AD-752 712

DETECTING A POINT RADIATOR IN THE PRESENCE OF NONGAUSSIAN INTERFERENCE

M. Kh. Reznik

and the state of the

Naval Intelligence Support Center Washington, D. C.

16 November 1972

DISTRIBUTED BY:

National Technical Information Service U. S. DEPARTMENT OF COMMERCE 5285 Port Royal Road, Springfield Va. 22151



TRANSLATOR: DWM

NATIONAL TECHNICAL INFORMATION SERVICE U S Deportment of Commettee Springfield VA 22151

NISC TRANSLATION NO. 3366

APPROVEI <u><u><u>p</u></u> K</u>

DATE _____ 16 Nov mber 1972____

DETECTING A POINT RADIA FOR IN THE PRESENCE OF NONGAUSSIAN INTERFERENCE

SCOPERSON STREET

[Reznik, M. Kh., Obnaruzheniye Tochechnogo Izlychatelya v Prisutstvii Fonovykh Pomekh Negaussovogo Tip., Optiko-mekhanicheskaya promyshlennost', No. 3, 1972, pp. 3-6, Rus ian]

An optimal method of detecting a pulse signal in the presence of background interference of the kind presented by cloudy sky radiation characterized by a nongaussian distribution law is described.

Many problems in optical lo ation, navigation direction finding, and the like, are connected with the detection of signal from a point radiator against a background of natural formation:.

In works known to date [1-3] the effect of background interference was taken into account within the framework of cor elation analysis, leading to linear methods of analyzing the electri al signal arising in the radiation receiver. Experience in the use f linear systems shows that in the case where signals are received rom point objects, i.e., signals having the form of Dirac of -functions, linear systems are least protected from the effects of interference whose values change spasmodically.

The construction of a statistically optimal s stem consists first of all in the selection of a background model that reflects its spasmodic structure. The simplest such model is a background model in the form of a population of spots (clouds) randomly distributed against a "clear sky" and having random (but constant within the limits of each spot) brightness values. We shall assume that in any cross section of such a background the distribution of boundaries (transitions from dark to light and back again) is similar, i.e., that the background is homogeneous and isotropic, and that the boundary distribution is Poissonian, i.e.,

$$p(n, r) = \frac{1}{n!} (\nu r)^n \exp(-\nu n), \qquad (1)$$

where p(n,r) is the probability that in the interval r w ll be equal to n intersections; v is the average frequency of intersoctions in a single segment.

It is useful to distinguish this model from the anisotropic twolevel background model in the form of a "chess board," examined in works [3-4].



[3

We shall first examine the interference signals that form during search scanning (in the case of uniform displacement of the radiation receiver along the x axis) when a background inhomogeneity boundary [4 is intersected in the assumption of ideal optics. The shortcomings of optics, or aberration distortions, we shall consider later.

THE REAL PROPERTY OF

なず

We shall assume that the radiation receiver is in the form of a Hxh (H>h) rectangle whose angular dimensions are much less than the mean angular dimensions of the background inhomogeneity, i.e.,

v// 1. (2)

Granting this, the inhomogeneity boundary can be prese ted in the form of a direct boundary with inclination θ to the scannin direction. The signal received from the interference φ (x), received as the result of boundary intersections, for a considerable part of its change is described as a linear function (as shown in figure 1, wit receiver displaced from position 1 to position 2). This makes it ossible to approximate its increase or decrease with a linear funct on threighout the entire sector of change. The signal from the poin object has the form of a rectangular pulse having width h.

We shall eliminate the constant component present in the interference either by means of linear transformation or by means of a second receiver located beside the first and connected in an opposite way. Therefore the resulting signal from the interference $\Phi_1(x) = \varphi(x) = \varphi(x - h)$ is realized in a form consisting of nonintersecting isosceles trapezoids located along the abscissa (capable of degenerating into triangles) of positive and negative polarity, random extension as i amplitude (but limited on top along the modulus, since maximal background brightness drops are limited) and randomly separate from each other. It is important to note that the frontal lengths of the trape oidal formations are precisely equal to the width of the rectangula pulse.

