THE AVERAGE SPEED OF A FAST VEHICLE MOVING IN A STREAM OF SLOW VEHICLES

by

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## REPORT DOCUMENTATION PAGE

<table>
<thead>
<tr>
<th>Field</th>
<th>Information</th>
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</thead>
<tbody>
<tr>
<td>1. REPORT NUMBER</td>
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<td>18. SUPPLEMENTARY NOTES</td>
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</tr>
<tr>
<td>19. KEY WORDS (Continue on reverse side if necessary and identify by block number)</td>
<td>TRAFFIC THEORY PROBABILISTIC MODELS</td>
</tr>
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20. Abstract continued.

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This paper studies the average speed of a fast test car moving in a stream of slow vehicles on a two-lane highway. Arrivals of the slow vehicles are assumed to follow a Poisson process and the test car arrives independently of the slow vehicles. The highway is assumed to consist of sections in which passing is possible and sections in which passing is impossible; the lengths of these sections are random variables. Two passing mechanisms are studied: the first assumes that the duration of a passing maneuver is a random variable while in the second passings are instantaneous.
1. Introduction

This paper studies the speed of a fast test car moving in a stream of slow vehicles in a two-lane two-way highway. Let "our direction" designate the direction in which the test-car is traveling. We assume that the highway consists of two alternating sections, sections of Type I in which passing is possible and sections of Type II in which passing is impossible. The highway is assumed to begin with a Type I section. Let $X_i$ and $Y_i$ denote the length of the $i$th Type I and Type II sections, respectively. We assume that $X_1, X_2, \ldots$ are i.i.d. distributed according to a continuous c.d.f. $G$ with an expectation $E[X]$. The random variables $Y_1, Y_2, \ldots$ are also assumed to be i.i.d.; they are distributed according to a continuous c.d.f. $H$ with an expectation $E[Y]$. The functions $G$ and $H$ depend on road conditions and on traffic moving in the opposite direction. We assume that the slow vehicles moving in our direction arrive at the highway according to a Poisson process with parameter $\lambda_1$. These zero size vehicles always maintain their free speed $v_1$. The test car arrives at the highway independently of the slow vehicles, and it has a free speed $v_2$ ($v_2 > v_1$).
As the test car travels along the road it occasionally comes up against slow vehicles. At these points the test car's driver may decide to pass the slow vehicle or otherwise to reduce his speed immediately to \( v_1 \) and to follow the slow vehicle for a while. The decisions to pass are made either at the points where the test car comes up against slow vehicles in Type I sections, or at the beginnings of the Type I sections to which the test car arrives traveling behind a slow vehicle. We assume that the driver's decision is dependent on the distance from the decision point to the end of the Type I section and independent of the distance he has already been following the slow vehicle. Two passing mechanisms are studied. The first mechanism assumes that at each decision point the test car's driver samples a required passing distance \( W_1 \) from a c.d.f. \( B \). If the distance to the end of the Type I section exceeds \( W_1 \) then the passing will take place \( W_1 \) units of distance from the decision point; otherwise the test car continues following the slow vehicle at least until the beginning of the next Type I section. In the second passing mechanism, the driver samples a r.v. \( W_2 \) from a c.d.f. \( C \). He passes instantaneously at the decision point if the distance to the end of the Type I section exceeds \( W_2 \); otherwise he follows the slow vehicle at least until the beginning of the next Type I section.


The paper is comprised of five sections. In Sections 2 and 3 we discuss the traveling of the test car in Type I sections under the first
and the second passing mechanisms, respectively. In Section 4 we determine the test car's average speed. Section 5 is a short summary.

2. The Movement of the Test Car in a Type I Section Under Passing Mechanism Number One

In this section we discuss the movement of the test car in the ith Type I section under passing mechanism number one. Let us assume temporarily, for the convenience of the analysis, that this section is infinitely long. Define that the test car is in state i (i=1,2) at a point along the road if it is moving there at speed \( v_i \). Denote by \( Z_n \) the distance the test car travels at a speed \( v_n \) for the jth time since entering the ith section. It was shown earlier in [1] that while moving at a speed \( v_2 \), the distance from the test car to the preceding slow vehicle is an exponential random variable with a parameter \( \lambda_1/v_1 \). From this result we obtain that \( Z_{1,1}, Z_{2,1}, \ldots \) are i.i.d. random variables distributed according to an exponential distribution function with parameter \( \alpha = \lambda_1(1/v_1 - 1/v_2) \). The random variables \( Z_1,1, Z_1,2, Z_1,3 \) are (by assumption) i.i.d. random variables distributed according to a c.d.f. \( B \).

