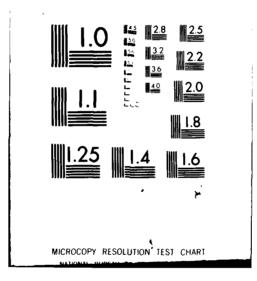
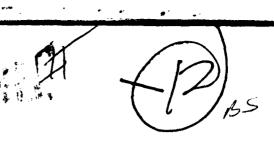
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# Final Report: COMPUTERIZED ADAPTIVE ABILITY TESTING

David J. Weiss

**April 1981** 



COMPUTERIZED ADAPTIVE TESTING LABORATORY PSYCHOMETRIC METHODS PROGRAM DEPARTMENT OF PSYCHOLOGY UNIVERSITY OF MINNESOTA MINNEAPOLIS MN 55455

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(3) investigation of the utility of a number of person-fit indices designed to identify lack of fit of individuals to item response theory models; and (4) investigation of the potential of several cognitive information-processing types of tasks for computerized adaptive administration. Research approaches and preliminary results are summarized for each objective. Additional research plans currently being implemented under Project NR 150-433 with which this project was combined, are described.

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# FINAL REPORT: COMPUTERIZED ADAPTIVE ABILITY TESTING

This research program was designed to study four areas relevant to adaptive ability testing:

- 1. Evaluation of adaptive testing branching strategies.
- 2. Use of different item formats and response modes.
- 3. Fit of individuals to item response theory models.
- 4. New types of ability tests designed specifically for computerized adaptive administration.

Research in pursuance of these objectives, originally scheduled for a three-year period, began on April 1, 1979, and continued through June 30, 1980, at which time the research objectives of this project were combined with those of Project NR 150-433, "Computerized Adaptive Achievement Testing."

This report summarizes the progress made during this 15-month period. No technical reports were completed during this period; technical reports begun in this project will be completed under Project NR 150-433. For each of the four objectives listed above, this report (1) describes the objective, (2) details the approaches used to study the objective, (3) summarizes results that were available at the completion of the reporting period, and (4) describes tentative plans for further research on the objective to be continued in Project NR 150-433.

#### Adaptive Testing Strategies

Previous simulation studies using adaptive testing have used relatively unrealistic item pools (e.g., Gorman, 1980; McBride, 1976; Reckase, 1976; Urry, 1970; Urry, 1971; Vale, 1975). These item pools have been unrealistic because they assumed that the item parameters describing the items in the pool were completely error free, as well as assuming item difficulty and distribution characteristics that did not reflect those of real ability tests. Previous studies have also been unrealistic in that they have assumed that the responses of the hypothetical testees to these items have conformed precisely to the one-dimensional latent trait model. When the results of previous simulation studies are extrapolated to real item pools constructed from real item parameters, they may not generalize, because real item pools are constructed from item parameters that include estimation error and may deviate substantially from unidimensionality.

#### Objective

The objectives of this research program were to evaluate the performance of adaptive testing strategies under conditions that more reasonably represent the conditions under which these strategies might occur in live-testing applications and to compare findings from selected simulation studies to those obtained in live testing. Research during the reporting period was concerned with (1) effects of errors in item parameters on the performance of adaptive testing strategies and (2) live-testing comparisons of adaptive testing strategies.

### Effects of Errors in Item Parameter Estimates on Adaptive Testing Strategies

Approach. This objective was pursued by means of monte carlo simulation studies that built upon empirical information regarding the nature and extent of errors in item parameter estimates due to the numbers of testees and items on which the item response theory (IRT) item parameters were estimated. Data on the kinds and degrees of error associated with IRT item parameterization techniques by different item parameterization methods were modeled in the monte carlo simulations. The kinds and degrees of errors observed in real data item parameterization were translated into the monte carlo simulation model and served as independent variables in a series of studies systematically varying the magnitude and kind of item parameterization estimation error for the difficulty, discrimination, and "guessing" parameters separately and in combination. Dependent variables in these studies were test information, bias, correlation of ability estimates with true ability, and other characteristics of the ability estimates derived from the application of selected adaptive testing strategies; two conventional tests were also included in the study for comparison purposes. The studies were also designed to use an item pool that realistically reflected the composition of real item pools used in actual ability tests, in terms of the distributions of the IRT item parameter estimates.

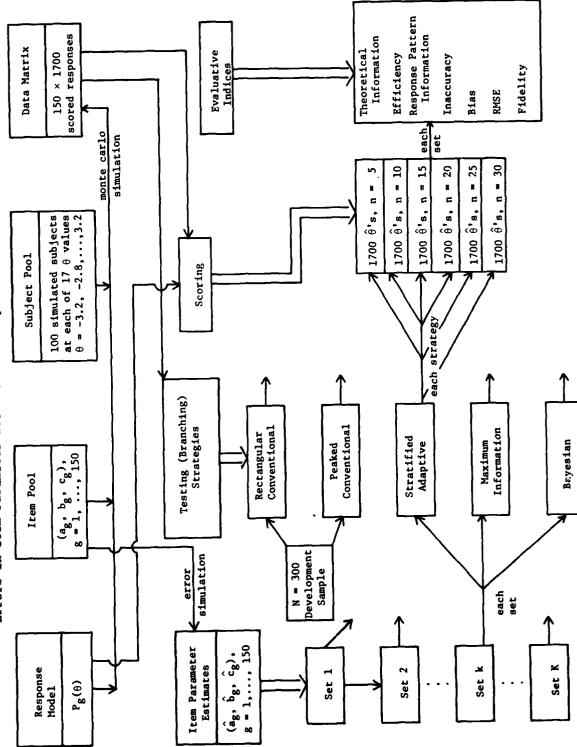
Figure 1 summarizes the design of this study. Using a three-parameter IRT model and an item pool designed to reflect an adaptive testing item pool that had been used in a live-testing study, monte carlo data were generated for 100 simulees at each of 17 levels of ability, ranging from  $\theta = -3.2$  to  $\theta = +3.2$ . Based on data available in this IRT item parameterization literature, varying degrees of error were added to the parameter estimates for item discrimination (a), difficulty (b), and "guessing" (c). Table 1 shows the item parameter sets used in this study: Set 1 was the baseline comparison data set in which there was no error in the item parameter estimates; in Sets 11, 12, and 13 varying amounts of error were added to the a parameter; in Sets 21, 22, 23, and 24 error was added to the b parameter; Sets 31 and 32 added errors to the c parameter; and in Sets 41 and 42 errors occurred in all three parameters simultaneously.

