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CAPACITY OF MISMATCHED GAUSSIAN CHANNELS

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Introduction

The capacity of the Gaussian channel without feedback, subject to a generalized energy constraint, is determined in [1/2]. In that work, the constraint is given in terms of the covariance of the channel noise process. However, there are many situations where one may wish to determine capacity subject to a constraint determined by a covariance that is different from that of the channel noise. An example is in jamming or countermeasures situations.

Channels where the covariance of the noise is the same as that of the constraint will be called matched channels; otherwise, we say that the channel is mismatched (to the constraint). In this paper, the capacity of the mismatched Gaussian channel is determined for two situations: the finite-dimensional channel, and the infinite-dimensional channel with a dimensionality constraint on the space of transmitted signals. Results on the infinite-dimensional mismatched channel without a dimensionality constraint on the signal are given elsewhere [2]. Various special cases of the mismatched channel have been treated previously [3]-[5].

The results for the mismatched channel differ significantly from those for the matched channel. A discussion of these differences follows the proof of the main result.

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Definitions and Structure

The channel is defined as in [1]. H₁ and H₂ are real separable Hilbert spaces with Borel σ fields β_1 and β_2 and inner products $\langle \cdot, \cdot \rangle_1$ and $\langle \cdot, \cdot \rangle_2$. The message process X in H₁ is represented by a probability μ_X on (H₁, β_1). The message is encoded into the transmitted signal A(X) in H₂ by a β_1/β_2 -measurable coding function A. To each sample function of the signal process, the channel adds a sample function from the noise process N, represented by a Gaussian measure μ_N on (H₂, β_2). The received signal (channel output) is then a sample function from the process Y = A(X) + N, represented by the measure μ_Y . As usual X and N are assumed independent, so that $\mu_Y(B) = \mu_X \vartheta \mu_N \{(X,Y): A(X) + Y \in B\}$ where $\mu_X \vartheta u_N$ is product measure. The channel probability μ_{XY} , which has marginal measures $\mu_X \vartheta u_N \{(X,Y): (X,A(X) + Y) \in C\}$. The average mutual information is then I[μ_{XY}], where I[μ_{XY}] $\equiv \infty$ if it is false that μ_{XY} is absolutely continuous with respect to $\mu_X \vartheta u_Y (\mu_{XY} << \mu_X \vartheta u_Y)$, and otherwise I[μ_{XY}] = $\int_{H_1 \times H_2} \log \left[\frac{d\mu_{XY}}{d\mu_X \vartheta u_Y} \right] (x,y) d\mu_{XY} (x,y)$.

The information capacity is then sup $I[\mu_{\chi\gamma}],$ where Q is a set of admissible pairs $(\mu_{\chi},A).$

For this paper, a covariance operator in a Hilbert space will be defined to be a symmetric and trace-class bounded linear operator. The constraint on the transmitted signal process A(X) will be given in terms of a covariance operator R_W in H_2 ; as is well-known, to every such covariance operator there corresponds a zero-mean Gaussian measure on (H_2, β_2) .

When $H_2 = L_2 [0,T]$ and R is a covariance operator, R can be represented as an integral operator with kernel R which is a covariance function. There is then a well-known isomorphism between range $(R^{1/2})$ and the reproducing kernel Hilbert space of R. All measures considered here will be assumedd WLOG to have zero mean. The capacity will he determined under the following assumptions:

(A-1) $R_N = R_W^{1/2}$ (I + S) $R_W^{1/2}$, where R_N is the covariance operator of the noise ³ measure μ_N , and S is a compact linear operator that does not have -1 as an eigenvalue;

(A-2) The admissible set Q is the set of all (μ_{χ}, A) $\int_{H_1} ||R_W^{-1/2} A(x)||_2^2 d\mu_{\chi}(x) \leq P$, where P>O is fixed.