Let us now denote by \( T(x) \) the time it takes the test car to arrive at a distance \( x \) from the beginning of the section; let \( M_1(x) \) denote the state of the car at that point, and define \( U(x) = x/v_1 - T(x) \). A realization of \( T(x) \) and the corresponding \( U(x) \) for \( M_1(0) = 2 \) is given in Figure 1. It is easy to analyze \( T(x) \) using the analysis of \( U(x) \).

Denote

\[
q_{2j}(x,u) = P[U(x) \leq u, M_1(x) = j \mid M_1(0) = 2], \quad j=1,2,
\]

\[
Q_{2j}(\theta, \xi) = \int_{x=0}^{\infty} \int_{u=0}^{\infty} e^{-\theta x} e^{-\xi u} d_u q_{2j}(x,u) dx,
\]

and

-3-
Figure 1. A typical realization of $T(x)$ and $U(x)$ for $M_1(0) = 2$. 

\[ \text{slope} = \frac{1}{v_1} \]

\[ \text{slope} = \frac{1}{v_2} \]

\[ \text{slope} = \left( \frac{1}{v_1} - \frac{1}{v_2} \right) \]
Denoting $\beta = 1/v_1 - 1/v_2$, we obtain
\[
q_{21}(x, u) = \begin{cases} 
\frac{u/\beta}{u - \alpha z} [1 - B(x-z)]dz + 0 \\
\int \int ae^{-\alpha z} q_{21}(x-z-y, u-z\beta)dB(y)dz, & 0 < u < x\beta \\
q_{21}(x, \cdot), & x\beta < u
\end{cases}
\]

hence,
\[
Q_{21}(\theta, \xi) = \frac{[1 - B^*(\theta)]\alpha}{\theta [\alpha + \theta + \xi\beta - \alpha B^*(\theta)]},
\]

where
\[
B^*(\theta) = \int_{x=0}^{\infty} e^{-\theta x} dB(x).
\]

To obtain $Q_{22}(\theta, \xi)$, we notice that
\[
q_{22}(x, u) = \begin{cases} 
\frac{u/\beta}{u - \alpha z} x-z \int \int ae^{-\alpha z} q_{22}(x-z-y, u-z\beta)dB(y)dz, & 0 < u < x\beta \\
\int \int ae^{-\alpha z} q_{22}(x-z-y, u-z\beta)dB(y)dz + e^{-\alpha z}, & u = x\beta \\
q_{22}(x, \cdot), & u > x\beta
\end{cases}
\]

hence,
\[
Q_{22}(\theta, \xi) = [\alpha + \theta + \xi\beta - \alpha B^*(\theta)]^{-1}.
\]

Now we denote
\[
p_{ij}(x, t) = P[T(x) \leq t, M_1(x) = j \mid M_1(0) = i], \quad i, j = 1, 2
\]

and
\[ P_{ij}(0, \xi) = \int_{x=0}^{\infty} \int_{t=0}^{\infty} e^{-\theta x} e^{-\xi t} d_1 p_{ij}(x,t) dx, \quad i,j = 1,2; \quad \theta, \xi > 0. \]

To determine \( P_{ij}^*(\theta, \xi) \) we realize that

\[ p_{2j}(x,t) = \]

\[ P\left[U(x) > \frac{x}{v_1} - t, M_1(x) = j \mid M_1(0) = 2\right] + P\left[U(x) = \frac{x}{v_1} - t, M_1(x) = j \mid M_1(0) = 2\right]; \]

hence we obtain from (2.1)

\[
P_{21}^*(\theta, \xi) = \frac{\alpha\left[1 - B^*\left(\theta + \frac{\xi}{v_1}\right)\right]}{\left(\theta + \frac{\xi}{v_1}\right)\left[\alpha + \theta + \frac{\xi}{v_2} - \alpha B^*\left(\theta + \frac{\xi}{v_1}\right)\right]}, \quad (2.3)
\]

and from (2.2) we get

\[
P_{22}^*(\theta, \xi) = \left[\alpha + \theta + \frac{\xi}{v_2} - \alpha B^*\left(\theta + \frac{\xi}{v_1}\right)\right]^{-1}. \quad (2.4)
\]