Using the error-laden item parameter sets, three types of adaptive tests (stratified adaptive, or stradaptive; maximum information; and Bayesian) were administered to each of the 1,700 simulees. All tests were scored by maximum likelihood at test lengths of 5 to 30 items, in increments of 5 items. In addition, both peaked and rectangular conventional tests were constructed using classical test construction procedures; and these tests, along with the adaptive tests using the error-free item pool, were also administered to the 1,700 simulees. Testing strategies were compared in terms of fidelity (the correlation of true and estimated  $\theta$  levels), observed and theoretical information, efficiency, inaccuracy, bias, and root mean square error (RMSE) for the  $\theta$  estimates.

<u>Results</u>. Table 2 presents a selection of the results for four of the dependent variables. The fidelity measure was computed on a normally distributed sample of 300 simulees; data for the other criterion measures were averaged across the 1,700 simulees. As Table 2 shows, with the exception of Item Set 42,

- 2 -

Figure 1 Design of Monte Carlo Simulation Study of Effects of Errors in Item Parameter Estimates on Adaptive Testing Strategies



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Iten		Spec	cified	RMSE	Obta	ined	RMSE	Obta	ined r	(p, ĵ)
Set	Description	a	<u>b</u>	<u>c</u>	<u>a</u>	b	<u>c</u>	<u>a</u>	<u>b</u>	<u>c</u>
1	Error-Free			-						
	Item Set	.00	.00	.00	.00	.00	.00	1.00	1.00	1.00
11	Small Error									
	in a	.20	.00	.00	.22	.00	.00	.75	1.00	1.00
12	Moderate Error									
	in a	.40	.00	.00	.39	.00	.00	.56	1.00	1.00
13	Large Error									
	in a	.60	.00	.00	.52	.00	.00	.37	1.00	1.00
21	Moderate Error									
	in b	.00	.10	.00	.00	.09	.00	1.00	.99	1.00
22	Large Error									
	in <u>b</u>	.00	.30	.00	.00	.30	.00	1.00	.98	1.00
23	Extreme Error									
	in <u>b</u>	.00	1.00	.00	.00	.89	.00	1.00	-88	1.00
24	Very Large									
	Error in <u>b</u>	.00	.50	.00	.00	.48	.00	1.00	.96	1.00
31	Moderate Error									
	in <u>c</u>	.00	.00	.04	.00	•00	.04	1.00	1.00	.70
32	Large Error									
	in <u>c</u>	.00	.00	•08	.00	•00	•08	1.00	1.00	.46
41	Worst Probable									
	Combined Error	.60	.30	.08	.51	.32	.08	.47	.98	.46
42	Extreme									
	<b>Combined Error</b>	.60	.00	.08	•28	.97	•08	.44	.88	.43

 Table 1

 Error Simulated in Item Parameter Estimate Sets

which represented extreme (and probably unrealistic) levels of error in all three item parameters, the adaptive tests with error-laden item parameters achieved higher fidelities at all test lengths than did the peaked (P) and rectangular (R) conventional tests, with larger differences occurring for shorter test lengths. There were virtually no differences in fidelities for the adaptive strategies at 20- or 30-item test lengths, with a tendency for the maximum information (MI) adaptive test to perform somewhat more poorly than the stratified adaptive (SA) or Bayesian (B) tests at 10-item test lengths. Results for the other dependent measures tended to support the fidelity analysis; that is, with the exception of Item Set 42, adaptive tests using error-laden item parameter estimates generally achieved scores with lower levels of inaccuracy, bias, and RMSE than did conventional tests of the same lengths using error-free item estimates.

Analyses of the data in terms of dependent measures conditioned on values of  $\theta$ --inaccuracy, bias, RMSE, the two information measures, and efficiency--supported the findings from the overall analysis. When errors occurred in the <u>a</u>, <u>b</u>, and <u>c</u> parameters separately, there was very little effect on these indices and the adaptive tests measured better than the conventional tests at virtually all levels of  $\theta$ . There was essentially no measurement degradation as the result of errors in <u>c</u> and <u>a</u>, with a slightly greater effect for <u>b</u>. For realistic val-

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			Conve	ent for M	Conventional and Maximum 1		Peaked Conv Information		. P	(P) Bay		and f (B) Ad	l for Stra Adaptive	Stradaptive ive Tests		(3),				ł
Measure			Iten	Item Set	13:	Ites	n Set	22:	Item	Set 2	23:	Item	Set	32:	Item	Set	41:	Item	Set	42:
hue			RMSE	<b>1</b>	.52	RMSI			RMSE	# ct	00.	RMSE		00.	RMSE	H rcj	.51	RMSE	60	.58
Number			RMSE	<b>ا</b> مرز	00.	RMSI		.30	RMSE	ہ م;	.89	RMSE	# م.	00.	RMSE	ء م	.32	RMSE	<b>ء</b>	.97
of			RMSE	ו ט י היו	0.	RMSE	။ (ပ) (ည	8.	RMSE	။  ပ	0.00	RMSE	။  ပ	•08	RMSE	• । अ	•08	RMSE	• اب	8
Items	R	Ъ.	SA	IJ	8	SA		m	SA	Ħ	m	SA	IW	m	SA	IJ	8	SA		m
Fidelity																	:	ł	Î	ä
10	.83	.73	.88	-84	.90	16.	.90	.93			.87	•93	-90	.93			- 89	.82	.78	81
20	.92	.89	-94	-95	.95	.96	.96	.96	-92	. 93	.93	.96	.96	.96	.94	<b>.</b> 95	.95	.89	.89	68.
30	-94	.93	.96	.97	.97	.97	.97	.97			.95	-97	.97	.97			•96	.93	.92	.93
Inaccuracy	•													ļ		!			c t	2
10	.53	.70	-40	-44	.40	.40	.42	•36			.48	• 35	.38	.37		4/	• • •		ø.,	0
20	.38	.48	.29	.28	.27	.29	.27	-26	.41	.46	.39	.26	.24	.28	.31	000	.32	- 62	••	4 ·
30	.33	.41	.24	.24	.24	.25	.23	.23			.38	.21	.20	.22		.25	.25		÷2•	.43
Bias																				
		18	•		07	05		03	I						1					01. -
		15	8.	64	01	02	5	03	.12	.10	60	02 -	- 07 -		- 02 -	- 6	- 80- -	- 10	- 13	01
	.02	10			.02	<u>8</u>		00.						03						01.
RMSE													ŗ	Ċ		ŗ	27	10	00	5
10	.67	.98	•55	.64	.56	• 55	•58	.47			-62	.48	/?.	<u>.</u>	<b>•</b> 0•	2	00.	ò ;	0 · 1	
20	.48	-69	.38	.37	.35	.37	.36	.35	.51	.57	.49	• 33	.32	.35	.41	.40	44.	: :	+ · ·	
30	-44	.59	.32	.31	.30	.33	.30	.31			.48	.28	.27	.29	•34	.32	.34	• • •	40.	<u>;</u>