It will be assumed WLOG [1] that range $(R_N) = H_2$, so that $R_N^{-1/2}$ exists. Assumption (A-1) then implies that $R_W^{-1/2}$ exists; in fact, that range $(R_N^{-1/2})$ =range $(R_W^{-1/2})$. Thus, there exists a unitary operator U in H₂ such that $R_N^{1/2} = R_W^{1/2}(I + S)^{1/2}U^*$, where U* is the adjoint of U.

The class of all zero-mean Gaussian measures μ_N with covariance operator as in (A-1) includes all those that are mutually absolutely continuous with respect to μ_W , where μ_W is zero-mean Gaussian with covariance $R_W[6]$.

From the results of [1], one can limit attention to cases where $\mu_{A(X)}$ is Gaussian with covariance operator

$$R_{A(X)} = \Sigma_n \tau_n [R_N^{\prime} u_n] \ge [R_N^{\prime} u_n]$$
 (2)

where $\tau_n \ge 0$ for $n \ge 1$, $\Sigma_n \tau_n < \infty, \{u_n, n \ge 1\}$ is a c.o.n. set and $(uav)x = \langle v, x \rangle u$. When $\mu_{A(X)}$ has (2) for covariance and is Gaussian then [1] $I[\mu_{XY}] = (1/2) \Sigma_n \log [1 + \tau_n].$ (3)

Moreover, $E_{\mu\nu} ||R_W^{-1/2} A(S)||_2^2 = \text{Trace } R_W^{-1/2} R_{A(X)} R_W^{-1/2}$

$$= \sum_{n} \tau_{n} || (I + S)^{1/2} || U^{*}u_{n} ||_{2}^{2} .$$
(4)

Defining $X_n^2 = \tau_n ||(I + S)^{1/2} ||U^*u_n||_2^2$, the capacity problem is thus reduced to maximizing $\binom{1}{2} \sum_n \log [1 + X_n^2 (1 + \gamma_n)^{-1}]$ (5) over all sequences (X_n^2) and c.o.n. sets $\{v_n, n \ge 1\}$ such that $\sum_n X_n^2 \le P$, where $\gamma_n \equiv \langle Sv_n, v_n \rangle_2, n \ge 1$.

The supremium of (5) subject to the stated constraint is the capacity and will be denoted as $C_W(P)$; the capacity for the matched channel ($R_W = R_N$) will be denoted by $C_N(P)$.

Preliminary Results

Lemma 1: Let (Y_n) , $n \le M$, be any non-decreasing sequence of strictly positive real numbers. Let (X_n) be any sequence of M real numbers. Fix P > 0, and define $g(M, P, \gamma) = \sup_{\substack{X: \\ Y \\ X}} M (Y_n + X_n^2)/Y_n.$

Then $g(M,P,\gamma) = \prod_{n=1}^{K} (\sum_{i=1}^{K} \gamma_i + P)/(K\gamma_n)$ where $K \leq M$ is the largest integer such that $\sum_{i=1}^{K} \gamma_i + P \geq K\gamma_K$. $g(M,P,\gamma)$ is uniquely attained by (X_n^2) such that $X_n^2 = (\sum_{i=1}^{K} \gamma_i + P)/K - \gamma_n$ $n \leq K$

Proof: Define
$$f_M: IR^M \rightarrow IR$$
 by $f_M(\underline{y}) = \sum_{n=1}^{M} \log [1 + y_n \gamma_n^{-1}]$. We seek to maximize $n=1$

n > K.

f_M subject to the constraints

$$g(\underline{y}) = \sum_{1}^{M} y_{n} - P \le 0$$

h_i(\underline{y}) = -y_i \le 0, i = 1,..., 1

= 0

This is a constrained optimization problem with objective function f_M which is concave over the convex set { \underline{Z} in IR^M ; $Z_i \ge 0$, i = 1, ..., M}. Moreover, each constraint function is linear. Thus, any solution to this problem will define a unique global maximum for f_M [7]. In order that \underline{y}^* be a solution, it is necessary and sufficient that the following set of equations be satisfied [7]:

$$\frac{1}{y_{i}^{*}+\gamma_{i}}+\beta-\alpha_{i}=0 \quad i=1,..,M$$
 (6)

$$\sum_{1}^{M} y_{n}^{\star} - P \leq 0, \quad \beta \left[\sum_{1}^{\Sigma} y_{n}^{\star} - P \right] = 0$$
(7)

 $-y_{i}^{\star} \leq 0, \ \alpha_{i} \ y_{i}^{\star} = 0, \ i = 1,..,M$ (8)

for some set of non-positive real numbers $\{\beta, \alpha_1, \dots, \alpha_M\}$.