Now we determine \( P_{ij}^*(\theta, \xi), j=1,2 \). For \( j=1 \) we have

\[
p_{11}(x,t) = \begin{cases} 
0 & , \quad t \leq x/v_2 \\
\frac{(t-x/v_2)/\beta}{\int_{y=0}^{x/v_2} p_{21}(x-y, t-y/v_1) dB(y)} & , \quad x/v_2 < t < x/v_1 \\
\frac{(t-x/v_2)/\beta}{\int_{y=0}^{x/v_2} p_{21}(x-y, t-y/v_1) dB(y) + 1-B(x)} & , \quad t = x/v_1 \\
p_{11}(x,*) & , \quad t > x/v_1
\end{cases};
\]

hence,

\[
P_{11}^*(\theta, \xi) = \frac{1 - B^*\left(\theta + \frac{\xi}{v_1}\right)}{\theta + \frac{\xi}{v_1}} + \frac{\alpha\left[1 - B^*\left(\theta + \frac{\xi}{v_1}\right)\right]B^*\left(\theta + \frac{\xi}{v_1}\right)}{\left(\theta + \frac{\xi}{v_1}\right)\left[\alpha + \theta + \frac{\xi}{v_2} - \alpha B^*\left(\theta + \frac{\xi}{v_1}\right)\right]}, \quad (2.5)
\]
and for $j=2$ we obtain

$$
\begin{align*}
p_{22}(x,t) &= \begin{cases} 
0, & t \leq \frac{x}{v_2} \\
\frac{(t-x/v_2)/\beta}{\int_{y=0}^{\infty} p_{22}(x-y, t-y/v_1)dB(y)} & x/v_2 < t < x/v_1 \\
p_{12}(x,\cdot), & x/v_1 < t
\end{cases}
\end{align*}
$$

hence,

$$
P_{12}(\theta, \xi) = \frac{B^*(\theta + \frac{\xi}{v_1})}{\alpha + \theta + \frac{\xi}{v_2} - \alpha B^*(\theta + \frac{\xi}{v_1})}.
$$

Let us now denote by $T_{ij}(x)$ the time spent by the test car in a Type I section of length $x$ given that $M_1(0)=i$ and $M_1(x)=j$. The c.d.f. of $T_{ij}(x)$ satisfies

$$
P[T_{ij}(x) \leq t] = \frac{P[T(x) \leq t, M_1(x)=j | M_1(0)=i]}{P[M_1(x)=j | M_1(0)=i]};
$$

hence, denoting by $R^*_i(x, \xi)$ the inverse Laplace transform of $P^*_i(\theta, \xi)$ and letting $r_{ij}(x) = P[M_1(x)=j | M_1(0)=i]$, we obtain

$$
E[T_{ij}(x)] = \left( \frac{3}{\alpha^2} R^*_{ij}(x,0) \right) / r_{ij}(x).
$$

The probabilities $r_{ij}(x), i,j = 1,2$, satisfy $r_{ij}(x) = R^*_{ij}(x,0)$. To invert $P^*_i(\theta, \xi)$ we have to specify the c.d.f. $B$. We assume that $B$ is the exponential distribution function with parameter $\eta$. The inversion can be carried out for other distribution functions, but the expressions are likely to be very messy. Using tables of Laplace transforms we obtain

$$
r_{21}(x) = \alpha [1 - \exp(-(\alpha+\eta)x)]/((\alpha+\eta)),
$$

$$
r_{11}(x) = [\alpha + \eta \exp(-(\alpha+\eta)x)]/((\alpha+\eta)).
$$
\[ r_{21}(x)E[T_{21}(x)] = \alpha \left[ x(\alpha/v_1 + \eta/v_2) + (\eta-\alpha)\beta \left( 1 - \exp\left\{ -(\alpha+\eta)x \right\} \right) \right] \frac{1}{(\alpha+\eta)}, \quad (2.10) \]
\[ r_{22}(x)E[T_{22}(x)] = \left[ x(\alpha/v_1 + \eta^2/v_2) + 2\left( 1 - \exp\left\{ -(\alpha+\eta)x \right\} \right) \alpha\beta/(\eta+\alpha) \right] \frac{1}{(\alpha+\eta)^2}, \quad (2.11) \]
\[ r_{11}(x)E[T_{11}(x)] = \left[ x(\alpha^2/v_1 + \alpha\eta/v_2) + 2\left( 1 - \exp\left\{ -(\alpha+\eta)x \right\} \right) \alpha\eta\beta/(\eta+\alpha) \right] \frac{1}{(\alpha+\eta)^2}, \quad (2.12) \]
\[ r_{12}(x)E[T_{12}(x)] = r_{21}(x)E[T_{21}(x)]n/\alpha. \quad (2.13) \]

