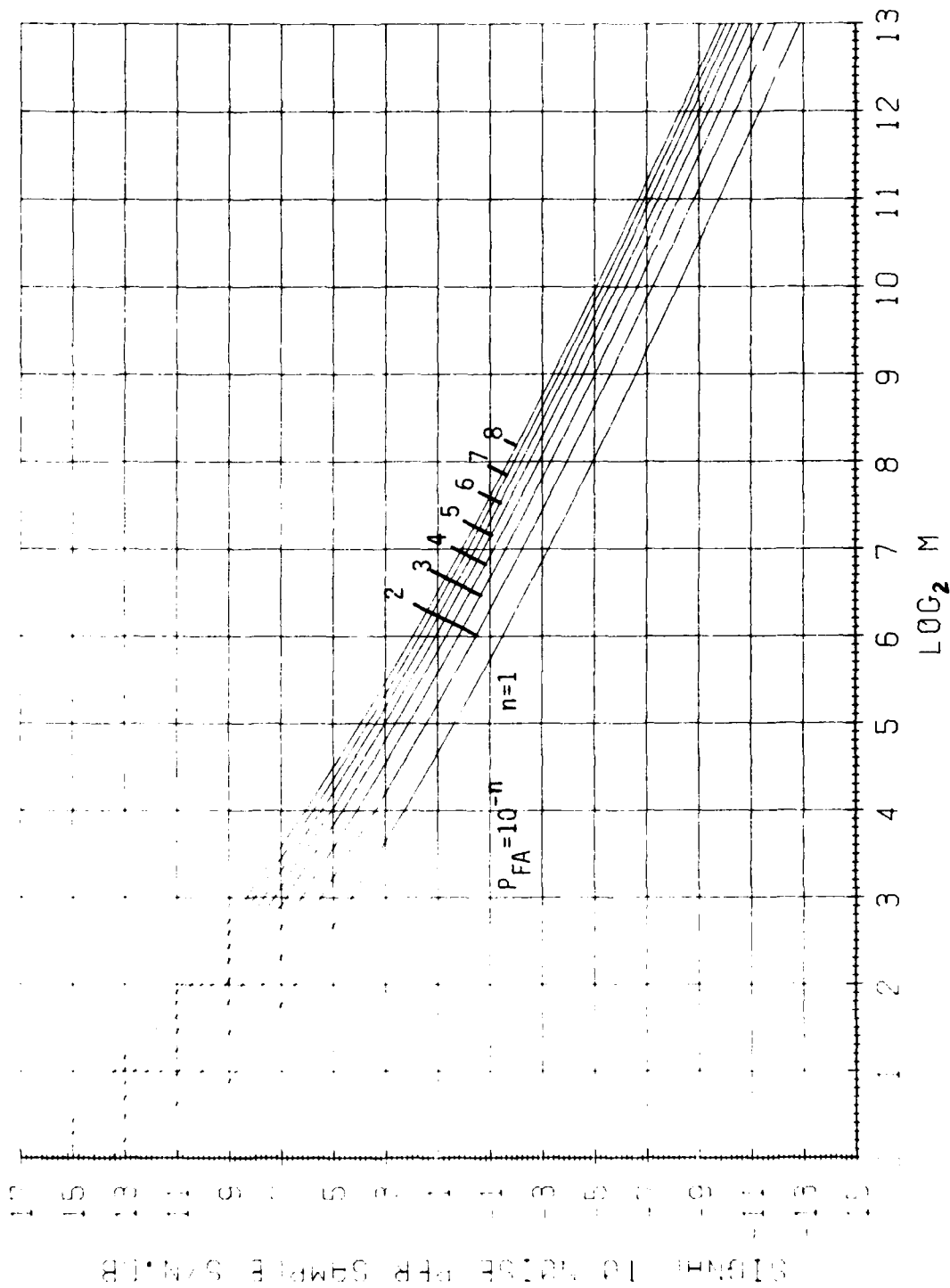


Figure 5. Required Input S/N for P_D = .999

The results in figures 1 through 5 only cover a selected set of detection and false alarm probability values. A more complete description is afforded by the receiver operating characteristics, namely detection probability vs. false alarm probability, with signal-to-noise ratio as a parameter. In figures 6 through 19 are given these operating characteristics for

$$M = 1, 2, 4, 8, 16, 32, 64, 128, 256, 512, 1024, 2048, 4096, 8192, \quad (33)$$

respectively. The false alarm probability covers the range 10^{-10} to .5, while the detection probability covers 10^{-10} to .999. Both abscissa and ordinate in these figures employ the inverse function to the Gaussian cumulative distribution function Φ defined in (19); thus, a truly Gaussian random variable would plot as a series of equally spaced parallel straight lines (with parameter α). Observe that the curves are nearly equally spaced with parameter α , except for very small α , where the nonlinear envelope operation causes small signal suppression and a crowding together of the curves.

If the decision variable x is presumed Gaussian, and the operating characteristics overlaid on the exact results in figures 6-19, it is found that the two sets of curves for $M=8192$ are virtually identical in the range of P_{FA} and P_D plotted. However, for $M=16$, the Gaussian approximation is somewhat optimistic; for example, the exact curve for $\alpha=2.75$ is well-approximated by the Gaussian approach for $\alpha=2.62$. For small M , the Gaussian approximation is overly optimistic for small P_{FA} ; however, the two sets cross near $P_{FA}=.5$, which is not a practical range of interest anyway.