We first attempt to obtain a solution by setting $\alpha_1 = \alpha_2 = \dots = \alpha_M = 0$. This requires $\beta(\gamma_i + y_i^*) = -1$ for $i = 1, \dots, M$; thus,

$$\begin{array}{cccc} M & M \\ \Sigma & y^{\star} & + & \Sigma & \gamma_{i} \\ 1 & i & 1 \end{array} ^{-1}, \text{ and so } y^{\star}_{n} = \begin{pmatrix} M & \star & M \\ \Sigma & y^{\star}_{i} & + & \Sigma & \gamma_{i} \\ 1 & 1 & 1 \end{array} ^{-1} / M - \gamma_{n}$$

for n = 1, 2, ..., M. This definition of y^* and the constraints (8) require that

$$\begin{array}{c} M & M \\ \Sigma & y_i^* + \Sigma & \gamma_i \geq M \gamma_n \\ 1 & 1 & 1 \end{array}$$

for $n \le M$; this inequality is satisfied for all $n \le M$ if and only if it is satisfied for n = M. Also, $\beta^{-1} = -(y_i^* + \gamma_i)$ for $i \le M$ implies $\beta < 0$, so that $\sum_{i=1}^{N} y_i^* = P$ by constraints (7). Thus, if $P + \sum_{i=1}^{N} \gamma_i \ge M \gamma_M$, an optimum solution is given by

$$y_{i}^{*} = (P + \sum_{i=1}^{N} p_{i} - M_{Y_{i}})/M, \quad i \leq M.$$

If there exists K<M such that $K_{Y_{K}} \leq P + \sum_{i=1}^{K} i_{i} < (K+1)_{Y_{K+1}}$, then constraints (6)-(8) are satisfied by choosing $\beta = -K[P + \sum_{i=1}^{K} j_{i}]^{-1}$,

$$K_{\Sigma y_{i}}^{K} = P$$

$$y^{*} = 0, i > K$$

$$y_{i}^{*} = K^{-1} \left[P + \sum_{1}^{K} \gamma_{n} - K \gamma_{i} \right], i \le K$$

$$\alpha_{i}^{K} = -K \left[P + \sum_{1}^{K} \gamma_{n} \right]^{-1} + \gamma_{i}^{-1} i > K$$

 $\alpha_1 = \alpha_2 = \ldots = \alpha_K = 0,$

Thus,

$$\sup_{\substack{M \ 2 \\ \{\underline{X}: \ \underline{\Sigma}X_n \le P\}}} \frac{M}{\prod_{n=1}^{M} (\gamma_n + X_n)/\gamma_n} = \frac{K}{n=1} \frac{K}{\binom{\Sigma}{1}} \frac{Y_i}{i} + \frac{P}{\binom{K}{n}} (K\gamma_n)$$

where $K \leq M$ is the largest integer such that $\sum_{i=1}^{K} \gamma_i + P \geq K \gamma_K$. The supremum is attained by \underline{y}^* as defined above, or for

$$X_{n}^{2} = \left[P + \sum_{i=1}^{K} \frac{1}{i}\right]/K - \gamma_{n} \qquad n \le K$$
$$= 0 \qquad n > K.$$