3. The Movement of the Test Car in a Type I Section under Passing Mechanism Number Two

In this section we derive results similar to those of Section 2 when the passing is instantaneous upon the test car driver's decision to pass. We start with the analysis of the case where \( M_1(0)=2 \) and assume that \( X_1=x \). The distribution of \( T_{22}(x) \) can easily be determined because \( M_1(x)=2 \) means here that the test car is unimpeded in the \( i \)th Type I section; hence,
\[ T_{22}(x) = x/v_2, \] with probability one. \( (3.1) \)

Now we assume without loss of generality that the test car arrives at the entrance of the \( i \)th Type I section at time zero. To obtain \( \{ T_{21}(x) < t \}, x/v_2 < t \leq x/v_1 \), the test car has to be unimpeded by slow vehicles arriving at this section during \( [-(x/v_1 - t), 0] \). Let \( J(x,t) \) denote the number of slow vehicles arriving at the \( i \)th Type I section during \( [-(x/v_1 - t), 0] \). It is known that given that \( J(x,t) = n \), then the epochs of the arrivals are independent and uniformly distributed on \( [-(x/v_1 - t), 0] \). Consequently we obtain
\[
P[T(x) \leq t \mid M_1(0) = 2, J(x,t) = n] = \left[ \frac{(x/v_1) - t}{t} \right]^{n} C\left\{ \frac{(t+y-x/v_2)}{\beta} \right\} \frac{dy}{(x/v_1) - t},
\]

where \( C\left\{ \frac{(x+y-x/v_2)}{\beta} \right\} \) is the probability that the test car passes immediately a slow vehicle that arrives at the section in \( \{-(x/v_1 - t) + y\} \).

Since \( J(x,t) \) is a Poisson random variable with parameter \( \lambda_1(x/v_1 - t) \), we obtain

\[
P[T(x) \leq t \mid M_1(0) = 2] = \exp \left[ -\lambda_1 \left\{ x/v_1 - t \right\} - \int_0^{1/v_1} C\left( \frac{(t+y-x/v_2)}{\beta} \right) dy \right].
\]  
(3.2)

From (3.2) we obtain

\[
r_{21}(x) = P[T(x) > x/v_2 \mid M_1(0) = 2] = 1 - \exp \left[ -\lambda_1 \left\{ x/v_1 - t \right\} - \int_0^{x/v_2} C\left( \frac{y}{\beta} \right) dy \right],
\]  
(3.3)

\[
r_{22}(x)E[T_{22}(x)] = \exp \left[ -\lambda_1 \left\{ x/v_1 - t \right\} - \int_0^{x/v_2} C\left( \frac{y}{\beta} \right) dy \right] x/v_2,
\]  
(3.4)

and

\[
r_{21}(x)E[T_{21}(x)] = \frac{x/v_1}{x/v_2} \int_{x/v_2}^{t} t \cdot d_{t} \{ P[T(x) \leq t \mid M(0) = 2] \}. \]  
(3.5)

Now we turn to determine the results associated with \( M_1(0) = 1 \). The derivation is based on the fact that

\[
P[T(x) \leq t \mid M_1(0) = 1] = \begin{cases} 
C(x)P[T(x) \leq t \mid M_1(0) = 2], & x/v_2 < t < x/v_1 \\
1, & t = x/v_1
\end{cases}
\]

from which we obtain

\[
r_{11}(x) = C(x)P[T(x) > x/v_2 \mid M_1(0) = 2] + (1 - C(x)),
\]  
(3.6)

\[
r_{11}(x)E[T_{11}(x)] = C(x)r_{21}(x)E[T_{21}(x)] + (1 - C(x))x/v_1,
\]  
(3.7)

and
\[ r_{12}(x)E[T_{12}(x)] = C(x)r_{22}(x)E[T_{22}(x)]. \] (3.8)

4. The Test Car's Average Speed

In this section we use the results of Sections 2 and 3 to obtain the test car's average speed under the two passing mechanisms. As we follow the test car's journey along the road we realize that its state at the beginnings of the Type I sections forms a Markov chain. The analysis is based on this property.