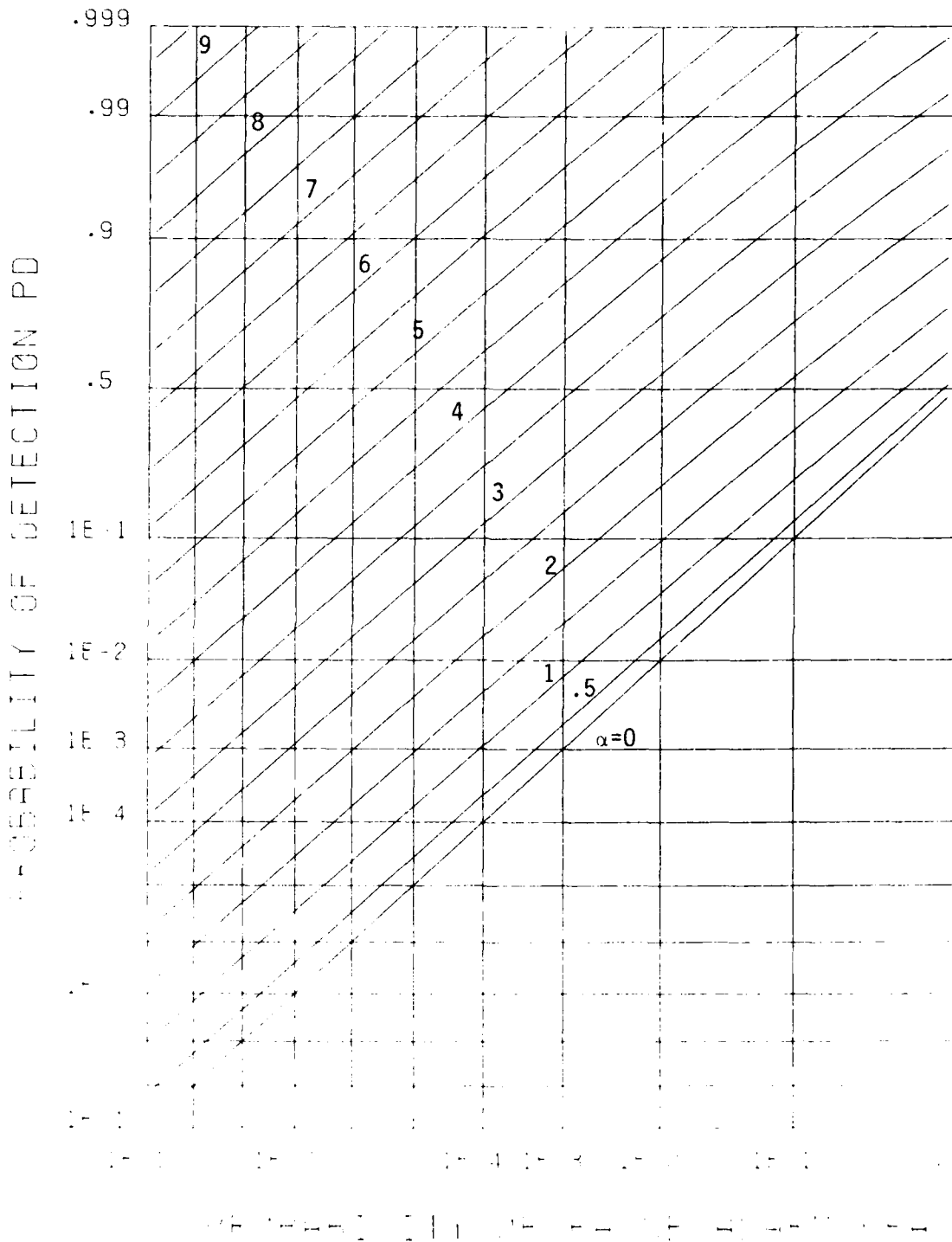


Figure 6. Receiver Operating Characteristics for $M=1$

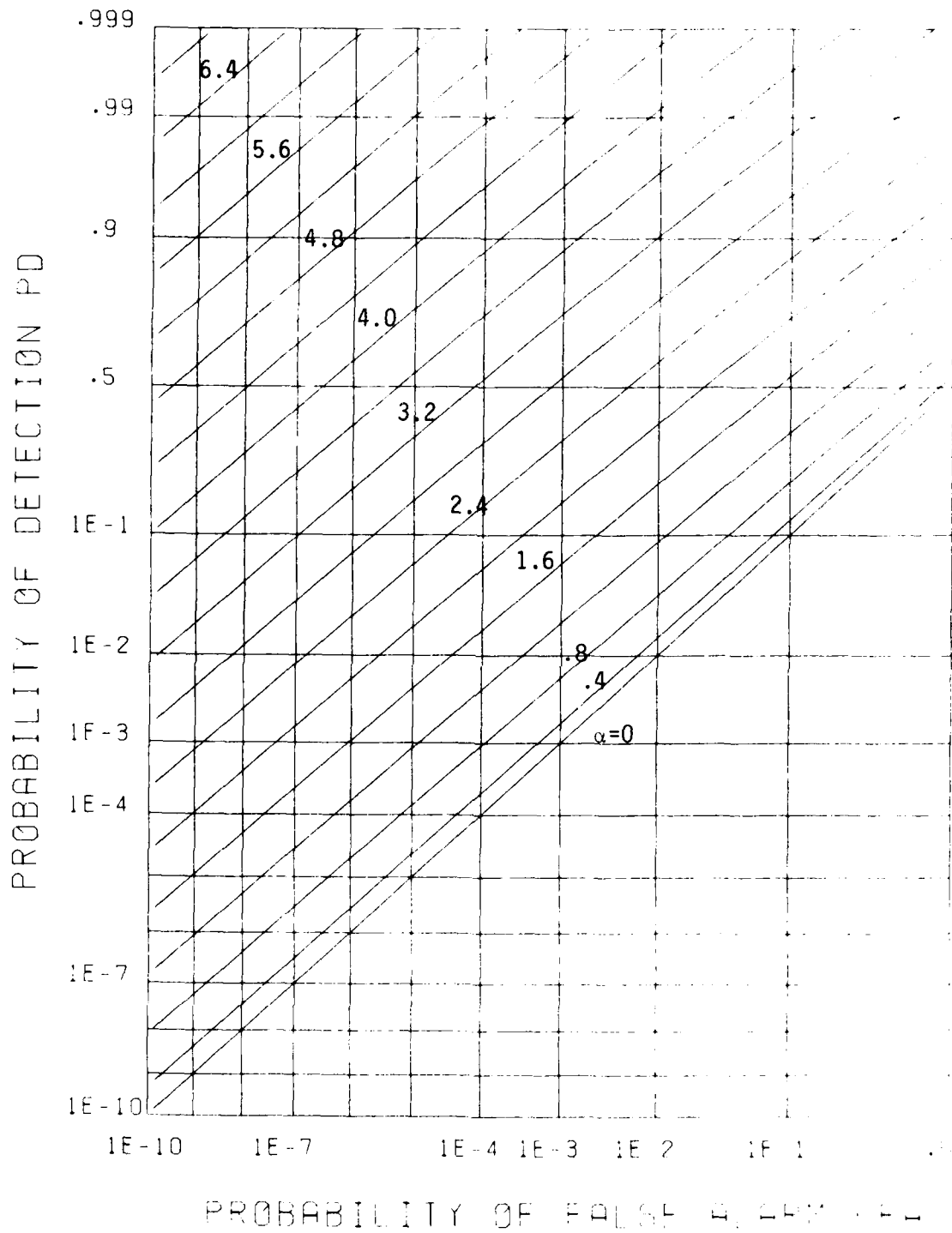


Figure 7. Receiver Operating Characteristics for $M=2$

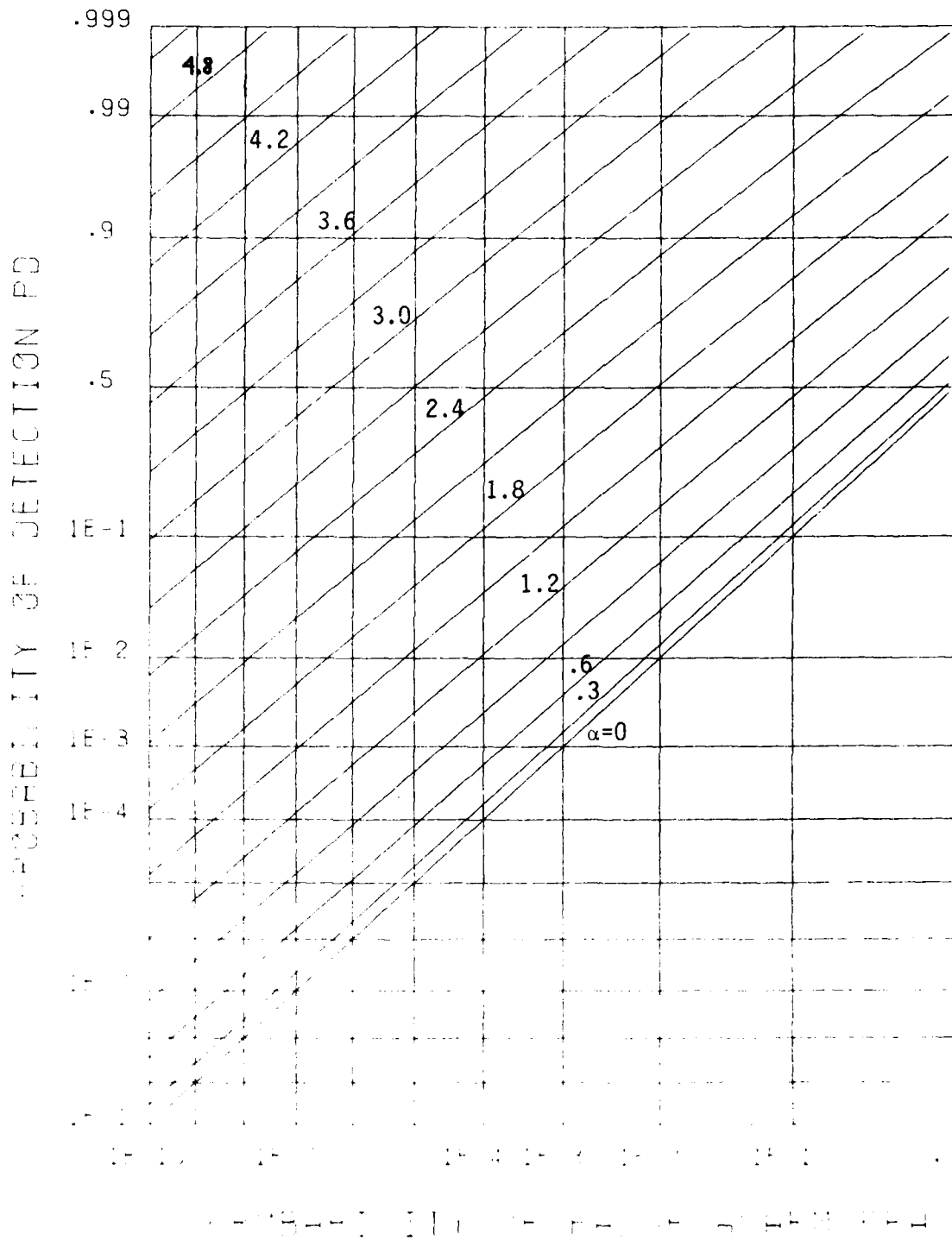


Figure 8. Receiver Operating Characteristics for $M=4$

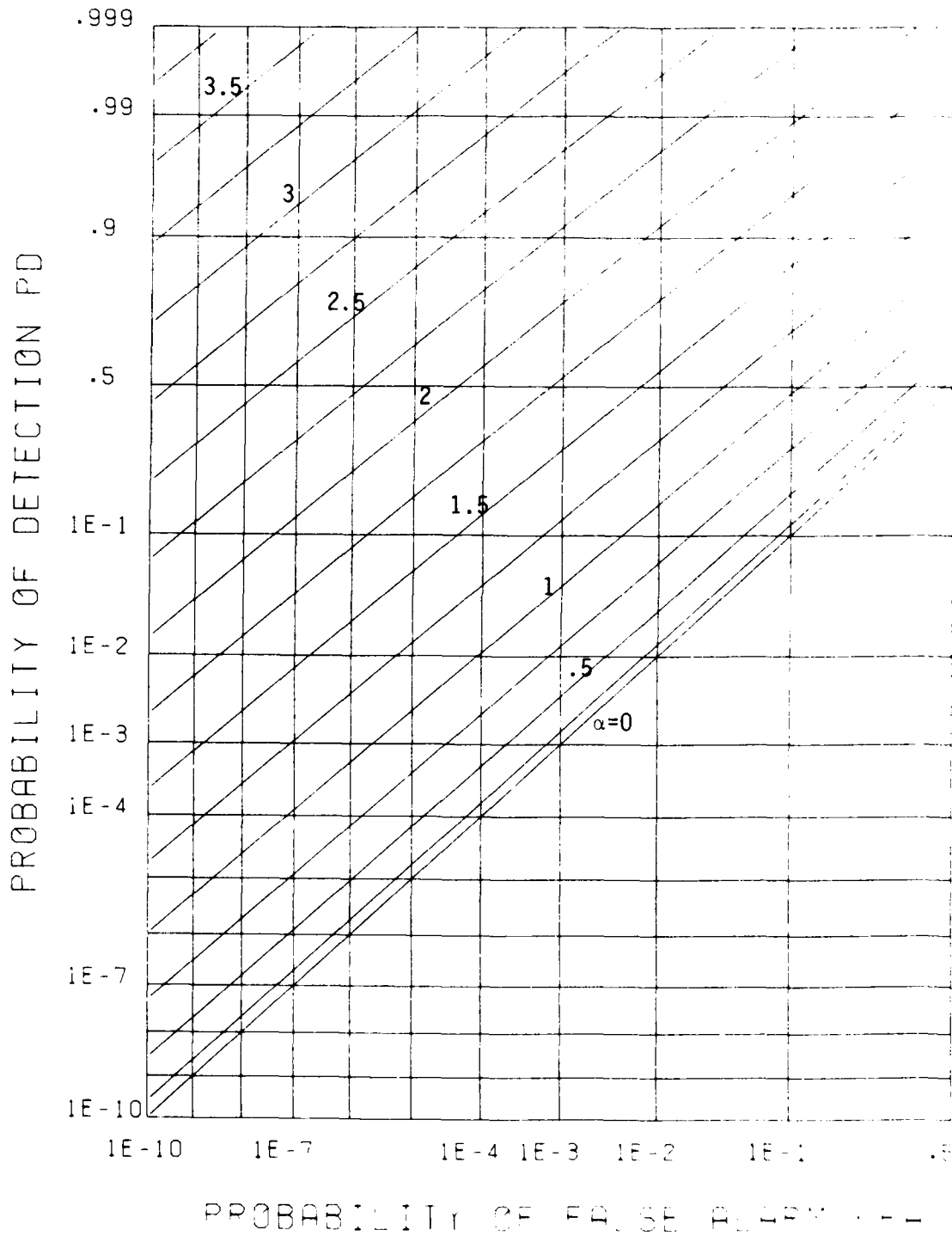


Figure 9. Receiver Operating Characteristics for M=8

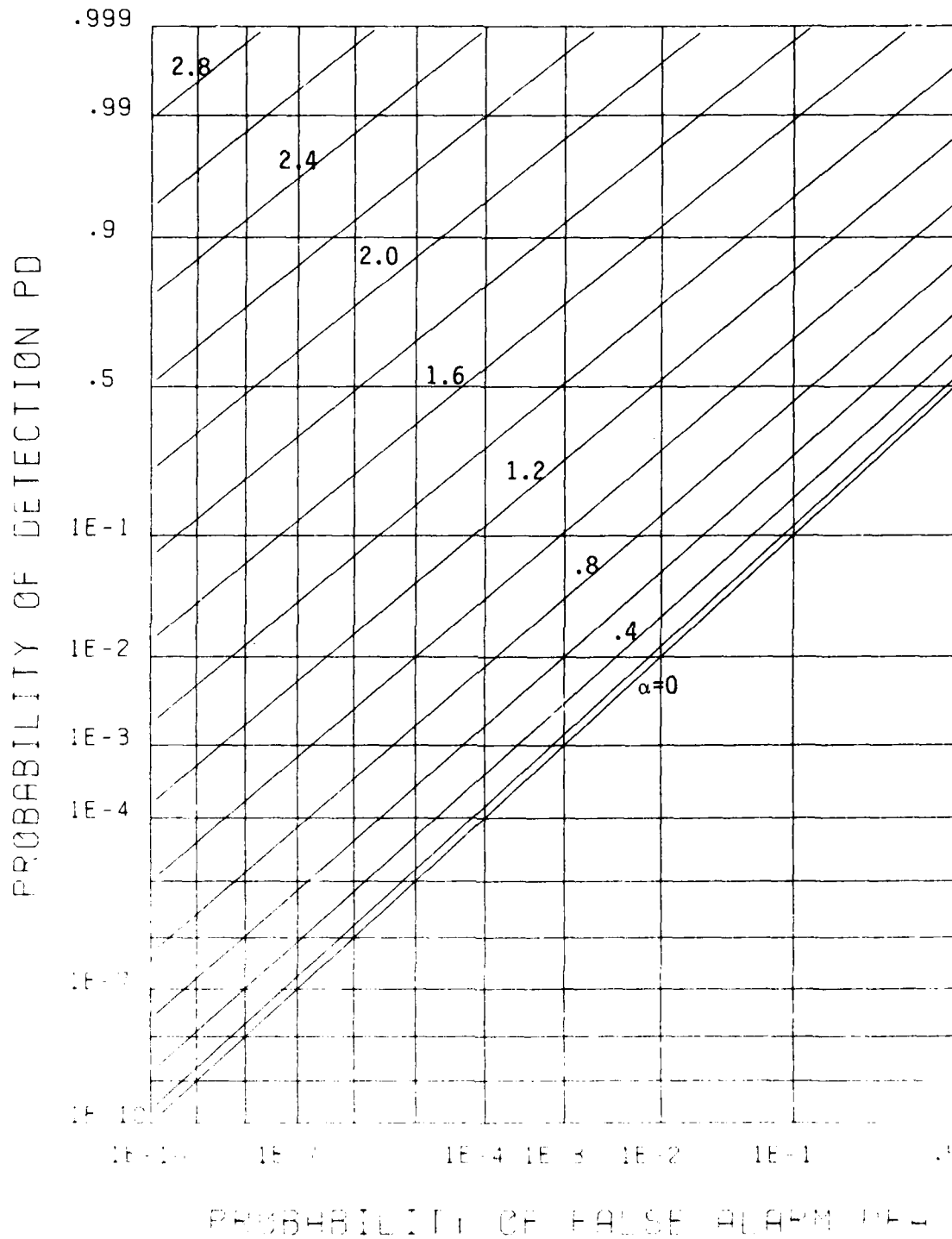


Figure 10. Receiver Operating Characteristics for $M=16$

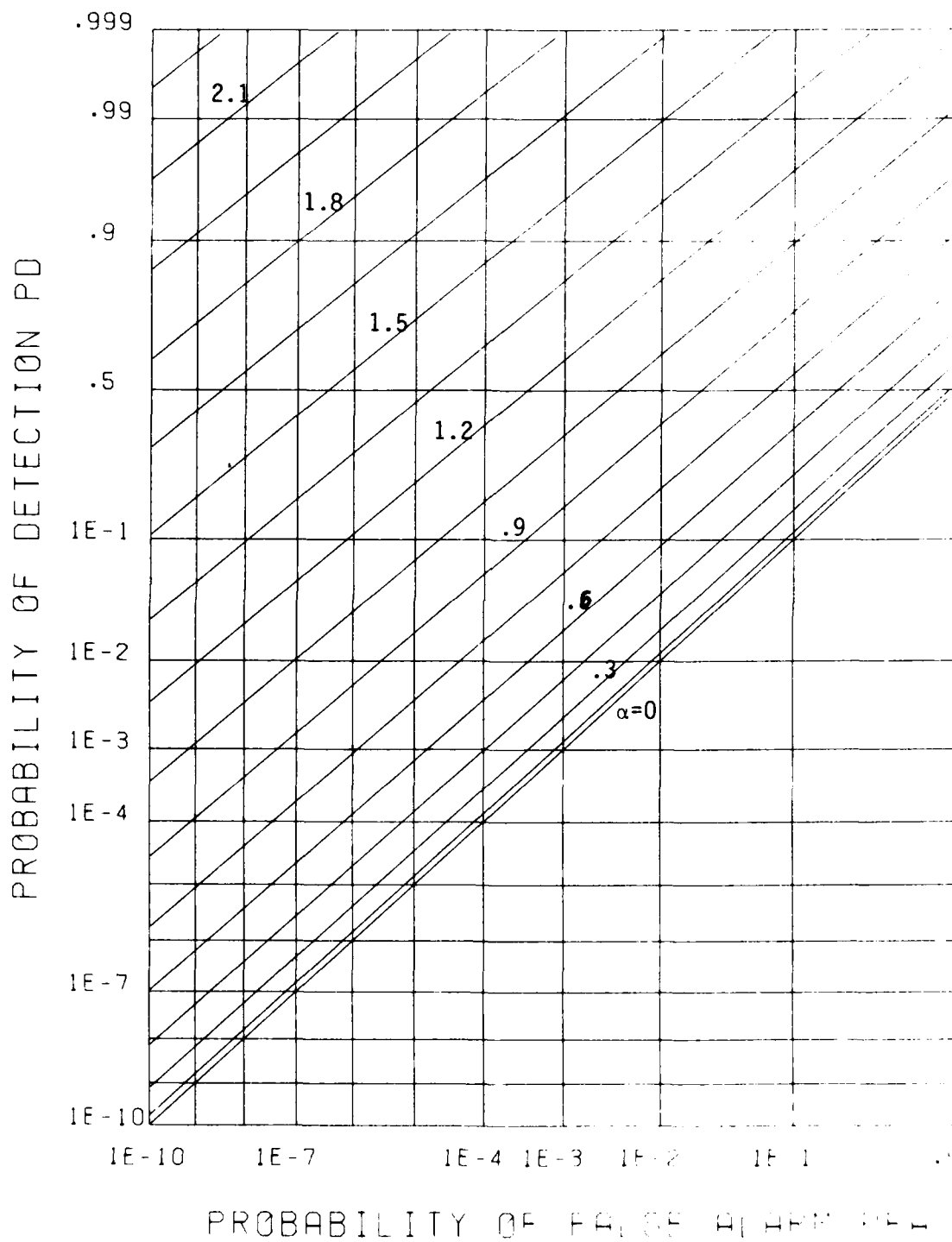


Figure 11. Receiver Operating Characteristics for $M=32$

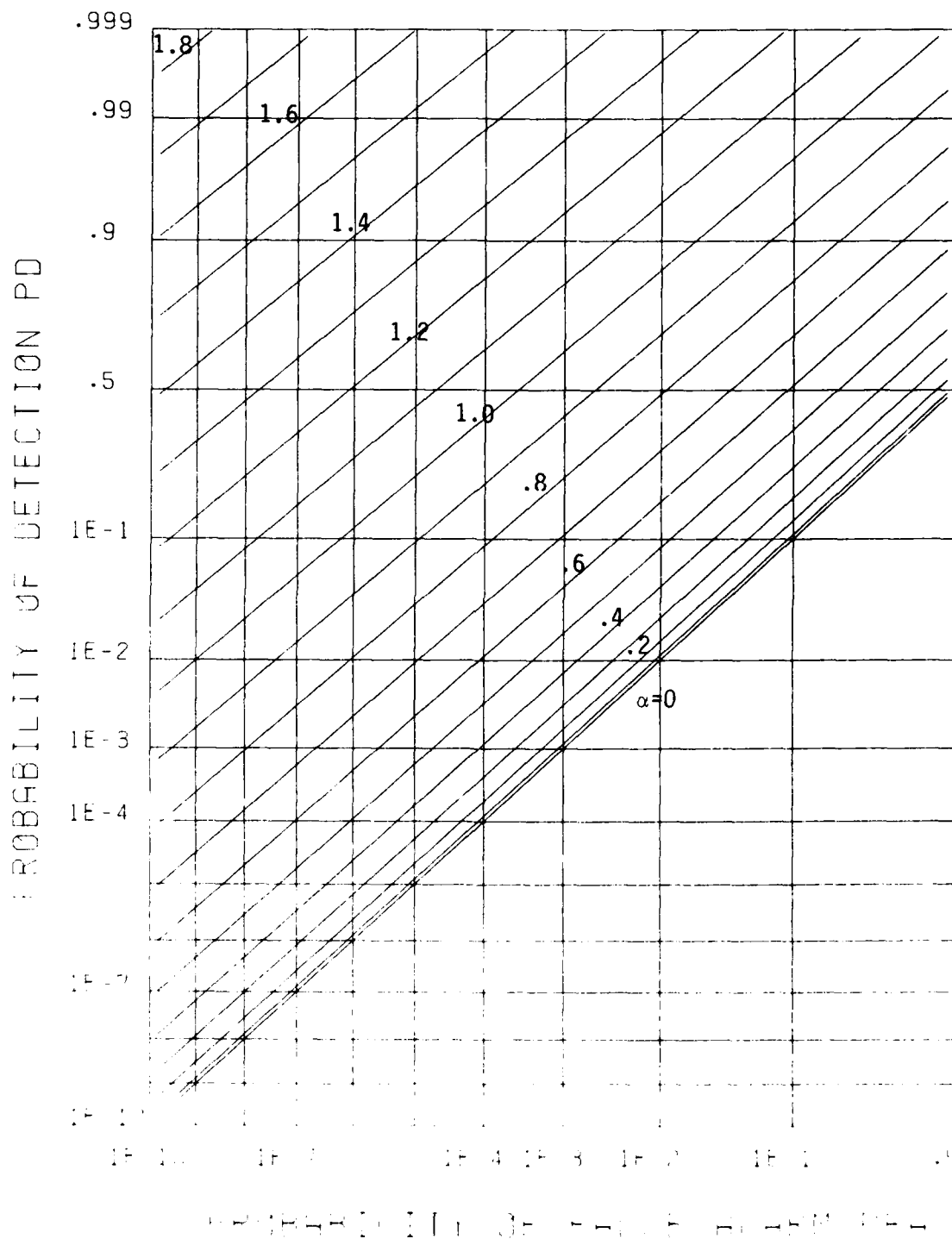


Figure 12. Receiver Operating Characteristics for M=64

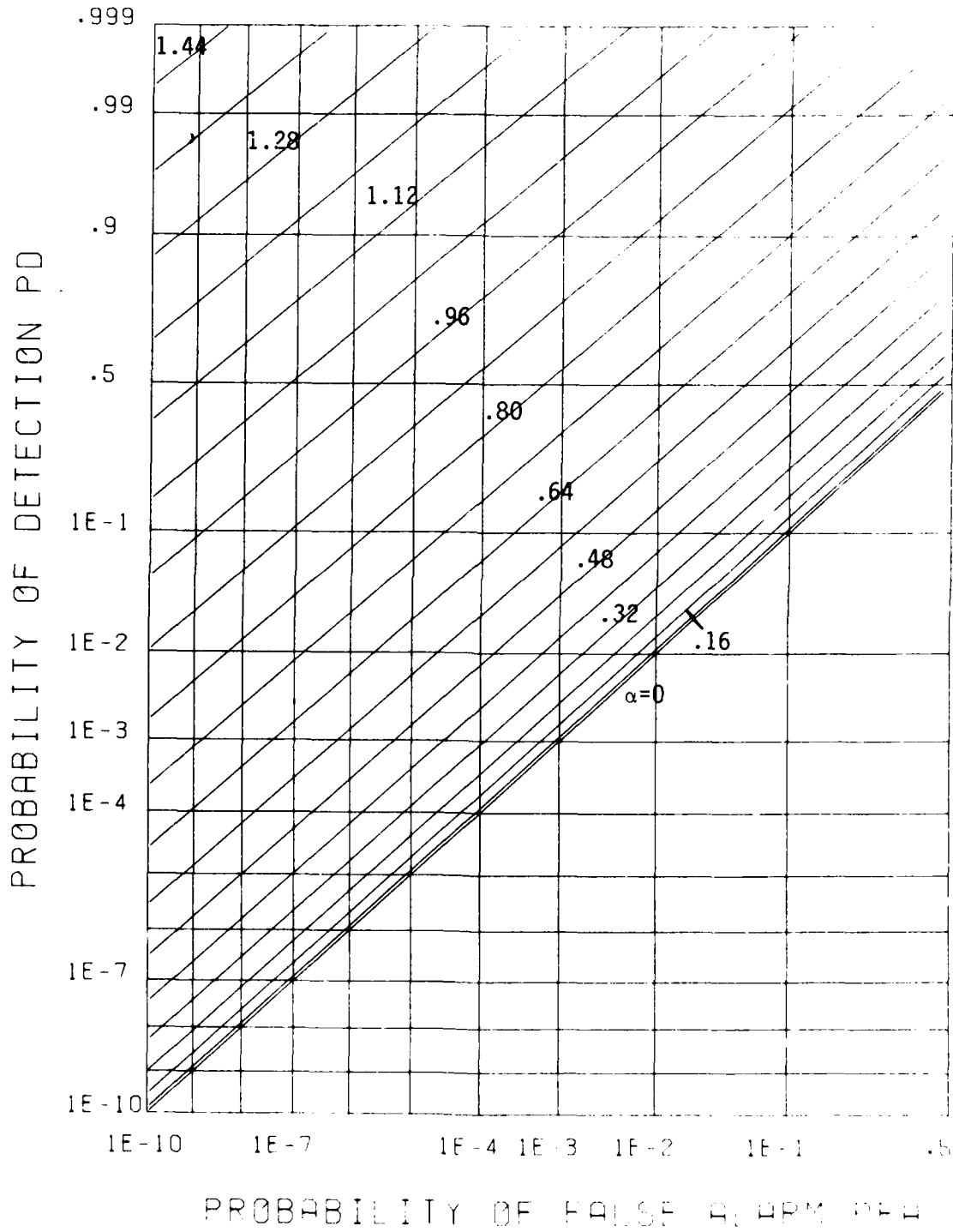


Figure 13. Receiver Operating Characteristics for $M=128$