Table 2

Overall Measures at 10-, 20-, and 30-Item Test Lengths for Varying Amounts of Error in a, b, and c, and All Three Simultaneously, for Rectangular (R) Conventional and Peaked Conventional (P) tests, and for Stradaptive (S).

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ues of combined error in the three item parameters, some measurement degradation occurred for the adaptive tests, but they still measured better than the rectangular conventional test at all  $\theta$  levels, and better than the peaked conventional test for about three-fourths of the  $\theta$  scale. There were few consistent differences in the performance of the different adaptive testing strategies.

Additional research in progress. Results of this study indicated very little effect of errors in item parameter estimates on the measurement performance of adaptive testing strategies. Since this study was the first to investigate this question, it was necessarily limited in a number of ways. Consequently, further simulations are planned that (1) vary the characteristics of the item pool used in order to determine the generality of the findings across item pools with different characteristics, in terms of levels of the three item parameters; that (2) allow correlated errors to occur in the item parameter estimates, since only uncorrelated errors were used in this study; and that (3) examine the effects of error in item parameter estimates separately for one-, two-, and threeparameter IRT models.

#### Live-Testing Comparison of Adaptive Testing Strategies

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<u>Approach</u>. Three testing strategies--peaked conventional, Bayesian adaptive, and maximum information adaptive--were compared on the basis of alternate forms reliability and observed information. The tests were composed of 60 fivechoice vocabulary items that were divided into two 30-item alternate forms. The conventional test was peaked in information values evaluated at  $\theta = 0.0$ . Items administered in the maximum information and Bayesian testing strategies were selected according to their adaptive item selection routines. There were 373 students in the conventional testing condition, 390 in the Bayesian testing condition, and 233 in the maximum information testing condition.

Testing strategy was the major independent variable of interest. Methods of scoring were also compared. These included logistic maximum likelihood scoring, Bayesian scoring, and (for the conventional test) proportion-correct scoring. Test length was a third independent variable of interest. Thirty test lengths were obtained by scoring each 30-item test at each test length from 1 to 30 items. Testing strategies were compared on the basis of alternate forms reliability by correlating corresponding ability estimates obtained from Forms A and B for a given testing strategy.

Since the test data were scored in at least two ways (Bayesian and maximum likelihood), a total of seven combinations of testing strategy and scoring method were compared on the basis of alternate forms reliability. Scoring strategy was compared on the basis of alternate forms reliability by comparing reliabilities of a single testing strategy scored by more than one method. Three of the alternate forms reliabilities paired the appropriate scoring method with each of the three testing strategies. These were proportion-correct scoring of conventional tests, maximum likelihood scoring of maximum information tests, and Bayesian scoring of Bayesian-administered tests. The remaining four alternate forms reliabilities were obtained by scoring the item response data by a scoring routine other than the appropriate one. In this way, reliabilities were obtained for the Bayesian scoring of the maximum information test, maximum likelihood scoring of the Bayesia test, Bayesian scoring of the conventional test, and maximum like hood sr \_\_ng of the conventional test. Reliabilities were calculated as a function of test length. Scoring method correlations were obtained by correlating estimates obtained from different scorings of the same testing strategy. These correlations were used to analyze the similarity of ability estimates obtained from different scoring methods applied to a single set of data.

The three testing strategies were also compared on the basis of their errors of measurement. This was assessed in two ways: (1) using estimated errors of measurement derived from maximum likelihood scoring and (2) using estimated errors of measurement from Bayesian scoring. In the first method, test item responses were scored by maximum likelihood methods, and the standard errors of measurement (SEM) associated with each ability estimate was calculated. These values are the reciprocal of the square root of test information at a given  $\hat{\theta}$  level and estimate the standard deviation of the estimated  $\theta$  values around the true  $\theta$  value; the larger the SEM, the more likely the estimate will be inaccurate. The posterior variance of the Bayesian ability estimate was the second index used to compare the testing strategies on the basis of measurement accuracy. Both the SEM and the posterior variances were examined as a function of estimated ability level.

<u>Results</u>. A preliminary report of the results of this study is in Johnson and Weiss (1980). Parallel forms reliabilities of the three testing strategies showed that after 11 items the peaked conventional test yielded higher reliabilities than either of the adaptive tests. The greatest difference between reliabilities was r = .09 between the adaptive and conventional tests at the 30-item test length; the reliabilities of the adaptive tests were r = .81, compared with the final reliability of r = .90 for the conventional test. The conventional test reliability was nearly identical to that of the Bayesian test up to the 10-item test length, but after that point the conventional test reliability increased more quickly than that of the adaptive tests. Although adaptive test reliabilities showed signs of leveling off toward the end of the test, the reliability of the conventional test appeared to increase steadily.