Lemma 2: Let (λ_i) , $1 \le i \le K$, be a non-decreasing sequence of strictly positive real numbers. Suppose that (γ_n) is a non-decreasing sequence such that $\sum_{i=1}^{r} \sum_{j=1}^{r} \sum_{i=1}^{r} \lambda_{i}$ for all $J \leq K$, and let P > 0 be such that $\sum_{i=1}^{r} \sum_{j=1}^{r} P \geq K_{Y_{K}}$. Define $f_{K}(\gamma) = \prod_{n=1}^{K} (P + \sum_{j=1}^{K} \gamma_{j})/(K\gamma_{n})$. Then $f_{K}(\gamma) \leq f_{K}(\lambda)$ with equality if and only if $\gamma_i = \lambda_i$ for all $i \leq K$. <u>Proof</u>: For any fixed n, $\partial f_{K}(\gamma)/\partial \gamma_{n}$ is negative, using $\sum_{i=1}^{K} + P \ge K_{\gamma_{K}}$. Thus $f_{K}(\gamma)$. increases for γ_n decreasing. One can now assume that $\sum_{n=1}^{\infty} \gamma_n = \sum_{n=1}^{\infty} \lambda_n$. To see this, suppose $\sum_{n=1}^{\infty} \gamma_n > \sum_{n=1}^{\infty} \lambda_n$. First assume that there exists $p \le K$ such that $\gamma_p > \gamma_{p-1}$ and $\gamma_p > \lambda_p$. Define a sequence (γ_n') by $\gamma'_n = \gamma_n$ if $n \neq p$, $\gamma'_p = \gamma_p - \epsilon$, $\epsilon = \min (\gamma_p - \gamma_{p-1}, \sum_{i=1}^{n} (\gamma_i - \lambda_i), \gamma_p - \lambda_p).$ Continuing to form new sequences in this manner, one will eventually obtain a non-decreasing sequence $\binom{\gamma}{n}$ with $\sum_{j} \frac{\gamma}{n} \geq \sum_{j} \frac{\lambda}{n}$ for all $j \leq K$, and either $\sum_{j} \frac{\kappa}{n} = \sum_{j} \frac{\kappa}{n}$ or else $\gamma_1 = \gamma_2 = \dots = \gamma_p^{\prime}$, where p is the largest integer i such that $\gamma_1 > \lambda_1$. If the latter case holds, define a new sequence (γ_n) , with $\gamma_n = \gamma_n$ for $2 \le n \le K$,

If the latter case holds, define a new sequence (γ_n) , with $\gamma_n = \gamma_n$ for $2 \le n \le K$, while $\gamma_1 = \gamma_1 - \varepsilon$, $\varepsilon = \min(\gamma_1 - \lambda_1, \sum_{i=1}^{K} (\gamma_i - \lambda_i))$. (γ_n) is non-decreasing and $\sum_{i=1}^{K} \gamma_i^{"} \ge \sum_{i=1}^{L} \lambda_n$. If $\varepsilon = \gamma_1 - \lambda_1$, the procedure is repeated for (γ_n) and γ_2 ; if $\varepsilon = \sum_{i=1}^{K} (\gamma_i - \lambda_i)$, the procedure is repeated for (γ_n) and $\gamma_1^{"}$. Continuing in this manner will eventually produce a sequence $(\gamma_n^{"})$ such that $\sum_{i=1}^{K} \gamma_i^{"} = \sum_{i=1}^{K} \lambda_i$.

Assume then that $\sum_{1}^{K} \gamma_n = \sum_{1}^{K} \lambda_n$. If γ and λ are not identical, let $p \leq K$ be the largest integer such that $\gamma_p \neq \lambda_p$; since $\sum_{1}^{p-1} \gamma_n \geq \sum_{1}^{p-1} \lambda_n$ and $\sum_{1}^{p} \gamma_n = \sum_{1}^{p} \lambda_n$, $\gamma_p < \lambda_p$. Let t < p be the largest integer such that $\gamma_t > \lambda_t$; such t must exist. Define a new sequence (γ'_n) by $\gamma'_n = \gamma_n$ if $n \neq t$, $n \neq p$, while $\gamma'_p = \gamma_p + \varepsilon$, $\gamma'_t = \gamma_t - \varepsilon$, $\varepsilon = \inf (\lambda_p - \gamma_p, \gamma_t - \lambda_t)$. $f_K(\gamma) < f_K(\gamma')$, since $(\gamma_t - \varepsilon)(\gamma_p + \varepsilon)$ $= \gamma_t \gamma_p - \varepsilon(\gamma_p - \gamma_t) - \varepsilon^2$, and $\gamma_p \geq \gamma_t$. This procedure is successively repeated; it will terminate when and only when one obtains a sequence (γ'_n) such that $\gamma'_n = \lambda_n$ for all $n \leq K$.

