STOCHASTIC MODELING OF BIOCHEMICAL REACTIONS

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Abstract

The most common theoretical approach to model the interactions in a biochemical process is through chemical reactions. Often for these reactions, the dynamics of the first $M$-order statistical moments of the species populations do not form a closed system of differential equations, in the sense that the time-derivatives of first $M$-order moments generally depend on moments of order higher than $M$. However, for analysis purposes, these dynamics are often made to be closed by approximating the needed derivatives of the first $M$-order moments by nonlinear functions of the same moments. These functions are called the moment closure functions.

This paper presents a systematic procedure to construct these moment closure functions. This is done by first assuming that they exhibit a certain separable form, and then matching time derivatives of the exact (not closed) moment equations with that of the approximate (closed) equations for some initial time and set of initial conditions. Using these results a stochastic model for gene expression is investigated. We show that in gene expression mechanisms, in which a protein inhibits its own transcription, the resulting negative feedback reduces stochastic variations in the protein populations.

1 Introduction

This paper presents a systematic procedure for constructing approximate stochastic models for chemical reactions used for modeling biochemical processes such as gene regulatory networks. Such models are motivated by the recent work [1, 2] which provide considerable experimental evidence for stochastic fluctuations in gene expression and regulation and may account for the large amounts of cell to cell variation observed in genetically identical cells exposed to the same environment conditions [3, 4]. Furthermore, studies of engineered genetic circuits designed to act as toggle switches or oscillators have revealed large stochastic effects. Stochastically is therefore an inherent feature of biological dynamics and developing stochastic models which capture this stochastically have become increasingly important. Analysis of such model not only helps us to discover the benefits that biology draws from stochasticity but also helps to explore its use for other applications, for example, design of stochastic bio-inspired decision algorithms to control the motion of networks of artificial mobile agents such as small autonomous micro-UAVs or UGVs involved in surveillance and/or tracking missions. Although our work focuses on biochemical reactions, the modeling tools developed in this paper can be applied to a very general class of stochastic systems, in particular Markovian systems whose state evolves due to events triggered by stochastic processes. Other examples of such systems include ecological systems where these events correspond to births, recruitment or deaths and in networked control systems where the events correspond to exchanges of messages, communication faults, etc.

The most common theoretical approach is to model the interactions in a biochemical process as chemical reactions. The stochasticity in the biochemical process is then captured using the stochastic formulation of chemical kinetics which treats the reactions as probabilistic events. The time evolution of the system is then described by a single equation for a probability density function, where time and species populations appear as independent variables, called the Chemical Master Equation (CME) [5, 6]. This equation can only be solved analytically for relatively few, highly idealized
**Stochastic Modeling Of Biochemical Reactions**

1. **REPORT DATE**  
   01 NOV 2006

2. **REPORT TYPE**  
   N/A

3. **DATES COVERED**  
   -

4. **TITLE AND SUBTITLE**  
   Stochastic Modeling Of Biochemical Reactions

5. **AUTHOR(S)**  
   Department of Electrical and Computer Engineering  
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6. **PERFORMING ORGANIZATION NAME(S) AND ADDRESS(ES)**  
   Department of Electrical and Computer Engineering  
   University of California, Santa Barbara, CA 93101

7. **SPONSORING/MONITORING AGENCY NAME(S) AND ADDRESS(ES)**  
   -

8. **PERFORMING ORGANIZATION REPORT NUMBER**  
   -

9. **DISTRIBUTION/AVAILABILITY STATEMENT**  
   Approved for public release, distribution unlimited

10. **ABSTRACT**  
    See also ADM002075.

11. **SUPPLEMENTARY NOTES**  
    See also ADM002075.

