

AD-A036 392

AIR FORCE CIVIL ENGINEERING CENTER TYNDALL AFB FLA
THE EFFECT OF NAVY AND AIR FORCE AIRCRAFT ENGINE TEST FACILITIE--ETC(U)
OCT 76 B C GREMS, D F NAUGLE
AFCEC-TR-76-36

F/G 21/5

UNCLASSIFIED

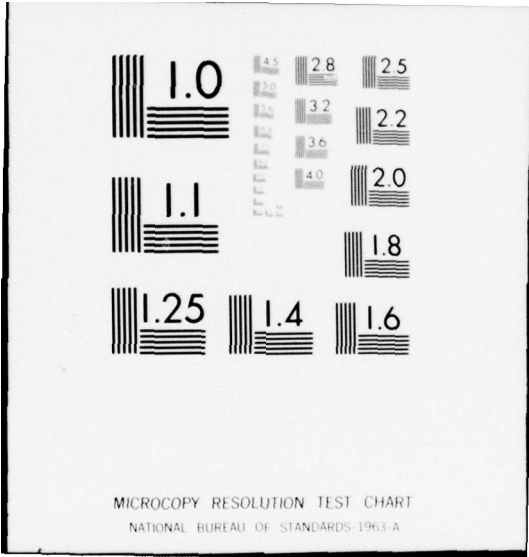
NL

1 of 1
ADA036392



END

DATE
FILMED
3-77



6

J

ADA 036392

THE INSTITUTE OF STATISTICS

THE CONSOLIDATED UNIVERSITY
OF NORTH CAROLINA



Approved for public release;
distribution unlimited.

DDC
 REFORMED
 MAR 4 1977
 REGISTERED
 C



DEPARTMENT OF BIostatISTICS
Chapel Hill, North Carolina

AIR FORCE OFFICE OF SCIENTIFIC RESEARCH (AFSC)
NOTICE OF TRANSMITTAL TO DDC

This technical report has been reviewed and is
approved for public release IAW AFR 190-12 (7b)
Distribution is unlimited.

A. D. BLOSE
Technical Information Officer

SECURITY CLASSIFICATION OF THIS PAGE (When Data Entered)

(19) REPORT DOCUMENTATION PAGE		READ INSTRUCTIONS BEFORE COMPLETING FORM	
1. REPORT NUMBER	2. GOVT ACCESSION NO.	3. RECIPIENT'S CATALOG NUMBER	
(18) AFOSR - TR - 77 - 0108			
4. TITLE (and Subtitle)		5. TYPE OF REPORT & PERIOD COVERED	
(6) WEAK CONVERGENCE OF SOME QUANTILE PROCESSES ARISING IN PROGRESSIVELY CENSORED TESTS,		Interim	
7. AUTHOR(s)		6. PERFORMING ORG. REPORT NUMBER	
(10) Pranab Kumar / Sen			
9. PERFORMING ORGANIZATION NAME AND ADDRESS		8. CONTRACT OR GRANT NUMBER(s)	
University of North Carolina Department of Statistics Chapel Hill, NC 27514		(15) ✓ AF - AFOSR 74-2736-74	
11. CONTROLLING OFFICE NAME AND ADDRESS		10. PROGRAM ELEMENT, PROJECT, TASK AREA & WORK UNIT NUMBERS	
Air Force Office of Scientific Research/NM Bolling AFB, Washington, DC 20332 (12) 31p.		(16) 61102F (17) 2304 A5	
14. MONITORING AGENCY NAME & ADDRESS (if different from Controlling Office)		12. REPORT DATE	
(14) mimeo Ser-1092		(11) October 1976	
		13. NUMBER OF PAGES	
		29	
		15. SECURITY CLASS. (of this report)	
		UNCLASSIFIED	
		15a. DECLASSIFICATION/DOWNGRADING SCHEDULE	
16. DISTRIBUTION STATEMENT (of this Report)			
Approved for public release; distribution unlimited.			
17. DISTRIBUTION STATEMENT (of the abstract entered in Block 20, if different from Report)			
18. SUPPLEMENTARY NOTES			
19. KEY WORDS (Continue on reverse side if necessary and identify by block number)			
Brownian motions, linear combinations of functions of order statistics (with stochastic coefficients), martingales, progressive censoring schemes, quantile processes, sampling from a finite population, time-sequential tests, weak convergence.			
20. ABSTRACT (Continue on reverse side if necessary and identify by block number)			
For progressive censoring schemes pertaining to a general class of (parametric as well as nonparametric) testing situations, one encounters a (partial) sequence of linear combinations of functions of order statistics where the coefficients are themselves stochastic variables. Weak convergence of such a quantile process to an appropriate Gaussian function is studied here, and the same is incorporated in the formulation of suitable (time -) sequential tests based on these quantile processes.			

* Work supported by the Air Force Office of Scientific Research, A.F.S.C.,
U.S.A.F., Grant No. AFOSR-74-2736.

ACQUISITION TX	White Section <input checked="" type="checkbox"/>
ADS	Bull Section <input type="checkbox"/>
BTC	<input type="checkbox"/>
UNANNOUNCED JUSTIFICATION	
BY	DISTRIBUTION AVAILABILITY CODES
Dist.	ATAIL and/or SPECIAL
A	

D D C
D D C
MAR 4 1977
D D C

WEAK CONVERGENCE OF SOME QUANTILE PROCESSES ARISING
IN PROGRESSIVELY CENSORED TESTS*

by

Pranab Kumar Sen

University of North Carolina at Chapel Hill

Institute of Statistics Mimeo Series No. 1092

October 1976

WEAK CONVERGENCE OF SOME QUANTILE PROCESSES ARISING
IN PROGRESSIVELY CENSORED TESTS*

by

Pranab Kumar Sen
University of North Carolina, Chapel Hill

ABSTRACT

For progressive censoring schemes pertaining to a general class of (parametric as well as nonparametric) testing situations, one encounters a (partial) sequence of linear combinations of functions of order statistics where the coefficients are themselves stochastic variables. Weak convergence of such a quantile process to an appropriate Gaussian function is studied here, and the same is incorporated in the formulation of suitable (time -) sequential tests based on these quantile processes.

AMS 1970 Classification NOS: 60B10, 62G30, 62L99.

Key Words and Phrases: Brownian motions, linear combinations of functions of order statistics (with stochastic coefficients), martingales, progressive censoring schemes, quantile processes, sampling from a finite population, time-sequential tests, weak convergence.

* Work supported by the Air Force Office of Scientific Research, A.F.S.C., U.S.A.F., Grant No. AFOSR-74-2736.

1. INTRODUCTION

Let X_1, \dots, X_N , the survival times of $N (> 1)$ items under life-testing, be independent random variables (rv) with continuous distribution functions (df) F_1, \dots, F_N , respectively, all defined on the real line $(-\infty, \infty)$. In a life-testing problem, the smallest observation comes first, the second smallest next, and so on, until the largest one emerges last. Thus, the observable random variables can be represented as

$$(1.1) \quad \{(Z_{N1}, Q_{N1}), \dots, (Z_{NN}, Q_{NN})\}$$

where Z_{Nj} is the j -th smallest observation among X_1, \dots, X_N ($1 \leq j \leq N$) and

$$(1.2) \quad Z_{Nj} = X_{Q_{Nj}}, \quad \text{for } j = 1, \dots, N;$$

by virtue of the assumed continuity of the F_i , ties among the X_i (and hence, the Z_{Ni}) may be neglected, in probability, so that the Q_{Nj} are uniquely (in probability) defined by (1.2) and (Q_{N1}, \dots, Q_{NN}) represents a permutation of $(1, \dots, N)$. Since a complete collection of (1.1) demands the span of the experimentation until Z_{NN} is observed, while practical considerations often set time and cost limitations on the duration of experimentation, the experiment may be terminated at the r -th failure Z_{Nr} , where

$$(1.3) \quad r = [Np] + 1 \quad \text{for some } 0 < p < 1$$

([s] being the largest integer contained in s). Thus, here, the observable random variables are

$$(1.4) \quad \tilde{Z}_N^{(r)} = (Z_{N1}, \dots, Z_{Nr}) \quad \text{and} \quad \tilde{Q}_N^{(r)} = (Q_{N1}, \dots, Q_{Nr}) .$$

[We also know the complementary set $Q_N^{(N)} - Q_N^{(r)}$, but without any idea of the order in which the elements do appear.] For testing suitable statistical hypotheses concerning the df's F_1, \dots, F_N , a terminal test based on $(Z_N^{(r)}, Q_N^{(r)})$ is termed a *censored test*; we denote the corresponding test statistic by T_{Nr} .