In comparisons of testing strategies scored by other than optimal scoring strategies, the Bayesian scoring of the conventional and maximum information testing strategies yielded higher reliabilities than the maximum likelihood scoring of the conventional and Bayesian testing strategies. These data indicate that Bayesian scoring of an adaptive test may yield more stable estimates of ability than maximum likelihood scoring. The data also illustrate the inappropriateness of scoring conventional tests with maximum likelihood scoring methods, since extremely low reliabilities (maximum of r = .75) were obtained at all test lengths. The correlations between scores on the same testing strategy scored by different methods showed that the highest correlations were obtained for Bayesian and proportion-correct scores of the conventional test, with most correlations between .97 and .99. The second highest level of correlation was between the Bayesian and maximum-likelihood-scored maximum information test, with most correlations between .93 and .95. When the maximum information adaptive test was scored by the Bayesian scoring method, reliabilities of short adaptive tests were higher than those of the conventional test, and differences in reliabilities were smaller at longer test lengths.

On the basis of the reliability data, few conclusions can be drawn about the relative merits of the three testing strategies. Limitations of the item pool might account in part for the lowered reliability of the adaptive tests in comparison to the conventional test, since adaptive tests depend heavily on the quality of the items in the item pool. The item pool used for the two adaptive

tests had fewer items at the extremes of the ability range, and these items had relatively lower discrimination parameters. Especially at abilities where there were fewer items, it is likely that the correlations between ability estimates would be attenuated and that the adaptive process would be at a disadvantage as testing progressed. The result would be that toward the end of testing there would be fewer and fewer items available at a given ability level.

Another factor that limits the comparison of the testing strategies in terms of alternate forms reliability correlations is the distribution of ability in the population. Since values of the Pearson product-moment correlations depend on the distributions of the ability estimates involved, different ability distributions can result in different levels of correlation. Thus, the reliability correlations confound the distribution of the ability estimates with the measurement precision of the testing strategies.

Errors of measurement derived from test information yield comparisons of testing strategies that are unconfounded by the distribution of the ability estimates. Comparisons of the testing strategies on the basis of SEMs and posterior variances showed that at no point on the ability continuum were the errors of measurement smaller in the conventional test than in the adaptive tests. In both error of measurement comparisons there was poorer measurement at the low end of the ability distribution, although the extremes--both positive and negative--were less precisely measured than the center of the ability continuum. The results indicate that the adaptive tests yielded about the same level of measurement precision and that these levels were greater than those obtained from the conventional test at all levels of ability. Thus, adaptive testing strategies yielded scores with greater precision/information (lower errors of measurement) than did the conventional testing strategy.

Additional research in progress. Since the reliability results of this study were contrary to expectations and conflicted with other research using a similar design but different tests (Kingsbury & Weiss, 1980; McBride, 1980), a fourth test was added to the study. To examine the effects of test difficulty on the results, this test was a second conventional test in which average item difficulty was higher than that of the first conventional test. Data were collected using this test from 530 students on a 60-item conventional test consisting of two embedded 30-item alternate forms. The alternate forms reliabilities of these tests will be computed at test lengths from 1 to 30 items, and the data will be further analyzed to permit direct comparisons with the three other testing strategies.

#### Future Research Plans

In addition to using item parameters that contain varying degrees of error, real adaptive testing item pools may deviate from the unidimensional IRT model that has been applied in all adaptive testing simulations. Since deviations from unidimensionality (e.g., Bejar, 1979, 1980; Bejar, Weiss, & Kingsbury, 1977; Reckase, 1978) can potentially affect the performance of adaptive testing strategies, a series of monte carlo simulation studies will be constructed around the degrees and types of dimensionality observed in ability test data.

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These studies will consist of the generation of testee responses using the underlying multidimensional structures observed in ability test items, but adaptive branching will occur by means of several adaptive testing strategies designed for unidimensional adaptive testing. Thus, the research question will be the effects of violation of the unidimensionality assumption on the performance of adaptive testing strategies. Again, the evaluative criteria will consist of information, bias, correlation of true ability and ability estimates, and other characteristics of the ability estimates derived from the unidimensional adaptive testing strategies.

#### Item Formats and Response Modes

The use of interactive computers to administer ability tests allows the design and use of test items that do not make use of the typical multiple-choice item (Bejar, 1975; Vale, 1977) suggests that there is considerable improvement possible in information utilization from adaptive testing by use of response modes other than multiple-choice items. Consequently, continued research in this area was indicated.

#### **Objective**

To evaluate the utility for adaptive testing of a number of response modes and item formats usable in adaptive testing.

#### Approach

This objective was pursued using the six item types shown in Figure 2. The studies were concerned with the following characteristics of these item types:

- 1. The relationship of responding in the various formats to ability levsls.
- 2. The reliability of test item responses and ability estimates obtained in the various formats.
- 3. Information characteristics of items and tests utilizing the various formats.
- 4. The relationships among test item responses using the various formats.
- 5. The relative validity of responses obtained from the various formats.
- 6. The generality of findings obtained from the different response formats to different populations and different ability dimensions.
- 7. The comparative factor structure of tests administered in the various formats.

The research was designed to study the characteristics of the six item formats in several ability areas. It is essentially a search for an item format that allows testees to express as much knowledge as they have available about a given question while minimizing the effects of guessing. The results obtained from this series of studies will be used to select several item formats to be used in computerized adaptive testing.

Two sets of 30 multiple-choice items were chosen from available item pools. One set of items was chosen from a pool of analogy items, and the second set was chosen from a pool of arithmetic reasoning items. Both item pools included item

## Figure 2 Description of Response Formats

1. <u>Multiple-choice items with conventional response format</u>. These items were conventional multiple-choice items with four alternatives. The examinee was asked to choose the correct answer.

Example. Procedure : Activity

- 1. Diplomacy : Tact
- \*2. Itinerary : Journey
- 3. Minutes : Committee
- 4. Index : Book
- 2. <u>Multiple-choice items with probabilistic response format</u>. These items were exactly the same as the conventional multiple-choice items, but the examinees were asked to assign 100 points among the four alternatives to indicate their confidence in the "correctness" of each alternative.