Main Result

Theorem:

(a) Suppose that H_2 has dimension $M < \infty$. The capacity is then

$$C_{W}(P) = \binom{1}{2} \sum_{n=1}^{K} \log \left[\frac{K}{\sum \Theta_{1}} + P + K}{\frac{1}{K(1 + \Theta_{n})}} \right]$$

where $\theta_1 \leq \theta_2 \leq \cdots \leq \theta_M$ are the eigenvalues of S, and K is the largest integer $\leq M$ such that $\sum_{l=0}^{K} \theta_l + P \geq K \theta_K$. The capacity is attained by a Gaussian $\mu_{A(X)}$ with covariance operator (2), where $u_n = Ue_n$ and $\tau_n = \begin{bmatrix} K \\ \Sigma \theta_i + P - K \theta_n \end{bmatrix} (1 + \theta_n)^{-1} K^{-1}$ for $n \leq K$, $\tau_n = 0$ for n > K, and $\{e_n, n \geq 1\}$ are o.n. eigenvectors of S corresponding to the eigenvalues (θ_n) . No other Gaussian $\mu_{A(X)}$ can attain capacity. The same result is obtained if H_2 has dimension $L < \infty$ and $\mu_{A(X)}$ is constrained to have support of dimension M < L.

- (b) Suppose that H₂ is infinite-dimensional and that support $(\mu_{A(X)})$ is restructed to have dimension $\leq M < \infty$. Let (λ_n) , $n \geq 1$, be the nondecreasing strictly negative eigenvalues of S.
- (i) If $\{\lambda_n, n \ge 1\}$ is empty, then $C_W(P) = (M/2) \log [1 + P/M]$. Capacity can be attained if and only if S has zero as an eigenvalue of multiplicity $\ge M$. In this case $C_W(P)$ is attained by a Gaussian $\mu_{A(X)}$ with covariance (2), where $u_i = Ue_i$ and $\tau_i = P/M$ for $i \le M$, with $\{e_1, \dots, e_M\}$ any o.n. set in the null space of S.
- (ii) If $K_{\lambda} \leq \sum_{i=1}^{K} i + P < K_{\lambda} + P < K_{K+1}$ for some $K \leq M$, then the capacity is as in (a), with $\theta_{i} = \lambda_{i}$, i = 1, ..., K, and can be similarly attained.

(iii) If S has K < M strictly negative eigenvalues, K > 0, and P $+\Sigma\lambda_i \ge K\lambda_K$, then the capacity is

$$C_{W}(P) = (M/2) \log \frac{P + M + \sum_{i=1}^{K} \lambda_{i}}{M}$$

$$- (\frac{1}{2}) \sum_{n=1}^{k} \log (1 + \lambda_n).$$

The capacity can be attained if and only if zero is an eigenvalue of S with multiplicity $\geq M-K$. The capacity is then achieved by a Gaussian μ_{AX} with covariance (2), where $u_n = Ue_n$ and $\tau_n = (\sum_{\lambda} i + P - M\lambda_n)(1 + \lambda_n)^{-1}M^{-1}$ for $n \leq K$, with $Se_n = \lambda_n e_n$ and $\{e_1, \dots, e_K\}$ an o.n. set; and with $u_n = Uv_n$ and $\tau_n = (P + \sum_{1}^{K} \tau_1)M^{-1}$ for $K + 1 \leq n \leq M$, where $Sv_n = 0$ and $\{v_{K+1}, \dots, v_M\}$ is an o.n. set. The sets $\{u_1, \dots, u_K\}$ and $\{\tau_1, \dots, \tau_K\}$ are uniquely defined for any maximizing Gaussian $\mu_A(X)$.

