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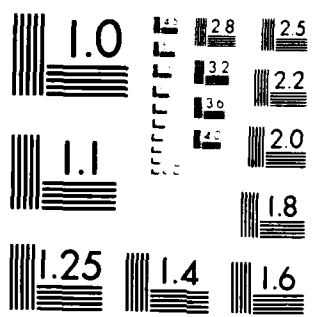
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EVALUATION OF AN OBSERVATION-BASED CLIMATOLOGY MODEL FOR PREDICTING VISIBILITY FOR DATA-VOID LOCATIONS IN GERMANY

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We should like to acknowledge the contribution of Jon Dittmer in the accomplishment of the results in this paper. Jon was responsible for writing several computer programs and made all of the numerous runs required for this report.

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Evaluation of an Observation-based Climatology Model for Predicting
Visibility for Data-void Locations in Germany

by

S. J. Bean and P. N. Somerville

University of Central Florida

1.0 INTRODUCTION

1.1 Review of Problem

A goal of the Air Weather Service has been to achieve a capability to determine the climatic probability of above-threshold conditions of the weather relative to the success of an Air Force flight mission, anywhere, at any time, expeditiously. Such a capability would materially heighten the effectiveness of a weapon system, since it is well-known that the environment can both degrade and enhance system effectiveness.

One method of summarizing or compacting the huge volume of historical records is by means of empirical cumulative distribution functions (cdf's). An empirical cumulative distribution function is simply the tabulated cumulative relative frequencies, or probabilities that a given variable will fall below specified values. Somerville and Bean (1979) have demonstrated that a number of climatological variables may be modeled with closed form distribution functions. For a given location and time (e.g., month and hour) the historical observations can be used to estimate specific model parameters.

Somerville and Bean (1981) used the Weibull distribution to model visibility in Germany for 30 stations. The cumulative distribution function for the Weibull is given by

$$F(x) = 1 - e^{-\alpha x^R}$$

where α and β are constants. Values of α and β are derived for each station, for each month and each of the eight 3-hour periods of the day. The probability of visibility less than x miles is then obtained by substitution in $F(x)$.

1.2 Methods of Extending Visibility Probabilities to Data-Void Regions

A more difficult problem is to develop models which can be used to estimate probabilities for locations where records presently do not exist. Somerville and Bean (1981) developed two models for Germany. Thirty stations for which visibility records were available were used in the development of the models. For each of the stations, for a specified month and hour period, the Weibull distribution was used to model visibility. That is, a value for each of α and β was obtained. Having obtained the values for α and β , these values were regressed on a set of variables which were thought to have a possible influence on visibility. These included elevation, elevation relative to the average elevation on a circle whose radius is 20 kilometers from the station, east-west and north-south elevation difference, population density, relative humidity, proximity to a major body of water, mean wind speed, mean precipitation, latitude, longitude, and functions of and interactions between the above. A stepwise regression program was used to select which of the variables could be used as predictors for a specified month and hour period. Finally, a specially designed least squares non-linear regression program was used to simultaneously determine the regression coefficient in the formulas for α and β . This model was named the "variables model." The coefficients and the regression models are given in the above referenced paper.

A second model, named the "constants model" was also developed. Here only month and time of day were considered. That is, no information regarding elevation, humidity, wind speed, etc., was used. For each month, and time of day, non-linear regression was used to determine the "best" values for α and β . These were also tabulated in Somerville and Bean (1981).

In either model, extension of the climatology was accomplished by obtaining values for α and β at the data-void station, and then using the Weibull distribution to determine the desired visibility probabilities.

In this manuscript we will restrict ourselves to the evaluation of the constants model. Evaluation of the variables model will be given in a future report.