Example.	Procedure : Activity	(Possible Answer)
	1. Diplomacy : Tact	0
	2. Itinerary : Journey	75
	3. Minutes : Committee	25
	4. Index : Book	0

3. Dichotomous items with a yes-no (dichotomous) response format. For these items only the item stem and one alternative were presented. The examinees were asked to respond with a "yes" if they thought the alternative provided was a correct answer to the question, and "no" if they thought it was not.

Example.	Q.	Procedure : Activity	(Possible Answer)
		Index : Book	Yes

4. Dichotomous items with a probabilistic response format. These items were identical to the dichotomous items with a yes-no response format, but, the examinees were asked to respond with a probability (a number from 0 to 100) which reflected their confidence that the alternative provided was a correct answer to the question.

Example.Q.Procedure : Activity(Possible answer)A.Index : Book10

5. Free-response items with a conventional response format. For these items only the item stem was presented. The examinees were asked to provide their own answers.

Example. Q. Procedure : Activity Itinerary :

6. Free-response items with a probabilistic response format. Once again only the item stem was presented, but this time the examinees were asked to provide an answer to the question and to assign a probability (a number from 0 to 100) to the answer they gave to indicate their confidence in the "correctness" of their response.

Example.	Q.	Procedure	:	Activity	(Possible answer)
		Itinerary			Trip 90

parameters calculated on large numbers of individuals using the three-parameter logistic model of LOGIST (Wood & Lord, 1976; Wood, Wingersky, & Lord, 1976). The items were chosen to represent a uniform range of difficulty and discrimination parameters. Each set of 30 items was then modified to conform to the item formats shown in Figure 2. For each item, the item stem remained the same, while the response formats were changed, resulting in the 6 sets of 30 items, each with a different response format. Tests were then constructed utilizing these items and the tests were administered in various combinations to a number of groups of several hundred college students.

The data collected will be used to determine the factor structure and the convergent and discriminant validity of the tests using the different response formats in the two ability domains and will be compared with similar data already available on vocabulary ability. The data will also be used to compare the item parameters, item and test information functions, internal consistency, and interrelationships among scores derived from the various response formats.

#### Results

Preliminary data analyses were completed on data collected on three of the six item formats using the analogies items. The formats analyzed were multiplechoice items with conventional response format (MCC), multiple-choice items with probabilistic response format (MCP), and dichotomous items with a dichotomous (MCD) response format (Types 1 through 3 in Figure 2).

Examination of Table 3 shows the average item scores for the 30 items in the three formats. The MCC and MCP items were rank ordered very similarly, as evidenced by the correlation of .91 between the average item scores for these two formats. The DD items ordered themselves somewhat differently, as indicated by the correlations between the DD average item scores and the MCC and MCP average item scores, which were .43 and .52, respectively. Although the present results reflect only one scoring system for the MCP items, several other scoring methods will also be investigated for this response format.

Table 4 shows validity and internal consistency reliability coefficients obtained for the three item formats. The validity coefficient reported is the correlation of total score with the reported grade-point average of the students (later analyses are planned using actual GPA rather than reported GPA). The validity coefficients for all three formats were not significantly different from each other, and were significantly different from zero. The MCP items had the highest internal consistency reliability, but the lowest validity. The MCC items had the second highest reliability, and highest validity; and the DD items had lowest reliability with moderate validity.

A number of factor analyses were performed on the data from the three item formats. Both principal axes and confirmatory analyses were used. The principal axes analyses showed that two orthogonal factors were extracted for each response format, and the pattern of positive and negative loadings was extremely similar for the MCC and MCP items but was very different for the DD items. In addition, the percent of total score variation accounted for by the two-factor solution varied with item format.

Confirmatory factor analysis was performed using the principal axes factor

	and Dichotomous	-Dichotomous (DD) F	ormat							
Item	Response Format									
Number	MCC*	MCP**	DD*							
1	.86	1.18	.87							
2	.64	.66	.42							
3	.78	.99	•65							
4	•52	•58	.81							
5	.73	1.02	.80							
6	.79	1.08	.52							
7	.28	.24	.64							
8	-48	.64	.61							
9	.61	.92	•51							
10	•58	.80	.73							
11	.60	.73	.44							
12	.89	1.19	.71							
13	.77	1.08	.79							
14	.62	.77	.32							
15	.75	1.29	.82							
16	.78	1.07	.92							
17	•64	.94	.70							
18	•84	1.48	.87							
19	.41	.63	-80							
20	•68	1.05	.84							
21	.87	1.14	.98							
22	•84	1.06	•65							
23	•87	1.32	.78							
24	.91	1.41	.98							
25	.68	1.08	.56							
26	.77	.88	.80							
27	.75	1.25	.83							
28	.86	1.29	.94							
29	.54	.60	•58							
30	.35	.42	.66							

Table 3 Average Item Scores for the Same 30 Analogies Items in Multiple-Choice Conventional Format (MCC), Multiple-Choice Probabilistic Format (MCP), and Dichotomous-Dichotomous (DD) Format

\*Proportion correct.

Provide a substantial second second second

\*\*Average score with range from 2.00 to -1.00.

solution for the MCC items as the basis for the model. The data showed better fit to the model for the MCP items than for the DD items, with the second factor a rather inconsequential factor for all three response formats.

Thus, the results obtained thus far advise against the use of the DD item format. This item format is less reliable than the other response formats, has only moderate validity, shows high levels of guessing, and does not appear to be consistently measuring analogies ability. The DD response format does not appear to be a viable alternative to the multiple-choice item. On the other hand, the MCP item format does appear to be a promising alternative to the traditional

•	Correla	cy Reliability tions with Rep sponse Formats	orted GPA
Response Format	N	Coefficient Alpha	Validity Coefficient
Multiple-Choice			
Conventional Multiple-Choice	486	•85	.23
Probabilistic Dichotomous-	299	.91	.17
Dichotomous	303	•59	.20

Table 4

multiple-choice item. It is more reliable, nearly equally as valid, and appears to be measuring analogies ability to a greater degree than the conventional multiple choice-items.