12. **SUBJECT TERMS**  
    -

13. **SECURITY CLASSIFICATION OF:**  
    a. REPORT  
    unclassified

    b. ABSTRACT  
    unclassified

    c. THIS PAGE  
    unclassified

14. **LIMITATION OF ABSTRACT**  
    UU

15. **NUMBER OF PAGES**  
    7

16. **NAME OF RESPONSIBLE PERSON**  
    -
cases and generally Monte Carlo simulation techniques are used which are also known as the Stochastic Simulation Algorithm (SSA) to study stochasticity in bio-chemical reactions [7, 8, 4, 9, 10]. Since one is often interested in only the first and second order statistical moments for the number of molecules of the different species involved, much effort can be saved by applying approximate methods to produce these low-order moments, without actually having to solve for the probability density function. Various such approximate methods have been developed, for example, using the Fokker-Plank approximation, expanding the Master equation, etc [5, 11]. In this paper, an alternative approximate method for estimating lower-order moments is introduced using moment closure techniques.

We show that the biochemical reactions can be conveniently modeled as a Stochastic Hybrid System (SHS), the state of which are the populations of different species involved in the reactions [12]. Then, the time evolution of the moments of the population is obtained using results from the SHS literature. It has been shown that for reactions with more than one reactants, time evolution of the first \( M \) order moments of the population is not closed, in the sense that it depends on moments of order higher than \( M \). For analysis purpose, the time evolution of the first \( M \) order moments is made to be closed by approximating these higher order moments as a nonlinear function of moments up to order \( M \), which we refer to as the moment closure function. This paper presents explicit formulas to construct these moment closure functions using the recently introduced techniques [13, 14] of matching time derivatives between the exact (not closed) moment equations with that of the approximate (closed) equations for some initial time and set of initial conditions. The striking feature of these formulas are that they are independent of the reaction parameters (reaction rates and stoichiometry), moreover, the accuracy of the approximation can be improved by increasing \( M \).

These closed moment equations provide time evolution of lower order moments for populations of species involved in the biochemical reactions. Apart from providing fast simulation times and lesser computation burden compared to Monte Carlo simulations these approximate models also open the doors to other types of analysis tools, for example, sensitivity analysis of chemical master equation with respect to reaction parameters. However, they provide lesser information about the probability distribution as compared to Monte Carlo simulations. To illustrate the applicability of our results, we investigate a stochastic model for gene expression. We show that a negative feedback mechanism caused by the expressed protein inhibiting its own transcription, reduces stochastic variations in the protein populations.

2 Stochastic Modeling of Chemically Reacting Systems

Consider a system of \( n \) species \( X_j, j \in \{1, \ldots, n\} \) inside a fixed volume \( V \) involved in \( K \) reactions of the form

\[
R_i : u_{i1}X_1 + \ldots + u_{in}X_n \xrightarrow{c_i} v_{i1}X_1 + \ldots + v_{im}X_m + \ast
\]

for all \( i \in \{1, \ldots, K\} \), where \( u_{ij} \in \mathbb{N}_{\geq 0} \) is the stoichiometry associated with the \( j^{th} \) reactant in the \( i^{th} \) reaction and \( v_{ij} \in \mathbb{N}_{\geq 0} \) is the stoichiometry associated with the \( j^{th} \) product in the \( i^{th} \) reaction, and \( \ast \) represents products other than the species \( X_j \). As all chemical reactions occur in a series of elementary reactions [15], which are generally uni- or bi-molecular, we assume

\[
u_{i1} + \ldots + u_{in} \leq 2, \ \forall i \in \{1, \ldots, K\},
\]

and hence, we only allow reactions which have the form given in the first column of Table 1. The reaction parameter \( c_i \) characterizes the reaction \( R_i \) and, together with the stoichiometry, defines the probability that a particular reaction takes place in an “infinitesimal” time interval \( [t, t+dt] \). This probability is given by the product \( c_i h_i dt \) where \( h_i \) is the number of distinct molecular reactant combinations present in \( V \) at time \( t \) for the reaction \( R_i \) and \( c_i dt \) is the probability that a particular reactant combination of \( R_i \) will actually react on \( [t, t+dt] \). The number, \( h_i \), depends both on the reactants stoichiometry \( u_{ij} \) in \( R_i \) and on the number of reactant molecules in \( V \). Table 1 shows the value of \( h_i \) for different reaction types [7]. In this table and in the sequel, we denote by \( x_j \), the number of molecules of the species \( X_j \) in the volume \( V \) and \( x := [x_1, \ldots, x_n]^T \in \mathbb{R}^n \).