The signal from the point object is obtained in the form of rectangular pulses of different polarity having width h a jacent in the base.

The effect of aberration distortions manifests itself in the "smoothing" of the processes obtained. As a result of this "smoothing," the rectangular pulse assumes a bell-shaped form. The function describing its change as a function of the time argument is designated s (t) and it is assumed that at $t_1 > h/2 \sin(t) = 0$ Aberration distortions of opt is can be treated as a kind of linear transformation, consequently, the "smoothed" interference fronts will as before have a length equal to that of signal d' and be described by functions that are integrals of s (t).



Non-one-section

1. N. I. I. I.



The random interference time function Φ (t) can be presented in the form Φ (t) = F(t)A(t), where function F(t), normed to absolute maximum value, is the same type of function as Φ (t), while A(t) > 0 is a spasmodic process with a uniform distribution of values in segment $[o, A_{\max}]$ that changes its values only at those time moments when process F(t) deviates from the zero state. For ease of later computations, we shall consider that process A(t) assumes only a finite number of values $0 < a_1 < a_2 < \ldots < a_n = A_{\max}$; $\mu \gg 1$ is a whole number and

$$P(A(t) = a_i) = \frac{1}{p}; i = 1, 2 \dots p.$$
 (3)

One of the possible realizations of interference $\phi(t)$:ogether with functions A(t), F(t), and s_i(t) is depicted in figure 2.

The signal from the point object will become bipolar and will be determined by function ε s(t), where ε is the signal amplitude unknown to the observer.

Radiation receiver noise is considered white and normal.

Since we are assuming that the radiator, subject to detection, is a point radiator, signals from the target and background will be additive and the total signal will be determined as

$$A(t)F(t) \rightarrow n(t) + :s(t),$$

y(t) = if there is a signal, and

A(t)F(t) + n(t),

if there is not a signal.



Figure 2: Realization of random functions $\phi(t)$, F(t), and A(t).

Without disturbing the generality of the examinat on, we solve the problem in the discrete form, assuming that the observ tion of process y(t) occurs at discrete time moments

$$t_1, t_2 \dots t_N; t_i = t_1 + (i-1) \Delta t_i$$

and

19-16518-000

$$y_m = y(t_m)$$

are values of the observed process at reference points. We shall select the separation interval so that there is a sufficient number of readings for the duration of the signal, i.e., so that $\delta / \Delta t = k \gg 1$. Receiver noise will be given by the sequence n_j of normal, mutually independent values with parameters $(0, \mathbf{C})$. In examining the problem in a timediscrete sense, and assuming that $k \gg 1$, it is natural to assume that both the signal arrival time and the time moments of transition of process F(f) from one set of states to another correspond with selected t, readings. Whence it follows that interference values F(t) are completely determined by the set of 4k-1 of f, numbers, where $i = -(2k - 1) \dots$ (2k - 1), which are computed with the formulas

$$f_{0} = 0,$$

$$f_{1} = \int_{-i.2}^{-i/2+1_{0}i} S_{0}(y) \, dy / \int_{-i.2}^{i.2} S_{0}(y) \, dy;$$

$$i = 1, 2 \dots k,$$

$$f = f_{2k-i}; \ i = k+1, \ k+2 \dots 2k-1; \ (5)$$

$$f_{1} = -f_{-i}, \ i < 0.$$

The vilues

$$f_{1}, f_{2} \dots f_{k} (f_{-1}, f_{-2} \dots f_{-(k-1)}) \text{ and } \\ f_{-1}, f_{k+2} \dots f_{2k-1} (f_{-1}, f_{-2} \dots f_{-(2k-1)})$$

correspond respectively to increasing (falling in the negative region) and fallin;; (increasing in the negative region) interference pulse fronts; f_k and f_{k} are extremal internal values.