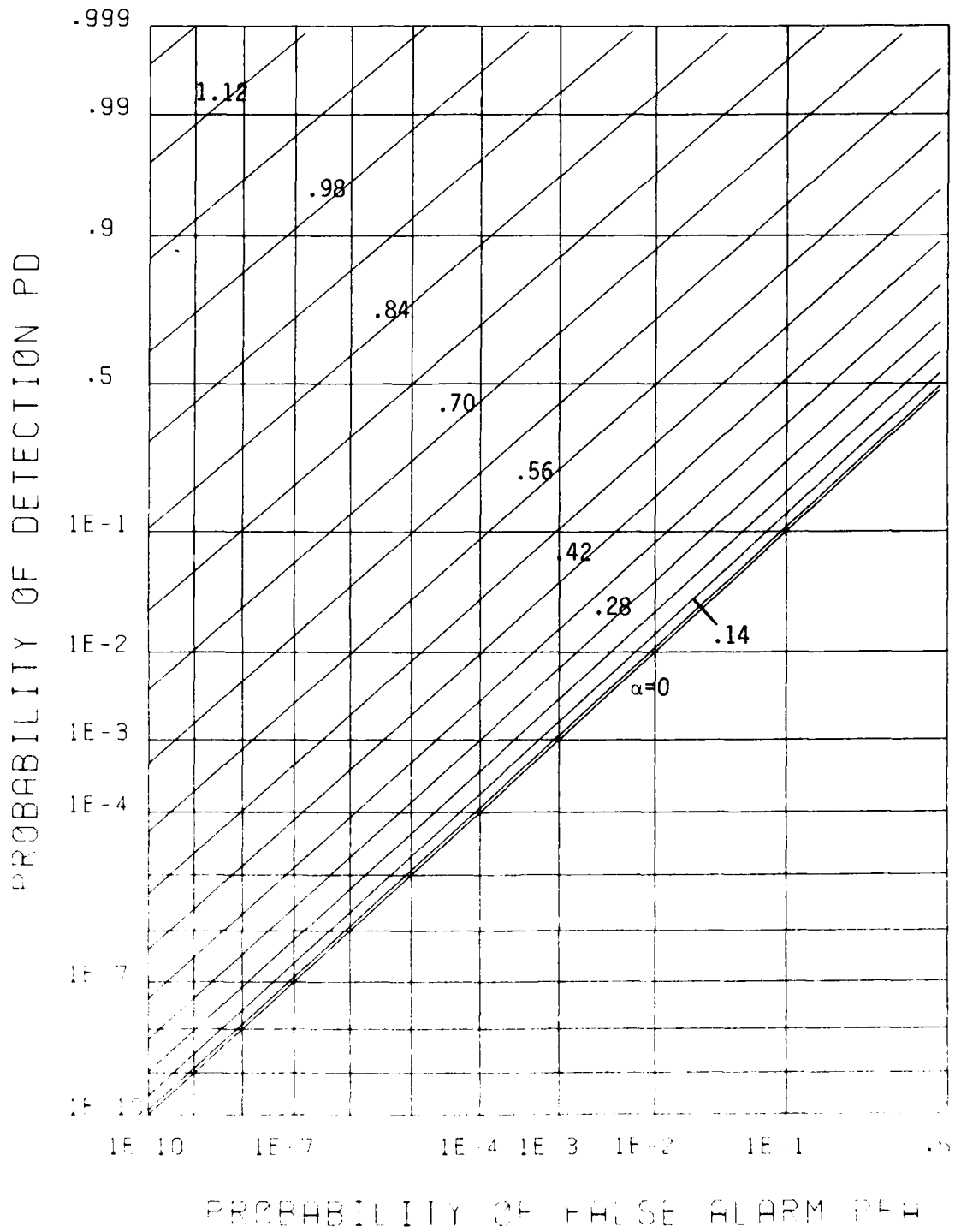


Figure 14. Receiver Operating Characteristics for $M=256$

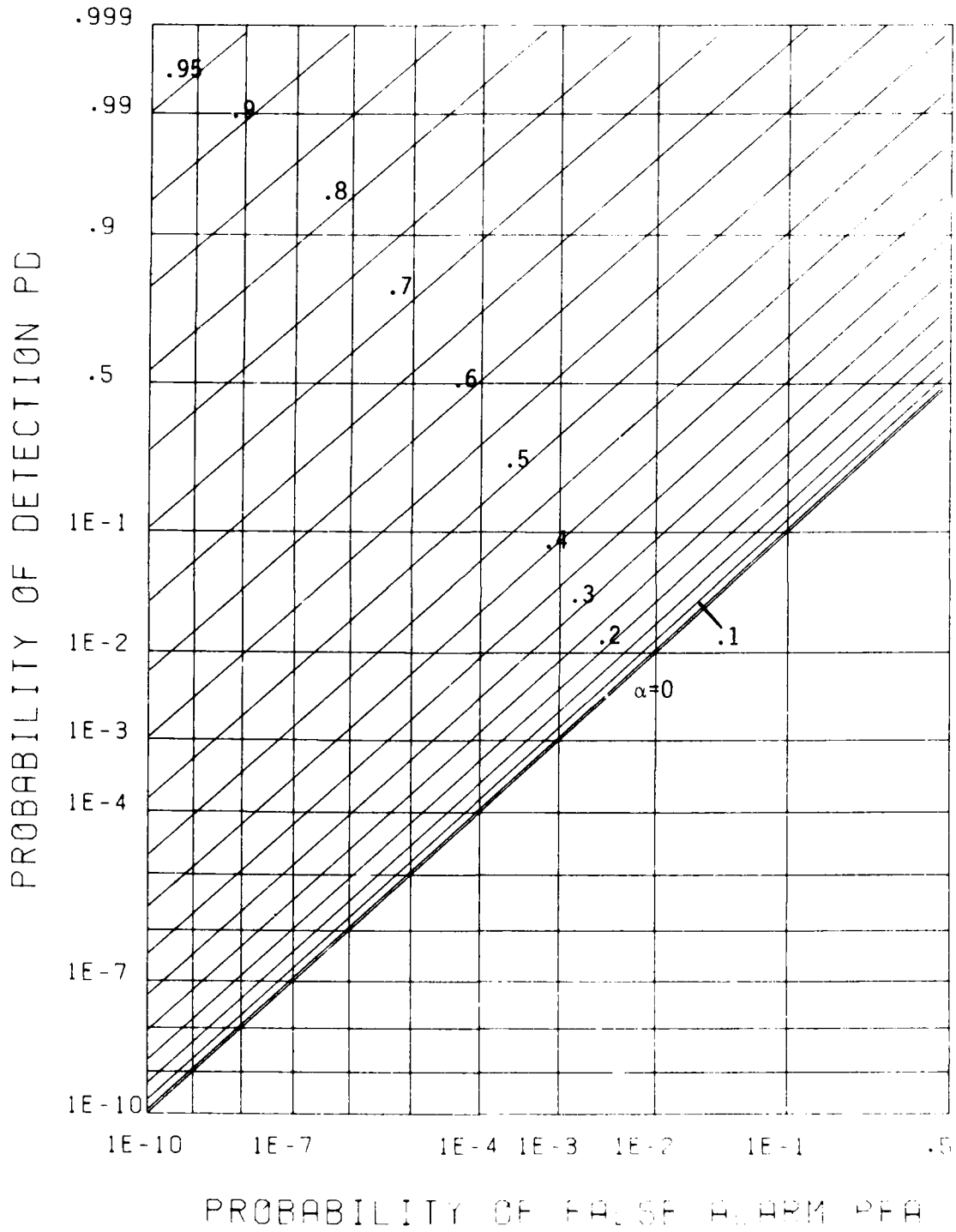


Figure 15. Receiver Operating Characteristics for M=512

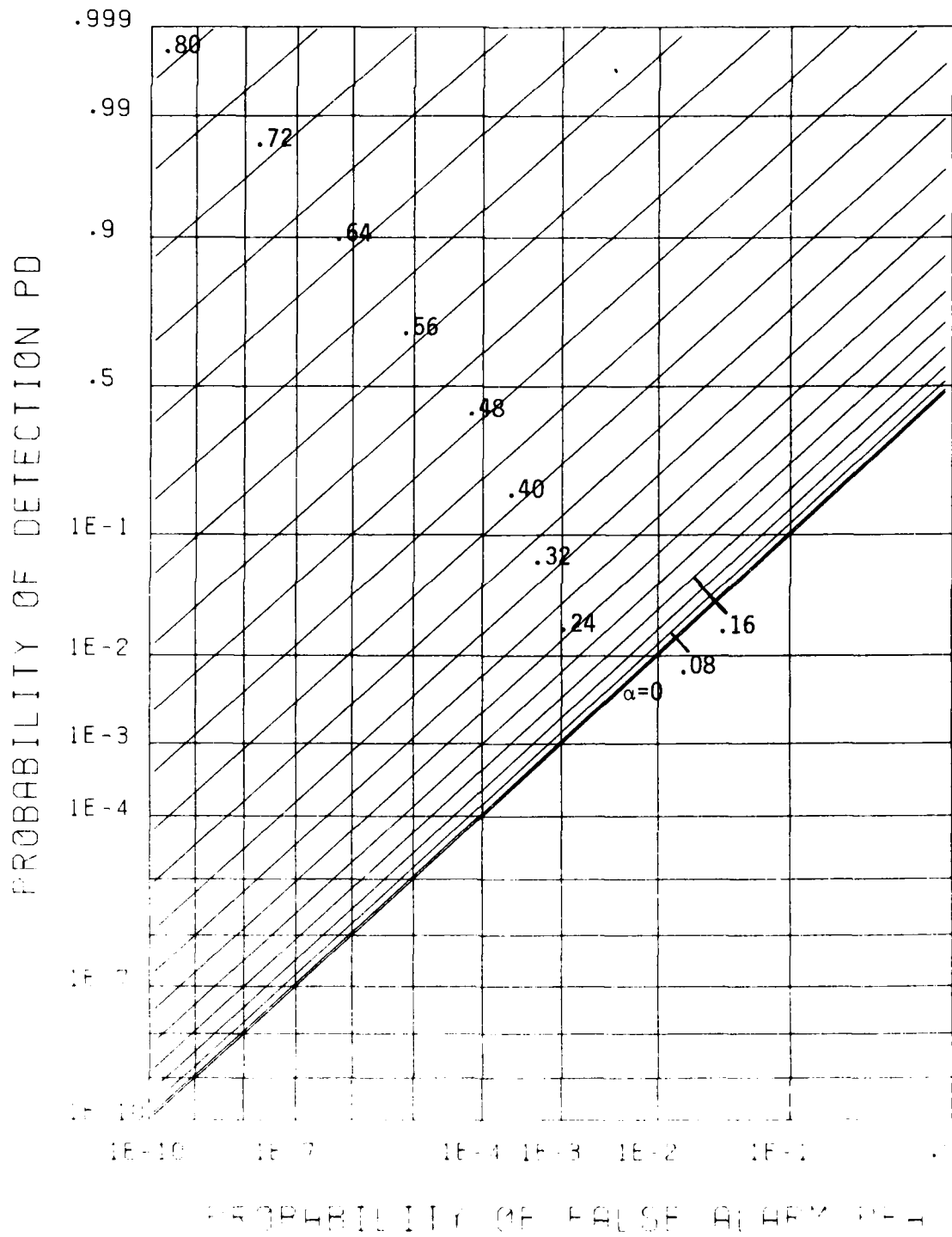


Figure 16. Receiver Operating Characteristics for M=1024

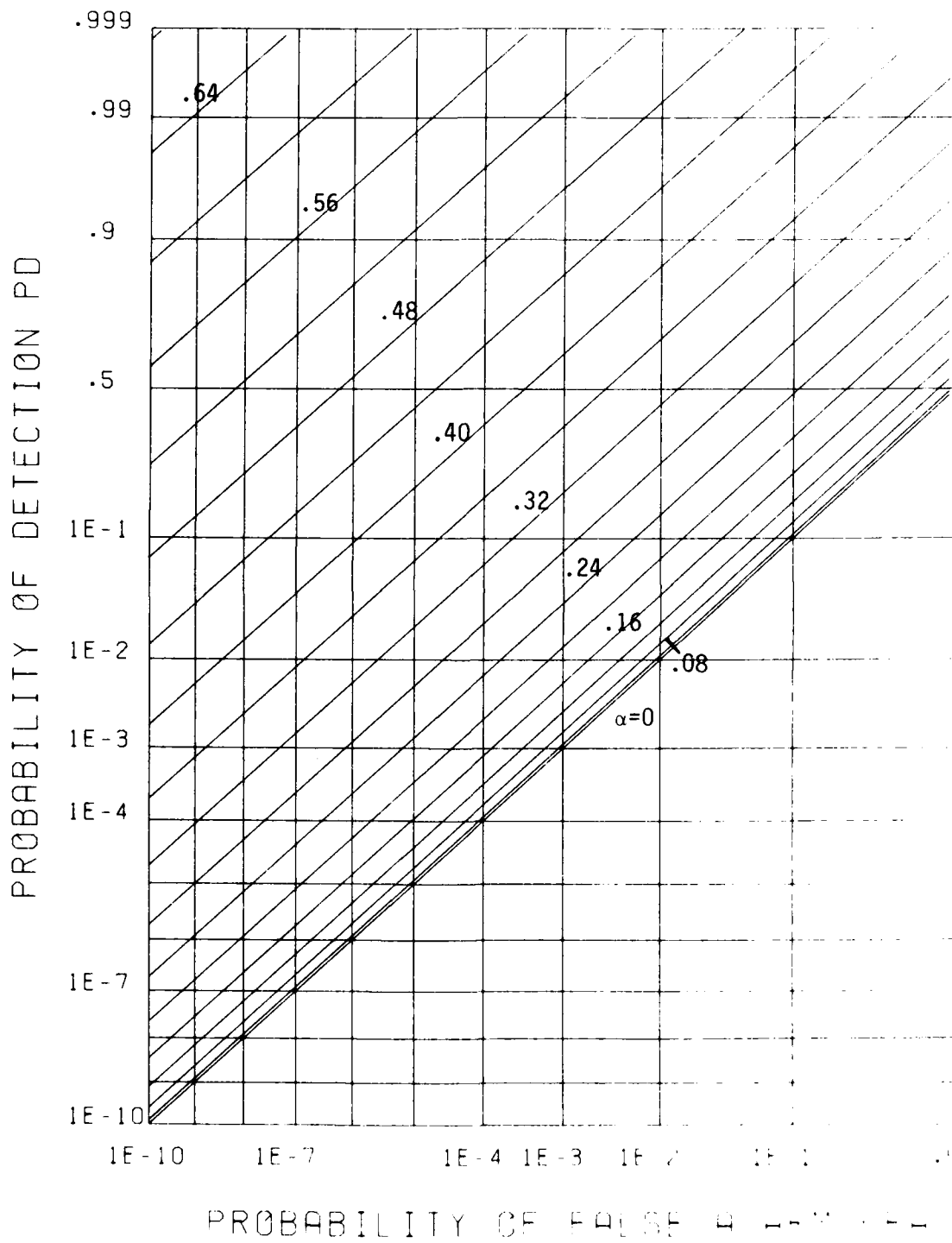


Figure 17. Receiver Operating Characteristics for M=2048

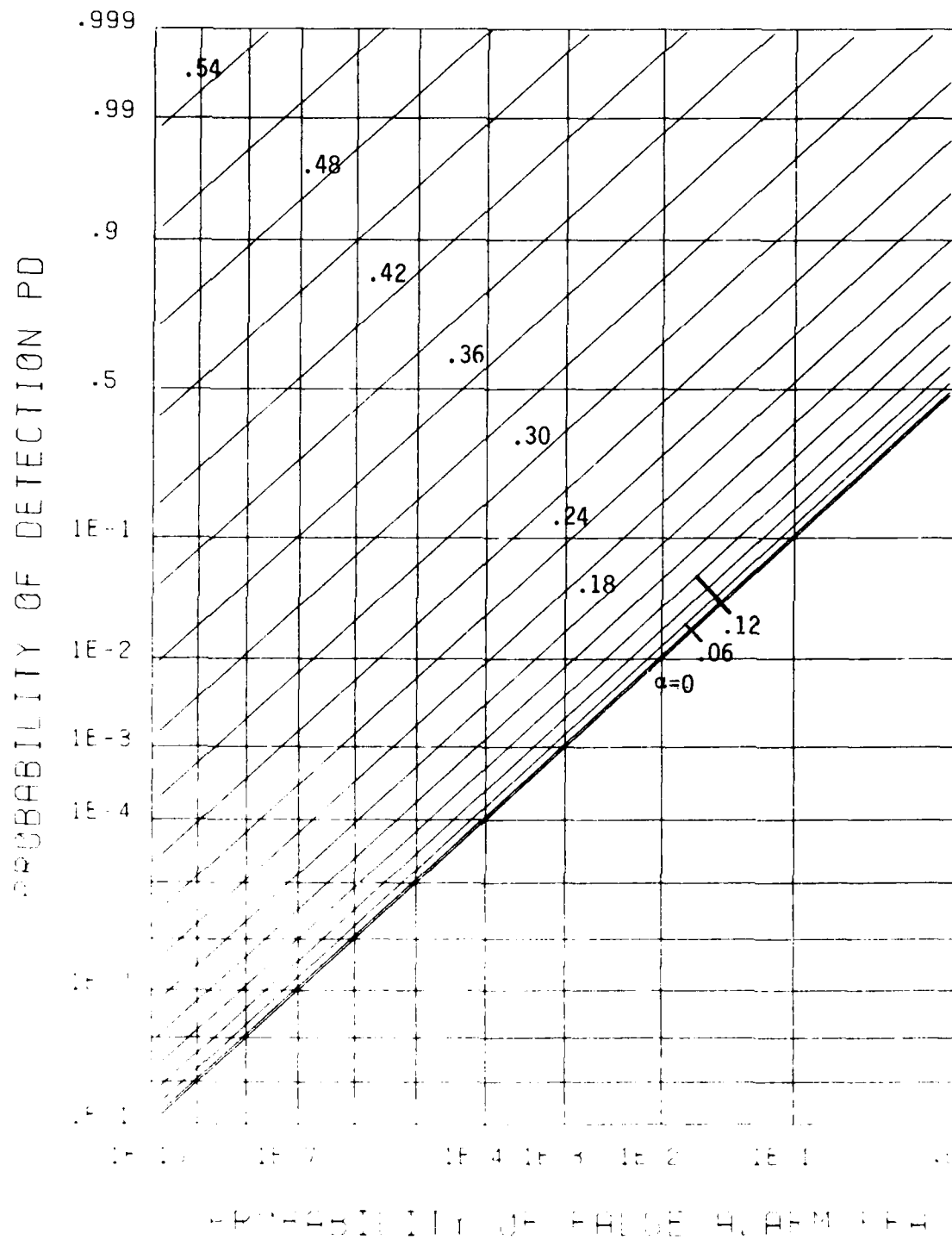


Figure 18. Receiver Operating Characteristics for M=4096

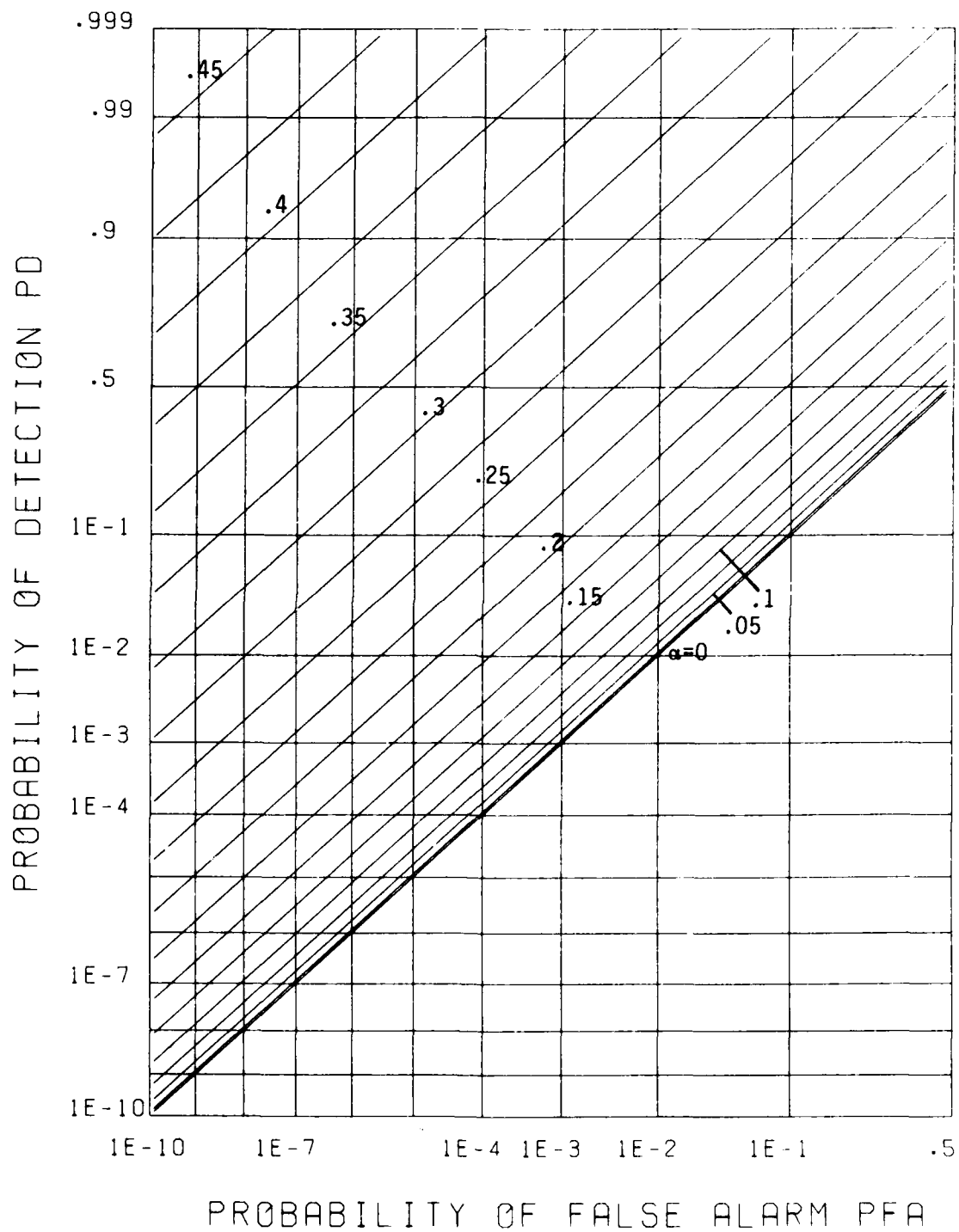


Figure 19. Receiver Operating Characteristics for M=8192

SUMMARY

A method for exact evaluation of the exceedance distribution function, of a linear sum of M envelopes of a narrowband Gaussian process and