To model the time evolution of the number of molecules \( x_1, x_2, \ldots, x_n \), a special class of Stochastic Hybrid Systems (SHS) were introduced in [12]. More specifically, to fit the framework of our problem, these system are characterized by trivial dynamics

\[
x = 0, \ x = [x_1, \ldots, x_n]^T,
\]

2
Table 1: \( h_i(x) \) for different reaction types.

<table>
<thead>
<tr>
<th>Reaction ( R_i )</th>
<th>( h_i(x) )</th>
</tr>
</thead>
<tbody>
<tr>
<td>( X_j \rightarrow * )</td>
<td>( x_j )</td>
</tr>
<tr>
<td>( X_j + X_i \rightarrow * )  ( (t \neq j) )</td>
<td>( x_j x_i )</td>
</tr>
<tr>
<td>( 2X_j \rightarrow * )</td>
<td>( \frac{1}{2}x_j(x_j - 1) )</td>
</tr>
</tbody>
</table>

a family of \( K \) reset maps

\[
x = \phi_i(x^-), \quad \phi_i : \mathbb{R}^n \rightarrow \mathbb{R}^n,
\]

and a corresponding family of \( K \) transition intensities

\[
\lambda_i(x), \quad \hat{\lambda}_i : \mathbb{R}^n \rightarrow [0, \infty)
\]

for all \( i \in \{1, \ldots, K\} \). Each of the reset maps \( \phi_i(x) \), and corresponding transition intensities \( \lambda_i(x) \) are uniquely defined by the \( i^{th} \) reaction and given by

\[
x \mapsto \phi_i(x) = \begin{bmatrix}
x_1 - u_{i1} + v_{i1} \\
x_2 - u_{i2} + v_{i2} \\
\vdots \\
x_n - u_{in} + v_{in}
\end{bmatrix}, \quad \lambda_i(x) = c_i h_i(x)
\]

for all \( i \in \{1, \ldots, K\} \). In essence, if no reaction takes place, the state remains constant and whenever the \( i^{th} \) reaction takes place, \( \phi_i(x) \) is “activated” and the state \( x \) is reset according to (6), furthermore, the probability of the activation taking place in an “infinitesimal” time interval \([t, t + dt]\) is \( \hat{\lambda}_i(x)dt \).

### 3. Moment Dynamics

Given a vector \( m = (m_1, m_2, \ldots, m_n) \in \mathbb{N}_{\geq 0}^n \) of \( n \) greater than equal to zero integers, we define the (uncentered) moment of \( x \) associated with \( m \) to be

\[
\mu^{(m)}(t) = \mathbb{E}[x^{(m)}(t)], \quad \forall t \geq 0
\]

where \( \mathbb{E} \) stands for the expected value and

\[
x^{(m)} := x_1^{m_1} x_2^{m_2} \cdots x_n^{m_n}.
\]

The sum \( \sum_{j=1}^n m_j \) is called the order of the moment \( m \).