2.0 EVALUATION OF THE CONSTANTS MODEL

Two methods were used to evaluate the model. First, 30 stations were used as a "calibration" set, and the resulting constants model was used on a second independent "evaluation" data set of 30 West German stations. Second, sample re-use was used to evaluate the constants model. From the 60 evaluations one "overall" RMS (Root Mean Square) of the method was obtained. Sample re-use (sometimes called cross-validation) is a relatively new technique which makes it possible to use the same set of data for "calibration" and "evaluation". Briefly, if there is a total of n stations, n separate solutions are obtained. For each solution, one station is used as the evaluation set with the remaining $n-1$ used as the calibration set. Using the method one can obtain the Root Mean Square (RMS) error of the modeling procedure for each station. The individual RMS errors may then be combined to

obtain an overall RMS error for the procedure. For future use the recommended model is the one using all the stations for the calibration set. For a good account of the sample re-use technique, papers by Stone (1974) and Geisser (1975) are recommended.

The first evaluation method gave some encouraging results, and it also indicated some ways we may improve the model. Exhibit 2.1 gives the RMS of the probability estimates for each station taken over all months and hour periods with the exception of hour periods 1 and 2. The data for these early pre-dawn hours are frequently missing, and they were, therefore, eliminated from the study. Stations 4, 13, and 26 stand out as very poor fits which inflate the overall RMS considerably. These three stations are much higher in elevation than the other stations in the study. Also, they are much higher than the surrounding area. These factors seem to give rise to much different visibility conditions than the other areas in the study.

Exhibit 2.2 gives the resulting RMS values averaged over all stations in the evaluation set for each month and hour periods 3 through 8. The overall RMS of .108 compares favorably with the calibration data set overall RMS of .063. The model obviously does not fit as well on the evaluation set, and we certainly could not expect that it would.

Station		RMS
1	AACHEN, DL	.074
2	BREMEN, GER	.055
3	LINGEN, GER	.050
4	KAHLER ASTEN, GER	.241
5	AIGEN ENNSTAL, GER	.109
6	PLEZEN / DOBRA, CZ	.110
7	BREMGARTEN, GER	.052
8	OBERSTDORF, GER	.107
9	KONSTANZ, GER	.054
10	INNSBRUCK, OS	.092
11	SALZBURG, OS	.071
12	PASSAU, GER	.055
13	FELDBERG, GER	.298
14	NEUHAUSEN, DL	.075
15	NURBURG, GER	.076
16	KOBLENZ, DL	.083
17	GIESSEN, GER	.085
18	HERSFIELD, DL	.068
19	KISSINLEN, DL	.063
20	COBURG, GER	.053
21	HOF, GER	.055
22	BERUS, GER	.061
23	KARLSRUHE, GER	.063
24	OHRINGEN, GER	.057
25	STAUBING, GER	.100
26	GROSSER FALK, GER	.249
27	LAHR, GER	.057
28	LAUDHEIM, GER	.054
29	KAUFBEUREN, DL	.066
30	MUHLIDORF, GER	.051

Exhibit 2.1

Overall RMS For Each of The 30 Stations
in The Evaluation Set

	06-08	09-11	12-14	15-17	18-20	21-23	all
Jan	.137	.130	.130	.134	.134	.149	.136
Feb	.136	.126	.123	.125	.124	.138	.129
Mar	.124	.119	.107	.105	.113	.127	.116
Apr	.124	.112	.096	.092	.099	.104	.105
May	.114	.097	.079	.067	.076	.092	.089
Jun	.112	.092	.071	.058	.070	.084	.083
Jul	.106	.083	.063	.051	.056	.075	.075
Aug	.099	.082	.064	.058	.068	.077	.076
Sep	.115	.098	.074	.067	.085	.102	.092
Oct	.113	.111	.097	.095	.101	.121	.107
Nov	.136	.132	.130	.131	.136	.146	.135
Dec	.131	.124	.124	.127	.128	.145	.130
all	.121	.110	.100	.097	.103	.116	.108

Exhibit 2.2
RMS Over All Stations in the Evaluation Set
By Month and Hour Period (LST)

The results for the sample re-use evaluation were comparable to those for the first study. The RMS values for each of the 60 stations averaged over all months and hour periods 3 through 8 are given in Exhibit 2.3. The RMS values over all 60 stations for all months and hour periods 3 through 8 are given in Exhibit 2.4. The RMS values corresponding to sample re-use (Exhibit 2.4) are generally smaller than the results using the second 30 stations for evaluation (Exhibit 2.2). This is mainly due to the larger sample size in the sample re-use evaluation. The sample re-use evaluation makes use of 59 stations to build a model, whereas the first procedure makes use of only 30 stations in the model estimation.