sinewave, has been utilized to determine the receiver operating characteristics for a wide range of values of M and signal-to-noise ratio. Also, the required input signal-to-noise ratio vs. M has been determined for a selected set of false alarm and detection probabilities. Programs are also supplied by which other values of the various parameters can be investigated by the user.

Agreement between the current results and those in [1,2] is very good over the range of common values plotted. For M larger than 8192, the approximation given in (27) and (28) is recommended, since the summation variable is then well represented by a Gaussian random variable.

APPENDIX A. DERIVATION OF RICE CHARACTERISTIC FUNCTION

The normalized probability density function of a Rice random variable was given in (2) as

$$p_e(u) = u \exp\left(-\frac{u^2 + \alpha^2}{2}\right) I_0(\alpha u) \quad \text{for } u \geq 0 \quad . \quad (\text{A-1})$$

The corresponding characteristic function is

$$\begin{aligned} f_e(\xi) &= \int_{-\infty}^{+\infty} du \exp(i\xi u) p_e(u) = \int_0^{+\infty} du u \exp\left(i\xi u - \frac{u^2 + \alpha^2}{2}\right) I_0(\alpha u) = \\ &= \exp(-r) \sum_{n=0}^{\infty} \frac{(r/2)^n}{(n!)^2} \int_0^{\infty} du u^{2n+1} \exp(i\xi u - u^2/2) \quad , \end{aligned} \quad (\text{A-2})$$

where we have expanded I_0 in a power series [5, 8.447 1] and defined power signal-to-noise ratio

$$r = \alpha^2/2 \quad . \quad (\text{A-3})$$

(If desired, a power series in ξ could be developed by expanding $\exp(i\xi u)$ in a power series instead of I_0 .)