Using results from the SHS literature, more specifically by applying Theorem 1 in [16] to the SHS (3)-(5), one can show that the time derivative of a moment \( \mu^{(m)} \) is

\[
\dot{\mu}^{(m)} = \mathbb{E} \left[ \sum_{i=1}^K (\phi_i(x^{(m)}) - x^{(m)}) \lambda_i(x) \right].
\]

If any of the reactions have more than one reactants, i.e, has the form \( X_j + X_i \xrightarrow{c_i} * \) for which \( \hat{\lambda}_i(x) \) is quadratic in \( x \), then, from (9) it can be shown that the time derivative of a moment of order \( m^* \), is given by a linear combination of moments of orders upto \( m^* + 1 \) [17]. Hence, if one stacks all moments in an infinite vector

\[
\mu_\infty = [\mu^{(m_1)}, \mu^{(m_2)}, \ldots]^T, \quad m_p \in \mathbb{N}_{\geq 0}, \quad \forall p \in \{1, 2, \ldots\}
\]

its dynamics can be written as

\[
\dot{\mu}_\infty = A_\infty \mu_\infty,
\]

for some infinite matrix \( A_\infty \). As the above infinite dimensional system cannot be solved analytically, we truncate (11) by creating a vector

\[
\mu = [\mu^{(m_1)}, \mu^{(m_2)}, \ldots, \mu^{(m_k)}]^T \in \mathbb{R}^k
\]

containing the top \( k \) elements of \( \mu_\infty \) which correspond to the lower-order moments of interest. Then, (11) can be rewritten as

\[
\dot{\mu} = I_k \otimes A_\infty \mu_\infty = A_\mu + B_\hat{\mu}
\]

where \( \mu \in \mathbb{R}^k \), \( I_k \otimes \) denotes a matrix composed of the first \( k \) rows of the infinite identity matrix and \( \hat{\mu} \in \mathbb{R}^r \) contains all the moments that appear in the first \( k \) elements of \( A_\infty \mu_\infty \) but that do not appear in \( \mu \).

In this paper we let the vector \( \mu \in \mathbb{R}^k \) contain all the moments of \( x \) of order up to \( M \in \mathbb{N}_{\geq 2} \), i.e., we consider an \( M^{th} \) order truncation. With this, the evolution of vector \( \mu \) can be written as (13) for some matrices \( A \) and \( B \) with \( \hat{\mu} \in \mathbb{R}^r \) being a vector of moments of order \( M + 1 \). Our goal now is to approximate (13) by a finite-dimensional nonlinear ODE of the form

\[
\dot{v} = Av + B\phi(v), \quad v = [v^{(m_1)}, v^{(m_2)}, \ldots, v^{(m_k)}]^T
\]

where the map \( \phi : \mathbb{R}^k \rightarrow \mathbb{R}^r \) should be chosen so as to keep \( v(t) \) close to \( \mu(t) \). This procedure is commonly referred to as moment closure. We call (14) the truncated moment dynamics and each element \( \phi^{(m)}(\mu) \) of \( \phi(\mu) \) the moment closure function for the corresponding element \( \mu^{(m)} \) in \( \hat{\mu} \).

The procedure we adopt to construct these moment closure functions is to first assume a certain separable form for each element \( \phi^{(m)}(\mu) \) of \( \phi(\mu) \) and then matching time derivatives of \( \mu \) and \( v \) at some initial time \( t_0 \), for every deterministic initial condition.
of the form $x(t_0) = \tilde{x}$ with probability one. The choice of initial conditions is justified by the fact that the class of deterministic distributions forms a natural basis for the infinite dimensional space $\Omega_\omega$ containing every possible state $\mu_\omega$ of (11). Referring the reader to [17] for further details the above procedure leads to the following moment closure functions. Assume that each entry $\phi^{(m)}(\mu)$ of $\phi$ has the following separable form

$$\phi^{(m)}(v) = \sum_{p=1}^{k} \left( \nu^{(m \rho)} \right)_{\gamma} = v^{(\gamma)} , \quad \gamma = (\gamma_1, \ldots, \gamma_k) \tag{15}$$

with $\gamma_p$ being the unique solution of the following system of linear equations\footnote{$C^{(m)}$ is defined as follows: $\forall \ell, h \in \mathbb{N}_{>0}$}

$$C^{(m)}_{(\nu s)} = \sum_{p=1}^{k} \gamma_p C^{(m \rho)}_{(\nu s)} , \quad \forall s = \{1, \ldots, k\} \tag{16}$$

where we define for vectors $\hat{m} = (\hat{m}_1, \ldots, \hat{m}_n)$ and $\check{m} = (\check{m}_1, \ldots, \check{m}_n)$

$$C^{(m)} := C^{(m \check{m})}_{\hat{m}_1 \check{m}_2 \cdots \check{m}_n} \tag{17}$$