The s.gnal will be given by vector

$$S_i = \varepsilon(S_i, S_2 \dots S_{rk}),$$

where

$$s_i = s_0 (-\delta/2 + i\Delta t); \quad i = 1, 2, 3 \dots 2k.$$

Let us study the statistical properties of process $F_m = F(t_m)$, $m = 1, 2 \dots N$. Condition (1) indicates that the telegraphic signal, received in a random one-dimensional background cross section, is a Markov process. Process F_m by virtue of its structure resembles the telegraphic signal and differs from it only in that its fronts are smoothed and in that it takes both positive and negative values. Therefore we shall also examine process F_m as a random Markov sequence with

possible f, states and with consequent transition probabilities. From

state f_0 the only possible transition is to states f_0 , f_{-1} , f_1 with probabilities α' , $\alpha/2$. $\alpha/2$ respectively ($\alpha' = 1 - \alpha$). From state f_k , process F_m with probability β turns into f_{k+1} and with probability $\beta' =$ $1 - \beta$ remains f_k . From states f_j , where $j \ge 1$; $j \ne k$; $f \ne 2k-1$, with probability 1 a transition takes place to state f_{j+1} . Finally, from state f_{2k-1} , the process with probability 1 turns into f_0 . Analogous relationships are true for states f_{-i} , where $j \ge 0$.

Since A and F in the aggregate form a Markov process, the 3-d:-menional vector

$$\vec{w}_r = (y_r, A_r, F_r)$$

also forms a Markov process. Therefore the detection problem obtained here in discrete form fits totally into the problem scheme used in the detection of a weakly determined signal, examined in reference [5]. In accordance with the result obtained in reference [5], the optimal re- [6 ceiver consists of circuits connected in series, located right after the unit to eliminate the constant component of interference.

1. Of a nonlinear transformation effecting "identification" of the inte ference and transforming input realization y into the sequence

$$y_{r}^{\bullet} := y_{r} \cdots E(\Phi_{r} | \hat{y}_{r}) := y_{r} - E(A_{r} F_{r} | \hat{y}_{r})$$
$$= y_{r} - \sum_{i=1}^{p} \sum_{j=-i(2k-1)}^{2k-1} a_{i} f_{j} c_{i}(a_{i}; f_{j} | \hat{y}_{r}), \quad (b)$$

Reproduced from best available copy.

where vector

C.U.S. C.C.C.C.C.

the estimate of the current interference value

$$E(\Phi_{i}|\mathbf{y}_{i})$$

 $\mathbf{y}_r = (\mathbf{y}_1, \mathbf{y}_2, \dots, \mathbf{y}_r);$

is the conditional expection value Φ_r , computed on the basis of all observations known to the current time moment and in the assumption that the signal is absent;

$$\omega(a_i; f_j | y_i) = P(A_r = a_i; F_r = f_i | y_i)$$

are a posteriori probabilities determining the conditional expectation

 $E(\Phi_r|\vec{y}_r).$

2. Of a linear filter consistent with the form of the signal received.

ŝ

3. Of a threshold unit.

The nonlinear transformation (6) plays the major role. Since vector (y_1, A_1, F_1) is a Markov process, the a posteriori probabilities that figure in (6)

$$\omega(a_i; f_j \ y_i); i = 1, 2 \dots p;$$

 $j = -(2k-1) \dots (2k-1)$

form a conditional Markov process [6]. This very important property makes it possible to determine the conditional probabilities without the need of remembering all the current values of vector y_{1} . The method of computing these probabilities has been quite well developed [6] consequently, omitting intermediate calculations, we shall give the final relationships