We define

$$C_n(\xi) = \frac{1}{2^n (n!)^2} \int_0^{\infty} du u^{2n+1} \exp(i\xi u - u^2/2) \quad \text{for } n \geq 0 \quad , \quad (\text{A-4})$$

and get the characteristic function series

$$f_e(\xi) = \exp(-r) \sum_{n=0}^{\infty} r^n C_n(\xi) \quad . \quad (\text{A-5})$$

In order to get a recurrence on $C_n(\xi)$, we also define

$$B_k(\xi) = \int_0^{\infty} dw w^k \exp(i\xi w - w^2/2) \quad \text{for } k \geq 0 \quad , \quad (\text{A-6})$$

for then

$$C_n(\xi) = \frac{B_{2n+1}(\xi)}{2^n (n!)^2} \quad . \quad (A-7)$$

By integrating by parts on (A-6), there follows

$$B_k = i\xi B_{k-1} + (k-1)B_{k-2} \quad \text{for } k \geq 1 \quad . \quad (A-8)$$

This recurrence can be started with [5, 3.896 3&4]

$$B_0 = \exp(-\xi^2/2) \left[\sqrt{\frac{\pi}{2}} + i\xi {}_1F_1\left(\frac{1}{2}; \frac{3}{2}; \frac{\xi^2}{2}\right) \right] \quad . \quad (A-9)$$

By looking at three adjacent terms of recurrence (A-8), we can generate the alternative recurrence

$$B_k = (2k-3-\xi^2)B_{k-2} - (k-2)(k-3)B_{k-4} \quad . \quad (A-10)$$

By means of (A-7), this translates into

$$C_n = \frac{1}{n^2} \left[\left(2n - \frac{1+\xi^2}{2}\right) C_{n-1} - \frac{n - \frac{1}{2}}{n-1} C_{n-2} \right] \quad \text{for } n \geq 2 \quad . \quad (A-11)$$

Starting values are (via manipulation of hypergeometric series and Kummer's transformation) expressible as

$$\begin{aligned} C_0 &= \exp(-\xi^2/2) \left[{}_1F_1\left(-\frac{1}{2}; \frac{1}{2}; \frac{\xi^2}{2}\right) + i\sqrt{\frac{\pi}{2}} \xi \right], \\ C_1 &= \exp(-\xi^2/2) \left[{}_1F_1\left(-\frac{3}{2}; \frac{1}{2}; \frac{\xi^2}{2}\right) + i\sqrt{\frac{\pi}{2}} (3-\xi^2)\frac{\xi}{2} \right]. \end{aligned} \quad (A-12)$$

Each of the series for ${}_1F_1$ consists of terms of the same polarity, except for one term, and are therefore useful for obtaining very accurate initial values. C_0 is the characteristic function of the Rayleigh probability

density function. Relations (A-11)-(A-12) constitute recurrences on both the real and imaginary parts of C_n .

It was found that the terms $\exp(-r) r^n$ in (A-5) became very large for large n , while the $C_n(\mathcal{F})$ terms became very small. In order to avoid overflow and underflow, we defined the total term

$$A_n = \exp(-r) r^n C_n(\mathcal{F}) \quad . \quad (A-13)$$

Reference to (A-11) readily yields the recurrence on A_n , and (A-12) furnishes corresponding obvious starting values for A_0 and A_1 .

APPENDIX B. DESCRIPTION OF PROGRAMS AND LISTINGS

Overview

Information obtained via evaluation of the Rice characteristic function may be displayed in three formats.

FORMAT 1: Display PD vs. PFA

The user defines the number of samples M and the range of values for α , a voltage signal-to-noise ratio measure. An algorithm then utilizes the Rice characteristic function for $\alpha=0$ and for the alphas specified by the user. This results in the production of a threshold vs. PFA and M ($\alpha=0$) and threshold vs. PD and M ($\alpha>0$) tables. These two tables are stored on an output file. For each user-defined M , a plot routine displays PD vs. PFA for the set of user-defined alphas.

FORMAT 2: Display SNR vs M

The user supplies the input which specifies a PD. The algorithm then solves for the threshold values corresponding to $PFA=10^{**(-IPFA)}$, ($IPFA=1, \dots, 8$) and $M=2^{**IM}$, ($IM=0, \dots, 13$) and $\alpha=0$. A root finding technique is then employed to solve for the SNR defined by a threshold value and user-defined PD. An SNR is found for each threshold value. The results are stored in an output file. A plot routine displays the required SNR vs. M for $PFA=10^{**(-IPFA)}$, ($IPFA=1, 2, \dots, 8$).

FORMAT 3: Print SNR

The user specifies a value for PD, PFA, M . The program solves for the threshold corresponding to PFA and M . A root finding technique is then employed to determine the SNR corresponding to this threshold and user-defined PD and M . The results are printed.

Description of Input

Inputs to the program consist of cards which either specify values (PARAMETER CARDS), activate the reading of tabularized values (TABLES), assign files (FILE NAME CARDS), process data (COMMAND CARDS), or specify a plot device (PLOT DEVICE CARDS). The basic format of a card is

$$\text{CARD NAME} = \text{value units}$$

where CARD NAME is an alphanumeric expression from Tables 2-6. The alphanumeric must begin in column 1, value is a floating point or integer number, and units is an alphanumeric.

Parameter cards, file names, and tables constitute the data upon which commands operate. If two cards with the same name specify different data, then the last entry overrides the other.

For the programmers convenience, FORTRAN variable names associated with file names or parameters may be located in the Tables 2 through 6. Since input and values stored represent the same physical quantity, it is convenient to refer to both in this paper by the same variable name. The convention adopted is to express the variable by the lower case letters and reserve upper case letters for constants.

Parameter Cards

Parameter cards are used to specify an axis length or assign a range of values to a parameter. These cards are shown in Table 2. For example,

```
NUMBER OF SAMPLES MINIMUM = 1.  
NUMBER OF SAMPLES MAXIMUM = 8192.  
NUMBER OF SAMPLES FACTOR = 2.
```

implies that the program will process data for $M=1,2,4,8,16,\dots,4096,8192$.

Table Cards

A table card contains the values that are to be assigned to a variable. The last card that must appear in a table is an EOF card. This card terminates the reading of the table. Table cards exist for PD and PFA only. A list of the table cards appears in Table 4. For example,

```
PROBABILITY OF DETECTION TABLE
.5
.7
.99
EOF
```

This table assigns values of .5, .7, .99 to PD.

Files Cards

A file card allows for dynamic assignment of all mass storage files. This is accomplished by linking internal FORTRAN unit numbers to files during execution. The file card is shown in Table 4. Two of the three algorithms use files. They are

```
Display PD vs PFA : A file is used to store output.
Display SNR vs M  : A file is used to store output.
```

For example,

```
OUTPUT FILE = PDFILE
```

directs the output of a program to a file called PDFILE.

Command Cards

Command cards are used to compute, plot, or terminate a run stream. Command cards are given in Table 5.

Plot Device Cards

Plot device cards direct the plot output to either a TEKTRONIX, FR80, or a CALCOMP plotter. The cards necessary for that operation are shown in Table 6.

Examples of Output

Example 1: Display PD vs PFA

The input deck for the first example appears in Table 7. This deck designates that PD vs. PFA data will be computed for $M=1$ and $\alpha=.5, 1.0, 1.5, \dots, 9.5$. The output is stored on a file called FILE1. The plot corresponding to the data is shown in figure 6. The second half of the run stream computes PD vs. PFA data for $M=2$ and $\alpha=0., .4, .8, \dots, 7.2$. The output is stored in FILE2. The plot of the data appears in figure 7.

Example 2: Display SNR vs. M

The input deck for the second example appears in Table 8. The first half of the input deck designates that the SNR vs. M plots will be computed for a value $PD=.5$. The output is displayed in figure 1. The parameter cards specify that the axis will be scaled as follows: -19 DB (minimum), 13 DB (maximum), 2 DB (increment), and 5 inches long for the SNR axis and 6.86 inches long for the number of samples axis. It should be noted that the limits for the number of samples axis are predefined by the program to be 1 (minimum), 8192 (maximum), 2 (factor). The output is stored in a file called PDFIL1. The second half of the run stream computes SNR vs. M for a value $PD=.9$. The axis limits for SNR were changed to -17 DB (minimum), 15 DB (maximum), 2 DB (increment). Alpha curves were computed for $\alpha=0., .4, .8, \dots, 7.2$. This output is stored in file PDFIL2. A plot of this data appears in figure 2.

Example 3: Print SNR

The input deck for the third example appears in Table 9. The output appears in Table 10.

TABLE 2. PARAMETER CARDS

INPUT CARDS	UNITS
SNR AXIS LENGTH = snraxs	IN
SAMPLE AXIS LENGTH = smpaxs	IN
PD AXIS LENGTH = pdaxs	IN
PFA AXIS LENGTH = pfaaxs	IN
SNR MINIMUM = snrmin	DB
SNR MAXIMUM = snrmax	DB
SNR INCREMENT = snrinc	DB
ALPHA MINIMUM = alpmin	
ALPHA MAXIMUM = alpmax	
ALPHA INCREMENT = alpinc	
NUMBER OF SAMPLES MINIMUM = smpmin	
NUMBER OF SAMPLES MAXIMUM = smpmax	
NUMBER OF SAMPLES FACTOR = smpfct	

TABLE 3. TABLE CARDS

INPUT CARDS	VARIABLE
PROBABILITY OF DETECTION TABLE	PD
PROBABILITY OF FALSE ALARM TABLE	PFA

TABLE 4. FILE CARDS

```

-----
INPUT CARDS
-----
OUTPUT FILE = name

```

TABLE 5. COMMAND CARDS

```

-----
INPUT CARDS
-----
RUN MAIN
COMPUTE PD VS PFA
COMPUTE SNR VS M
PLOT PD VS PFA
PLOT SNR VS M
END

```

TABLE 6. PLOT DEVICE CARDS

INPUT CARDS	OPTIONS
BAUD RATE = 960. PLOT DEVICE = device RESET PLOT DEVICE	FR80, TEKTR0, CALCOMP

TABLE 7. SAMPLE INPUT DECK FOR PD VS PFA

```

RUN MAIN
BAUD RATE = 960.
PLOT DEVICE = TEKTR0
RESET PLOT DEVICE
PD AXIS LENGTH = 6.86 IN
PFA AXIS LENGTH = 5. IN
OUTPUT FILE = FILE1
NUMBER OF SAMPLES MINIMUM = 1
ALPHA MINIMUM = .5
ALPHA MAXIMUM = 9.5
ALPHA INCREMENT = .5
COMPUTE PD VS PFA
PLOT PD VS PFA
OUTPUT FILE = FILE2
NUMBER OF SAMPLES MINIMUM = 2
ALPHA MINIMUM = 0.
ALPHA MAXIMUM = 7.2
ALPHA INCREMENT = .4
COMPUTE PD VS PFA
PLOT PD VS PFA
END

```

TABLE 8. SAMPLE INPUT DECK FOR SNR vs M

```
RUN MAIN
BAUD RATE = 960.
TEMPORARY FILE = FALSE
PLOT DEVICE = TEKTR0
RESET PLOT DEVICE
OUTPUT FILE = PDFIL1
SNR MINIMUM = -19. DB
SNR MAXIMUM = 13. DB
SNR INCREMENT = 2. DB
SNR AXIS LENGTH = 5. IN
SAMPLE AXIS LENGTH = 6.86 IN
PROBABILITY OF DETECTION TABLE
.5
EOF
COMPUTE SNR VS M
PLOT SNR VS M
OUTPUT FILE = PDFIL2
SNR MINIMUM = -17. DB
SNR MAXIMUM = 15. DB
SNR INCREMENT = 2. DB
PROBABILITY OF DETECTION TABLE
.9
EOF
COMPUTE SNR VS M
PLOT SNR VS M
END
```

TABLE 9. SAMPLE INPUT DECK FOR PRINTING SNR

```
RUN MAIN
PROBABILITY OF DETECTION TABLE
.5
.9
EOF
PROBABILITY OF FALSE ALARM TABLE
.1
.001
EOF
NUMBER OF SAMPLES MINIMUM = 1.
NUMBER OF SAMPLES MAXIMUM = 2048.
NUMBER OF SAMPLES FACTOR = 2.
PRINT SNR
END
```

TABLE 10. PRINT OUT OF SNR VS M

PD =	0.500	PFA = 0.100D+00	PD =	0.900	PFA = 0.100D+00
M = 1	SNR = 2.50		M = 1	SNR = 7.18	
M = 2	SNR = 0.68		M = 2	SNR = 5.03	
M = 4	SNR = -1.08		M = 4	SNR = 2.99	
M = 8	SNR = -2.77		M = 8	SNR = 1.05	
M = 16	SNR = -4.41		M = 16	SNR = -0.79	
M = 32	SNR = -6.02		M = 32	SNR = -2.55	
M = 64	SNR = -7.59		M = 64	SNR = -4.24	
M = 128	SNR = -9.15		M = 128	SNR = -5.89	
M = 256	SNR = -10.69		M = 256	SNR = -7.50	
M = 512	SNR = -12.22		M = 512	SNR = -9.08	
M = 1024	SNR = -13.74		M = 1024	SNR = -10.64	
M = 2048	SNR = -15.26		M = 2048	SNR = -12.18	

PD =	0.500	PFA = 0.100D-02	PD =	0.900	PFA = 0.100D-02
M = 1	SNR = 8.06		M = 1	SNR = 10.76	
M = 2	SNR = 5.74		M = 2	SNR = 8.29	
M = 4	SNR = 3.63		M = 4	SNR = 6.01	
M = 8	SNR = 1.69		M = 8	SNR = 3.89	
M = 16	SNR = -0.14		M = 16	SNR = 1.92	
M = 32	SNR = -1.87		M = 32	SNR = 0.06	
M = 64	SNR = -3.54		M = 64	SNR = -1.72	
M = 128	SNR = -5.16		M = 128	SNR = -3.42	
M = 256	SNR = -6.75		M = 256	SNR = -5.07	
M = 512	SNR = -8.31		M = 512	SNR = -6.68	
M = 1024	SNR = -9.86		M = 1024	SNR = -8.26	
M = 2048	SNR = -11.39		M = 2048	SNR = -9.82	

Listing of Program

This section contains a listing of three master programs and associated subroutines. Subroutines which read input and plot the output have been omitted. Table 11 contains a list of the subroutine names and a brief description of the pertinent subroutines.