Then, when $\mu(t_0) = v(t_0)$, we have

$$\frac{d^i v(t_0)}{dt^i} = \frac{d^i v(t_0)}{dt^i} + e_i(\tilde{x}), \quad \forall i \geq 1$$

for every deterministic initial conditions of the form $x(t_0) = \tilde{x}$ with probability one for every integer $\tilde{x}$. In the equation above, each entry of the vector $\frac{d^i v(t_0)}{dt^i}$ is a polynomial in $\tilde{x}$, whose degree exceeds by at least $M$ the degree of the corresponding entry of the error vector $e_i(\tilde{x})$. Thus, with increasing $M$, the truncated moment dynamics $v(t)$ provides a more accurate approximation to the moments in $\mu(t)$. The striking feature of the moment closure constructed is that they are independent of the reaction parameters (reaction rates and stoichiometry) and moreover the dependence of higher-order moment on lower order ones is consistent with $x$ being jointly lognormally distributed, in spite of the fact that the procedure used to construct $\phi$ did not make any assumption on the distribution of the population. For a three-specie reaction ($n = 3$) and a second order truncation $M = 2$, i.e.,

$$\mu = \left[ \mu^{(1,0,0)}, \mu^{(0,1,0)}, \mu^{(0,0,1)}, \mu^{(2,0,0)}, \mu^{(0,2,0)}, \mu^{(0,0,2)} \right],$$

$$\begin{array}{c} 
\mu^{(1,1,0)}, \mu^{(1,0,1)}, \mu^{(0,1,1)} \end{array} \tag{18}$$

Table 2 lists moment closure functions obtained from (16), which approximate different third order moments of $x$ in terms of the first two.

<table>
<thead>
<tr>
<th>$\mu^m$</th>
<th>$\phi^{(m)}(\mu)$</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\mu^{(3,0,0)}$</td>
<td>$\left( \begin{array}{c} \mu^{(2,0,0)} \ \mu^{(1,0,0)} \end{array} \right)^3$</td>
</tr>
<tr>
<td>$\mu^{(2,1,0)}$</td>
<td>$\left( \begin{array}{c} \mu^{(2,0,0)} \ \mu^{(1,0,0)} \end{array} \right)^2 \mu^{(1,0,0)}$</td>
</tr>
<tr>
<td>$\mu^{(1,1,1)}$</td>
<td>$\mu^{(1,0,1)} \mu^{(1,0,0)} \mu^{(0,1,0)}$</td>
</tr>
</tbody>
</table>

### 4 Modeling of Gene Expression with negative feedback

In this section we consider negative feedback in gene expression caused by the protein inhibiting its own transcription. Such auto-regulatory networks are common means of stabilizing protein levels in biochemical pathways [18]. This negative feedback is realized by first assuming that there is an activator that activate a gene to transcribe a mRNA. Proteins are translated from the gene at a constant rate. These proteins can then interact with the activator to change its shape, making it incapable of activation. Hence, more protein implies lesser number of activators, and hence, lesser transcription.