To complement the results on the movement of the test car in a Type I section we need to have similar results on the travel in Type II sections. For this purpose we denote by \( S_k(y), k=1,2, \) the time it takes the test car to travel \( y \) units of length along a Type II section given that \( M_2(0)=k \) \([M_2(u) \text{ designates the state of the test car } u \text{ units of length from the beginning of the Type II section}]. \) We also denote

\[ a_{ij}(y) = P[M_2(y)=j \mid M_2(0)=i]. \]

Here again we use the property that while the test car is in state 2 the distance between it and its preceding slow vehicle is an exponential random variable (parameter \( \lambda_1/v_1 \)) and obtain

\[ S_1(y) = y/v_1 \text{ with probability one,} \] (4.1)

\[ a_{11}(y) = 1, \] (4.2)

\[ a_{22}(y) = \exp(-\lambda_1 y) \] (4.3)

\[ P[S_2(y) \leq s] = \exp[-\lambda_1 (y/v_1 - s)], \quad y/v_2 \leq s \leq y/v_1. \] (4.4)

For a more detailed derivation of (4.3) and (4.4), see [1]. From (4.4) we obtain

\[ E[S_2(y)] = y/v_1 - [1 - \exp(-\lambda_1 y)]/\lambda_1. \] (4.5)
We are now in a position where we can sum up the results derived so far to obtain the test car's average speed. To this end, let $N_{ik}$ be the number of Type I sections to which the test car arrives at a speed $v_k$, $k=1,2$, while traveling up to the end of the $i$th Type II section. Denote by $\tau_k(j)$, $j=1,\ldots,N_{ik}$, the time it takes the test car to travel along a road section that consists of a Type I section and its following Type II section, given that the Type I section is the $j$th to which it arrives at speed $v_k$. The average speed at the end of the $j$th Type II section is given by

$$\bar{V}_j = \frac{\sum_{i=1}^{N_{ij}} \sum_{k=1}^{j} \tau_{ij}(k)}{\sum_{i=1}^{j} (X_i + Y_i)}.$$ (4.6)

We will determine

$$\lim_{j \to \infty} \bar{V}_j.$$ 

From (4.6) we obtain

$$\lim_{j \to \infty} \bar{V}_j = \lim_{j \to \infty} \frac{1}{j} \sum_{k=1}^{j} \tau_{ij}(k) \frac{2}{N_{ij}} \lim_{j \to \infty} \frac{1}{j} \sum_{i=1}^{N_{ij}} \frac{1}{j} \sum_{k=1}^{j} (X_i + Y_i).$$ (4.7)

The RHS of (4.7) calls for the use of the strong law of large numbers. Applying this law yields

$$\lim_{j \to \infty} \frac{1}{j} \sum_{k=1}^{j} (X_i + Y_i) = \frac{1}{E[X] + E[Y]},$$ (4.8)
\[
\lim_{k=1}^{N} \frac{\tau_{i}(k)}{N} = \frac{2}{\beta} \{ E_{X}[r_{1k}(X)E[T_{1k}(X)] \} + E[r_{1k}(X)] E_{Y}[E[S_{k}(Y)] \},
\]
(4.9)

and

\[
\lim_{j \to \infty} \frac{N_{ji}}{j} = \pi_{i},
\]
(4.10)

where \( \pi = (\pi_{1}, \pi_{2}) \) is the invariant distribution for the Markov chain of the state of the test car upon arrivals at Type I sections. The vector \( \pi \) is obtained as follows. Define

\[ Y_{ij} = E[r_{ij}(X)]; \quad \Gamma = \{Y_{ij}\} \]

and

\[ \phi_{ij} = E[a_{ij}(Y)]; \quad \phi = \{\phi_{ij}\}, \]

then \( \pi \) satisfies

\[ \pi(\Gamma \phi) = \pi \]

and

\[ \sum_{i=1}^{2} \pi_{i} = 1. \]

5. Summary

In the present paper we determined the average speed of a fast test car that is moving in a stream of slow vehicles. Two passing mechanisms were studied. The first mechanism assumes that after a driver decides to pass he still spends some time before reaching his free speed. The second mechanism, on the other hand, assumes instantaneous passing upon making the decision to pass. The first mechanism seems more realistic when the \( X' \)s are small with respect to the \( Y' \)s. Here the passings occur mainly at the beginnings of Type I sections and the test car has to accelerate before passing. The second mechanism seems more realistic.
when the X's are large with respect to the Y's and the test car comes up against slower vehicles mainly in Type I sections. The realism of the two mechanisms may also depend on road conditions and traffic congestion. We could have easily added a third passing mechanism that is a combination of the first two--instantaneous passing inside Type I sections and passing according to mechanism number one upon arrival at the beginning of a Type I section. However, we do not think that this addition makes a substantial contribution on top of the other two.

Finally, we would like to note that the current model cannot be used in cases where traffic is heavy because in these cases the assumption that slow vehicles arrive according to a Poisson process is not suitable.
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