TABLE 11. DESCRIPTION OF SUBROUTINES

NAME -----	DESCRIPTION -----
CMPDVA	MASTER PROGRAM FOR COMPUTING PD VS PFA
CMPSVS	MASTER PROGRAM FOR COMPUTING SNR VS M
PRTSNR	MASTER PROGRAM FOR COMPUTING AND PRINTING SNR
FFT	COMPUTES THE FAST FOURIER TRANSFORM OF A FUNCTION
RDC	COMPUTES AN APPROXIMATE S/N FOR A GIVEN PD, PFA, M (SEE REF 7)
FNPD	COMPUTES THE PROBABILITY OF DETECTION FOR A GIVEN M, S/N, AND THRESHOLD
FNPF	COMPUTES THE PROBABILITY OF FALSE ALARM FOR A GIVEN M AND THRESHOLD
FNF11	COMPUTES THE CONFLUENT HYPERGEOMETRIC FUNCTION
RICE	COMPUTES THE CHARACTERISTIC FUNCTION OF A RICE VARIATE
FNIPHI	COMPUTES THE INVERSE OF THE CUMULATIVE GAUSSIAN DISTRIBUTION
DIST	COMPUTES THE EXCEEDANCE DISTRIBUTION FUNCTION FOR A GIVEN M AND S/N

```

SUBROUTINE FFT(N,X,Y)
IMPLICIT DOUBLE PRECISION (A-H,O-Z)
DIMENSION C(0:256),X(0:1023),Y(0:1023),L(0:9)
DATA PI/3.14159265358979324D0/

T=2.D0*PI/N
J1=N/4
DO 100 J=0,J1
C(J)=DCOS(T*DFLOTJ(J))
100 CONTINUE

N1=N/4
N2=N1+1
N3=N2+1
N4=N3+N1
L2=JIDINT(1.4427D0*BLOG(DFLOTJ(N))+.5D0)
DO 600 I1=1,L2
I2=2**(L2-I1)
I3=2D0*I2
I4=N/I3

DO 500 I5=1,I2
I6=I4*(I5-1)+1
IF( I6.LE.N2 ) GO TO 350
V6=-C(N4-I6-1)
V7=-C(I6-N1-1)
GO TO 375
350 V6=C(I6-1)
V7=-C(N3-I6-1)
375 L3=N-I3

DO 400 I7=0,L3,I3
I8=I7+I5
I9=I8+I2
V8=X(I8-1)-X(I9-1)
V9=Y(I8-1)-Y(I9-1)
X(I8-1)=X(I8-1)+X(I9-1)
Y(I8-1)=Y(I8-1)+Y(I9-1)
X(I9-1)=V6*V8-V7*V9
Y(I9-1)=V6*V9+V7*V8
400 CONTINUE
500 CONTINUE
600 CONTINUE

I1=L2+1
DO 700 I2=1,10
L(I2-1)=1.D0
IF( I2.GT.L2 ) GO TO 700
L(I2-1)=2**(I1-I2)
700 CONTINUE

```

```
IC0=L(0)
IC1=L(1)
IC2=L(2)
IC3=L(3)
IC4=L(4)
IC5=L(5)
IC6=L(6)
IC7=L(7)
IC8=L(8)
IC9=L(9)
```

```
      K=1
      DO 1900 I1=1,IC9
      DO 1800 I2=I1,IC8,IC9
      DO 1700 I3=I2,IC7,IC8
      DO 1600 I4=I3,IC6,IC7
      DO 1500 I5=I4,IC5,IC6
      DO 1400 I6=I5,IC4,IC5
      DO 1300 I7=I6,IC3,IC4
      DO 1200 I8=I7,IC2,IC3
      DO 1100 I9=I8,IC1,IC2
      DO 1000 I10=I9,IC0,IC1
      J=I10
      IF( K.GT.J ) GO TO 900
      A=X(K-1)
      X(K-1)=X(J-1)
      X(J-1)=A
      A=Y(K-1)
      Y(K-1)=Y(J-1)
      Y(J-1)=A
900   K=K+1
1000  CONTINUE
1100  CONTINUE
1200  CONTINUE
1300  CONTINUE
1400  CONTINUE
1500  CONTINUE
1600  CONTINUE
1700  CONTINUE
1800  CONTINUE
1900  CONTINUE
```

```
      RETURN
      END
```



```

SUBROUTINE FNPD(ALPHA,U,AM,AL,AD,ABS,PD)
IMPLICIT DOUBLE PRECISION (A-H,O-Z)
DATA PI/3.14159265358979324D0/

```

```

SNR=.5D0*ALPHA*ALPHA
CALL FNF11(1.5D0,1.D0,SNR,F11)
FAC=DSQRT(.5D0*PI)*EXP(-SNR)*F11
AMUY=AM*FAC+ABS
AM2=AM/2.D0
VD=V*AD
EXC=.5*AD*AMUY
NS1=JIDINT(AL/AD)
DO 100 NS=1,NS1
XI=AD*NS
CALL RICE(XI,SNR,FR,FI)
A=DATAN2(FI,FR)
FYI=DSIN(AM*A+ABS*XI)*(FR*FR+FI*FI)**AM2
ADD=FYI*DCOS(VD*DFLOTJ(NS))/DFLOTJ(NS)
EXC=EXC+ADD
100 CONTINUE
PD=2.D0*EXC/PI

RETURN
END

```

```

SUBROUTINE FNPF(U,AM,AL,AD,ABS,PF)
IMPLICIT DOUBLE PRECISION (A-H,O-Z)
DATA PI/3.14159265358979324D0/

```

```

FAC=DSQRT(.5D0*PI)
AMUY=AM*FAC+ABS
AM2=AM/2.D0
VD=V*AD
EXC=.5*AD*AMUY
NS1=JIDINT(AL/AD)
DO 100 NS=1,NS1
XI=AD*NS
X2=.5D0*XI*XI
E=EXP(-X2)
CALL FNF11(-.5D0,.5D0,X2,F11)
FR=E*F11
FI=E*FAC*XI
A=DATAN2(FI,FR)
FYI=DSIN(AM*A+ABS*XI)*(FR*FR+FI*FI)**AM2
ADD=FYI*DCOS(VD*DFLOTJ(NS))/DFLOTJ(NS)
EXC=EXC+ADD
100 CONTINUE
PF=2.D0*EXC/PI

RETURN
END

```

```

SUBROUTINE RDC(AM,FF,PD,ALPHA)
IMPLICIT DOUBLE PRECISION (A-H,O-Z)

```

```

A=DLOG(.62D0/FF)
B=DLOG(PD/(1.D0-PD))
FACT=6.2D0 + 4.54D0/DSQRT(AM+.44D0)
SNRDB=-5.D0*DLOG10(AM) + DLOG10(A+.12D0*A*B+1.7D0*B)*FACT
ALPHA=DSQRT(2.D0*10.D0**(.1D0*SNRDB))
RETURN
END

```

```

SUBROUTINE FNF11(A,B,X,F11)
IMPLICIT DOUBLE PRECISION (A-H,O-Z)

```

```

F11=1.D0
T=1.D0
DO 100 K=1,300
U=K-1
T=T*(A+U)*X/((B+U)*K)
F11=F11+T
IF( DABS(T).LE.DABS(F11)*1.D-18 ) GO TO 200
100 CONTINUE
PRINT 101
101 FORMAT(2X,'300 TERMS IN FNF11')
200 CONTINUE

RETURN
END

```

```

SUBROUTINE FNIPHI(X,PHI)
IMPLICIT DOUBLE PRECISION (A-H,O-Z)

```

```

100 Y=DMAX1(X,1.D-12)
Y=DMIN1(Y,1.D0-1.D-12)
D=X-.5D0
IF( DABS(D).GT..01D0 ) GO TO 250
PHI=2.50662827463D0*D*(1.D0+D*D*1.04719755120D0)
GO TO 300
250 PHI=Y
IF( Y.GT..5D0 ) PHI=.5D0-(Y-.5D0)
PHI=DSQRT(-2.D0*DLOG(PHI))
T=1.D0+PHI*(1.432788D0+PHI*(.189269D0+PHI*.001308D0))
PHI=PHI-(2.515517D0+PHI*(.802853D0+PHI*.010328D0))/T
300 RETURN
END

```

```

SUBROUTINE DIST(AM,ALPHA,MF,X,Y)
IMPLICIT DOUBLE PRECISION (A-H,O-Z)
DIMENSION X(0:1023), Y(0:1023)
COMMON /PDVPF/AL,AD,ABS
DATA PI/3.14159265358979324D0/

SNR=.5D0*ALPHA*ALPHA
CALL FNF11(1.5D0,1.D0,SNR,F11)
AMU=DSQRT(.5D0*PI)*DEXP(-SNR)*F11
AMUS=AM*AMU+ABS
AM2=AM/2.D0

DO 100 I=0,1023
X(I)=0.D0
Y(I)=0.D0
100 CONTINUE

X(0)=.5D0*AMUS*AD
NS1=JIDINT(AL/AD)
DO 1000 NS=1,NS1
XI=AD*NS
CALL RICE(XI,SNR,U,V)
T=DATAN2(V,U)
FI=DSIN(AM*T+ABS*XI)*(U*U+V*V)**AM2
MS=JMOD(NS,MF)
X(MS)=X(MS)+FI/NS
1000 CONTINUE

CALL FFT(MF,X,Y)

FAC=2.D0/PI
KS1=MF/2.D0
DO 2000 KS=0,KS1
T=X(KS)*FAC
X(KS)=1.D0-T
Y(KS)=T
2000 CONTINUE

RETURN
END

```

```

SUBROUTINE RICE(X,SNR,FR,FI)
IMPLICIT DOUBLE PRECISION (A-H,O-Z)
DATA PI/3.14159265358979324D0/

```

```

X2=.5D0*X*X
E=DEXP(-X2-SNR)
CALL FNF11(-.5D0,.5D0,X2,F11)
AOR=E*F11
AOI=E*DSQRT(.5D0*PI)*X
CALL FNF11(-1.5D0,.5D0,X2,F11)
ANR=E*SNR*F11
ANI=SNR*(1.5D0-X2)*AOI
FR=AOR+ANR
FI=AOI+ANI
BR=DMAX1(DABS(AOR),DABS(FR))
BI=DMAX1(DABS(AOI),DABS(FI))
T=.5D0+X2