In a progressive censoring scheme (PCS), the experiment is monitored from the very beginning with the objective of an early termination (prior to Z_{Nr}) whenever statistically feasible, i.e., one observes T_{Nk} at each failure time Z_{Nk} ($1 \leq k \leq r$), and, if for some k ($\leq r$), T_{Nk} provokes a clear statistical decision in favor of one of the hypotheses, experimentation is terminated at that time-point; if no such k ($< r$) exists, the experimentation is stopped at the r -th failure Z_{Nr} along with an appropriate statistical decision. Thus, by constitution, a PCS test is based on the entire partial sequence

$$(1.5) \quad \{Z_N^{(k)}, Q_N^{(k)}, 1 \leq k \leq r\},$$

and is time-sequential in nature. Since the updated sequence $\{T_{Nk}, 1 \leq k \leq r\}$ involves dependent random elements and the PCS involves repeated testing on these dependent statistics, statistical analysis of such a problem, often, becomes complicated. In this context, suitable invariance principles for $\{T_{Nk}, 1 \leq k \leq r\}$ provide us with convenient tools for formulating a PCS test and studying its (asymptotic) properties.

In the context of *nonparametric life testing*, Chatterjee and Sen (1973) have studied PCS tests based on a general class of linear rank statistics; the theory rests on an invariance principle for PCS linear rank statistics. For the case of $F_1 = \dots = F_N = F$ involving an unknown parameter θ (form

of F assumed to be specified), Sen (1976) has constructed PCS tests for $H_0: \theta = \theta_0$ vs. $H_1: \theta \neq$ (or $>$ or $<$) θ_0 based on PCSLR (likelihood ratio-) statistics; here also, the theory is based on an invariance principle for the PCSLR. The object of the present investigation is to focus on a general class of location, scale and regression models where the PCPLR statistics yield suitable *quantile processes* (QP) and to develop suitable invariance principles for such PCQP's. These models are introduced in Section 2 and the corresponding PCSLR statistics are derived and incorporated in the construction of appropriate PCQP's. By nature, such a PCQP involves a partial sequence of linear combinations of functions of order statistics with stochastic coefficients depending on the various censoring stages. Invariance principles for the PCQP are studied in Section 3. The last section deals with some application of the main theory to some time-sequential tests.

2. A CLASS OF PCQP'S

Let Θ be an open interval containing 0 and let $\{f(x;\theta), \theta \in \Theta\}$ be a family of absolutely continuous probability density functions (pdf), and for every $f: -\infty < x < \infty$, let us denote by

$$(2.1) \quad g(x) = -(\partial/\partial\theta)\log f(x;\theta) \Big|_0$$

$$G(x) = [1-F(x;0)]^{-1} \int_x^\infty g(z) dF(x;0)$$

where $F(x;0) = \int_{-\infty}^x f(z;0) dz$. We conceive of the model where the df F_i

admits of the pdf f_i and

$$(2.2) \quad f_i(x) = f(x; \Delta(c_i - \bar{c}_N)) , \quad -\infty < x < \infty , \quad 1 \leq i \leq N ,$$

where c_1, \dots, c_N are given constants (not all equal), $\bar{c}_N = N^{-1} \sum_{i=1}^N c_i$ and Δ is an unknown parameter. We intend to test

$$(2.3) \quad H_0: \Delta = 0 \quad \text{vs.} \quad H_1: \Delta \neq (\text{or } > \text{ or } <) 0 .$$

Let us also denote by

$$(2.4) \quad C_N^2 = \sum_{i=1}^N (c_i - \bar{c}_N)^2 \quad \text{and} \quad c_{Ni}^* = C_N^{-1} (c_i - \bar{c}_N) , \quad 1 \leq i \leq N ,$$

so that $\sum_{i=1}^N c_{Ni}^* = 0$ and $\sum_{i=1}^N (c_{Ni}^*)^2 = 1$. Then, the likelihood function for $(Z_N^{(k)}, Q_N^{(k)})$ is given by

$$(2.5) \quad L_{N,k}(Z_N^{(k)}, Q_N^{(k)}) = \prod_{i=1}^k f_{Q_{Ni}}(Z_{Ni}) \prod_{i=k+1}^N (1 - F_{Q_{Ni}}(Z_{Nk})) \\ = \prod_{i=1}^k f(Z_{Ni}; \Delta(c_{Q_{Ni}} - \bar{c})) \prod_{i=k+1}^N [1 - F(Z_{Nk}; \Delta(c_{Q_{Ni}} - \bar{c}))] .$$

Simple computations leads us to

$$(2.6) \quad T_{Nk} = C_N^{-1} \left\{ (\partial/\partial\Delta) \log L_{N,k}(Z_N^{(k)}, Q_N^{(k)}) \Big|_{\Delta=0} \right\} \\ = \sum_{i=1}^k c_{NQ_{Ni}}^* [g(Z_{Ni}) - \bar{G}(Z_{Nk})] , \quad k = 1, \dots, N$$

and $T_{N0} = 0$. Note that

$$T_{NN} = \sum_{i=1}^N c_{NQ_{Ni}}^* g(Z_{Ni}) = \sum_{i=1}^N c_{Ni}^* g(X_i) .$$

Thus, the LMP (locally most powerful) test statistic based on $(Z_N^{(k)}, Q_N^{(k)})$

is T_{Nk} , and in the setup of Progressive censoring, the sequence $\{T_{Nk}; 0 \leq k \leq r\}$ relates to a sequence of linear combinations of functions of order statistics with the coefficients $\{c_{NQ_{Ni}}^*\}$ all stochastic in nature. By reference to Hájek and Sidák (1967, pp. 70-71), we may remark that the model (2.2) includes as special cases, the classical two-sample location and scale models as well as the so called regression model in location and scale

We are primarily concerned here with weak convergence of suitable stochastic processes constructed from the partial sequence $\{T_{Nk}; 0 \leq k \leq r\}$ where r satisfies (1.3). In statistical applications, often, we face some related PCQP's which we pose below.

Note that under the usual Cramér-regularity conditions, $\int_{-\infty}^{\infty} g(x)dF(x;0) = 0$, so that by (2.1)

$$(2.8) \quad \bar{G}(x) = -\{1-F(x;0)\}^{-1} \int_{-\infty}^x g(z)dF(z;0), \quad -\infty < x < \infty,$$

Let now $u(t)$ be equal to 1 or 0 according as t is \geq or $<$ 0, and let

$$(2.9) \quad S_N(x) = N^{-1} \sum_{i=1}^N u(x-X_i), \quad -\infty < x < \infty$$

be the empirical df. We define

$$(2.10) \quad \bar{G}_N(x) = \begin{cases} -\{1-S_N(x)\}^{-1} \int_{-\infty}^x g(x)dS_N(x), & x < Z_{NN}, \\ g(Z_{NN}), & x \geq Z_{NN}. \end{cases}$$

Then, in (2.6), we replace $\bar{G}(Z_{nk})$ by $\bar{G}_N(Z_{Nk})$, and obtain a related sequence

$$(2.11) \quad T_{Nk}^* = \begin{cases} 0, & k = 0 \\ \sum_{i=1}^k c_{NQ_{Ni}}^* [g(Z_{Ni}) - \bar{G}_N(Z_{Nk})], & 1 \leq k \leq N-1 \\ T_N, & k = N. \end{cases}$$

Note that, one can rewrite T_{Nk}^* ($1 \leq k \leq N-1$) as

$$(2.12) \quad \begin{aligned} T_{Nk}^* &= \sum_{i=1}^k c_{NQ_{Ni}}^* [g(Z_{Ni}) + \frac{1}{N-k} \sum_{i=1}^k g(Z_{Ni})] \\ &= \sum_{i=1}^k g(Z_{Ni}) [c_{NQ_{Ni}}^* + \frac{1}{N-k} \sum_{\alpha=1}^k c_{NQ_{N\alpha}}^*], \end{aligned}$$

so that it is a linear combination of a function of order statistics with stochastic coefficients depending on the censoring stage.

We conclude this section with the asymptotic stochastic equivalence of the two sequences $\{T_{Nk}; 0 \leq k \leq r\}$ and $\{T_{Nk}^*; 0 \leq k \leq r\}$. We specifically provide the proof for the null hypothesis situation and we shall see in Section 3 that the conclusion remains true for contiguous alternatives. We denote by $F_0(x) = F(x;0)$ and consider first the following.