(c) In (b-ii) and (b-iii), the capacity is strictly greater than for the case $R_N = R_W$; i.e., $C_W(P) > C_N(P)$. In (b-i), $C_W(P) = C_N(P)$. In (a), $C_W(P) > C_N(P)$ if

$$\sum_{i=1}^{M} \frac{\zeta}{2} 0, \text{ or if } P + \sum_{i=1}^{K} \frac{\zeta}{2} 0. \quad C_{W}(P) < C_{N}(P) \text{ if } 0 \leq \theta_{1} < \theta_{M}.$$

 $\begin{array}{lll} \underline{Proof}: & (a) & \text{From (5),} \\ \hline C_W(P) &= \sup \binom{l}{2} \sum\limits_{1}^{\Sigma} \log \left[1 + \chi \frac{2}{n} \gamma_n^{-1}\right], & \text{where } \gamma_n = 1 + \langle Sv_n, v_n \rangle_2, \{v_n, n \leq M\} \\ \text{is a c.o.n. set, and the supremum is over all such c.o.n. sets and all } (\chi_n^2) \\ \text{such that } \sum\limits_{1}^{X} \chi_n^2 &\leq P. & \text{Since } \theta_1 \leq \theta_2 \leq \cdots \leq \theta_M \text{ are the non-decreasing eigenvalues of} \\ \text{S, } \sum\limits_{1}^{\Sigma} \left[1 + \langle Sv_n, v_n \rangle_2\right] &\geq \sum\limits_{1}^{J} \left[1 + \theta_n\right] \text{ for all } J \leq M \text{ and any fixed c.o.n. set} \\ \{v_n, n \leq M\}. & \text{The expression of } C_W(P) \text{ in (a), and the unique covariance of the} \\ \text{maximizing Gaussian } \mu_A(\chi), \text{ both now follow from Lemma 1 and Lemma 2. The same} \end{array}$

maximizing Gaussian $\mu_{A(X)}$, both now follow from Lemma 1 and Lemma 2. The same result holds if dim(H₂) = L < ∞ and dim [supp($\mu_{A(X)}$)] < M < L, since in this case S again has M smallest eigenvalues.

(b) If S is strictly positive, then S being compact implies that S does not have M smallest eigenvalues. However, given any $\varepsilon > 0$, one can find eigenvalues $\gamma_1^{\varepsilon}, \ldots, \gamma_M^{\varepsilon}$ such that $0 < \gamma_i^{\varepsilon} < \varepsilon$ for $i \leq M$. Using this in (3) one obtains

$$\begin{aligned} & [\gamma_{\chi\gamma}] = (1/2) \sum_{n=1}^{M} \log [1 + \tau_{n}] \\ &= (1/2) \sum_{n=1}^{M} \log [1 + \chi_{n}^{2} (1 + \gamma_{n}^{\epsilon})^{-1}] \\ &\geq (1/2) \sum_{n=1}^{M} \log [1 + \chi_{n}^{2} (1 + \epsilon)^{-1}]. \end{aligned}$$

The expression on the right of the inequality is maximized, over all (X_n^2) such that $\begin{array}{l} M \\ \Sigma X_n^2 \leq P \\ 1 \\ n \leq P \end{array}$, by defining $X_n^2 = P/M$, $n \leq M$. Thus, $C_W(P) \geq \frac{1}{2} \sum_{\Sigma}^{M} \log \left[1 + (1 + \varepsilon)^{-1} P/M\right]$ for all $\varepsilon > 0$, and so $C_W(P) \geq \frac{1}{2} \sum_{\Gamma}^{M} \log \left[1 + P/M\right]$. For the reverse inequality, one notes that under the constraint $E_{\mu_X} ||R_N^{-1/2} A(X)||_2^2 \leq P$, it is shown in [1] that $C_N(P) = (M/2) \log (1 + P/M)$. For $S \geq 0$, $||R_N^{-1/2} A(X)||_2^2 \leq ||R_W^{-1/2}||A(X)||_2^2$, so that the solution for $C_N(P)$ is the supremum over a larger set than for $C_W(P)$;