```

```

SNR2=SNR**2
DO 100 N=2,200
FO=N**2
F1=SNR*(N+N-T)/FO
F2=SNR2*(N-.5D0)/((N-1)*FO)
R=F1*ANR-F2*ADR
V=F1*ANI-F2*AOI
ADR=ANR
AOI=ANI
ANR=R
ANI=V
FR=FR+R
FI=FI+V
BR=DMAX1(BR,DABS(FR))
BI=DMAX1(BI,DABS(FI))
IF( DABS(V).LE.5.D-19*DABS(FI) .AND. DABS(R).LE.5.D-19*DABS(FR))
1 GO TO 200
100 CONTINUE
PRINT 101
101 FORMAT(2X,'200 TERMS IN RICE')
200 DR=18.-DLOG10(DABS(BR/FR))
DI=18.-DLOG10(DABS(BI/FI))
RETURN
END

```

```

SUBROUTINE CMPDVA
PARAMETER MF=2**10
PARAMETER PBDNUM=18
DOUBLE PRECISION AL,AD,ABS,BSA,AM,ALPHA,ALFA,X(0:1023),Y(0:1023)
PARAMETER (NUMFIL=30)
CHARACTER*6 FILES(NUMFIL)
COMMON /FILEC/FILES
CHARACTER*6 PBDNAM
EQUIVALENCE
1 (PBDNAM,FILES(18))
PARAMETER (NUMPAR=200)
COMMON /PARAMC/PARAMS(NUMPAR)
EQUIVALENCE
1 (SMPMIN,PARAMS(187)),
1 (SNMIN,PARAMS(184)), (SNMAX,PARAMS(185)), (SNDEL,PARAMS(186))
COMMON/PDVPF/AL,AD,ABS
DOUBLE PRECISION PI
DATA PI/3.14159265358979324D0/

C
C OPEN THE FILE
C CALL OPNFIL(PBDNUM,PBDNAM)
C
C COMPUTE THE NUMBER OF SNR CURVES
C
C NSN=(SNMAX-SNMIN)/SNDEL + 1
C
C
C STORE HEADER INFO
C WRITE(PBDNUM) SMPMIN,SNMIN,SNMAX,SNDEL,NSN
C
C AM = SMPMIN
C AL = DMIN1(9.D0,17.D0/DSQRT(AM))
C AD = .12D0/DSQRT(AM)
C BSA = -DSQRT(PI/2.D0)*AM + 6.D0*DSQRT(AM)
C ABS = DMIN1(0.D0,BSA)
C
C COMPUTE SNR VS PFA
C ALFA=0.D0
C CALL DIST(AM,ALFA,MF,X,Y)
C
C STORE THE SNR VS PD
C WRITE(PBDNUM) (Y(I),I=0,512)
C
C DO 1000 ISN=1,NSN
C SNR= SNMIN + SNDEL*(ISN-1)
C
C ALPHA = SNR
C CALL DIST(AM,ALPHA,MF,X,Y)
C
C STORE THE SNR VS PD
C WRITE(PBDNUM) (Y(I),I=0,512)
C
1000 CONTINUE
2000 CONTINUE
RETURN
END

```

```

SUBROUTINE PRTSNR
IMPLICIT DOUBLE PRECISION (A-H,O-Z)
DIMENSION PFA(10),PD(10),V(14,8),SNR(14,8)
DATA PI/3.14159265358979324D0/
REAL SMPMIN,SMPMAX,SMPFCT,PARAMS
PARAMETER NUMPAR=200
COMMON/PARAMC/PARAMS(NUMPAR)
EQUIVALENCE
1 (SMPMIN,PARAMS(187)),(SMPMAX,PARAMS(188)), (SMPFCT,PARAMS
COMMON/PDFF/NPD,NPFA,PD,PFA

MMAX=ALOG10(SMPMAX/SMPMIN)/ALOG10(SMPFCT) + 1

F1=DSQRT(.5D0*PI)
F2=DSQRT(2.D0-.5D0*PI)
DO 1000 IM=1,MMAX
AM=SMPMIN*SMPFCT**(IM-1)
AL = DMIN1(9.D0,17.D0/DSQRT(AM))
AD = .12D0/DSQRT(AM)
BSA = -DSQRT(PI/2.D0)*AM + 6.D0*DSQRT(AM)
ABS = DMIN1(0.D0,BSA)
AMU=F1*AM
SIG=F2*DSQRT(AM)
DO 900 IPF=1,NPFA
PF=PFA(IPF)
IF( AM.GT. 1.D0 ) GO TO 250
VN=DSQRT(-2.*DLOG(PF))
GO TO 750
250 CALL FNIPHI(PF,YF)
V1=AMU-SIG*YF+ABS
IF( IPF.GT.1 ) V1=DMAX1(V1,VN)
V2=V1+.5D0
IF( V1.NE.VN ) GO TO 300
P1=PN
GO TO 325
300 CALL FNPFF(V1,AM,AL,AD,ABS,P1)
325 CALL FNPFF(V2,AM,AL,AD,ABS,P2)
IF( DABS(P1-PF).LT.DABS(P2-PF) ) GO TO 350
V0=V1
P0=P1
VN=V2
PN=P2
GO TO 400
350 V0=V2
P0=P2
VN=V1
PN=P1
400 CALL FNIPHI(P0,Y0)
GO TO 550
500 CALL FNPFF(VN,AM,AL,AD,ABS,PN)
550 CALL FNIPHI(PN,YN)
IF( DABS(PN-PF).LE.1D-9*PF ) GO TO 750
T=(V0*(YN-YF)+VN*(YF-Y0))/(YN-Y0)

```

```

      V0=VN
      Y0=YN
      VN=T
      GO TO 500
750   V(IM,IPF)=VN
900   CONTINUE
1000  CONTINUE

      DO 4000 IPD=1,NPD

      CALL FNIPHI(PD(IPD),YD)
      DO 3000 IM=1,MMAX
      AM=SMPMIN*SMPFCT**(IM-1)
      AL = DMIN1(9.DO,17.DO/DSQRT(AM))
      AD = .12DO/DSQRT(AM)
      BSA = -DSQRT(PI/2.DO)*AM + 6.DO*DSQRT(AM)
      ABS = DMIN1(0.DO,BSA)
      DO 2900 IPF=1,NPFA
      PF=PPA(IPF)
      CALL RDC(AM,PF,PD(IPD),A1)
      A2=A1*1.01DO
      VV=V(IM,IPF)
      CALL FNPD(A1,VV,AM,AL,AD,ABS,P1)
      CALL FNPD(A2,VV,AM,AL,AD,ABS,P2)
      IF( DABS(P1-PD(IPD)).LT.DABS(P2-PD(IPD)) ) GO TO 2350
      A0=A1
      P0=P1
      AN=A2
      PN=P2
      GO TO 2400
2350  A0=A2
      P0=P2
      AN=A1
      PN=P1
2400  CALL FNIPHI(P0,Y0)
      GO TO 2550
2500  CALL FNPD(AN,VV,AM,AL,AD,ABS,PN)
2550  CALL FNIPHI(PN,YN)
      IF( DABS(PN-PD(IPD)).LE.1D-6*PD(IPD) ) GO TO 2750
      T=(A0*(YN-YD)+AN*(YD-Y0))/(YN-Y0)
      A0=AN
      Y0=YN
      AN=T
      GO TO 2500
2750  SNR(IM,IPF)=10.*DLOG10(.5DO*AN*AN)
2900  CONTINUE
3000  CONTINUE

```

TR 7117

```
      DO 3200 IPF=1,NPFA
      PRINT 3001
3001  FORMAT(2(/))
      PRINT 3011, PD(IPD),PFA(IPF)
3011  FORMAT(2X,'PD =',F10.3,5X,'PFA =',D10.3)
      DO 3100 IM=1,MMAX
      M=SMPMIN*SMPFCT**(IM-1)
      PRINT 3021, M,SNR(IM,IPF)
3021  FORMAT(2X,'M =',I5,5X,'SNR =',F7.2)
3100  CONTINUE
3200  CONTINUE

4000  CONTINUE
```

```
      RETURN
      END
```

```
      SUBROUTINE CMPSVS
      IMPLICIT DOUBLE PRECISION (A-H,O-Z)
      PARAMETER MMAX=14
      PARAMETER NUMFIL=30, PBDNUM=18
      CHARACTER*6 FILES(NUMFIL)
      COMMON/FILEC/FILES
      CHARACTER*6 PBDNAM
      EQUIVALENCE (PBDNAM,FILES(18))
      DIMENSION PFA(10),PD(10),V(14,8),ALPHA(14,8)
      DIMENSION THRS(14,8)
      DATA PI/3.14159265358979324D0/
      COMMON/PDPF/NPD,NPFA,PD,PFA
```

```
      CALL OPNFIL(PBDNUM,PBDNAM)
```



```

F1=DSQRT(.5D0*PI)
F2=DSQRT(2.D0-.5D0*PI)
DO 1000 IM=1,MMAX
AM=2.** (IM-1)
AL = DMIN1(9.D0,17.D0/DSQRT(AM))
AD = .12D0/DSQRT(AM)
BSA = -DSQRT(PI/2.D0)*AM + 6.D0*DSQRT(AM)
ABS = DMIN1(0.D0,BSA)
AMU=F1*AM
SIG=F2*DSQRT(AM)
DO 900 IPF=1,8
PF=10.**(-DFLOTJ(IPF))
IF( AM.GT. 1.D0 ) GO TO 250
VN=DSQRT(-2.*DLOG(PF))
GO TO 750
250 CALL FNIPHI(PF,YF)
V1=AMU-SIG*YF+ABS
IF( IPF.GT.1 ) V1=DMAX1(V1,VN)
V2=V1+.5D0
IF( V1.NE.VN ) GO TO 300
P1=PN
GO TO 325
300 CALL FNPF(V1,AM,AL,AD,ABS,P1)
325 CALL FNPF(V2,AM,AL,AD,ABS,P2)
IF( DABS(P1-PF).LT.DABS(P2-PF) ) GO TO 350
V0=V1
P0=P1
VN=V2
PN=P2
GO TO 400
350 V0=V2
P0=P2
VN=V1
PN=P1
400 CALL FNIPHI(P0,Y0)
GO TO 550
500 CALL FNPF(VN,AM,AL,AD,ABS,PN)
550 CALL FNIPHI(PN,YN)
IF( DABS(PN-PF).LE.1D-9*PF ) GO TO 750
T=(V0*(YN-YF)+VN*(YF-Y0))/(YN-Y0)
V0=VN
Y0=YN
VN=T
GO TO 500
750 V(IM,IPF)=VN
THRS(IM,IPF)=(VN-ABS)/AM
900 CONTINUE
1000 CONTINUE

```

```

WRITE(PBDNUM)  NPD,(PD(I),I=1,10)
DO 4000 IPD=1,NPD

CALL FNIPHI(PD(IPD),YD)
DO 3000 IM=1,MMAX
AM=2.DO**(IM-1)
AL = DMIN1(9.DO,17.DO/DSQRT(AM))
AD = .12DO/DSQRT(AM)
BSA = -DSQRT(PI/2.DO)*AM + 6.DO*DSQRT(AM)
ABS = DMIN1(0.DO,BSA)
DO 2900 IPF=1,8
PF=10.DO**(-DFLOTJ(IPF))
CALL RDC(AM,PF,PD(IPD),A1)
A2=A1*1.01DO
VV=V(IM,IPF)
CALL FNPD(A1,VV,AM,AL,AD,ABS,P1)
CALL FNPD(A2,VV,AM,AL,AD,ABS,P2)
IF( DABS(P1-PD(IPD)).LT.DABS(P2-PD(IPD)) ) GO TO 2350
A0=A1
P0=P1
AN=A2
PN=P2
GO TO 2400
2350 A0=A2
P0=P2
AN=A1
PN=P1
2400 CALL FNIPHI(P0,Y0)
GO TO 2550
2500 CALL FNPD(AN,VV,AM,AL,AD,ABS,PN)
2550 CALL FNIPHI(PN,YN)
IF( DABS(PN-PD(IPD)).LE.1D-6*PD(IPD) ) GO TO 2750
T=(A0*(YN-YD)+AN*(YD-Y0))/(YN-Y0)
A0=AN
Y0=YN
AN=T
GO TO 2500
2750 ALPHA(IM,IPF)=AN
2900 CONTINUE
3000 CONTINUE

WRITE(PBDNUM) ((ALPHA(IM,IPF),IPF=1,8),IM=1,MMAX)

4000 CONTINUE

RETURN
END

```

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