Lemma 2.1. Let $\{d_{Ni}, 1 \leq i \leq N; N \leq 1\}$ be a triangular array of real numbers satisfying

$$(2.13) \quad \sum_{i=1}^N d_{Ni} = 0 \quad \text{and} \quad \sum_{i=1}^N d_{Ni}^2 = 1.$$

Also, let $q = \{q(t): 0 < t < 1\}$ be a continuous, non-negative, U-shaped and square integrable function inside $I = [0,1]$. Finally, let $Q = (Q_1, \dots, Q_N)$ takes on each permutation of $(1, \dots, N)$ with equal probability $(N!)^{-1}$. Then

$$(2.14) \quad P\left\{ \max_{1 \leq k \leq N-1} q(k/N) \left| \sum_{i=1}^k d_{NQ_i} \right| \geq 1 \right\} \leq \int_0^1 q^2(t) dt.$$

Proof. Let P_N denote the uniform probability measure over the set of $N!$ permutations of $(1, \dots, N)$. Then, $E[d_{NQ_i} | P_N] = 0$ and

$$(2.15) \quad E(d_{NQ_i} d_{NQ_j} | P_N) = \begin{cases} \frac{1}{N} \sum_{\alpha=1}^N d_{N\alpha}^2 = \frac{1}{N}, & i = j \\ \frac{1}{N(N-1)} \sum_{\alpha \neq \beta=1}^N d_{N\alpha} d_{N\beta} = \frac{-1}{N(N-1)}, & i \neq j. \end{cases}$$

Thus, if we let

$$(2.16) \quad U_{Nk} = (N-k)^{-1} \sum_{i=1}^k d_{NQ_i}, \quad 1 \leq k \leq N-1,$$

we have

$$(2.17) \quad E[U_{Nk} | P_N] = 0 \quad \text{and} \quad E[U_{Nk}^2 | P_N] = k\{N(N-1)(N-k)\}^{-1}.$$

Further, under P_N ,

$$(2.18) \quad E[d_{NQ_{k+1}} | Q_1, \dots, Q_k] = (N-k)^{-1} \sum_{j=k+1}^N d_{NQ_j} \\ = -(N-k)^{-1} \sum_{j=1}^k d_{NQ_j} = -U_{Nk}, \quad 0 \leq k \leq N-1,$$

and hence, by (2.16) and (2.18)

$$(2.19) \quad E[U_{Nk+1} | Q_1, \dots, Q_k] = \frac{1}{N-k-1} [U_{Nk} + E\{d_{NQ_{k+1}} | Q_1, \dots, Q_k\}] \\ = U_{Nk}, \quad \text{for } k = 0, 1, \dots, N-2,$$

so that under P_N , $\{U_{Nk}\}$ is a martingale. Let

$$(2.20) \quad h_{Nk} = (N-k)q(k/N), \quad 1 \leq k \leq N-1.$$

Then, by the U-shapedness of q , there exists an $\alpha: 0 < \alpha < 1$, such that $(N-k)q(k/n)$ is \searrow in k for $1 \leq k \leq N\alpha$. Hence, by the Chow (1960) extension of the Hájek-Rényi inequality,

$$\begin{aligned}
 (2.21) \quad & P\left\{ \max_{1 \leq k \leq N\alpha} q(k/N) \left| \sum_{i=1}^k d_{NQ_i} \right| \geq 1 \right\} = P\left\{ \max_{1 \leq k \leq N\alpha} h_{Nk} |U_{Nk}| \geq 1 \right\} \\
 & \leq \left\{ h_{N1}^2 E(U_{N1}^2) + \sum_{k=2}^{[N\alpha]} h_{Nk}^2 [E(U_{Nk}^2) - E(U_{Nk-1}^2)] \right\} \\
 & = \left\{ N^{-1} q(1/N) + \sum_{k=2}^{[N\alpha]} q^2(k/N) [(N-k)/(N-1)(N-k+1)] \right\} \\
 & \leq N^{-1} \sum_{k=1}^{[N\alpha]} q^2(k/N) \quad (\text{as } (N-k)/(N-1)(N-k+1) \leq N^{-1}, k \geq 1), \\
 & \leq \int_0^\alpha q^2(t) dt, \quad \text{as } q \text{ is U-shaped.}
 \end{aligned}$$

Since $\sum_{i=1}^k d_{NQ_i} = -\sum_{i=k+1}^N d_{NQ_i}$, $1 \leq k \leq N-1$, the case of $N\alpha < k \leq N-1$ can be reduced by reflection and an inequality for this complementary part be obtained in the same manner. Q.E.D.

In particular, if we let $q(t) = K^{-1}$, $0 \leq t \leq 1$ where $0 < K < \infty$ and choose K large, we obtain from (2.14) that

$$(2.22) \quad \max_{1 \leq k \leq N} \left| \sum_{i=1}^k d_{NQ_i} \right| = O_p(1), \quad \text{uniformly in } N.$$

Note that if the X_i are i.i.d. with df F_0 , then by the Glivenko-Cantelli Lemma, as $N \rightarrow \infty$,

$$(2.23) \quad \max_{1 \leq k \leq N} |F_0(Z_{Ni}) - k/N| \rightarrow 0 \quad \text{almost surely (a.s.)}.$$

Lemma 2.2. *If the X_i are i.i.d. with df F_0 , then under (1.3) and $\int |g| dF_0 < \infty$,*

$$(2.24) \quad \max_{1 \leq k \leq r} |\bar{G}(Z_{Nk}) + (N-k)^{-1} \sum_{i=1}^k g(Z_{Ni})| \rightarrow 0 \quad \text{a.s., as } N \rightarrow \infty.$$

Proof. First note that under the hypothesis of Lemma 2.2,

$$(2.25) \quad \sup_{-\infty < z < \infty} \left| \int_{-\infty}^z g(x) dS_N(x) - \int_{-\infty}^z g(x) dF_0(x) \right| \rightarrow 0 \quad \text{a.s., as } N \rightarrow \infty;$$

the proof is straightforward [see, for example, Basu and Borwankar (1971)], and hence, is omitted. Secondly, under (1.3), $r/N \rightarrow p$; $0 < p < 1$,

$$(2.26) \quad \max_{1 \leq k \leq r} | \{1 - F_0(Z_{Nk})\} N / (N-k) - 1 | \rightarrow 0 \quad \text{a.s., as } N \rightarrow \infty; \quad \text{by (2.23).}$$

The rest of the proof follows from (2.8) and (2.24)-(2.26). Q.E.D.

Note that for i.i.d. X_1, \dots, X_N (with df F_0), $Q_N^{(N)} = (Q_{N1}, \dots, Q_{NN})$ takes on each permutation of $(1, \dots, N)$ with equal probability $1/N!$ Thus, by (2.6), (2.12), (2.22) and (2.24), we obtain that under H_0 and (1.3)

$$(2.27) \quad \max_{1 \leq k \leq r} | T_{Nk} - T_{Nk}^* |$$

$$\leq \left\{ \max_{1 \leq k \leq r} \left| \sum_{i=1}^k c_{NQ_{Ni}}^* \right| \right\} \left\{ \max_{1 \leq k \leq r} \left| \bar{G}(Z_{Nk}) + \frac{1}{N-k} \sum_{\alpha=1}^k g(Z_{N\alpha}) \right| \right\}$$

$$\rightarrow 0, \quad \text{in probability.}$$

With this results at our disposal, we are tempted to consider a more general class of PCQP's and then to study invariance principles for this class, leading to similar results for $\{T_{Nk}\}$ as special cases.

3. AN INVARIANCE PRINCIPLES FOR PCQP

Instead of considering PCQP's derivable from some PCSLR statistics, we study here a broader class of PCQP's.

Let $J = \{J(x), -\infty < x < \infty\}$ be absolutely continuous (on finite intervals) and be a difference of two non-decreasing and square integrable (with

respect to F_0 functions, so that

$$(3.1) \quad \delta^2 = \int_{-\infty}^{\infty} J^2(x) dF_0(x) (< \infty) .$$

Further, let $\{d_{N1}, \dots, d_{NN}; N \geq 1\}$ be a triangular array of real numbers satisfying the conditions:

$$(3.2) \quad \sum_{i=1}^N d_{Ni} = 0, \quad \sum_{i=1}^N d_{Ni}^2 = 1 \quad \text{and} \quad \max_{1 \leq i \leq N} |d_{Ni}| \rightarrow 0 \quad \text{as} \quad N \rightarrow \infty .$$

Finally, let $\{Z_N^{(k)}, Q_N^{(k)}, 1 \leq k \leq r\}$ and r be defined as in (1.3) and (1.4), and let

$$(3.3) \quad L_{Nk} = \begin{cases} 0, & k = 0 \\ \sum_{i=1}^k J(Z_{Ni}^{(k)}) [d_{NQ_{Ni}} + \frac{1}{N-k} \sum_{\alpha=1}^k d_{NQ_{N\alpha}}] , & 1 \leq k \leq N-1 \\ L_{NN-1}, & k = N . \end{cases}$$

Our primary concern is to develop an invariance principle for $\{L_{Nk}; 0 \leq k \leq r\}$, and we consider first the case of the null hypothesis (H_0) where the X_i are i.i.d.r.v with an absolutely continuous df F_0 . We denote the expectation and variance under H_0 by E_0 and V_0 , respectively. Let

$$(3.4) \quad \delta_{Nk}^2 = E_0(L_{Nk}^2), \quad 0 \leq k \leq N ,$$

and for every $N (> r \geq 1)$, we consider a stochastic process $W_N = \{W_N(t), t \in I\}$ ($I = [0,1]$) by introducing a sequence of non-decreasing, right-continuous and integer-valued functions $\{k_N(t), t \in I\}$, where

$$(3.5) \quad k_N(t) = \max\{k: \delta_{Nk}^2 \leq t\delta_{Nr}^2\}, \quad t \in I ,$$

and then letting

$$(3.6) \quad W_N(t) = \delta_{Nr}^{-1} L_{Nk_N}(t), \quad t \in I.$$

Note that W_N belongs to the $D[0,1]$ space endowed with the Skorokhod J_1 -topology (for $N=0,1$, $W_N(t)=0, \forall t \in I$). Our primary concern is to show that under suitable regularity conditions,

$$(3.7) \quad W_N \xrightarrow{D} W, \quad \text{in the } J_1\text{-topology on } D[0,1],$$

where $W = \{W(t), t \in I\}$ is a standard Brownian motion in I .