i.e. $C_N(P) \ge C_W(P)$. Thus $C_W(P) \le (M/2) \log [1 + P/M]$, so that $C_W(P) = (M/2) \log [1 + P/M]$.

If $S \ge 0$, with zero an eigenvalue of multiplicity K, the above argument is modified in an obvious way $(\gamma_i^{\epsilon} = 0 \text{ for } i = 1, ..., \min (K,M))$ to again obtain $C_{W}(P) = (M/2) \log [1 + P/M].$

To prove (b-ii), the proof of (a) is repeated after substituting λ_i for θ_i , i \leq M.

Now suppose that S has K < M strictly negative eigenvalues $\lambda_1 \leq \cdots \leq \lambda_K$, and that $\sum_{i=1}^{K} \sum_{i=1}^{K} P \geq K_{\lambda_K} \cdot C_W(P) = \sup_{\substack{(P_1, \underline{v}) \\ P_1, \underline{v}}} C_W(P_1, \underline{v})$ where

$$\begin{split} & C_{W}(P_{1},\underline{v}) = \sup \left[\frac{1}{2} \prod_{n=1}^{M} \log \left[1 + X_{n}^{2} \left(1 + \langle Sv_{n}, v_{n} \rangle_{2} \right)^{-1} \right] \right], \\ & \underline{v} = \{v_{n}, n \leq M\} \text{ is any o.n. set, } 0 \leq P_{1} \leq P, \text{ and the supremum is over all } (X_{n}^{2}) \\ & \text{such that } \sum_{1}^{K} X_{n}^{2} \leq P_{1}, \sum_{1}^{K} X_{n}^{2} \leq P. \quad \text{Bepeating the analysis of (a) and (b-i), one finds that} \\ & C_{W}(P_{1},\underline{v}) = (1/2) \sum_{n=1}^{J} \log \left[\frac{1}{2} \frac{1}{\lambda_{1}} + P_{1} = J}{J(1 + \lambda_{n})} \right] \end{split}$$

+
$$(\frac{1}{2})(M-K)$$
 log [1 + $\frac{P-P_1}{M-K}$]

where $J \leq K$ is the largest integer such that $\sum_{i=1}^{J} \lambda_i + P_1 \geq J\lambda_J$. Since this result holds for any o.n. set $\{v_n, n \leq M\}$, it remains only to determine the value of P_1 that maximizes $C_W(P_1, \underline{v})$ (a differentiable function of P_1 in [0,P]). Differentiating, one sees that $C_W(P_1, \underline{v})$ is increasing with P_1 so long as $P_1 < [JP - (M-K) \sum_{i=1}^{J} \lambda_i] (M-K + J)^{-1}$. Since $P_1 < J\lambda_j - \sum_{i=1}^{J} \lambda_i$, this inequality is satisfied

if $(M-K + J)\lambda_J - \sum_{i}^{K}\lambda_i < P$, which always holds because $P + \sum_{i}^{K}\lambda_i \ge K\lambda_K$ and

 $\lambda_{J} \leq \lambda_{K} < 0.$ It thus follows that $C_{W}(P_{1}, \underline{v})$ is an increasing function of P_{1} for $P_{1} < -\sum_{\substack{K \\ 1 \\ i}}^{K} + K\lambda_{K}$. Assumming that $P_{1} \geq -\sum_{\substack{\Sigma\lambda \\ 1 \\ i}}^{K} + K\lambda_{K}$, the maximum of $C_{W}(P_{1}, \underline{v})$ is attained uniquely by $P_{1} = [KP - (M-K)\sum_{\substack{\Sigma\lambda \\ 1 \\ i}}^{K}]^{-1}$. Using the value of P_{1} in the expression for $C_{W}(P_{1}, \underline{v})$, one obtains $C_{W}(P)$ as in (b-iii). The statement on attaining capacity follows from the results of (b-i) and (b-ii).