#### Additional Research in Progress

Further research and analysis, of course, are needed to determine whether or not these results are generalizable across ability domains, and whether any of the item formats not yet analyzed will also show promise as alternatives to the standard multiple-choice item. Considerable amounts of additional data will be obtained on these item formats using both the analogies and numerical reasoning items, and the resulting data will be analyzed to investigate the questions raised earlier. In addition, further evidence of the generality of the findings will be sought using vocabulary ability items that are being analyzed by similar methodologies. Once remaining item types are identified as replacements for the multiple-choice item, adaptive tests using these item types will be designed and their characteristics investigated.

#### Fit of Individuals to Item Response Theory Models

Previous research on the person response curve (Trabin & Weiss, 1979) and related research on person fit (Levine & Drasgow, 1980; Levine & Rubin, 1979) promises the capability of identifying individuals who, on a given test, are not responding in accordance with a given IRT model. This lack of fit to the model can derive from a number of possible causes, including the following:

- 1. Lack of motivation to respond appropriately to the test;
- 2. Inappropriate or nonrandom guessing, or lack of guessing;
- 3. Responses that are not in accord with the unidimensionality assumption of ICC theory.

Knowledge that an individual is not responding according to the model for any of these reasons would be appropriate information to be used in applied situations suggesting that the scores of that person be carefully considered before decisions are made on the basis of those scores. It may also be an important moderator variable for use in prediction studies.

To further investigate the utility of the person response curve (PRC) concept and to identify the correlates of deviations of an individual from the IRT model.

#### Approach

To properly investigate the usefulness of the PRC approach to the problem of person fit, its characteristics were investigated within the context of other indices designed for similar purposes. Based on a thorough review of the relatively small literature on person fit (also known as "appropriateness" measurement), the following indices were identified:

- 1. Trabin and Weiss's (1979) chi-square index of the fit of observed and expected PRCs.
- 2. Reckase's (1977) mean square deviation (MSD) index averaged over items,

$$MSD_{j} = \sum_{i=1}^{N} (u_{ij} - p_{ij})^{2} / N$$
[1]

where

 $MSD_{i}$  = the mean squared deviation for person j,

 $\underline{u}_{1i}$  = the actual response to item <u>i</u> by person <u>j</u>,

 $\underline{p_{ij}}$  = the probability of a correct response as predicted by the three-parameter IRT model, and

N = the number of items in the test.

This index is a special case of the fit of the observed PRC to the expected PRC.

- 3. A variation of Reckase's index in which only improbable responses are scored. It is difficult to argue that those responses that are in the predicted direction should be included in a person-fit statistic. Hence, only where  $(\underline{u}_{ij} - \underline{p}_{ij})^2$  is greater than .25 is it included in the statistic. The divisor of the statistic is still the number of items in the test.
- 4. The likelihood ratio index of person fit. For a given  $\hat{\theta}$  value, a probability distribution for the possible response vectors can be generated. The probability of an individual's actual response vector is divided by the probability of the most likely response vector to produce the likelihood ratio.
- 5. Wright (1968) proposed an index of person or item fit, which is the sum of the standardized squares of the residual after fitting the model. For the Rasch model it is  $e^{(p-1)}$  for an incorrect answer and  $e^{(1-p)}$  for a correct answer, where e = 2.71, p is  $\hat{\theta}$  for an individual, and <u>i</u> is the difficulty of the item.

- 6. Response pattern information, which reflects the flatness of the likelihood function.
- 7. The three appropriateness indices,  $L_0$ ,  $L_n$ , and  $\hat{\sigma}$  described by Levine and Drasgow (1980) and Levine and Rubin (1979).
- 8. The difference between the difficulties of the easiest item answered incorrectly and the most difficult item answered correctly.
- 9. Donlon and Fisher's (1968) "personal biserial correlation" and Jacobs' (1963) variation of it.
- 10. The posterior variance of IRT-based Bayesian ability estimates.

The utility of these indices for identifying individual nonfit to IRT models is first being studied in simulation. As a first step, the null distributions of these person-fit statistics are being examined, to serve as reference points for the later studies to determine how well each statistic identifies person nonfit. Using a set of 2,500 simulees rectangularly distributed at 25 equally spaced levels of  $\theta$ , 625 items were administered to each simulee. Items were rectangularly distributed in b with 25 items at each of 25 levels of b corresponding to the 25 levels of  $\theta$ . Within each level of b, items varied in discrimination at .07 intervals. Simulation data were generated separately for the 2,500 simulees for c = 0.0, .20, and .25 in order to examine the effects of guessing on the null distributions. To examine the effects of test length and discrimination, shorter tests were selected from the item pool at differing levels of <u>a</u>. For each of these test configurations, null distributions were computed for each of the person-fit indices.

After distributions of these indices are known for model-conforming data, increasing amounts of random responses will be added to the response vectors to simulate random responding, inattention or low motivation, and guessing. Changes in the distributions that result from increasingly random responding will be noted. In this phase, percent of random response is an additional independent variable. Those person-fit statistics that are affected most strongly by random responding will be retained as good candidates for further live-data research. Results of the data analysis are not yet available from these studies.

#### Additional Research Plans

The effect of multidimensionality on the person-fit indices will be examined. This will be studied by generating response data from a  $\theta$  known to correlate to varying degrees with the original  $\theta$  and by inserting these responses into the response vectors. Degree of correlation as well as number of dimensions will be manipulated. Dependent variables studied will be the ability of the person-fit indices to identify the existence of the multidimensional response patterns.

Once promising person-fit indices are identified in the simulation studies, live-testing studies will be designed in which the fit of persons to the IRT model is empirically tested on a given pool of items. Experimental studies will be designed to attempt to induce deviations from the model in groups of individuals and to observe whether the person-fit indices are sufficiently sensitive to identify those deviations when they exist. In addition, the existence of natu-

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rally occurring groups in which deviation from the IRT model might occur will be studied. For example, it can be hypothesized that on certain kinds of subtests (e.g., verbal ability tests) students from a non-English speaking culture would likely show significant deviations from unidimensionality. It could also be hypothesized that students who are not "test-wise" or who have a lack of familiarity with multiple-choice tests would show specific deviations from IRT models in their test performance. In addition, the generality of person-fit variables across ability dimensions will be studied in order to determine whether deviations of fit to the model for an individual are specific to an item pool or occur across different kinds of item domains.