Denoting the mRNA, protein and the activator by $X_1$, $X_2$ and $X_3$, respectively, the above interactions can be written as

$\text{Gene} + X_3 \xrightarrow{K_a} \text{Gene} + X_3 + X_1,$

$X_1 \xrightarrow{d_p} *, \quad X_1 \xrightarrow{K_p} X_1 + X_2, \quad X_2 \xrightarrow{d_p} *$

$* \xrightarrow{K_a} X_3, \quad X_3 \xrightarrow{d_a} *$

$X_3 + X_2 \xrightarrow{k_{10}} X_2 + *.$


Hence the mRNA $X_1$ is transcribed from the gene at a rate $K_x x_3$, where $x_3$ denotes the number of molecules of the activator. The protein $X_2$ is translated from the mRNA at a constant rate $K_p$. Both mRNA and the protein decay at rates $d_x$ and $d_p$ respectively. $K_a$ and $d_a$ denote the birth and death rate of the species $X_3$, respectively. The rate at which the protein $X_2$ reacts with the activator $X_3$ and takes it away from the transcription process is given by $k_1$. Note that increasing values of $k_1$ denote larger negative feedback. If $k_1 = 0$ then there is no negative feedback. The deterministic chemical rate equations are then given by

\[ \begin{align*}
    x_{1d} &= K_x x_3 - d_x x_{1d}, \quad x_{2d} = K_p x_{1d} - d_p x_{2d} \quad (19a) \\
    x_{3d} &= d_a + k_1 x_{2d} x_3 d_3 \quad (19b)
\end{align*} \]

where $x_{1d}$, $x_{2d}$ and $x_{3d}$ are continuous deterministic approximates of $x_1$, $x_2$ and $x_3$, respectively. At steady state we have

\[ \frac{K_a}{d_a} = [1 + \frac{k_1}{d_a} x_{2d}(\infty)] x_{1d}(\infty). \quad (20) \]

As this deterministic model does not provide us with any information about the stochasticity in the protein population, we turn to a stochastic formulation. The above interactions can be modeled as a SHS with reset maps

\[ x \mapsto \phi_1(x) = \begin{bmatrix} x_1 + 1 \\ x_2 \\ x_3 \end{bmatrix}, \quad x \mapsto \phi_2(x) = \begin{bmatrix} x_1 - 1 \\ x_2 \\ x_3 \end{bmatrix} \]

\[ x \mapsto \phi_3(x) = \begin{bmatrix} x_1 \\ x_2 + 1 \\ x_3 \end{bmatrix}, \quad x \mapsto \phi_4(x) = \begin{bmatrix} x_1 \\ x_2 - 1 \\ x_3 \end{bmatrix} \]

\[ x \mapsto \phi_5(x) = \begin{bmatrix} x_1 \\ x_2 \\ x_3 + 1 \end{bmatrix}, \quad x \mapsto \phi_6(x) = \begin{bmatrix} x_1 \\ x_2 \\ x_3 - 1 \end{bmatrix} \]

\[ x \mapsto \phi_7(x) = \begin{bmatrix} x_1 \\ x_2 \\ x_3 - 1 \end{bmatrix} \]

and corresponding transition intensities given by

\[ \lambda_1(x) = K_x x_3, \quad \lambda_2(x) = d_x x_1, \quad \lambda_3(x) = K_p x_1, \quad (22) \]

\[ \lambda_4(x) = d_p x_2, \quad \lambda_5(x) = K_a, \quad \lambda_6(x) = x_3 d_a \quad (23) \]

\[ \lambda_7(x) = k_1 x_2 x_3. \quad (24) \]

Using (9) the time derivative of the the vector $\mu$ which is given by (18) is

\[ \dot{\mu} = \bar{A} \mu + A \lambda \mu + B \dot{\mu}. \quad (25) \]

for some matrices $\bar{A}$, $A$, $B$ and

\[ \mu = [\mu^{(0,1,2)}, \mu^{(0,2,1)}, \mu^{(1,1,1)}]^T. \]