For absolutely continuous F_0 , the α -quantile ξ_α is defined (uniquely) by

$$(3.8) \quad F_0(\xi_\alpha) = \alpha: 0 < \alpha < 1.$$

Let then

$$(3.9) \quad v_\alpha^2 = \int_{-\infty}^{\xi_\alpha} J^2(x) dF_0(x) + (1-\alpha)^{-1} \left(\int_{-\infty}^{\xi_\alpha} J(x) dF_0(x) \right)^2, \quad 0 < \alpha < 1;$$

by (3.1), $v_\alpha^2 < \infty$ for every $0 < \alpha < 1$. First, we consider the following.

Lemma 3.1. Under (3.1), (3.2) and H_0 , as $N \rightarrow \infty$,

$$(3.10) \quad \left[\frac{k}{N} \rightarrow \alpha \right] \Rightarrow E_0(L_{Nk}^2) \rightarrow v_\alpha^2, \quad \forall 0 < \alpha < 1.$$

Proof. Let \mathcal{Q}_N be the set of all possible $(N!)$ permutations of $(1, \dots, N)$. Then, under H_0 ,

$$(3.11) \quad L_{N,N}(z_N^{(N)}, Q_N^{(N)}) = \prod_{i=1}^N f_0(z_{Ni}), \quad \forall Q_N^{(N)} \in \mathcal{Q}_N,$$

and hence $z_N^{(N)}, Q_N^{(N)}$ are stochastically independent with $Q_N^{(N)}$ assuming each permutation of $(1, \dots, N)$ with the same probability $(N!)^{-1}$. This insures that for each $k(=1, \dots, N)$, $Q_N^{(k)}$ is independent of $z_N^{(N)}$ (and

hence, of $Z_N^{(k)}$ when H_0 holds. Thus, proceeding as in the proof of Lemma 2.1, we obtain that

$$(3.12) \quad E_0(L_{Nk} | Z_N^{(k)}) = E_0(L_{Nk}) = 0, \quad 1 \leq k \leq N;$$

$$(3.13) \quad E_0(L_{Nk}^2) = E_0\left\{E_0(L_{Nk}^2 | Z_N^{(k)})\right\} = E_0\left\{V_0(L_{Nk} | Z_N^{(k)})\right\} \\ = E_0\left\{V_0\left[\sum_{i=1}^k J(Z_{Ni}) d_{NQ_{Ni}} + \sum_{i=k+1}^N d_{NQ_{Ni}} \left(-\frac{1}{N-1} \sum_{\alpha=1}^k J(Z_{N\alpha})\right)\right]\right\} \\ = \left\{\frac{N}{N-1} \sum_{i=1}^N \left[d_{Ni} - \frac{1}{N} \sum_{\alpha=1}^N d_{N\alpha}\right]^2\right\} E_0\left\{\frac{1}{N} \sum_{i=1}^k J^2(Z_{Ni}) + \frac{1}{N(N-k)} \left[\sum_{i=1}^k J(Z_{Ni})\right]^2\right\} \\ = \frac{N}{N-1} E_0\left\{\int_{-\infty}^{Z_{Nk}} J^2(x) dS_N(x) + \frac{N}{N-k} \left(\int_{-\infty}^{Z_{Nk}} J(x) dS_N(x)\right)^2\right\}, \quad \text{by (2.9) and (3.2).}$$

Now, $k/N \rightarrow \alpha: 0 < \alpha < 1 \Rightarrow N/(N-k) \rightarrow (1-\alpha)^{-1} < \infty$, and by (2.23)-(2.25),

$$(3.14) \quad \int_{-\infty}^{Z_{Nk}} J^r(x) dS_N(x) \rightarrow \int_{-\infty}^{\xi_\alpha} J^r(x) dF_0(x) \quad \text{a.s., as } N \rightarrow \infty \quad (r=1,2).$$

Finally, for $r=1,2$,

$$(3.15) \quad \left[\int_{-\infty}^{Z_{Nk}} J^r(x) dS_N(x)\right]^{2/r} \leq \left[\int_{-\infty}^{\infty} |J^r(x)| dS_N(x)\right]^{2/r} \\ \leq \int_{-\infty}^{\infty} J^2(x) dS_N(x) = N^{-1} \sum_{i=1}^n J^2(X_i)$$

where under (3.1), $N^{-1} \sum_{i=1}^n J^2(X_i)$ (being a reverse-martingale) is uniformly (in N) integrable. Hence, (3.10) follows from (3.13)-(3.15) and the Dominated Convergence Theorem [cf. Loeve (1963, p. 124)]. Q.E.D.

Let $B_{Nk}^* = B(Z_N^{(k)}, Q_N^{(k)})$ be the σ -field generated by $(Z_N^{(k)}, Q_N^{(k)})$ and $B_{Nk}^* = B(Z_N^{(N)}, Q_N^{(k)})$ be the σ -field generated by $(Z_N^{(N)}, Q_N^{(k)})$, for $k=1,2,\dots,N$.

Lemma 3.2. Under H_0 , $\{L_{Nk}, B_{Nk}^*; 0 \leq k \leq N\}$ (and hence, $\{L_{Nk}, B_{Nk}; 0 \leq k \leq N\}$) are martingales for every $N(\geq 1)$.

Proof. Note that $L_{NN} = L_{NN-1}$, while for $k \leq N-2$, by (3.2),

$$\begin{aligned}
 (3.16) \quad L_{Nk+1} - L_{Nk} &= \sum_{i=1}^k J(Z_{Ni}) \left[\frac{1}{N-k-1} \sum_{\alpha=1}^{k+1} d_{NQ_{N\alpha}} - \frac{1}{N-k} \sum_{\alpha=1}^k d_{NQ_{N\alpha}} \right] \\
 &+ J(Z_{Nk+1}) \left[d_{NQ_{Nk+1}} + \frac{1}{N-k-1} \sum_{\alpha=1}^{k+1} d_{NQ_{N\alpha}} \right] \\
 &= \frac{(N-k)}{(N-k-1)} \left[d_{NQ_{Nk+1}} + \frac{1}{N-1} \sum_{\alpha=1}^k d_{NQ_{N\alpha}} \right] \left[J(Z_{Nk+1}) + \frac{1}{N-k} \sum_{\alpha=1}^k J(Z_{N\alpha}) \right].
 \end{aligned}$$

Since, under H_0 , $Q_N^{(k+1)}$ is independent of $Z_N^{(k+1)}$, while as in (2.18), $E_0[d_{NQ_{Nk+1}} | B_{Nk}^*] = E_0[d_{NQ_{Nk+1}} | Q_N^{(k)}] = -(N-k)^{-1} \sum_{\alpha=1}^k d_{NQ_{N\alpha}}$, it follows from (3.16) and the above that

$$(3.17) \quad E_0[L_{Nk+1} - L_{Nk} | B_{Nk}^*] = 0, \quad \forall 0 \leq k \leq N-2.$$

Thus, $E_0[L_{Nk+1} | B_{Nk}^*] = L_{Nk}$, $\forall 0 \leq k \leq N-2$. Further by (3.17),

$$\begin{aligned}
 (3.18) \quad E_0[L_{Nk+1} | B_{Nk}] &= E_0\{E_0[L_{Nk+1} | B_{Nk}^*] | B_{Nk}\} \\
 &= E_0\{L_{Nk} | B_{Nk}\} = L_{Nk}, \quad \forall 0 \leq k \leq N-2. \quad \text{Q.E.D.}
 \end{aligned}$$

Lemma 3.3. Under (3.1), (3.2) and H_0 , $k/N \rightarrow \alpha: 0 < \alpha < 1$ insures that

$$(3.19) \quad \sum_{s=0}^k E\{(L_{Ns+1} - L_{Ns})^2 | B_{Ns}\} \xrightarrow{P} \nu_\alpha \text{ as } N \rightarrow \infty.$$