To prove (c) for (b-ii) and (b-iii), it suffices to note that one can obtain a capacity of $C_W(P) = (M/2)\log(1 + P/M)$ by specifying that $\{u_n, n \le M\}$ be orthogonal to the subspace spanned by $\{Ue_n, n \le 1\}$, where $\{e_n, n \ge 1\}$ are o.n. eigenvectors of S correpsonding to all the strictly negative eigenvalues (λ_n) ; to see this, one can apply (b-i). The statement on (a) for $0 \le \theta_1 < \theta_M$ is clear. If $\sum_{n=1}^{\infty} \sum_{i=1}^{\infty} 0$,

then choosing $\tau_n = P/M$ for $n \le M$ gives $\sum_{l=1}^{M} n(l + \theta_n) \le P$; thus the constraint (A-2) is satisfied. Together with any choice of an o.n. set $\{u_n, n \le M\}$, this definition of $R_{A(X)}$ as in (2) with Gaussian $\mu_{A(X)}$ gives $I[\mu_{XY}] = (M/2)\log[1 + P/M]$. Since the maximizing Gaussian measure is unique, and is given as in (a), $C_W(P) > (M/2) \log [1 + P/M] = C_N(P)$. If $P + \sum_{l=1}^{K} \theta_l \le 0$, then $C_W(P) > P/2 > C_N(P)$ follows by $\log x^{-1} \ge 1 - x$.

Discussion

In the matched channel, $C_N(P) = (M/2) \log (1 + P/M)$ if H_2 is M-dimensional, or if H_2 is infinite dimensional with dim[support $(\mu_A(\chi))] \leq M$ [1, Theorem 1]; moreover, the capacity can be attained in each case by choosing $\mu_A(\chi)$ to be Gaussian with covariance (2), with $(u_{1,...,u_M})$ any M o.n. vectors in H_2 , and with $\tau_n = P/M$ for $n \leq M$.

Thus, it is seen that the mismatched channel differs in several ways. First, the value of the capacity is different unless $S \ge 0$. Secondly, the problem of attaining capacity is much more significant. Even in the finite-dimensional channel the vectors $u_{1,...,} u_{M}$ must be a specific set of vectors, not just any o.n. set. If H_2 is infinite-dimensional with dim $[supp(\mu_{A(X)})] \le M$, the situation is even worse except for (b-ii). That is, capacity can be attained only if S has zero as an eigenvalue of multiplicity $\ge M$ if $S \le 0$, or of multiplicity $\ge M$ -K if K has K < M strictly negative eigenvalues $\lambda_1 \le ... \le \lambda_K$ and $P + \sum_{i=1}^{K} \ge K\lambda_K$. Otherwise, in order to approach capacity, one will need to put part of the available "energy" P in elements (U_n) where (e_n) are eigenvectors of S corresponding to successively smaller eigenvalues. In practical applications, this usually corresponds to eigenfunctions at higher and higher frequencies.

Thus, we conclude that not only does the capacity $C_W(P)$ have a different value than $C_N(P)$ (except when S ≥ 0), but attaining or approaching capacity is significantly more difficult in the mismatched channel than in the matched channel.

It may be noted that the results given in (a) and (b-ii) of the Theorem are similar to those obtained in [4, p. 170], although the developments are quite different. However, these previous results are given in terms of a constraint on $E||A(S)||_2^2$, and assume that the noise variance components can be arranged in ascending order. This can only be done if the channel is finite-dimensional. In this case, one can take $R_W = I$, the identity, and thereby use a true power constraint. The assumption (A-1) becomes $R_N = I + S$, and the capacity is as given in (a); this agrees with the referenced results in [4].

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