Using Table 2 the above system can be closed using the following approximations

\[ \begin{align*}
    \mu^{(0,1,2)} &= \left( \frac{\mu^{(0,0,2)}}{\mu^{(0,1,0)}} \right)^2 \\
    \mu^{(0,2,1)} &= \left( \frac{\mu^{(0,0,2)}}{\mu^{(1,0,0)}} \right)^2 \\
    \mu^{(1,1,1)} &= \frac{\mu^{(1,1,0)} \mu^{(0,1,1)}}{\mu^{(0,1,0)}}. \quad (26c)
\end{align*} \]

The statistical variations in the protein population can now be obtained from the closed system (25)-(26). In this paper, we use the steady state coefficient of variation defined by

\[ CV := \frac{\sqrt{E[X_2^2(\infty)]} - (E[X_2(\infty)])^2}{E[X_2(\infty)]} \]

\[ = \frac{\sqrt{\mu^{(0,2,0)}(\infty)} - (\mu^{(0,1,0)}(\infty))^2}{\mu^{(0,1,0)}(\infty)} \]

(28)

to quantify noise strength in the protein population. As analytical solution of the closed moment equations are too complicated to be of any use we investigate them numerically by taking

\[ K_r = 1, \quad K_p = 5 \text{ sec}^{-1}, \quad d_r = .1 \text{ sec}^{-1}, \quad d_p = .001 \text{ sec}^{-1}, \quad d_a = 100 \text{ sec}^{-1}. \]

For comparison purpose, we now vary parameters $k_1$ and $K_a$ so as to keep $x_{2d}(\infty) = 500$ fixed and see its effect on the coefficient of variation. This implies from (20) that $x_{3d}(\infty) = .01$ and for a given $k_1, K_a$ is always chosen such that

\[ K_d = d_a \left[ 1 + \frac{k_1}{d_a} x_{2d}(\infty) \right] x_{3d}(\infty). \quad (29) \]

Table 3 lists the values of $CV$ for different values of $k_1$ obtained from the steady states of (25)-(26). Note that $CV$ is lower with non-zero values of $k_1$ than with when there is no feedback, i.e $k_1 = 0$. Hence, we conclude that such negative feedback reduce stochastic variations in the protein numbers.
Table 3: Steady state coefficient of variation in the protein population CV obtained from (25)-(26) for different values of $k_1$.

<table>
<thead>
<tr>
<th>$k_1$</th>
<th>CV</th>
</tr>
</thead>
<tbody>
<tr>
<td>0</td>
<td>.32</td>
</tr>
<tr>
<td>.2</td>
<td>.28</td>
</tr>
<tr>
<td>.8</td>
<td>.26</td>
</tr>
<tr>
<td>1.8</td>
<td>.23</td>
</tr>
</tbody>
</table>

5 Conclusions and Future work

An approximate stochastic model for chemically reacting systems was presented in this paper. This was done by representing the population of various species involved in a set of chemical reactions as the continuous state of a SHS. With such a representation, the dynamics of the infinite vector containing all the statistical moments of the continuous state are governed by an infinite-dimensional linear system of ODEs, which we approximate by finite-dimensional nonlinear ODEs.

This typically involved approximating higher order statistical moments of the population in terms of the lower order moments. Using derivative matching techniques, explicit analytical formulas to construct these approximations were provided. Using then we showed that in gene expression, where a protein inhibits its own transcription, the resulting negative feedback reduces stochastic variations in the protein populations.

An interesting line of future work would be look at gene expression and regulation where the mRNA transcribed from the gene is not immediately accessible for translation. This would corresponds to gene expression in Eukaryotic cells where the mRNA is transported from the nucleus to the cytoplasm where translation takes place. This can be incorporated in our model by introducing another species that corresponds to the inactive mRNA and is being converted into an active mRNA at some constant rate. Another possible direction of future work would be to investigate stochasticity in gene cascade activation network where protein expressed by one gene activates another gene to express a second protein and negative feedback schemes within this network.

6 Acknowledgements

The authors gratefully acknowledge funding by the Institute for Collaborative Biotechnologies through grant DAAD19-03-D-0004 from the U.S. Army Research Office to support this work.

References