Proof. Note that by (3.16) and the stochastic independence of $Q_N^{(N)}, Z_N^{(N)}$, we have for $0 \leq s \leq N-2$,

$$(3.20) \quad E\{(L_{Ns+1} - L_{Ns})^2 | B_{Ns}\} \\ = \left\{ (N-s)/(N-s-1) \right\}^2 \left\{ \sum_{i=s+1}^N \left[d_{NQ_{Ni}} - \frac{1}{N-s} \sum_{\ell=s+1}^N d_{NQ_{N\ell}} \right]^2 \right\} \cdot \\ E_0 \left\{ \left[J(Z_{Ns+1}) + \frac{1}{N-s} \sum_{i=1}^s J(Z_{Ni}) \right]^2 | Z_N^{(s)} \right\} .$$

Now, by (3.1), for every $\eta > 0$, there exists an $\epsilon > 0$, such that

$$(3.21) \quad \int_{-\infty}^{\xi_\epsilon} |J(x)|^r dF_0(x) < \eta \quad \text{for } r = 1, 2 .$$

For $s \leq N\epsilon$, we note that $\sum_{i=s+1}^N \left[d_{NQ_{Ni}} - \frac{1}{N-s} \sum_{\ell=s+1}^N d_{NQ_{N\ell}} \right]^2 \leq \sum_{i=s+1}^N d_{NQ_{Ni}}^2 \leq \sum_{i=1}^N d_{Ni}^2 = 1$, so that

$$(3.22) \quad \sum_{s=0}^{[N\epsilon]} E\{(L_{Ns+1} - L_{Ns})^2 | B_{Ns}\} \\ \leq \frac{1}{N} \left[\sum_{s=0}^{[N\epsilon]} E_0 \left\{ \left[J(Z_{Ns+1}) + \frac{1}{N-s} \sum_{i=1}^s J(Z_{Ni}) \right]^2 | Z_N^{(s)} \right\} \right] \left[1 + O(N^{-1}) \right] ,$$

and proceeding as in (3.14)-(3.15), it can be shown on using (3.21) that the right hand side of (3.22) can be made arbitrarily small, in probability (when $N \rightarrow \infty$).

Note that the conditional pdf of Z_{Ns+1} given $Z_N^{(s)}$ is

$$(3.23) \quad (N-s)f_0(z) [1-F_0(z)]^{N-s-1} / [1-F_0(z_{Ns})]^{N-s}, \quad z_{Ns} \leq z < \infty .$$

It is easy to show that $E[J(Z_{Ns+1}) | Z_N^{(s)}]$ exists for all $0 \leq s \leq N-1$ [under (3.1)] and further by the absolute continuity of $J(x)$ (on finite intervals) and the a.s. convergence of $|z_{Ns} - \xi_{s/N}|$ to 0 for every $N\epsilon \leq s \leq N\alpha$, $\alpha < 1$, it follows as in Theorem 3.1 of Sen (1961) that

$$(3.24) \quad \max_{s: \epsilon \leq \frac{s}{N} \leq \alpha} |E[J(Z_{Ns+1}) | Z_N^{(s)}] - J(z_{Ns})| \rightarrow 0 \quad \text{a.s., as } N \rightarrow \infty .$$

Let us then denote by

$$(3.25) \quad U_{Ns} = (N-s)^{-1} \sum_{i=1}^N d_{NQ_{Ni}} = -(N-s)^{-1} \sum_{i=1}^s d_{NQ_{Ni}}, \quad 1 \leq s \leq N-1,$$

$$(3.26) \quad \tilde{U}_{Ns} = (N-s)^{-1} \sum_{i=1}^N d_{NQ_{Ni}}^2, \quad s = 1, \dots, N-1.$$

It follows from (2.19) that $\{U_{Ns}, B(Q_N^{(s)}); 1 \leq s \leq N-1\}$ is a martingale when H_0 holds and as in (2.14) and (2.26),

$$(3.27) \quad \max_{s \leq N\alpha} |U_{Ns}| = O_p(N^{-1}) \quad \text{for every } 0 < \alpha < 1.$$

Also, note that

$$\begin{aligned} (3.27) \quad E_0[\tilde{U}_{Ns+1} | B(Q_N^{(s)})] &= (N-s-1)^{-1} \{ (N-s)\tilde{U}_{Ns} - E_0[d_{NQ_{Ns+1}}^2 | B(Q_N^{(s)})] \} \\ &= (N-s-1)^{-1} \{ (N-s)\tilde{U}_{Ns} - \frac{1}{N-s} \sum_{\alpha=s+1}^N d_{NQ_{N\alpha}}^2 \} \\ &= \tilde{U}_{Ns}, \quad 1 \leq s \leq N-2. \end{aligned}$$

Using the martingale property in (3.27) and the Kolmogorov inequality, we obtain that under H_0 , for every $\epsilon > 0$

$$(3.28) \quad P\left\{ \max_{1 \leq s \leq k} |N\tilde{U}_{Ns} - 1| > \epsilon \right\} \leq \epsilon^{-2} E_0\{N\tilde{U}_{Nk} - 1\}^2,$$

where

$$\begin{aligned} (3.29) \quad E_0\{N\tilde{U}_{Nk} - 1\}^2 &= E_0\left\{ \frac{N}{N-k} \sum_{s=k+1}^N \left[d_{NQ_{Ns}}^2 - \frac{1}{N} \right] \right\}^2 \\ &= \frac{N^2}{(N-k)^2} \frac{k(N-k)}{N(N-1)} \sum_{i=1}^N \left(d_{Ni}^2 - \frac{1}{N} \right)^2 \\ &= [Nk/(N-1)(N-k)] \left\{ \sum_{i=1}^N d_{Ni}^4 - \frac{1}{N} \right\} \\ &\leq [Nk/(N-1)(N-k)] \left\{ \left(\max_{1 \leq i \leq n} d_{Ni}^2 \right) \sum_{i=1}^N d_{Ni}^2 + \frac{1}{N} \right\} \end{aligned}$$

$\rightarrow 0$ by (3.2) and the fact that $k/N \rightarrow \alpha: 0 < \alpha < 1$, as $N \rightarrow \infty$.

Thus, from (3.28) and (3.29), we have under H_0 ,

$$(3.30) \quad \max_{1 \leq s \leq k} |\tilde{U}_{Ns} - \frac{1}{N}| = o_p(1), \text{ as } N \rightarrow \infty.$$

From (3.20) through (3.30), we obtain that for $k/n \rightarrow \alpha: 0 < \alpha < 1$,

$$\begin{aligned} (3.31) \quad & \sum_{s=1}^k E_0 \{ (L_{Ns+1} - L_{Ns})^2 | \mathcal{B}_{Ns} \} \\ &= N^{-1} \left[\sum_{s=1}^k \left\{ J(Z_{Ns}) + \frac{1}{N-s} \sum_{i=1}^s J(Z_{Ni}) \right\}^2 \right] [1 + o_p(1)] \\ &= N^{-1} \left[\sum_{s=1}^k J^2(Z_{Ns}) (N-s+1)/(N-s) + \right. \\ & \quad \left. \sum_{i=1}^k \sum_{j=1}^k J(Z_{Ni}) J(Z_{Nj}) ((N-ivj)^{-1} + \sum_{s=ivj}^k (N-s)^2) \right] [1 + o_p(1)] \\ &= N^{-1} \left\{ \sum_{s=1}^k J^2(Z_{Ns}) + \frac{1}{N-k} \left(\sum_{i=1}^k J(Z_{Ni}) \right)^2 \right\} \left\{ 1 + o_p(N^{-1}) \right\} \left\{ 1 + o_p(1) \right\} \\ &= \left[\int_{-\infty}^{Z_{Nk}} J^2(x) dS_N(s) + \frac{N}{N-k} \left[\int_{-\infty}^{Z_{Nk}} J(x) dS_N(x) \right]^2 \right] \left\{ 1 + o_p(1) \right\}, \end{aligned}$$

while

$$(3.32) \quad E_0(L_{N1}^2 | \mathcal{B}_{N0}) = E_0(L_{N1}^2) = N^2(N-1)^{-2} \frac{1}{N} E_0 J^2(Z_{N1}) \rightarrow 0 \text{ as } N \rightarrow \infty.$$

Hence, by (2.23) and (3.14), the right hand side of (3.31) converges (in probability) to v_α^2 , as $N \rightarrow \infty$, and the proof of the lemma then follows from (3.31) and (3.32). Q.E.D.

Remark. Note that in (3.32), $N^{-1} E_0 J^2(Z_{N1})$ follows from the fact that

$$\begin{aligned} (3.33) \quad N^{-1} E_0 J^2(Z_{N1}) &\leq \max_{1 \leq i \leq N} N^{-1} E_0 J^2(Z_{Ni}) \\ &\leq \max_{1 \leq i \leq N} N^{-1} E_0 J^2(X_i) \rightarrow 0, \text{ as } N \rightarrow \infty, \end{aligned}$$

where the last step follows by standard arguments under (3.1).

Let now $I(A)$ denote the indicator function of the set A .