#### New Types of Ability Tests

Psychometric attempts to measure individual differences in cognitive abilities during the last 60 years have produced tests of global abilities such as general reasoning, verbal and quantitative abilities, as well as more specific ability "factors" such as speed and flexibility of closure, spatial orientation, and word fluency (French, Ekstrom, & Price, 1963). Research in adaptive testing has shown that the precision and validity of ability tests can be improved by adaptive testing procedures. At the same time, cognitive psychologists have developed several standard tasks or paradigms that have been used to study the mechanisms and structure of aspects of memory, attention, and cognition (e.g., Sperling, 1960; Sternberg, 1966). Attempts to assess quantitative differences between individuals in such information-processing abilities and to relate them to more traditional psychometric measures is a fairly recent phenomenon (Chiang & Atkinson, 1976; Day, 1977; Hunt, Frost, & Lunneborg, 1973; Hunt, Lunneborg, & Lewis, 1975; Lunneborg, 1977; Rose, 1974). Carroll (1976) has provided an additional framework for relating traditional psychometric factors to their cognitive information-processing requirements. Some of this research has suggested that individual differences in such information-processing abilities can be measured reliably by interactive computers and that these abilities may add incremental validity in predicting external job criteria (Cory, 1977; Cory, Rimland, & Bryson, 1977).

In the past, tasks of the type that have been used by cognitive psychologists to measure information-processing abilities and by psychometricians to measure perceptual and spatial factors have been administered as blocks of fixed numbers of trials or replications. As a result, little is known about employing the important parameters of these tasks for adapting the difficulty level of replications to converge upon an individual's ability level. A major emphasis of the research, therefore, involved (1) studying some of these tasks from the point of view of how computerized adaptive administration could be meaningfully achieved and (2) evaluating the measurement benefits of adaptive administration of ability tests designed to utilize the unique capabilities of computerized administration.

#### Objective

The objectives were to investigate the application of adaptive testing techniques to improving the measurement characteristics of several cognitive information-processing tasks (e.g., short-term memory span, capacity of visual sensory memory). Computerized administration of these ability tests will be investigated as a means of modifying task presentations over time in a way that would not be possible in paper-and-pencil testing.

#### Approach

Two types of ability test items that utilized the unique capabilities of computer administration--Memory for Patterns and Digit Span--were studied to investigate the feasibility of adaptive administration of information-processing tests. These tasks were studied from the point of view of enabling the computer to adapt the difficulty of the tasks presented to the ability level of the testee during the process of testing.

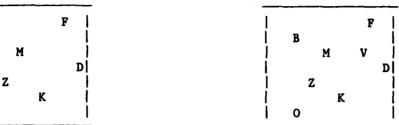
Memory for Patterns. This test was based on a procedure devised by Sperling (1960) for determining the capacity of visual sensory memory. The procedure is based on presentation of arrays of letters that must be studied by the testee. The procedure will be modified to adapt to a testee's recall on previous screen presentations in order to obtain precise quantitative estimates of individual differences in visual sensory memory capacity with fewer replications. For example, the size of the array (and thus the memory capacity demands) can be made larger or smaller based on an individual's previous performance; and/or the duration of the array presentation could be lengthened or shortened to adapt the task difficulty to the testee's ability level during the course of testing. Prior to implementing such an adaptive approach, however, psychometric research is needed to determine how much of an increase or decrease in stimulus array size constitutes a meaningful increment or decrement in difficulty and what ranges of array size are needed to adequately span differences existing in various populations of interest.

Data collection on two tests of short-term spatial and perceptual memory was therefore designed to allow a preliminary evaluation of potential adaptive testing parameters. The experimental Memory for Patterns items consisted of bounded two-dimensional arrays containing 3 to 10 letters. Each item consisted of a pair of successive screen presentations. The stimulus display was presented for a brief timed period and then was erased from the cathode ray terminal (CRT) and replaced with the recall display. The recall display contained a bounded letter pattern that was identical to the first pattern (the stimulus display) except that one or two letters had changed position. The recall display was untimed and accepted the student's response, indicating which letter(s) she/he thought had moved. Figure 3 shows sample Memory for Patterns stimulus and response display, constituting one test item. Data were also collected on a related set of items designed to measure Space Memory, which were similar to the Memory for Patterns items except that the letter patterns presented were unbounded and thus spread about the entire CRT screen.

In the initial data collection for the Memory for Patterns and Space Memory items, subtests varied in the number of letters that could change position from the first pattern to the second. Three 10-item subtests were administered to each student in a number of different experimental groups. The first subtest was composed of patterns in which one letter changed position, the second subtest was composed of patterns in which two letters changed position, and the third subtest included patterns in which either one or two letters changed position. Students were assigned to one of four conditions that varied the order in which the items were presented and the presentation time, in seconds, of the first pattern of each item pair. In the Memory for Patterns test the two presentation times were 5 and 10 seconds, and in the Space Memory test they were 7 and 12 seconds. The two orders in which items were presented were (1) from lowest to highest difficulty and (2) from moderate to highest to lowest difficulty,

#### Figure 3 A Sample Memory for Patterns Item, with Stimulus Display on the Left and Recall Display on the Right

Type the letter which has changed position, or a question mark (?). Then press "RETURN".



B

0

V

where difficulty was indexed by the number of letters in the pattern. Reactions to the tests were also obtained from each student.

The analysis of the data was directed at scaling the difficulty of items and at studying the effects of pattern density and presentation times on item difficulty indices. A comparison of the item order conditions was made to examine the effects of practice and proactive inhibition on indices of item difficulty. A comparison of presentation time conditions was directed at determining reasonable item exposure times and at investigating the possibility of using presentation time along with pattern density and number of pattern changes as adaptive parameters for future adaptive testing.