$$\begin{split} & \omega \left(a_{i}; f_{v} | \vec{y}_{r} \right) = \frac{Q \left(y_{r} - a_{l} f_{0} \right)}{I_{r}} \left[x'^{\alpha_{0}} \left(a_{i}; f_{v} | \vec{y}_{r-1} \right) + \frac{1}{2} \right] \\ & + \omega \left(a_{i}; f_{(2k-1)} | \vec{y}_{r-1} \right) + \omega \left(a_{i}; f_{2k-1} | \vec{y}_{r-1} \right) \right]; \\ & \omega \left(a_{i}; f_{\pm 1} | \vec{y}_{r} \right) = \\ & = \frac{aQ \left(y_{r} - a_{l} f_{0} \right)}{2pI_{r}} \sum_{l=1}^{p} \omega \left(a_{j}; f_{0} | \vec{y}_{r-1} \right), \quad (7) \\ & \omega \left(a_{i}; f_{\pm 1} | \vec{y}_{r} \right) = \frac{Q \left(y_{r} - a_{l} f_{-1} \right)}{I_{r}} \omega \left(a_{i}; f_{\pm (l-1)} | \vec{y}_{r-1} \right); \quad (7) \\ & \omega \left(a_{i}; f_{\pm 1} | \vec{y}_{r} \right) = \frac{Q \left(y_{r} - a_{l} f_{-1} \right)}{I_{r}} \omega \left(a_{i}; f_{\pm (l-1)} | \vec{y}_{r-1} \right); \quad (7) \\ & \omega \left(a_{i}; f_{-k} | \vec{y}_{r} \right) = \frac{Q \left(y_{r} - a_{l} f_{\pm k} \right)}{I_{r}} \omega \left(a_{i}; f_{\pm (l-1)} | \vec{y}_{r-1} \right); \\ & \omega \left(a_{i}; f_{-k} | \vec{y}_{r} \right) = \frac{Q \left(y_{r} - a_{l} f_{\pm k} \right)}{I_{r}} \left[\omega \left(a_{i}; f_{\pm (k-1)} | \vec{y}_{r-1} \right) + \frac{3}{2} \omega \left(a_{i}; f_{\pm k} | \vec{y}_{r-1} \right) \right]; \\ & \omega \left(a_{i}; f_{\pm (k+1)} | \vec{y}_{r} \right) = \frac{3Q \left(y_{r} - Q_{l} f_{\pm (k+1)} \right)}{I_{r}} \omega \times \\ & \times \left(u_{i}; f_{\pm k} | \vec{y}_{r-1} \right); \\ & \text{where} \qquad Q \left(x \right) = \left(2\pi z^{2} \right)^{-1/2} \exp - \frac{x^{2}}{2z^{2}}, \\ & & & \\ \hline & & \\ \hline$$

while the norming factor I_r is determined ; rom the equation

Red Barberry Barberry Contraction

$$\sum_{i=1}^{p} \sum_{j=-(2k-1)}^{2k-1} o(a_j; f_j | \vec{y}_j) = 1.$$
 (8)

As seen from (8) and (9), the a poste iori probabilities

$$\omega(a_i; f_j|\mathbf{y}_r)$$

are computed recursively from the current observation of y, and from the values of these same probabilities found for the preceding time moment r-1.

Analysis of the sensitivity and anti-interference capabilities of the detection procedure proposed is beyond the scope of the present article.

REFERENCES

1. Раковский Ю. Н., ОМП, 1966, № 4, гр. 1. 2. Шестов Н. С. Выделение оплических си налов на фоне случайных помех. М., «Советское ради» . 1967. 3. Левшии В. Л. Пространственная фил грация в оптических системах пеленгации. М., «Советсь зе ра-тися 1971.

дно», 1971. 4. Деньщиков К. К., Приборостроение, 1969.

№ 10, стр. 38.
5. Резник М. Х., Радиотехника и электроника.
1971, № 7, стр. 1282
6. Хазен Э. М. Методы оптимальных статисти ческих решений. М., «Советское радио», 1958

Submitted for editing 21 September 1971.



ţ