Lemma 3.4. For every $\varepsilon' > 0$, $k/N \rightarrow \alpha$: $0 < \alpha < 1$, as $N \rightarrow \infty$

$$(3.34) \quad \sum_{s=0}^k E_0 \{ (L_{Ns+1} - L_{Ns})^2 I(|L_{Ns+1} - L_{Ns}| > \varepsilon') | B_{Ns} \} \xrightarrow{P} 0 .$$

Proof. We breakup the sum into two subsets $\{s \leq N\varepsilon\}$ and $\{N\varepsilon < s \leq k\}$. Then, by arguments similar to that in (3.22),

$$(3.35) \quad \begin{aligned} & \sum_{s=0}^{[N\varepsilon]} E_0 \{ (L_{Ns+1} - L_{Ns})^2 I(|L_{Ns+1} - L_{Ns}| > \varepsilon | B_{Ns}) \} \\ & \leq \sum_{s=0}^{[N\varepsilon]} E_0 \{ (L_{Ns+1} - L_{Ns})^2 | B_{Ns} \} \xrightarrow{P} 0 . \end{aligned}$$

Since $J(x)$ is the difference of two non-decreasing, absolutely continuous and square integrable functions, it can be shown easily that for every $0 < \varepsilon < \alpha < 1$, there exists a $C = C(\varepsilon, \alpha)$ ($< \infty$), such that

$$(3.36) \quad \max_{s: \varepsilon \leq \frac{s}{N} \leq \alpha} |J(Z_{Ns+1}) + \frac{1}{N-s} \sum_{i=1}^s J(Z_{Ni})| \leq C \text{ a.s., as } N \rightarrow \infty .$$

On the other hand, by (3.2), for every $Q_N^{(N)} \in Q_N$,

$$(3.37) \quad \begin{aligned} & \max_{1 \leq i \leq N} \left\{ \left| d_{NQ_{Ni}} + \frac{1}{N-i} \sum_{s=1}^i d_{NQ_{Ns}} \right| = \max_{1 \leq i \leq N} \left| d_{NQ_{Ni}} - \frac{1}{N-i} \sum_{s=i+1}^N d_{NQ_{Ns}} \right| \right\} \\ & \leq 2 \left\{ \max_{1 \leq i \leq N} |d_{Ni}| \right\} \rightarrow 0 \text{ as } N \rightarrow \infty . \end{aligned}$$

Hence, for every C and $\varepsilon' > 0$, there exists an integer $N_0 (= N_0(\varepsilon', C))$, such that

$$(3.38) \quad P \left\{ \max_{s \leq N\alpha} \left| d_{NQ_{Ns}} + \frac{1}{N-s} \sum_{i=1}^s d_{NQ_{Ni}} \right| < \varepsilon'/C \right\} = 1, \quad \forall N \geq N_0 .$$

From (3.16), (3.36) and (3.38), it follows that for $N \geq N_0$,

$$(3.39) \quad \sum_{s=[N\epsilon]+1}^k E_0 \{ (L_{Ns+1} - L_{Ns})^2 I(|L_{Ns+1} - L_{Ns}| > \epsilon) | \mathcal{B}_{Ns} \} \stackrel{P}{\rightarrow} 0,$$

and (3.34) follows from (3.35) and (3.39), Q.E.D.

We are now in a position to formulate and prove our main theorem of this section.

Theorem 1. Under (1.3), (3.1) and (3.2), when the X_i are i.i.d.r.v with an absolutely continuous df F_0 , (3.7) holds.

Proof. By virtue of the martingale property of $\{L_{Nk}\}$, when H_0 holds, we are in a position to use Theorem 2 of Scott (1973), and to prove the theorem, all we need to show that

$$(3.40) \quad \delta_{Nr}^{-2} \sum_{i=1}^{k_N(t)} V_0 [L_{Ni} | \mathcal{B}_{Ni-1}] \stackrel{P}{\rightarrow} t, \quad \text{as } N \rightarrow \infty \quad (0 < t < 1),$$

$$(3.41) \quad \delta_{Nr}^{-2} \sum_{i=1}^r E_0 \{ [L_{Ni} - L_{Ni-1}]^2 I(|L_{Ni} - L_{Ni-1}| > \epsilon) | \mathcal{B}_{Ni-1} \} \stackrel{P}{\rightarrow} 0 \quad (\forall \epsilon > 0),$$

where $k_N(t)$ and r are defined by (3.5) and (1.3), respectively. Now, (3.41) follows directly from Lemmas 3.1 and 3.4 (where we note that (1.3) insures that $0 < p = \alpha < 1$). Since, by (3.4), (3.5) and Lemma 3.1,

$$(3.42) \quad \delta_{Nk_N}^2(t) \rightarrow tv_p^2, \quad \delta_{Nr}^2 \rightarrow v_p^2, \quad \text{as } N \rightarrow \infty,$$

the proof of (3.40) follows from (3.42) and Lemma 3.3. Q.E.D.

Remarks. The condition that J is the difference of two non-decreasing functions, through quite general, can be dispensed at the cost of strengthening (3.1) to

$$(3.43) \quad \int_{-\infty}^{\infty} |J(x)|^m dF_0(x) < \infty \quad \text{for some } m > 2.$$

In that case, in Lemma 3.4, a Liapounoff-type condition can be obtained (which implies (3.41), and the rest of the proof remains the same. Secondly, by (1.3), we have limited ourselves to $0 < p < 1$. Though p may be arbitrarily close to 1, there are a few technical barriers for allowing p to be equal to one. Note that v_p^2 may tend to ∞ as $p \rightarrow 1$ (viz., $J(x) = 1$, $\forall x \Rightarrow v_p^2 = p/(1-p)$). If, however, we impose the additional condition [as before (2.8)] that

$$(3.44) \quad \int_{-\infty}^{\infty} J(x) dF_0(x) = 0,$$

we obtain from (3.9) and (3.44) that for every $0 < p < 1$,

$$(3.45) \quad v_p^2 = \int_{-\infty}^{\xi_p} J^2(x) dF_0(x) + (1-p)^{-1} \left\{ \int_{\xi_p}^{\infty} J^2(x) dF_0(x) \right\}^2 \\ \leq \int_{-\infty}^{\xi_p} J^2(x) dF_0(x) + \int_{\xi_p}^{\infty} J^2(x) dF_0(x) = \delta^2 \quad \text{and} \quad \lim_{p \rightarrow 1} v_p^2 = \delta^2.$$

Even so, $\int_{-\infty}^{\infty} J(x) dS_N(x)$ is not necessarily equal to 0; in fact, it is $O_p(N^{-1/2})$. Further, $\left\{ \max_{1 \leq i \leq N} J^2(Z_{Ni}) \right\} = E_0 \left\{ \max_{1 \leq i \leq N} J^2(X_i) \right\} = o(N)$, under (3.9), while under (3.9) and (3.44), $E_0 \left\{ \left(\int_{-\infty}^{\infty} J(x) dS_N(x) \right)^2 \right\} = N^{-1} \delta^2$. Hence, it follows from (3.13) that for every (fixed $s \geq 1$), as $N \rightarrow \infty$,

$$(3.46) \quad E_0(L_{NN-s}^2) \rightarrow \delta^2 + s^{-1} \delta^2 = (1+s^{-1}) \delta^2 > \delta^2 = v_1^2.$$

This apparent anomaly can be straightened out with the help of (2.8) and (2.10). Note that if g satisfies (3.1), then for every (fixed $s \geq 1$), as $N \rightarrow \infty$,

$$(3.47) \quad \bar{G}(Z_{NN-s}) = o_p([N/s]^{1/2}) \quad \text{while} \quad \left| \int_{-\infty}^{Z_{NN-s}} g(x) dS_N(x) \right| = O_p(N^{-1/2}).$$

Thus, whereas (2.8) relates to a term (for $x = Z_{NN-s}$) $o_p([N/s]^{1/2})$, (2.10)

leads us to a term $O_p((N/s)N^{-1/2}) = O_p(N^{1/2}/s)$, and hence, $\bar{G}(Z_{NN-s})$ and $\bar{G}_N(Z_{NN-s})$ are of different (stochastic) orders of magnitude, and the stochastically larger order of magnitude from $\bar{G}_N(Z_{NN-s})$ pushes up the variance of L_{NN-s} ; in fact, here (2.24) and hence, (2.27) may not hold. But, if $s = s(N)$ be any (slowly varying) function with $\lim_{N \rightarrow \infty} s(N) = \infty$, it can be shown that

$$(3.48) \quad E_0(L_{NN-s}^2) = \delta^2(1 + 1/s(N) + o(1)) \rightarrow \delta \quad \text{as } N \rightarrow \infty.$$