Exhibit 2.5 gives the values of α and β for the constants model for each month and hour period where all 60 stations are used for the calibration set.

Exhibit 2.6 shows the location in Germany which were used in developing the models evaluated in this report.

IMO	Station	Lat	Long	RMS
10501	AACHEN, DL	50.78	-6.12	.091
10224	KREMEN, GER	53.05	-8.80	.040
10305	LINDEN, GER	52.52	-7.33	.127
10427	KAHLER ASTEN, GER	51.18	-8.50	.274
11157	AIDEN ENNSTADT, OS	47.53	-14.15	.101
11446	FLEZEN/DORRA, CZ	49.67	-13.30	.145
10900	OKFEGARTEN, GER	47.90	-7.63	.050
10948	UBERSTADT, GER	47.40	-10.30	.107
10929	KONSTANZ, GER	47.68	-9.20	.051
11120	LINDBROUN, OS	47.27	-11.37	.081
11150	SALZBURG, OS	47.80	-13.02	.059
10893	PASSAU, GER	48.58	-13.50	.066
10908	FELDFERS, GER	47.87	-8.02	.101
10971	NEUHAUSEN, DL	47.99	-8.92	.108
10510	MURBURG, GER	50.33	-6.97	.134
10515	MULENZ, DL	50.35	-7.60	.102
10530	GIESSEN, GER	50.57	-8.72	.076
10542	HENSFELD, DL	50.87	-9.72	.135
10658	NISSNEN, DL	50.20	-10.10	.126
10671	COBURG, GER	50.27	-10.97	.123
10685	HOF, GER	50.32	-11.90	.067
10704	BERG, GER	49.27	-6.70	.120
10717	MARK SAUHE, GER	49.02	-8.40	.077
10742	ORRINGEN, GER	49.20	-9.53	.092
10788	STACHING, GER	48.87	-12.60	.112
10791	GROSSER FALK, GER	49.08	-13.30	.075
10805	LAHR, GER	48.37	-7.85	.061
10877	LADDEIM, GER	48.22	-9.93	.094
10953	KAUFHUFEN, DL	47.87	-10.63	.065
10875	MUHLDOFF, GER	48.25	-12.55	.091
10616	HAHN AB	49.95	-7.27	.048
10610	BILBURG AB	49.95	-6.57	.056
10614	KANSTEIN AB	49.43	-7.58	.064
10607	SEANGAHEIM AB	49.97	-6.70	.067
10384	TEMPELHOFF ARPT	52.47	-13.40	.076
10755	ANSBACH AAF	49.32	-10.63	.071
10544	FULDA AAF	50.53	-9.63	.062
10869	ERDING AS	48.32	-11.93	.060
10765	FEUCHT AAF	49.38	-11.18	.055
10618	BAUMHULDER AAF	49.65	-7.30	.126
10626	BAD KREUZNACH AAF	49.87	-7.9	.111
10971	BAD TOLZ AAF	47.77	-11.67	.123
10714	ZWEIBRUCKEN AB	49.22	-7.40	.147
10633	WIESBADEN AB	50.05	-8.33	.050
10643	FINTHEN AAF	49.97	-8.15	.057
10763	LORCH AAF	49.50	-10.95	.090
10642	HANAU AAF	50.17	-8.95	.047
10852	SARLIMEN AAF	48.45	-10.67	.109
10653	GIEFELSTADT AUX AF	49.67	-9.88	.080
10687	GEAUFENWOHR AAF	49.70	-11.95	.073
10734	HEIDELBERG AAF	49.40	-8.65	.041
10752	ILLESHEIM AAF	49.47	-10.38	.149
10659	NITZINGEN AAF	49.75	-10.20	.060
10763	MURNBERG	49.50	-11.08	.076
10729	SOLEMAN AAF	49.57	-8.47	.114
10657	WERTHEIM AAF	49.77	-9.48	.104
10745	SCHWAERISCH HALL AA	49.17	-9.78	.077
10712	SEHRACH AB	49.52	-7.87	.037
10862	SIEGENBERG GUNNERY	48.77	-11.80	.177
10738	ECHTERDINGEN ARPT	48.68	-9.22	.057