The results of preliminary analyses of the Memory for Patterns items were used to design new items and to modify the experimental conditions under which they were administered. The major conclusions from the preliminary analysis was that item difficulty was more a function of type of pattern configuration than either rate or order condition. For this reason, types of Memory for Patterns items were hypothesized in a systematic manner. Eight Memory for Patterns item types, which can be separated into two groups, were developed. One group of items is composed of patterns taking a geometric form, such as a line, triangle, square, or pentagon. Four item types of this nature were developed:

- 1. An item that had a geometric form was changed to a nongeometric form through a small move,
- 2. An item that had a geometric form was changed to a nongeometric form through a large move,
- 3. An item that had a nongeometric form was changed into a geometric form through a small move, and
- 4. An item that had a nongeometric form was changed into a geometric form through a large move.

Within the four item types, there were various degrees of nongeometric form that the patterns could have.

A second group of items was composed of nongeometric forms, but the patterns varied in terms of pattern configuration definition. For example, some patterns, although not geometric in form, are better defined and thus easier to remember. The four item types were as follows:

- 1. Well-defined pattern configuration with a small change in letter configuration,
- 2. Well-defined pattern configuration with a large change in letter configuration,
- 3. Poorly defined pattern configuration with a small change in letter position, and
- 4. Poorly defined pattern configuration with a large change in letter position.

Hypotheses were made with regard to item type and resultant item difficulty. Ninety new items were written based on these eight basic Memory for Patterns item types.

Of the 90 items, 42 were administered under various experimental conditions. Students were assigned sequentially to one of 15 testing conditions. The 42 Memory for Patterns items were presented under three order and five rate conditions. Values of pattern densities, or the number of letters in a configuration were 3, 4, 5, 6, 7, 8, 9, and 10 letters. The 42 items were presented in three orders: (1) ordered from 3 to 10 letters in a pattern; (2) ordered from 6 to 10 letters, then 3 to 5 letters in a pattern; and (3) ordered from 8 to 10, then 3 to 7 letters in a pattern. The five rate conditions were 3, 5, 7, 10, and 13 seconds in duration. Analyses of these data will be oriented toward the identification of adaptive parameters for the items and the effects of administration conditions (e.g., sequence effects) on the adaptive parameters.

Digit Span. Contrasting with the Memory for Patterns tests, which tap both spatial abilities and short-term memory capacity, the Digit-Span test is primarily a test of short-term memory. In this test a series of numbers is presented in rapid succession. The respondent is asked to recall the numbers in the order they were presented by typing them into the computer terminal in a serial string. Since the time interval between presenting the numbers, clearing the screen, and asking for a response was very short, the test was essentially an indicator of short-term memory capacity.

To identify possible adaptive parameters for this type of test, presentation rate was varied experimentally so that items were presented at one of three rates--.2 seconds, .3 seconds, and .5 seconds--corresponding to fast, moderate, and slow rates. Series length (the number of stimulus values to be recalled), a second potential adaptive testing parameter, varied from 4 to 10 stimuli. Six items at each test length were administered to yield a total of 42 digit series in the test. Length of digit series and of presentation rate will be investigated as possible test and item parameters.

The items were presented in three different orders: (1) from easy to difficult (where difficulty was defined in terms of numbers of stimulus values to recall), (2) from moderately difficult to difficult to easy, and (3) from difficult to easy to moderately difficult. The order of item presentation will be analyzed to determine if there are practice or prohibitive effects from one item to the next, since such effects would be undesirable in adaptive test administration, and to determine the effects of series length and display time on item difficulties.

#### Analyses and Future Plans

Data analyses and future data collection will be oriented toward identifying potential adaptive parameters in terms of factors influencing the difficulty of test items. In addition, the influence of undesirable factors--such as sequence effects, inhibitive effects, and other factors that would interfere with adaptive administration--will be investigated in order to permit the design and evaluation of adaptive tests of information-processing abilities.

Other measures of more traditional psychometric factors, such as flexibility of closure, may also benefit from application of a computerized adaptive testing framework. A common measure of this construct has been variants of the Hidden or Embedded Figures test. The unique capabilities of computerized administration may allow increased validity to be achieved by inducing movement into either the stem figure and/or the alternative response figures. For example, the testee may be required to selectively attend to and to articulate the stem figure in a more complex figure, which not only contains distracting lines but also grows, shrinks, translates, and/or rotates over time. Adaptive administration can be achieved by modifying the amount of "noise" in the figure and by dynamically adapting the amount and speed of movement in the complex figure. Several psychometric questions can be studied. For example, what size increments in "noise" and movement will allow the computer to most efficiently converge upon the testee's ability level? Are amounts of "noise" and movement independent dimensions of difficulty to be manipulated? How is performance under varying degrees of noise and movement to be scored?

The design and implementation of adaptive tests of these information-processing kinds of ability tests raises a host of new questions and problems to be investigated. Beyond the identification of adaptive parameters for these kinds of items, and ruling out extraneous factors such as sequence effects (which can reduce the effectiveness of the adaptive procedures), new questions will need to be addressed with regard to the design and scoring of the adaptive tests. Design questions will include the identification of the functions relating display time to item difficulty and identification of the procedures for using this continuous function in adapting display time on each item to each individual's test performance on an item-by-item basis. Since it may be observed that this function is different for items of different difficulties based on stimulus characteristics (e.g., pattern density, length of span string), adaptive testing procedures will need to be developed that will jointly take into account the combination of discrete and continuous difficulty factors. New scoring procedures may also need to be designed for these kinds of tests, since each testee will receive a set of items selected to match his/her ability levels; and/or the applicability of IRT-based scoring procedures will need to be investigated.

Finally, comparisons of computerized adaptive, computerized nonadaptive and traditional measures of these abilities should be made to determine if more precise and efficient measurement can be achieved through adaptive administration. Where appropriate external criteria are available, predictive validity comparisons should also be made. If the findings from traditional ability testing generalize to these kinds of ability tests, the resulting tests will be shorter, more precise, and more valid and will permit more meaningful measurement of the range of human abilities.

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