Thus, as regards (2.27), we can proceed as follows. First, doing the same line (of proof) as in Lemma 3.1 of Sen (1976), it can be shown that under H_0 ,

$$(3.49) \quad \{T_{Nk}, B_{Nk}; 0 \leq k \leq N\} \text{ is a martingale.}$$

while for $\eta > 0$, arbitrarily small, on letting $r_N = [N(1-\eta)]$, it can be shown that $E_0(T_N^2) = v_1^2$ and $E_0(T_{Nr_N}^2) \rightarrow v_{1-\eta}^2$, so that by the Kolmogorov-inequality for martingales, for every $\epsilon > 0$,

$$(3.50) \quad P\left\{ \max_{r_N \leq k \leq N} |T_{Nk} - T_{Nr_N}| > \epsilon \right\} \leq \epsilon^{-2} (T_N - T_{Nr_N})^2 \\ = \epsilon^{-2} [v_1^2 - v_{1-\eta}^2 + o(1)],$$

which can be made smaller than any given $\delta (> 0)$ by choosing $\eta (> 0)$ sufficiently small [and noting that as $\int gdF_0 = 0$, by (3.4.5), $\lim_{\eta \rightarrow 0} v_{1-\eta}^2 = v_1^2$]. On the other hand, for $r \leq r_N = [N(1-\eta)]$, $\eta > 0$, we are in a position to use (2.27), so that the invariance principle for $\{T_{Nk}^*; 0 \leq k \leq r_N\}$ leads us to the same for $\{T_{Nk}; 0 \leq k \leq r_N\}$, and this along with (3.50) yields the desired result for the entire sequence $\{T_{Nk}; 0 \leq k \leq N\}$. In a similar manner, by the

martingale property (Lemma 3.2) of $\{L_{Nk}\}$ and (3.48), for any slowly varying $\{s(N)\}$, we can replace $\{L_{Nk}; 0 \leq k \leq N-s(N)\}$ by an appropriate $\{L_{Nk}; 0 \leq k \leq N(1-\eta)\}$ ($\eta > 0$) and apply our Theorem 1. In view of the fact that $E_0(L_{NN} - L_{NN-s(N)})^2 \rightarrow \delta^2$, (not to 0), we are, however, unable to replace $N-s(N)$ by N in this case. In actual practice, PCS mostly involves a terminal censoring number (r) corresponding to a value of p quite below 1, and hence, this technicality is not of much concern to us.

Let us now proceed on to the non-null case. We shall confine ourselves to local (contiguous) alternatives where parallel results can be derived and these will be incorporated in the next section for the study of asymptotic power of some PCS tests based on such PCQP's. Consider a triangular array $\{X_{Ni}, 1 \leq i \leq N; N \leq 1\}$ of (row-wise) independent rv's and assume that X_{Ni} has an absolutely continuous df F_{Ni} with an absolutely continuous pdf f_{Ni} and

$$(3.51) \quad f_{Ni}(x) = F(x; \Delta c_{Ni}^*) , \quad -\infty < x < \infty , \quad i = 1, \dots, N .$$

where f , Δ and c_{Ni}^* are all defined as in the beginning of Section 2. Note that, in (3.51), Δ is regarded as fixed while by (2.4), the c_{Ni}^* all go to 0 as $N \rightarrow \infty$. We denote such a sequence of alternative hypotheses by $\{H_N\}$, while H_0 relates to $\Delta = 0$. Our concern is to study the weak convergence of $\{W_N\}$, defined by (3.5)-(3.6), when $\{H_N\}$ hold.

We define the d_{Ni} as in (3.2), the c_{Ni}^* as in (2.4), and assume that they satisfy the limits

$$(3.52) \quad \lim_{N \rightarrow \infty} \sum_{i=1}^N d_{Ni} c_{Ni}^* = \rho^* (-1 \leq \rho^* \leq 1) , \quad \max_{1 \leq i \leq N} |c_{Ni}^*| \rightarrow 0 ;$$

in fact, for $d_{Ni} = c_{Ni}^*$, $1 \leq i \leq N$, $\rho^* = 1$ (by (2.4)), For every $t \in [0,1]$ and $0 < p < 1$, we define

$$(3.53) \quad \alpha(t,p) = \max\{\alpha: v_\alpha^2 \leq tv_p^2\}, \quad t \in [0,1],$$

where v_α^2 is defined by (3.9). Note that $v_{\alpha(t,p)}^2$ is \nearrow in $t \in (0,1)$ and $v_{\alpha(0,p)}^2 = 0$, $v_{\alpha(1,p)}^2 = v_p^2$, so that $\alpha(0,p) = 0$ and $\alpha(1,p) = p$. Here also, we denote $F(x;0)$ and $f(x;0)$ by F_0 and f_0 , respectively, and $g(x)$ and $\bar{G}(x)$ as in (2.1). Further, we define

$$(3.54) \quad J^*(x) = [1 - F_0(x)]^{-1} \int_{-\infty}^x J(y) dF_0(y), \quad -\infty < x < \infty;$$

$$(3.55) \quad \zeta_t^{(p)} = \left[\int_{-\infty}^{\xi_\alpha} J(x)g(x)dF_0(x) - (1-\alpha)^{-1}J^*(\xi_\alpha)\bar{G}(\xi_\alpha) \right]_{\alpha=\alpha(t,p)}, \quad t \in I.$$

We also assume that the pdf $f(x;\theta)$ is absolutely continuous in $\theta \in (0,1)$ for almost all x , $(\partial/\partial\theta)f(x;\theta) = f'_\theta(x;\theta)$ exists and converges to $f'_\theta(x;0)$ as $\theta \rightarrow 0$, and further, defining $g(x)$ as in Section 2 and letting $F_0(x) = F(x;0)$, we assume that

$$(3.56) \quad \lim_{\theta \rightarrow 0} \int_{-\infty}^{\infty} [f'_\theta(x;\theta)]^2 [f(x,\theta)]^{-1} dx = \int_{-\infty}^{\infty} g^2(x) dF_0(x) < \infty.$$

Finally, let us denote by

$$(3.57) \quad \mu = \{\mu(t) = \Delta\rho^*\zeta_t^{(p)}/v_p, t \in I\}$$

and note that by assumptions made on J and g , $\mu \in C[0,1]$ space. Then, we have the following.

Theorem 2. Let $\{W_N\}$ and W be defined as in (3.5)-(3.7). Then, under (1.3), (3.2), (3.52), (3.56) and $\{H_N\}$ in (3.51), as $N \rightarrow \infty$,

$$(3.58) \quad W_N - \mu \xrightarrow{D} W, \text{ in the } J_1\text{-topology on } D[0,1]$$

Proof. Let P_N and P_N^* be respectively the joint df of $(Z_N^{(N)}, Q_N^{(N)})$ when H_0 (i.e. $F_i = F_0, \forall i = 1$) and H_N in (3.51) hold. Then, under (3.2), (3.52) and the assumed regularity conditions on f , it can be shown [cf. Hájek and Sidák (1967, pp. 239-240)] that $\{P_N^*\}$ is *contiguous* to $\{P_N\}$. For $x \in D[0,1]$ and $\delta \in (0,1)$, let us define

$$(3.59) \quad \omega_\delta(x) = \sup\{\min\{|x(t)-x(s)|, |x(s)-x(u)|\}: 0 \leq u < \delta < t \leq u + \delta \leq 1\}.$$

Since, $W_N(0) = 0$, with probability 1, and, by Theorem 1, under $H_0, \{W_N\}$ is *tight*, it follows that

$$(3.60) \quad \lim_{\delta \rightarrow 0} \overline{\lim}_N P\{\omega_\delta(W_N) > \varepsilon | H_0\} = 0, \quad \forall \varepsilon > 0.$$

Also, W_N is a mapping of $(Z_N^{(N)}, Q_N^{(N)})$ into the space $D[0,1]$. Hence, by the contiguity of $\{P_N^*\}$ to $\{P_N\}$ and (3.60), we conclude that

$$(3.61) \quad \lim_{\delta \rightarrow 0} \overline{\lim}_N P\{\omega_\delta(W_N) > \varepsilon | H_N\} = 0, \quad \forall \varepsilon > 0,$$

that is $\{W_N\}$ remains tight under $\{H_N\}$. Thus, to prove (3.58), we need to establish only the convergence of the finite dimensional distributions of $\{W_N - \mu\}$ to the corresponding ones of W .