Exhibit 2.3

Overall RMS For Each of the Stations

(Sample Re-use)

	06-08	09-11	12-14	15-17	18-20	21-23	all
Jan	.126	.119	.114	.113	.124	.166	.128
Feb	.117	.105	.099	.099	.111	.152	.115
Mar	.103	.097	.087	.083	.094	.116	.097
Apr	.101	.089	.074	.072	.080	.089	.085
May	.093	.076	.060	.051	.060	.077	.071
Jun	.094	.073	.054	.045	.056	.070	.067
Jul	.088	.066	.049	.039	.047	.065	.061
Aug	.093	.070	.051	.046	.055	.080	.068
Sep	.106	.085	.064	.053	.074	.097	.082
Oct	.100	.097	.082	.078	.100	.134	.100
Nov	.112	.108	.104	.106	.121	.151	.118
Dec	.113	.107	.106	.108	.115	.167	.121
all	.104	.092	.082	.079	.090	.119	.096

Exhibit 2.4

RMS Over All Stations by Month and Hour Period

(Sample Re-use)

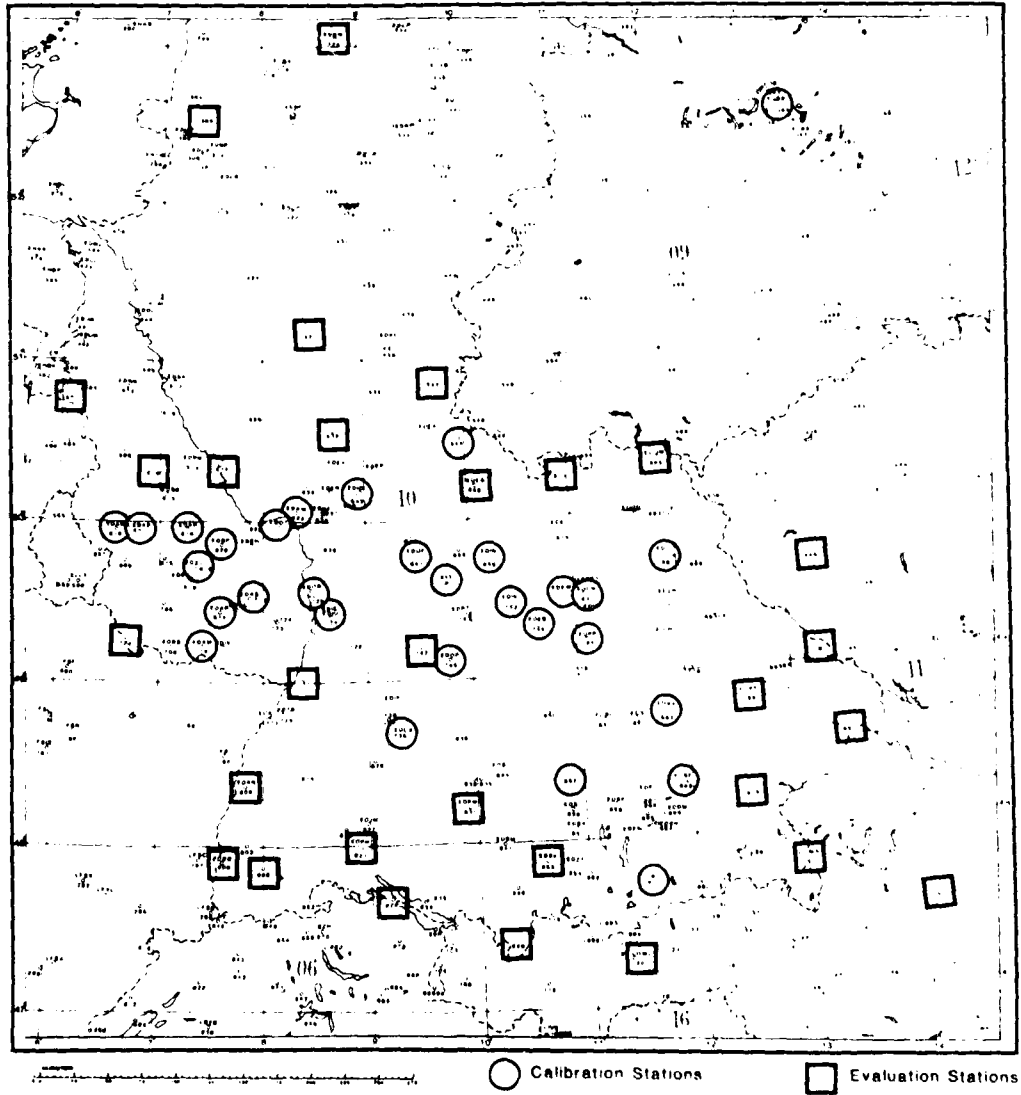
	06-08	09-11	12-14	15-17	18-20	21-23
Jan alpha	.248	.253	.182	.172	.190	.217
beta	.794	.799	.895	.891	.909	.866
Feb alpha	.258	.239	.140	.123	.141	.162
beta	.784	.834	.957	.943	.946	.887
Mar alpha	.185	.131	.066	.056	.068	.072
beta	.855	.954	1.078	1.036	1.043	1.089
Apr alpha	.128	.067	.041	.039	.044	.045
beta	.967	1.082	1.000	.901	.945	1.084
May alpha	.100	.039	.029	.022	.029	.036
beta	.953	1.134	.934	.939	.914	1.049
Jun alpha	.093	.032	.021	.017	.022	.026
beta	1.038	1.251	1.122	1.127	1.104	1.290
Jul alpha	.079	.027	.016	.013	.016	.019
beta	1.147	1.338	1.201	1.208	1.239	1.414
Aug alpha	.137	.039	.016	.014	.018	.027
beta	.952	1.341	1.408	1.342	1.352	1.370
Sep alpha	.261	.091	.024	.018	.027	.047
beta	.698	1.052	1.416	1.394	1.377	1.304
Oct alpha	.351	.208	.078	.064	.095	.154
beta	.561	.740	1.045	1.077	1.028	.849
Nov alpha	.254	.223	.133	.129	.145	.178
beta	.661	.727	.859	.846	.861	.761
Dec alpha	.221	.225	.163	.169	.183	.210
beta	.784	.798	.894	.857	.865	.808

Exhibit 2.5

Values of α and β for the Constants Model

by Month and Hour

Exhibit 2.6 West German Stations



V.D. CONCLUSIONS AND RECOMMENDATIONS

The constants model has been shown to give good results for estimating visibility probabilities with the exception of some higher elevation areas. Because the variables model makes use of a number of characteristics of the data-void location, including that of elevation, it is expected to improve the fit over the constants model.

One problem with the variables model, however, is that many of the input variables may be as hard to obtain as information about visibility. Two other models should be investigated to see if it is possible to improve on the constants model and at the same time not require too much information: one is a model that uses only topographical variables such as elevation and average elevation of the surrounding area; another is a model based on cluster analysis. That is, the known (v,s) parameters might be used to determine regions of homogeneous visibility characteristics. Constants models could then be used on the individual regions and, as before, sample re-use could be used to evaluate the results.

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