For this purpose, for any $k: k/N \rightarrow \alpha: 0 < \alpha \leq p < 1$, we rewrite L_{Nk} as

$$(3.62) \quad L_{Nk} = \sum_{i=1}^N d_{Ni} J(X_i) I(X_i \leq Z_{Nk}) = \left\{ \frac{1}{N-k} \sum_{i=1}^k J(Z_{Ni}) \right\} \left\{ \sum_{i=1}^N d_{Ni} I(X_i \leq Z_{Nk}) \right\},$$

where $I(A)$ stands for the indicator function of the set A . Defining ξ_α and J^* as in (3.8) and (3.54), we introduce

$$(3.63) \quad L_{Nk}^* = \sum_{i=1}^N d_{Ni} J(X_i) I(X_i \leq \xi_\alpha) + J^*(\xi_\alpha) \sum_{i=1}^N d_{Ni} I(X_i \leq \xi_\alpha).$$

If we write $S_N(\xi_\alpha) = N^{-1}k_N$, then by (3.62), (3.63) and the definition of $Q_N^{(k)}$, we have

$$(3.64) \quad L_{Nk}^* = L_{Nk_N} + \left\{ J^*(\xi_\alpha) - \frac{1}{N-k_N} \sum_{i=1}^{k_N} J(Z_{Ni}) \right\} \left\{ \sum_{i=1}^{k_N} d_{NQ_{Ni}} \right\}.$$

Note that $N^{-1}k_N \rightarrow \alpha$, in probability, under H_0 [viz., (2.23)], so that by the same technique as in Lemma 2.2,

$$(3.65) \quad \left| J^*(\xi_\alpha) - (N-k_N)^{-1} \sum_{i=1}^{k_N} J(Z_{Ni}) \right| \xrightarrow{P} 0, \text{ under } H_0,$$

while by (2.22), $\left| \sum_{i=1}^{k_N} d_{NQ_{Ni}} \right| = O_p(1)$, under H_0 . Hence, the second term on the right hand side of (3.64) converges in probability to 0 as $N \rightarrow \infty$ when H_0 holds. Further, by the martingale property (Lemma 3.2), we have by the Kolmogorov-inequality,

$$(3.66) \quad P \left\{ \max_{q: |k-q| < \delta N} |L_{Nq} - L_{Nk}| > \varepsilon \mid H_0 \right\} \rightarrow 0 \text{ as } \delta \downarrow 0,$$

and hence, noting that $|N^{-1}k_N - \alpha| \xrightarrow{P} 0$ and $k/N \rightarrow \alpha: 0 < \alpha < 1$, we obtain from the above that

$$(3.67) \quad L_{Nk}^* - L_{Nk} \xrightarrow{P} 0 \text{ as } N \rightarrow \infty \text{ when } H_0 \text{ holds.}$$

Again, by virtue of the contiguity of $\{P_N^*\}$ to $\{P_N\}$ and (3.67), we conclude that as $N \rightarrow \infty$,

$$(3.68) \quad L_{Nk}^* - L_{Nk} \xrightarrow{P} 0 \text{ under } \{H_N\} \text{ as well.}$$

Thus, for finitely many k 's, say, $k_1 \leq \dots \leq k_m$, $m(\geq 1)$ given, satisfying

$$(3.69) \quad N^{-1}k_j \rightarrow \alpha(t_j, p), \quad 0 \leq t_1 < \dots < t_m \leq 1,$$

to study the joint distribution of $W_N(t_1), \dots, W_N(t_m)$, it suffices to consider the joint df of $L_{Nk_1}^*, \dots, L_{Nk_m}^*$. Since the L_{Nk}^* involve a sum over independent random variables, by the classical (multivariate version of the) central limit theorem, it follows that under (3.1), (3.2), (3.51), (3.52) and (3.56), $[L_{Nk_1}^*, \dots, L_{Nk_m}^*]$ converges in law to a multivariate normal distribution with mean vector $[\mu(t_1), \dots, \mu(t_m)]v_p$ and dispersion matrix $v_p^2((t_j \wedge t_\ell))_{j, \ell=1, \dots, m}$ which conforms to the desired pattern. Q.E.D.

Remarks. In (2.27), we have proved the stochastic equivalence of $\{T_{Nk}; 0 \leq k \leq r\}$ and $\{T_{Nk}^*; 0 \leq k \leq r\}$ when H_0 and (1.3) hold. Here also, we can proceed on the sameline as in (3.62)-(3.64) and use the contiguity of $\{P_N^*\}$ to $\{P_N\}$ to show that (2.27) remains true when $\{H_N\}$ [in (3.51), (3.52) and (3.56)] holds along with (1.3).

One could have extended the results of Sen (1976) to the current setup of non-identically distributed rv's. However, that would have induced more complications in the proof along with some extra (mild) regularity conditions [viz., (2.8) and (3.36) of Sen (1976)]. The current approach provides an alternative and simple solution.

4. APPLICATIONS TO TIME-SEQUENTIAL TESTS BASED ON PCQP

A variety of rank based PCS tests is available in the literature. Hájek (1963) has developed the asymptotic theory of Kolmogorov-Smirnov (KS-) type tests for regression alternatives, and his results can be adapted readily in a PCS provided we let $r/N \rightarrow 1$. The simple limiting null distributions of these KS-type statistics [viz., (3) and (4) on

page 189 of Hájek and Sidák (1967)] are not valid if $r/N \rightarrow p$; $0 < p < 1$. However, some recent tabulations of the critical values of the truncated KS-type tests by Koziol and Byar (1975) provides us with these (approximate) critical values. In Chapters V and VI of Hájek and Sidák (1967), some related tests are also considered; in particular, the Rényi-type and Cramer-von Mises type tests for regression alternatives deserve mention and they can also be adapted in a PCS when we let $r/N \rightarrow 1$. Again, for $r/n \rightarrow p$; $0 < p < 1$, the limiting distributions of these statistics are no longer very simple and extensive simulation studies are being made to provide approximate critical values in such cases. Chatterjee and Sen (1976) have studied the weak convergence of PC linear rank statistics to a Brownian motion and their procedure can be used for any $r/N \rightarrow p$; $0 < p \leq 1$ with simple limiting null distribution theory provided by them. Usually, their procedure is better than the Hájek's ones. All these procedures share one common feature: namely, they are based solely on the vector $Q_N^{(r)}$, disregarding any information contained in the vector $Z_N^{(r)}$ of associated order statistics. Thus, it is quite intuitive to extract this information and in Section 2, we have shown that a PCPLR statistics sequence relates to PCQP's which again can be approximated by more convenient linear combinations of functions of order statistics with stochastic coefficients. Thus, in the same spirit as in Chatterjee and Sen (1973) and Sen (1976), we may be interested in employing the process W_N , defined by (3.5)-(3.6) and use as a test-statistic

$$(4.1) \quad M_N = M(W_N)$$

where $M(x) = M(x(t); 0 \leq t \leq 1)$ is a suitable functional. For example, we

may take the KS-type statistics as

$$(4.2) \quad M_N^+ = \sup_{0 \leq t \leq 1} W_N(t) = \max_{0 \leq k \leq r} L_{Nk} / \nu_p,$$

$$(4.3) \quad M_N^- = \sup_{0 \leq t \leq 1} |W_N(t)| = \max_{0 \leq k \leq r} |L_{Nk}| / \nu_p$$

and obtain the limiting distributions of M_N^+ or M_N^- with the aid of our Theorem 1 and the well known distributional results on $\sup_{0 \leq t \leq 1} W(t)$ or $\sup_{0 \leq t \leq 1} |W(t)|$. Theorem 2 provides us with the asymptotic power of such a test. We may also consider other functional (such as the Rényi-type or Cramér-von Mises type) of W_N and propose the same as test-statistics. This leads us to the study of the asymptotic behavior of different functionals of PCQP's with different $\{J\}$, and will be studied in a subsequent paper.

REFERENCES

- [1] Basu, A.P., and Borwonkar, J.D. (1970). Some asymptotic results, based on censored data, for the maximum likelihood estimators of parameters subject to restraints. *Sankhyā, Ser. A* 33, 35-44.
- [2] Chatterjee, S.K., and Sen, P.K. (1973). Nonparametric testing under progressive censoring. *Calcutta Statist. Assoc. Bull.* 22, 13-50.
- [3] Chow, Y.S. (1960). A martingale inequality and the law of large numbers. *Proc. Amer. Math. Soc.* 11, 107-111.
- [4] Hájek, J. (1963). Extension of the Kolmogorov-Smirnov test to regression alternatives. Proc. Bernoulli-Bayes-Laplace Seminar, Berkeley (ed. L. LeCam).
- [5] Hájek, J., and Sidák, Z. (1967). *Theory of Rank Tests*. Academic Press, New York.
- [6] Koziol, J.A., and Byar, D.P. (1975). Percentage points of the asymptotic distributions of the one and two-sample K-S tests for truncated or censored data. *Technometrics* 17, 507-510.

- [7] Loève, M. (1963). *Probability Theory*. Van Nostrand, Princeton (2nd Ed.)
- [8] Scott, D.J. (1973). Central limit theorems for martingales and for processes with stationary increments using a Skorokhod representation approach. *Adv. Appl. Prob.* 5, 119-137.
- [9] Sen, P.K. (1961). A note on the large sample behavior of extreme values from distributions with finite end points. *Calcutta Statist. Asso. Bull.* 10, 106-115.
- [10] Sen, P.K. (1976). Weak convergence of progressively censored likelihood ratio statistics and its role in asymptotic theory of life testing. *Ann. Statist.* 4, No. 6, in press.